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Genome-wide association study for tumour stage, grade, size, and age at diagnosis of non-muscle-invasive bladder cancer

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1 Genome-wide association study for tumour stage, grade, size, and age at diagnosis of non-
2 muscle-invasive bladder cancer

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26 **ABSTRACT**

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28 ***Background***

29 Non-muscle-invasive bladder cancer (NMIBC) causes a considerable health burden due to
30 the high recurrence and progression rates. Past studies have identified multiple candidate loci
31 associated with NMIBC prognosis, albeit lacking validation. Moreover, scarce reports exist
32 on genetic susceptibility to independent prognostic predictors of NMIBC, such as stage or
33 grade.

34 ***Objective***

35 To investigate genetic associations with NMIBC tumour and patient characteristics at the
36 time of diagnosis.

37 ***Design, Setting, and Participants***

38 A sample of 653 NMIBC cases come from the Bladder Cancer Prognosis Programme
39 (BCPP). Replication of the significant findings was conducted in the Nijmegen Bladder
40 Cancer Study (NBCS) cohort (N=1470).

41 ***Outcome Measurements and Statistical Analysis***

42 Genome-wide association study (GWAS) was carried out for outcomes of tumour size (as
43 continuous variable in centimetres), stage (Tis and T1 vs Ta), grade (G3 vs G2 and G1), and
44 age (as continuous (years) and dichotomous (70.2 years as a cut-off) variables).

45 ***Results and Limitations***

46 Significant ($P < 5E-08$) associations (N=61) with tumour size, stage, grade, and age were
47 identified in the GWAS discovery stage. None of the variants were independently
48 significantly associated in the replication cohort. A meta-analysis of both cohorts suggests
49 rs180940944 (13q13.3 locus, *NBEA*) was associated with tumour size as a continuous

50 variable ($\beta=0.9$ cm, $p=2.92E-09$). However, other SNPs in this region did not show evidence
51 of association in the meta-analysis.

52 ***Conclusions***

53 Our study suggests rs180940944 (*NBEA*) is associated with an increased NMIBC tumour size
54 at the time of diagnosis. Given study limitations, further replication is essential to validate the
55 finding.

56 ***Patient Summary***

57 Current study reports on a genome-wide association study on non-muscle-invasive bladder
58 cancer tumour and patient characteristics. We suggest *NBEA* gene might be associated with
59 increased tumour size at the time of diagnosis. The result must be replicated to establish
60 validity.

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75 INTRODUCTION

76 Urinary bladder cancer (UBC) accounts for 430 000 new cases worldwide annually, with 70-
77 80% of new cases presenting as non-muscle-invasive bladder cancer (NMIBC) [1]. NMIBC
78 causes significant burden on healthcare systems due to high recurrence and progression rates
79 (5-year recurrence rate: 50-70%, 5-year progression rate: 10-30%) [1]. Considerable clinical
80 improvements could be made by better, even personalised, prognostication and risk
81 stratification [1]. There have been several attempts to apply different approaches for accurate
82 disease prognostication, and although descriptive on a population-level, a substantial lack of
83 precision of individual outcomes remains [2], requiring ongoing improvement.

84 Few candidate-gene studies of UBC prognosis exist, with limited successful replication [3-5].
85 A recent study reported that out of 114 reported loci for UBC progression and prognosis,
86 only six single nucleotide polymorphisms (SNPs) showed significant associations in an
87 independent cohort, namely: NMIBC progression (rs6678136 (RGS4), rs11585883 (RGS5)),
88 recurrence among Bacillus Calmette–Guérin (BCG)-treated NMIBC patients (rs1799793
89 (ERCC2), rs187238 (IL18)), and muscle-invasive bladder cancer (MIBC) overall survival
90 (rs12035879 (RGS5), rs2075786 (TERT)) [3]. Powerful GWAS studies on NMIBC
91 prognosis show promise, but are still ongoing [6].

92 A previous attempt to include genetic variation failed to increase prognostic tool performance
93 [7], suggesting the issue is more complex. However, latter study utilised a relatively small
94 panel of SNPs (170,000), which has lower power of discovering significant loci in
95 comparison to genotype-imputed sets harbouring millions of variants for analysis [8]. The
96 inter-study lack of consensus might be due to several reasons: spurious findings, lack of
97 statistical power, and variation in outcome definition.

98 Other studies also suggest significant genetic signals might be only present for tumours of
99 certain grade or stage [9, 10]. However, reports on genetic associations for characteristics that

100 directly influence NMIBC outcome are scarce, precluding further investigations on their
101 relevance for NMIBC prognostication.

102 To provide more evidence on potential genetic associations, we have performed a GWAS on
103 key NMIBC characteristics (stage, grade, size of the tumour, EORTC risk category), as well
104 as age at the time of diagnosis within the West Midlands' Bladder Cancer Prognosis
105 Programme (BCPP) cohort including replication in the Nijmegen Bladder Cancer Study
106 (NBCS).

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125 MATERIAL AND METHODS

126 Participants and genotyping

127 BCPP is a prospective cohort that initially recruited 1,544 eligible patients and is described in
128 more detail elsewhere [11]. Clinical data on stage, grade, and size of tumours and
129 demographic information (age, gender) were gathered with bespoke case report forms.

130 Tumour size of the largest tumour was established visually while performing cystoscopy.

131 Blood samples of 888 participants with confirmed UBC were genotyped on the Illumina
132 Infinium OmniExpress-24 BeadChip array at deCODE Genetics (Iceland).

133 Tumours of stages pTa, pT1, or pTis were included to limit our analyses to NMIBC, resulting
134 in a dataset of 712 cases.

135 Quality control (QC)

136 QC procedures were carried out using PLINK v1.90 [12]. The exact thresholds applied and
137 number of exclusions per step are outlined in **Figure 1**.

138 Generic QC procedures per individual excluded those with an inconclusive gender call,
139 excessive genotype missingness rate, increased or reduced genotype heterozygosity rate,
140 duplicate samples, and related individuals.

141 To avoid any bias introduced by population stratification, a principal component analysis
142 (PCA) was carried out. Investigation of PCA plots resulted in exclusion of clear population
143 outliers. Genomic inflation factor (λ) value was estimated for all outcomes of interest; none
144 of the values exceeded 1.03.

145 Marker-specific QC procedures covered excluding SNPs deviating from the Hardy-Weinberg
146 equilibrium, exceeding acceptable missing rate, and rare variants.

147 In total, a dataset consisting of 653 individuals and 597,764 markers remained for further
148 analyses.

149 Imputation

150 Imputation utilised a two-step approach: haplotype phasing by Eagle v2.3.2 [13], followed by
151 genotype imputation with IMPUTE2 [14], using 1000 Genomes Phase 3 [15] as a reference
152 panel in the genome build 19 (GRCh37/hg19). Once imputed, the dataset was filtered for
153 SNPs with info values (an imputation accuracy measure) of >0.3 and MAFs of $>1\%$,
154 resulting in a dataset containing 11,914,228 markers available for genetic association
155 analyses.

156 **Statistical analysis**

157 Statistical analyses were performed using SNPtest v2.5.2 [8] and R statistical package
158 (v3.3.2) [16].

159 To establish the relation between germline variation and tested outcomes, linear regression
160 was used for continuous variables and logistic regression for all binary endpoints. Age was
161 tested as a continuous (years) and binary variable (mean was considered as a cut-off value for
162 categorisation (resulting in strata of $</\geq 70$ years)). Tumour size (cm) was tested as a
163 continuous and categorical variable ($</\geq 3$ cm [17]). Stage (Tis and T1 versus Ta) and grade
164 (G3 versus G2 and G1) were treated as binary variables. In addition, low-, intermediate-, and
165 high-risk EORTC categories were assigned to each NMIBC case and were tested as a
166 dichotomous variable of high- versus low- and intermediate-risk groups [17].

167 All analyses were adjusted for participant gender and first five genetic principal components
168 to increase estimate precision and to adjust for any potential residual population stratification
169 bias. An association was held significant if p-value $<5E-08$, and promising if below $5E-06$.

170 Post-GWAS power calculations were carried out in web-based GAS Power Calculator [18].

171 Manhattan and QQ graphs were plotted for each tested outcome. For significant hits, regional
172 association plots were constructed using LocusZOOM tool [19], except for hits that have not
173 yet been assigned an ID (rsID).

174 **Functional annotation**

175 Identified significant SNPs were mapped using a web-based SNPnexus tool [20], with
176 Ensembl [21] (Version 74) as a functional annotation system.

177 **Replication**

178 Genome-wide significant hits were attempted to replicate in a sample of 1470 NMIBC cases
179 from the NBCS [22] (**Figure 1**). Briefly, the NBCS recruited UBC patients via the
180 population-based cancer registry in the Nijmegen region. Eligible cases were diagnosed
181 during 1995-2006 and were under the age of 75; additional data was collected via linkage
182 with hospital-patient records [22], including tumour size, which was reported after visual
183 evaluation during cystoscopy. Details of genotype data cleaning and initial analysis is
184 provided in detail elsewhere [22].

185 We used META [23] software to perform meta-analysis on association results of both
186 cohorts and calculated a combined p-value per SNP. An inverse-variance method was used,
187 assuming a random-effects model. I^2 index and p-value were calculated to evaluate potential
188 heterogeneity between the estimates of the two cohorts [23].

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200 RESULTS

201 Baseline clinical characteristics of the discovery and replication cohorts are shown in **Table**
202 **1**.

203 Majority of cases in BCPP were male (78.1%), with an average age of 70 years. Tumour size
204 mean was 2.5 cm, and most of the participants were diagnosed with stage Ta (68%) and T1
205 (30.5%) tumours. More than a third of cases presented as G2 (37.5%), followed by G3
206 (31.7%) and G1 (29.2%) NMIBC. The distribution of variable categories and measures were
207 similar between the BCPP and NBCS cohorts.

208 In the discovery-stage analysis, a total of 61 SNPs, corresponding to 29 different regions,
209 showed genome-wide statistically significant associations with at least one of the outcomes.
210 Out of those, 20 loci were mapped to genes (all intronic regions) (**Table 2**). Significant
211 associations were observed for size and age as continuous variables, as well as for binary
212 outcomes of stage, grade, and age.

213 Most of the SNPs (N=47) were found to be associated with tumour size, the effect sizes
214 ranging from 0.65 (rs35225990 in *FAM194B*, $p=2.85E-08$) to 2.6 (rs370572716 in 9p13.1,
215 $p=4.04E-09$) centimetres (**Table 2**).

216 One SNP in 9q22.32, rs142492877, showed statistically significant association with
217 decreased age at diagnosis of almost one year ($\beta=-0.95$, $SE=0.16$, $p=1.05E-08$). Age as a
218 binary trait showed associations in the same direction, although in a different genomic region
219 (7q31.33) with an odds ratio (OR) ranging between 2.46 (rs17149580, $p=2.18E-08$) and 2.51
220 (rs17149636, $p=1.62E-08$) across eight SNPs.

221 The 14q11.2 locus showed strong associations with being diagnosed with a higher grade of
222 NMIBC (rs15091489 in the *TRAV16* gene (OR=3.42, 95%CI: 2.11-5.55, $p=5.13E-09$) and
223 rs116923391 (OR=3.86, 95% CI: 2.38-6.26, $p=2.07E-10$)).

224 Several protective variants for tumour stage were observed, namely: rs117248430 in *ANKS6*
225 (OR=0.003, 95%CI=1.71E-09-3895.6, p=3.73E-08), and two markers in the *SLCO1B1* gene
226 (rs76497895 (OR=0.03, 95%CI=0.001-0.83, p=4.18E-08); rs116946525 (OR=0.03,
227 95%CI=0.001-0.83, p=4.23E-08)). The strength of the effect and corresponding confidence
228 intervals in *ANKS6* might be explained by a very low MAF (<0.01%) among cases.

229 A Manhattan plot for tumour size as a continuous outcome (**Figure 2**) also shows there are
230 several polymorphisms in linkage disequilibrium (LD) with the leading SNP (Manhattan
231 plots for all other tested outcomes are available in the **Supplementary Figures 1-6**).

232 Regional association plot of 13q13.3 (**Figure 3**) in the BCPP confirms high LD with
233 surrounding variants, all mapping to the *NBEA* gene (although they did not reach the
234 statistical significance). Regional association plots for the remaining SNPs identified in the
235 discovery stage are presented in **Supplementary Figures 7-33**.

236 In the replication stage, 50 out of 61 SNPs were available to test in NBCS (**Table 2**). None
237 of these SNPs were significantly associated with the same outcomes in NBCS. A meta-
238 analysis of both cohorts showed variant rs180940944 in 13q13.3 locus to be associated with
239 increased tumour size at diagnosis ($\beta=0.96$, SE=0.16, p=2.92E-09), although the effect is
240 likely driven by BCPP data. Nevertheless, low I^2 estimate ($I^2=0\%$, p(heterogeneity)=0.75)
241 indicated there was no significant heterogeneity between the two cohorts for the replicated
242 SNP. A conditional association analysis on rs180940944 showed the associations in the
243 *NBEA* gene are likely to be driven by the top SNP, as none of the variants have reached
244 genome-wide significance when controlled for the effect of rs180940944 (**Supplementary**
245 **Figure 34**). Nevertheless, the analysis also suggests there is a region in the *NBEA* gene of
246 mildly inflated p-values, independent of the rs180940944.

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249 **DISCUSSION**

250 We have investigated genetic associations with NMIBC tumour (size, stage, grade) and
251 patient (age, EORTC risk category) characteristics at the time of diagnosis within the BCPP
252 cohort.

253 Multiple loci were identified in the discovery stage that are novel in the context of NMIBC.
254 One SNP, rs180940944, has reached statistical significance in a meta-analysis of two
255 NMIBC cohorts, mapping to the intronic region of the *NBEA* gene on 13q13.3. However,
256 associations of other SNPs in the *NBEA* have failed to be reproduced.

257 *NBEA* proteins have been mostly observed to play a significant role in synapse development
258 and function [24]. *NBEA* dysregulation does not affect the establishment of synapses *per se*,
259 but rather their intra-cellular organisation [24]. An in-depth analysis revealed impaired
260 synaptic ability was mostly due to the inappropriate distribution of actin, a protein essential
261 for synapse cytoskeleton structure [24]. The effect is most likely present due to alterations in
262 the Golgi-dependent processes of inter- and intra-cellular compound trafficking, including
263 actin and neural receptors [24].

264 The synaptic alterations are likely to be the contributing cause of autism spectrum disorders
265 [24]; however, the Golgi-related pathway may have a wider phenotypic manifestation [25],
266 including cancer. The prognostic utility of *NBEA* has been investigated in gastric cancer [26]
267 and oropharyngeal squamous cell carcinomas (OPSCC) [27], with promising results.
268 Collectively, these observations implicate the pleiotropic nature of *NBEA* effect across a
269 variety of traits.

270 In our study, we suggest there is an association between *NBEA* and increased NMIBC tumour
271 size. The role of Golgi complex in cancer progression has been reported independently, and
272 disruptions in normal protein transportation can contribute to increased tumour size and,
273 eventually, progression [25].

274 Our findings should be interpreted cautiously. Substantial sample sizes of specific
275 phenotypes such as ours are rare, and suffer from limited power to capture true genetic
276 associations, and spurious associations due to random effects cannot be ruled out. Our post-
277 hoc power calculations [18], underscore the importance of current analysis being ran on
278 bigger cohorts (e.g. association rs150914897 (14q11.2) of an OR=3.42 had power of 79%,
279 but it drops to only 16% for an OR=2.5, hence we may have missed existing associations of
280 more modest effect size).

281 Furthermore, tumour size measurements are subject to variability, degree of which is difficult
282 to establish. The lack of any genome-wide significant associations for categorised tumour
283 size (\leq / \geq 3 cm [17]) adds substantial caution in consideration of our main findings and study
284 power. However, clinically-relevant tumour size categories may not be adequate in a genetic
285 context, and different categorisation may be used in future analyses.

286 Our study only focused on NMIBC instead of a merged group of UBC, and we are unable to
287 comment on whether these genetic loci are relevant for advanced UBC. Given considered
288 limitations, we see this study as true to the GWAS design of hypothesis-generating nature,
289 instead of one offering conclusive findings. Hence, further replication is of essence to
290 establish validity of described results.

291 The 13q13.3 locus has not been observed in prior studies on NMIBC. It might be due to us
292 using an independent prognostic marker of NMIBC (i.e. tumour size) instead of recurrence
293 and/or progression as an outcome. Larger tumour indicates a worse disease course [17], but
294 there are other components that contribute to NMIBC prognosis. In a clinical setting, each
295 tumour characteristic (e.g. size) carries a different weighting [17], collectively contributing to
296 an endpoint (e.g. recurrence).

297 Importantly, powerful studies on UBC risk have already shown some signals to only be
298 associated with MIBC (UBC of T2-T4) [10]. Furthermore, a genome-wide methylation

299 investigation on high-grade NMIBC cases revealed epigenetic changes different from their
300 low-grade counterparts [9]. Direct comparability of these reports is limited, but we see the
301 unravelling genetic complexity within UBC being a connecting thread between all studies.
302 We therefore believe it is likely separate genetic relationships are present for NMIBC
303 determinants, rather than overall prognostic outcomes.

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324 CONCLUSIONS

325 Our study suggests variations in 13q13.3 locus may contribute to an increased NMIBC
326 tumour size in a European population. Further studies are warranted to confirm the
327 association.

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479