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Sor11 as an Alzheimer's Disease Predisposition Gene?

Jennifer A. Webster^a Amanda J. Myers^b John V. Pearson^a David W. Craig^a
Diane Hu-Lince^a Keith D. Coon^a Victoria L. Zismann^a Thomas Beach^{c, d}
Doris Leung^e Leslie Bryden^e Rebecca F. Halperin^a Lauren Marlowe^e Mona Kaleem^e
Matthew J. Huentelman^a Keta Joshipura^a Douglas Walker^{c, d} Christopher B. Heward^f
Rivka Ravidⁱ Joseph Rogers^{c, d} Andreas Papassotiropoulos^{a, j} John Hardy^b
Eric M. Reiman^{a, d, g, h} Dietrich A. Stephan^{a, d}

^aNeurogenomics Division, Translational Genomics Research Institute, Phoenix, Ariz.; ^bDepartment of Psychiatry and Behavioral Sciences, University of Miami, Miller School of Medicine, Miami, Fla.; ^cSun Health Research Institute, Sun City, Ariz.; ^dArizona Alzheimer's Consortium, Phoenix, Ariz.; ^eLaboratory of Neurogenetics, National Institute on Aging, National Institutes of Health, Bethesda, Md.; ^fKronos Science Laboratories, ^gBanner Alzheimer's Institute, Phoenix, Ariz., and ^hDepartment of Psychiatry, University of Arizona, Tucson, Ariz., USA; ⁱRoyal Dutch Academy of Sciences, Amsterdam, The Netherlands; ^jDivision of Psychiatry Research, University of Zurich, Zurich, Switzerland

Key Words

Sor11 · Alzheimer's disease · Predisposition gene · *APOE* gene · Sortilin-related receptor

large case-control whole-genome scan at over 500,000 polymorphisms which presents weak evidence for association and potentially narrows the association interval.

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Abstract

Alzheimer's disease (AD) is a neurodegenerative disorder characterized by progressively disabling impairments in memory, cognition, and non-cognitive behavioural symptoms. Sporadic AD is multifactorial and genetically complex. While several monogenic mutations cause early-onset AD and gene alleles have been suggested as AD susceptibility factors, the only extensively validated susceptibility gene for late-onset AD is the apolipoprotein E (*APOE*) ϵ 4 allele. Alleles of the *APOE* gene do not account for all of the genetic load calculated to be responsible for AD predisposition. Recently, polymorphisms across the neuronal sortilin-related receptor (*SORL1*) gene were shown to be significantly associated with AD in several cohorts. Here we present the results of our

Alzheimer's disease (AD) is a neurodegenerative disorder characterized by progressively disabling impairments in memory and other cognitive domains, but also by non-cognitive behavioural symptoms. AD preferentially affects individuals over 60 years with prevalence rates as high as 40% in nonagenarians [1]. Sporadic AD is multifactorial and genetically complex. Twin studies suggest that genetic factors may account for as much as 80% of the disease risk [2]. While several monogenic mutations cause early-onset AD and gene alleles have been suggested as AD susceptibility factors, the only extensively validated susceptibility gene for AD is the apolipoprotein E (*APOE*) ϵ 4 allele [3, 4]. Alleles of the *APOE* gene do not account for all of the genetic load calculated to be responsible for AD predisposition. Recently, Rogava et al. [5] reported polymorphisms across the neuronal sor-

J.A.W. and A.J.M. contributed equally.

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Fax +41 61 306 12 34
E-Mail karger@karger.ch
www.karger.com

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Accessible online at:
www.karger.com/nddDietrich A. Stephan
Neurogenomics Division, TGen, The Translational Genomics Research Institute
445 N. Fifth Street
Phoenix, AZ 85004 (USA)
Tel. +1 602 343 8727, Fax +1 602 343 8448, E-Mail dstephan@tgen.org

tilin-related receptor (*SORL1*) gene to be significantly associated with AD in several cohorts.

A genome-wide association study was performed using 502,627 SNPs on DNA samples extracted from brain tissue of donors who were at least 65 years of age at the time of their death. The donors included 664 patients who satisfied clinical and neuropathological criteria for the diagnosis of AD and 422 persons that did not meet clinical or neuropathological criteria for AD as controls. Their brain tissue and neuropathological diagnoses were supplied by investigators from 20 National Institute on Aging Alzheimer's Disease Centers (ADCs) in accordance with an agreement with these Centers, the National Institute on Aging, and the National Alzheimer's Coordinating Center (NACC). Additional post-mortem samples were received from Sun Health Research Institute and The Netherlands Brain Bank. In addition to a neuropathologist's diagnostic decision, 18% of the patients and 16% of controls were assessed using CERAD criteria, which provide information about neuritic plaque density [6]. 59% of the patients and 75% of the controls were also assessed using Braak and Braak staging [7], which assesses the distribution of neurofibrillary tangles [6]. The patients included 362 females and 302 males, and averaged 82 ± 7.7 years of age at the time of death. The controls included 255 females and 167 males, and averaged 79 ± 11.0 years of age at the time of death. The software program STRUCTURE was used to test for underlying genetic stratification, using 5,000 randomly selected SNPs with at least 100 per chromosome. Initial analysis yielded empirical evidence of three populations. One, which contained 14 samples and was far removed from the rest of the study population, was removed from further analyses. STRUCTURE was re-run with $K = 2$, and each sample was assigned an admixture of the resulting populations. Comparison of the resulting case and control populations defined by these admixture vectors yielded a silhouette score of 0.06. This indicates that while the samples in this study are likely to be the result of admixture of two populations, the distribution of those populations is equivalent in cases and controls [8]. As previously reported [9], the most significant SNP identified through Fisher's exact tests through the scan, rs4420638 (uncorrected p value = 1.06×10^{-39} , Bonferonni corrected p value for multiple hypothesis testing = 5.30×10^{-34} based on 500,000 comparisons), which was located on chromosome 19, 14 kb from the *APOE* $\epsilon 4$ variant, speaking to the technical robustness of the whole-genome association strategy at this SNP density.

The hypothesis that a specific allele of *SORL1* is associated with increased AD risk was tested. The 46 SNPs located in the *SORL1* genomic locus on the Affymetrix 500K Mapping Array were directly targeted for significance analysis in our whole-genome data set, and thus no multiple-testing correction was applied to the significance values. Calculation of a SNP's allelic frequency differences between cases and controls was based on the Fisher's exact test using each contiguous SNP within the gene sequentially. Table 1 identifies four SNPs with $p < 0.05$ between chromosome 11 SNPs rs668387 (120,873,131 bp) and rs1790208 (120,932,469 bp). This region overlaps the Rogaeva et al. significant region from rs668387 (120,873,131 bp) to rs12285364 (120,898,436 bp). The risk-associated alleles are presented in table 1, and articulate a 59-kb interval, which encompasses 12 exons (exons 7 through 18) of the *SORL1* gene. Of the 9 significant SNPs in Rogaeva et al., 3 were directly measured in our study and each of the remaining 6 are in LD with at least 1 SNP on the Affymetrix Array (fig. 1) [5]. Thus we have captured all of the significant SNPs from the Rogaeva et al. study, including those we have not directly measured. Potentially problematic is that SNPs rs7131432 and rs2101756 have very rare minor allele frequencies (1.6 and 0.5% respectively) and thus the p values are very sensitive to subtle distribution differences, residual genotype error, and subtle stratification between cases and controls and should be interpreted with caution. We do not see significant association with the SNPs that were reported in Rogaeva et al. that were present on the Affymetrix 500k SNP array (in italics in table 1, significant SNPs from Rogaeva et al. underlined), nor do we see significance at the 3' region of the locus noted in the original report [5]. The 5' region of the locus shows nominally significant replication with low odds ratio (OR) of 1.551 at rs11218313 (maximum value of the four associated SNPs with MAF $> 2\%$, and in contrast to an OR of 4 for *APOE* $\epsilon 4$ heterozygotes, and 15–20 for homozygotes). Of note, we would not see these effects if we corrected for multiple testing by Bonferoni. Replication of only one region of the *SORL1* locus is consistent with the findings of Rogaeva et al. Significance of this locus in our population using a different series of SNPs than the original report potentially gives insight into *SORL1*-mediated biology, and refining the association signal to 59 kb could help to narrow the search for disease causing alleles. Whole-genome association studies using well-phenotyped and well-powered study designs may provide disparate results for subtle allelic association findings, such as *SORL1*. The importance of the *SORL1* finding is also that it articulates a common prob-

Table 1. Fisher's exact p values across the *SORL1* locus from a whole-genome association scan using 1,086 AD cases and controls

dbSNP RS ID	Position	p value	Risk allele	MAF	OR (95% CI)
<i>rs4935774</i>	<i>120826964</i>	<i>0.8094</i>		<i>0.209</i>	
<i>rs17125349</i>	<i>120829417</i>	<i>0.2355</i>		<i>0.077</i>	
<i>rs610188</i>	<i>120834606</i>	<i>0.415</i>		<i>0.038</i>	
<i>rs1784934</i>	<i>120843203</i>	<i>0.5133</i>		<i>0.331</i>	
<i>rs676160</i>	<i>120845443</i>	<i>0.3994</i>		<i>0.094</i>	
<i>rs676759</i>	<i>120864475</i>	<i>0.6479</i>		<i>0.383</i>	
<i>rs560573</i>	<i>120866094</i>	<i>0.119</i>		<i>0.38</i>	
<i>rs2298525</i>	<i>120866225</i>	<i>0.8829</i>		<i>0.098</i>	
<i>rs985421</i>	<i>120867526</i>	<i>0.343</i>		<i>0.012</i>	
<i>rs9665981</i>	<i>120869213</i>	<i>0.6092</i>		<i>0.382</i>	
<i>rs12364988</i>	<i>120872836</i>	<i>0.3524</i>		<i>0.486</i>	
<u><i>rs668387</i></u>	<u><i>120873131</i></u>	<u><i>0.5226</i></u>		<u><i>0.422</i></u>	
rs2101756	120874460	0.01923	C	0.005	1.19 (0.866–1.74)
rs11218313	120888081	0.02026	C	0.118	1.55 (0.756–2.12)
<i>rs17125423</i>	<i>120909761</i>	<i>0.3375</i>		<i>0.003</i>	
rs626885	120914166	0.03616	T	0.479	1.19 (0.968–1.42)
<i>rs2276346</i>	<i>120919686</i>	<i>0.09332</i>		<i>0.355</i>	
<i>rs10502262</i>	<i>120920522</i>	<i>0.106</i>		<i>0.283</i>	
rs7131432	120932080	0.03869	A	0.016	1.73 (0.998–3.67)
<i>rs1790208</i>	<i>120932469</i>	<i>0.3479</i>		<i>0.001</i>	
<i>rs11218340</i>	<i>120936564</i>	<i>0.6076</i>		<i>0.143</i>	
<i>rs10892756</i>	<i>120939766</i>	<i>0.7295</i>		<i>0.075</i>	
<i>rs11218343</i>	<i>120940797</i>	<i>0.6357</i>		<i>0.082</i>	
<i>rs11218346</i>	<i>120944391</i>	<i>0.4664</i>		<i>0.081</i>	
<i>rs1792124</i>	<i>120946730</i>	<i>0.6149</i>		<i>0.017</i>	
<i>rs1790213</i>	<i>120947399</i>	<i>0.2949</i>		<i>0.385</i>	
<i>rs11218347</i>	<i>120951758</i>	<i>0.7815</i>		<i>0.077</i>	
<i>rs1699103</i>	<i>120957136</i>	<i>0.8596</i>		<i>0.42</i>	
<i>rs7116734</i>	<i>120957150</i>	<i>0.1016</i>		<i>0.408</i>	
<i>rs1792127</i>	<i>120957244</i>	<i>0.1566</i>		<i>0.001</i>	
<i>rs11218350</i>	<i>120957861</i>	<i>0.7873</i>		<i>0.227</i>	
<i>rs10892759</i>	<i>120969298</i>	<i>0.1122</i>		<i>0.341</i>	
<i>rs7127359</i>	<i>120970156</i>	<i>0.6789</i>		<i>0.358</i>	
<i>rs11218360</i>	<i>120978601</i>	<i>0.9359</i>		<i>0.031</i>	
<i>rs7128608</i>	<i>120978808</i>	<i>0.946</i>		<i>0.034</i>	
<i>rs1629493</i>	<i>120982306</i>	<i>0.2664</i>		<i>0.381</i>	
<i>rs2282648</i>	<i>120983705</i>	<i>0.2016</i>		<i>0.339</i>	
<u><i>rs2282649</i></u>	<u><i>120984168</i></u>	<u><i>0.97</i></u>		<u><i>0.285</i></u>	
<i>rs726601</i>	<i>120986617</i>	<i>0.19</i>		<i>0.308</i>	
<i>rs1784931</i>	<i>120988148</i>	<i>0.1377</i>		<i>0.397</i>	
<u><i>rs1010159</i></u>	<u><i>120988611</i></u>	<u><i>0.2545</i></u>		<u><i>0.355</i></u>	
<i>rs1503413</i>	<i>120992024</i>	<i>0.5687</i>		<i>0.057</i>	
<i>rs1614735</i>	<i>120998211</i>	<i>0.4347</i>		<i>0.471</i>	
<i>rs17125558</i>	<i>120999237</i>	<i>0.7932</i>		<i>0.025</i>	
<i>rs17125561</i>	<i>120999299</i>	<i>0.4623</i>		<i>0.004</i>	
<i>rs10892761</i>	<i>121003094</i>	<i>0.4663</i>		<i>0.412</i>	

SNPs highlighted in bold are significant in our population, SNPs italicized are those reported in the Rogava et al. study populations, and those that are underlined are reported as significant by Rogava et al.

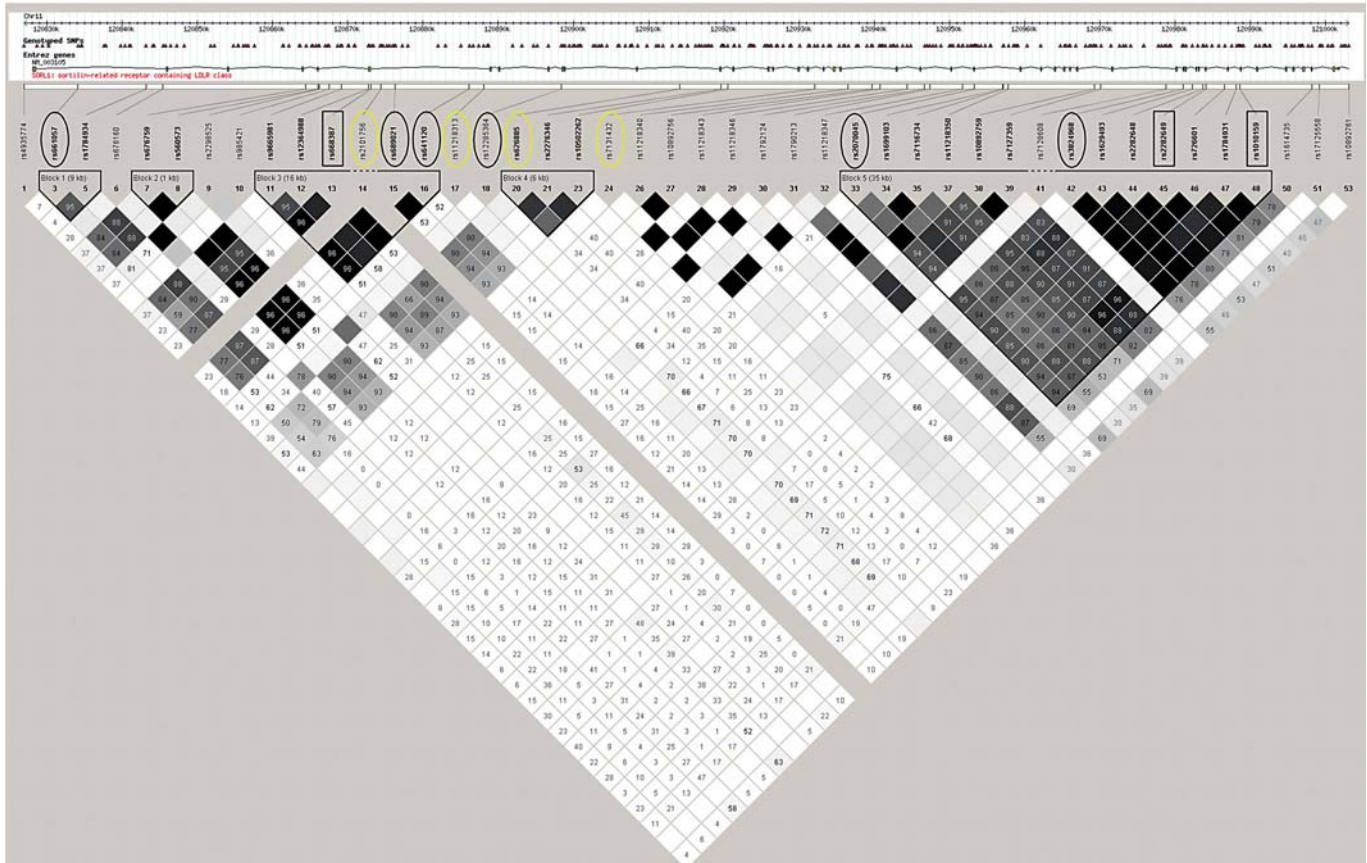


Fig. 1. Out of the nine markers in *SORL1* found to be significantly associated with disease risk in Rogaeva et al. [5], we have typed three of those markers (black rectangles). To examine whether we were able to capture the remaining 6 significant SNPs typed from the Rogaeva screen (black circles), we mapped the LD relationships between our SNPs and the Rogaeva SNPs using the HapMap database and the program Haploview. For each of those significant SNPs, there is at least one marker within our screen which is in complete or almost complete linkage disequilibrium (as measured by r^2) with the Rogaeva et al. markers. Thus, we have captured if not typed all of the significant Rogaeva markers. Yellow circles indicate significant SNPs from our screen demonstrating that we have potentially replicated one block of the association (rs668387 to rs7131432 from our data set) but not the other (rs2070045 to rs1010159).

lem with low OR associations in that power, population structure and history, allelic heterogeneity and SNP coverage across the locus may confound the ability to tease out subtle allelic associations in complex traits. If *SORL1* does demonstrate allelic heterogeneity, the presence of the admixture of two populations that STRUCTURE revealed for our study population could dilute any association signals that were present in the two original populations. Clearly our results can be interpreted either as a failure to replicate or as a partial replication depending on one's point of view and this illustrates a current major problem in association studies. Even analyses in large,

well-characterized studies such as ours fail to give signals clear enough to distinguish between no signal and variable signals in different cohorts.

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