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### **ABSTRACT**

We explore the question of how political institutions and particularly democracy affect economic growth. Although empirical evidence of a positive effect of democracy on economic performance in the aggregate is weak, we provide evidence that democracy influences productivity growth in different sectors differently and that this differential effect may be one of the reasons of the ambiguity of the aggregate results. We provide evidence that political rights are conducive to growth in more advanced sectors of an economy, while they do not matter or have a negative effect on growth in sectors far away from the technological frontier. One channel of explanation goes through the beneficial effects of democracy and political rights on the freedom of entry in markets. Overall, democracies tend to have much lower entry barriers than autocracies, because political accountability reduces the protection of vested interests, and entry in turn is known to be generally more growth-enhancing in sectors that are closer to the technological frontier. We present empirical evidence that supports this entry explanation.

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## 1 Introduction

There are two main open questions in the debate on the relationship between economic development and democratic institutions. One is whether or not countries become democratic only at high levels of per capita income, and the second is whether or not (or when) democracy enhances economic development, captured by per capita GDP growth.

On both questions there are disagreements amongst scholars. Most observers would agree that richer countries are democracies and therefore economic development favors transitions towards political freedom, as argued by Lipset (1959) and Barro (1999). However, Acemoglu, Johnson, Robinson, and Yared (2005) suggest that this cross-sectional evidence is not robust to controlling for factors affecting simultaneously income and political institutions. When they add country fixed effects in repeated cross-country regressions they find no effect of the level of income on democracy. Admittedly this procedure looses all the information arising from cross country comparisons, which seems crucial in this context. Huntington (1991) argues that the progress towards democratization is not linear and there are back and forth waves of democratization.

On the second question the evidence is mixed at best<sup>1</sup>. Perhaps democratic institutions increase redistributive pressures<sup>2</sup> that may be harmful to growth, especially for middle income countries, and this is one of the reasons why Barro (1996) argues that countries cannot become democracies too "soon" in terms of their level of income per capita. However, Mulligan, Sala-i-Martin and Gil (2003) do not find much of a difference in public policies between democracy and non democracies, questioning therefore any link between form of government and economic performance<sup>3</sup>. Glaeser et al. (2004) also question the causal effects of political institutions on economic growth.

Very few studies address both questions jointly. Two exceptions are Przewroski et al. (2000) and de Mesquita et al. (2003). Persson and Tabellini (2006) use the transition from democracies to dictatorship and the concept of the accumulated stock of democratic capital, concluding that the latter is conducive to growth and consolidation of democratic institutions at the same time<sup>4</sup>. Giavazzi and Tabellini (2005) analyze

<sup>&</sup>lt;sup>1</sup>See for instance Helliwell (1994), Barro (1996) and Papaioannou and Siourounis (2004) and the refereces cited therein.

<sup>&</sup>lt;sup>2</sup> For some results consistent with this observation see Aghion Alesina and Trebbi (2004) and Boix (2003).

<sup>&</sup>lt;sup>3</sup>They consider education spending, spending on pensions and welfare, trade openness amongst other variables.

<sup>&</sup>lt;sup>4</sup>On a related topic see also Rodrik and Wacziarg (2005).

the timing of economic and political liberalizations.

In this paper we will focus mostly on the second question of whether or when democracy enhances economic growth. Most of the existing papers on the subject employ aggregate data on income levels and economic growth. A first point of departure of this paper is its use of disaggregated data to shed new light on the debate. A second point of departure is the idea that political institutions, and democracy in particular, may have different effects on different sectors of the economy, possibly depending on the specific characteristics of the technology and the industry's market. Our results can be summarized in three points. First, we find that democratic institutions and political rights enhance growth of more advanced sectors, namely in sectors close to the technological frontier in the sense of Acemoglu, Aghion, and Zilibotti (2006). Second we uncover that an important channel of this effect is freedom of entry in markets. Political rights are associated with freedom of entry and the latter is especially important for sectors close to the technological frontier. In fact entry of new firms and competition are especially beneficial in spurring innovation at high levels of technological development, whereas it may discourage innovation in more backward sectors as argued by Aghion et al. (2006). Third, and this is an implication of the first two results, more advanced economies benefit more from democratic institutions and therefore the demand for democracy should increase with the level of per capita income in a country, calling into question, indirectly at least, the result by Acemoglu et al. (2005) suggesting that development does not bring about democratic institutions.

That democracy may entails both, positive and negative effects, on growth, has already been stressed for example by Acemoglu (2007). There, higher democracy tends to be good for growth because it reduces the extent to which existing oligarchies can prevent entry by potential competitors. On the other hand, democracy leads to higher tax rates in equilibrium, which in turn tends to discourage innovation everything else remaining equal. However, unlike in our analysis below, in Acemoglu (2007) the comparison between the costs and benefits of democracy does not interact with the economy's proximity to the world technological frontier.

The paper is organized as follows. Section 2 presents a model of democracy and entry which blends the entry model of Aghion, Burgess, Redding, and Zilibotti (2006), henceforth ABRZ, with the model of vested interests in Acemoglu, Aghion, and Zilibotti (2006), henceforth AAZ. Section 3 discusses the empirical

evidence. The last section concludes.

## 2 The model

The point of the model is to show that freedom of entry is especially growth enhancing in sectors closer to the technological frontier. Freedom of entry is correlated with political rights, as we assume in the model and we show below in the empirical part. Therefore the model derives the relationship between sectorial growth and democratic institutions which we then test.

#### 2.1 Production and profits

Time is discrete and all agents live for one period. One final good is produced competitively using a continuum of intermediate inputs, indexed from 0 to 1, according to the following technology:

$$y_t = \frac{1}{\alpha} \int_0^1 A_{it}^{1-\alpha} x_{it}^{\alpha} di,$$

where  $x_{it}$  is the quantity of intermediate input produced in sector i at date t,  $A_{it}$  is a productivity parameter that measures the quality of the intermediate input i in producing the final good, and the parameter  $\alpha \in (0,1)$ . The final good can be used either for consumption, or as an input in the process of production of intermediate goods, or for investments in innovation. We normalize the price of the final good to one.

In each intermediate sector i only one firm (a monopolist) is active in each period, and produces intermediate input i using final output as input, one for one. Since the final good sector is competitive, the intermediate monopolist i sells its intermediate good to the final good sector at a price equal to its marginal cost. The first order condition for the final sector producers reads as follows:

$$p_{it} = (A_{it}/x_{it})^{1-\alpha}. (1)$$

Then profit maximization for intermediate-good producers yields the choice:

$$x_{it} = A_{it}(\frac{1}{\alpha})^{-\frac{1}{1-\alpha}},$$

and the equilibrium profit:

$$\pi_{it} = \delta A_{it},\tag{2}$$

where  $\delta \equiv (\frac{1}{\alpha} - 1)(\frac{1}{\alpha})^{-\frac{1}{1-\alpha}}$ .

Substituting for  $x_{it}$  in the production function for final output, we also get:

$$y_t = \zeta A_t, \tag{3}$$

where

$$A_t = \int_0^1 A_{it} di$$

is the average productivity at date t and  $\zeta = (\frac{1}{\alpha})^{1-\frac{\alpha}{1-\alpha}}$  is a constant.

#### 2.2 Entry and incumbent innovation

Let p denote the probability of a potential entrant appearing in any intermediate sector. In the model the government directly chooses this parameter in a way described below. This is of course a short cut that captures different policies that may favor or hinder competition and entry. There are only two types of intermediate firms: advanced firms, with productivity  $A_{it} = \overline{A}_t$ , and backward firms with productivity  $A_{it} = \frac{1}{\gamma} \overline{A}_t$ , and we focus on technologically advanced entry; accordingly, each potential entrant arrives with the leading-edge technology parameter  $\overline{A}_t$ , which grows by the factor  $\gamma$  with certainty each period. An advanced firm can use a first-mover advantage to block entry and retain his monopoly. But if the firm is backward, then entry will occur, Bertrand competition will ensue, and the technologically dominated incumbent will be eliminated and replaced by the entrant.

A new innovation costs  $c_{it}A_{it-1}$  and allows the innovating incumbent firm to improve productivity by factor  $\gamma$ . We assume  $c_{it}$  to be random and i.i.d. across intermediate sectors with support  $\{0, \overline{c}\}$ , with

$$\Pr(c_{it} = 0) = \Pr(c_{it} = \overline{c}) = \frac{1}{2}.$$

The effect of entry threat p on incumbent innovation will depend on the marginal benefit  $v_{it}$  which the incumbent expects to receive from an innovation. Note that firms know their cost of innovation when they decide whether to innovate (i.e. there's no uncertainty).

Consider first an incumbent who was an advanced firm last period. If he innovates then he will remain on the frontier, and hence will be immune to entry. His profit will then be  $\delta \overline{A}_t = \delta \gamma \overline{A}_{t-1}$ . If he fails to innovate

then with probability p he will be eliminated by entry and earn zero profit, while with probability 1-p he will survive as the incumbent earning a profit of  $\delta \overline{A}_{t-1}$ . Dividing through by  $\overline{A}_{t-1}$ , an advanced firm with innovation cost  $c_{it}$  will innovate whenever

$$\delta(\gamma - (1 - p)) = \delta(\gamma - 1 + p) > c_{it}. \tag{4}$$

In particular an increase in entry threat encourages this incumbent to innovate. Intuitively, a firm close to the frontier responds to increased entry threat by innovating more in order to escape the threat.

Next consider an incumbent who was a backward firm last period, and who will therefore remain behind the frontier even if he manages to innovate, since the frontier will also advance by the factor  $\gamma$ . For this firm, profits will be zero if entry occurs, whether he innovates or not, because he cannot catch up with the frontier. Thus it will innovate whenever:

$$\delta\left(1-p\right)\left(\gamma-1\right) > c_{it}.\tag{5}$$

where the left hand side is the profit gain from innovation that will be realized with probability (1-p), the probability that no potential entrant shows up. Thus in this case innovation incentives depend negatively on the entry threat p. Intuitively, the firm that starts far behind the frontier is discouraged from innovating as much by an increased entry threat because he is unable to prevent the entrant from destroying the value of his innovation.

We assume that initially the entry rate p is equal to zero, and that

$$A1:\delta(\gamma-1)<\overline{c},$$

so that, absent any entry threat, no firm with innovation cost equal to  $\overline{c}$  ever innovates.

Using these assumptions, one can determine the steady-state fraction of advanced firms conditional upon p = 0. Following ABRZ, suppose that an advanced firm that successfully innovates at date t, starts out in period t + 1 as an advanced firm. All other firms start out as backward firms. Moreover, with exogenous probability h, a backward firm at the end of period t is replaced by a new, advanced, firm at date t + 1. If  $a_t$  denotes the fraction of advanced firms at t, then it satisfies the dynamic equation (see ABRZ):

$$a_{t+1} = z_A a_t + h (1 - z_A a_t),$$

where

$$z_A = \Pr(c = 0) = \frac{1}{2}$$

is the probability that an advanced firm innovates if p = 0. Thus the steady-state fraction of advanced firms is:

$$a^* = \frac{h}{1 - \frac{1}{2}(1 - h)} = \frac{2h}{1 + h}.$$

#### 2.3 Politics and the equilibrium probability of entry

Suppose that entry policy, p, is determined each period by a politician that cares about the current consumption, but may also respond to bribes. Following AAZ, we assume that the politician's pay-off is equal to  $H\overline{A}_{t-1}$ , where H > 0, if she chooses the policy that maximizes current output  $y_t$ , and to  $B_t$  otherwise, where  $B_t$  denotes the bribe that the politician may receive from private firms to limit entry. The parameter H reflects the aggregate welfare concerns of politicians, or the effectiveness of check-and-balances that the political system imposes on politicians. It is our proxy for democracy in this model and we discuss this below. To compute the equilibrium bribe incumbent firms are willing to pay to prevent moving from initial entry probability  $p_0 = 0$  to p > 0, we need to compute the equilibrium payoffs for each type of firm (advanced or backward) and for each cost realization  $c_{it} = 0$  or  $\overline{c}$ , as a function of p.

Consider first an advanced firm. If this firm's innovation cost is zero, then the firm will always innovate and its post-innovation profit  $\delta \gamma \bar{A}_t$  is independent of the entry probability. However, if the innovation cost is  $\bar{c}$ , this firm will lose from higher entry threat only if the threat p becomes sufficiently high that condition (4) holds. Then, indeed, the firm will lose the amount

$$\bar{A}_{t-1}[\bar{c} - \delta(\gamma - 1)]$$

which is positive by Assumption (A1). Thus the maximum bribe advanced firms would be ready to pay as a whole to prevent an increase in entry threat from zero to p, is given by:

$$B_a(p) = \bar{A}_{t-1} a^* \frac{1}{2} [\bar{c} - \delta(\gamma - 1)] . 1_{(\delta(\gamma - 1 + p) > \bar{c})},$$

where  $\frac{1}{2}$  is the probability that  $c_{it} = \overline{c}$  and  $1_{(\delta(\gamma-1+p)>\overline{c})}$  is equal to one whenever (4) holds and to zero otherwise.<sup>5</sup>

Now, consider a backward firm. Such a firm will innovate if and only if c = 0 by Assumption (A1), no matter the entry probability. And it will lose from a higher threat of entry, whether it innovates or not, just because this reduces her probability of survival. The maximum bribe backward firms will be willing to pay to prevent an increase in entry threat from zero to p, is equal to:

$$B_b(p) = \bar{A}_{t-1}(1 - a^*)[\frac{1}{2}\delta p + \frac{1}{2}\frac{\delta p}{\gamma}],$$

where the first (resp. second) term in the bracket is the expected loss incurred by backward firms with low (resp. high) innovation cost<sup>6</sup>.

Altogether, incumbent firms will successfully prevent the increase in entry threat from zero to p whenever p is greater than  $p^*$  such that:

$$B(p^*) = B_a(p^*) + B_b(p^*) = \bar{A}_{t-1}H.$$
(6)

Given that B(p) is strictly increasing in p, equation (6) defines the equilibrium entry probability  $p^*$  as an increasing function of H.

$$(\delta \gamma - \overline{c})\overline{A}_{t-1}$$

whereas it is just

$$\delta \overline{A}_{t-1}$$

if the firm does not innovate. The difference between the latter and the former is just equal to

$$(\overline{c} - \delta(\gamma - 1))\overline{A}_{t-1}$$
.

 $^6\mathrm{A}$  backward firm's ex ante expected profit is:

$$(\frac{1}{2}\delta(1-p) + \frac{1}{2}\delta\gamma(1-p))\overline{A}_{t-1}$$

whenever the entry probability is p.

<sup>&</sup>lt;sup>5</sup>Indeed the firm's expected profit if the entry probability is p, is equal to

### 2.4 Democracy, distance to frontier, and growth

The above analysis tells us that more democracy, as measured by a higher H, increases the probability of entry. But this in turn, encourages innovation by advanced firms whereas it discourages innovation by backward firms (from (4) and (5) above). The overall effect on aggregate innovation and growth will then depend upon the fraction of advanced firms  $a^*$  in the economy, which also measures the economy's proximity to the world technology frontier: for  $a^*$  sufficiently close to one, the overall effect of entry and therefore democracy on innovation will be positive, whereas it will be negative if  $a^*$  (i.e. h) is close to zero.

#### 2.5 Main prediction and discussion

Combining (6) with (4) and (5), yields the prediction that an increase in "democracy" as measured by H, will stimulate innovation by advanced firms but not by backward ones. One should thus expect a higher impact of democracy on productivity growth in sectors that are closer to the world technological frontier.

We identify with democratic institutions the parameter H in our model. This parameter captures the relative benefits of investing in a new technology relative to that of bribing a policy maker to make him raise barrier to entry. Bribes are certainly not unknown in democracies, but since they are more open, associated with more freedom of the press and involve alternation in power of different groups, democratic leaders are generally less likely to be permanently captured by incumbents. A different way of interpreting the evidence that we present below is to say that democracy enhance growth in technologically advanced sectors when democracy is associated with policies that allow more competition and more freedom of entry. In the empirical section which follows, we do indeed review evidence that imply a strong correlation between freedom of entry and democratic institutions as measured by indicators of political rights.

# 3 Empirical analysis

#### 3.1 Data

We use industry employment and value added data from the Industrial Statistics database collected by the UNIDO (2005, revision 2).<sup>7</sup> The data include 28 ISIC (level 3) manufacturing sectors at the three digit level for 180 countries for the period 1963 to 2003. UNIDO is arguably the most comprehensive source of sectorial data available in a cross-country-sector panel format<sup>8</sup>. The UNIDO data provide the most extensive coverage of developing countries sectorial output data available, a very important sub-sample for our study as most of them are partially free or not free from a political standpoint. However, a drawback of this source is the extensive use of national statistics, which tend to be especially noisy for industry-level data of developing countries.

For our empirical analysis two are the key independent variables: democracy and the distance to the technological frontier. In measuring the former we use standard, well-established measures of democracy in political economics, such as the aggregate indicators from the Polity IV database (2005) and the Freedom House measures of civil liberties and political rights. The Polity IV project records various regime characteristics for every independent state above half million total population. In particular we make use of the combined polity index<sup>9</sup> ranging from -10 to 10 (-10 = high autocracy; 10 = high democracy). Such index incorporates more specific sub-indexes concerned with constraints on the executive, open political competition, effectiveness of legislature, etc. The Freedom in the World data by Freedom House report annual analysts' assessments of both political dimensions<sup>10</sup> and civil dimensions<sup>11</sup> since 1972 for more than 192

<sup>&</sup>lt;sup>7</sup>Rajan and Zingales (1998) is one of the early seminal papers employing sectorial growth rates analysis.

<sup>&</sup>lt;sup>8</sup>Other well-recognized sources of country-sector panel include OECD and ILO. However UNIDO employs OECD data for the relevant subsample, while ILO produces employment data aggregated to the ISIC level-2, therefore more coarse than the data we employ here.

<sup>&</sup>lt;sup>9</sup>We exclude special polity conditions such as transitions, interruptions, and interregnums as coded in the Polity IV database -66; -77; -88.

<sup>&</sup>lt;sup>10</sup>Participate freely in the political process; Vote freely in legitimate elections; Have representatives that are accountable to them.

<sup>&</sup>lt;sup>11</sup>Exercise freedoms of expression and belief; Be able to freely assemble and associate; Have access to an established and equitable system of rule of law; Have social and economic freedoms, including equal access to economic opportunities and the

countries. For the Political Rights and Civil Liberties indexes, respectively; 1 represents the most free and 7 the least free rating. For consistency we re-scale both indexes in order to have 7 as the most free rating.

As for distance to the technological frontier we consider the logarithm of the value added per worker (VA/EMP) of a sector divided by the maximum of the log of the same variable in the same sector across all countries in each year and take one minus this ratio as a proxy for distance to frontier. Specifically we define:

$$DISTANCE_{ict} = 1 - \frac{(VA_{ict}/EMP_{ict})}{\max_{c'} (VA_{ic't}/EMP_{ic't})}$$

$$i = industry; c = country; t = year.$$

We also substituted log value added per worker with the logarithm of output per worker in order to construct an alternative measure of distance to frontier. We do not report results employing this variable, however, as they tend to be similar.

Finally, we make use of measures of entry barriers and costs of entry estimates as constructed by Djankov et al. (2002) and of real gross domestic product per capita in international dollars as reported in the Penn World Tables, mark 6.2. Summary statistics for the sample are reported, classified by year, in the Appendix Table A1.

### 3.2 Specification

On the left hand side of our regression we have the growth rate of either output (Y), value added (VA) or employment (EMP) in the industrial sector. On the right hand side we have a measure of democracy (and other measures of political/civil rights), a measure of distance to the technological frontier, and the interaction term between the two. This term allows the marginal effect of democracy to vary with the proximity to the world's most productive technology. We also include time and country-industry fixed effects in order to account for contemporaneous shifts in world growth rates and unobserved heterogeneity across countries and industries. Thus, the empirical specification is as follows, where with  $y_{ict}$  we mean either output, value added or employment in industry i of country c at time t:

right to hold private property.

$$\Delta_{s} \log y_{ict} = \beta_{0} + \beta_{1}DISTANCE_{ict} + \beta_{2}POL_{ct}$$

$$+\beta_{3}DISTANCE_{ict} * POL_{ct} + \alpha X_{ict} + u_{ict}$$

$$\Delta_{s} \log y_{ict} = \log y_{i,c,t+s} - \log y_{ict}$$

$$u_{ict} = \delta_{t} + \gamma_{i,c} + \varepsilon_{i,c,t+s}$$

$$s = \{5,10\}$$

$$i = industry; c = country; t = year.$$

$$(7)$$

In the analysis we study 5-year and 10-years output, value added and employment sectorial growth rates. We compute growth rates over non-overlapping periods and in particular 5-years growth rates are computed over the periods 1970-75, 1975-80, 1980-85, 1985-90, 1990-95, and 1995-2000. For the 10-year growth rates we use alternatively the 1975-85, 1985-95 (odd) years and the 1980-90, and 1990-2000 (even years). Given the somewhat measurement-error prone nature of the data, we perceive this sample segmentation as a conservative approach. By employing this segmentation approach we are able to exclude issues of mechanical serial correlation present in a moving average setting. As we show below, indeed the results are sometimes influenced by the specific assumptions on sample periods. In a series of results that we do not report we also reestimate our main specification with yearly growth rates (s = 1). By and large the main implications of our discussion hold for this sampling choice as well. However, the degree of measurement error is substantial in UNIDO data at the yearly level, mostly because idiosyncrasies at the national statistical level are present, and results tend to be sometimes (and expectably) sensitive in this respect. By employing growth rates over longer periods of time signal to noise ratios tend to improve and ameliorate measurement error issues.

All the standard errors are robust and clustered at the country or country-industry level whenever possible, in order to account for general variance-covariance structures at the country-industry level within the panel setup. This approach is useful in order to address arbitrary forms of serial correlation in the error terms by allowing general variance-covariance structures within any industry-country error term.

The set of controls that we systematically include in any industry-country regression,  $X_{ict}$ , include the level of real GDP per capita and its interaction with distance to frontier. The empirical model (7) is

essentially a reduced-form specification and we do not intend to over-emphasize any channel of causation. A two-way causation mechanism is surely at work within a growth-democracy (or income-democracy) aggregate specification. The aim of this paper is to emphasize a significant differential role in the correlation between growth and democracy, this does not require the reader assigning causal interpretation to our finding.

#### 3.3 Basic Results

Table 1 shows, in accordance to Acemoglu et al. (2005), that there is no effect of democracy on growth rates for manufacturing in a fixed-effects regression at the country level (obtained aggregating the UNIDO sectorial data at country-year level). If any, the effect of democracy on growth appears to be negative, since a few of the coefficients in the various regressions are significantly negative (even though most coefficients on democracy are insignificantly different from zero). Specifically, the upper panel (5-years growth rates) presents a split picture between value added and employment growth rates regressions. In the notation for this table (and the following ones) we employ the Lj operator notation for j—years lagged values. The middle panel (10-year growth for odd years) appears mostly positive and the lower panel (10-years even) presents mostly negative coefficients. These results are therefore consistent with what is found in the literature: no robust effects of democracy on aggregate growth.

Table 2 reports results on the effect of democracy on growth at the country-industry disaggregated level. The unit of observation becomes now a particular industry in a country at a specific moment in time. Still no effect of democracy is statistically or quantitatively significant in the data. Once again some of the coefficients of democracy (or political/civil rights) are negative and statistically different from zero. This table then shows that adding an extra dimension (industry level data) does not produce a departure from standard fixed effects country-level results per se. Notice that Table 2 introduces a control for the level of real gross domestic product per capita. We include a control for income levels (and its interactions when necessary) in order to partial out the effects of political institutions and of the level of economic development.

Table 3 introduces in specification at the country-industry level (7) an interaction term between democracy and distance from the technological frontier. We report results concerning output, value added and employment for all sample periods (5-year; the even and odd 10-year). This interaction matters statistically

and quantitatively in most specifications and enters with a negative and significant coefficient. The table shows that introducing a differential effect of distance to the technological frontier and democracy explains the data better in terms of fit. The sign of the interaction between distance to frontier and democracy is usually negative, while the level of democracy is positive. The result indicates that when close to the technological frontier, the effect of democracy on growth is positive. However, far away from the technological frontier the effect of democracy may be growth-diminishing. This can be interpreted as the basic reduced-form implication of the theoretical model we built. Finally, we interpret the positive sign of the coefficient associated with the regressor DISTANCE as evidence of convergence, a robust and recurrent finding in the empirical economic growth literature. Notice also that a component of convergence is also captured (predictably) by the level of GDP per capita which present a negative sign (growth rates are lower in countries with higher levels of GDP per capita). The interaction term between GDP per capita and distance to frontier is usually negative, although mostly insignificant.

Concerning quantitative implications of Table 3, let us begin from the first panel (5-year growth rates), recalling that the Polity index ranges from -10 to +10, increasing in the level of democratic freedom. The statistically and quantitatively significant results concern the value added and the employment growth regressions (columns 4-9) and the political rights and civil liberty indexes. In general output growth regressions tend to appear more noisy and so do regressions employing the Polity IV index as a proxy for democracy in the right-end side.

Column (5) implies that the effect of an increase of a unit point on the Polity IV scale of democracy for a sector that operates with the most advanced technology (i.e. distance from the frontier is 0) has the effect of increasing the growth rate of value added of 1.2 percent. However at the mean distance to frontier in the sample (0.16) the effect of a similar increase in democracy on the growth rate of value added decreases to -1 percent. Notice that this figure is roughly consistent with the results of Table 2. Considering a technology half the productivity of the world frontier we would find a really substantial effect of an increase of a Polity IV unit of democracy: a drop of -5.7 percent. The unit measure for political rights and civil liberties is different from the Polity IV index, ranging from 1 to 7 (rescaled so that 7 indicates full democracy and 1 full autocracy). Once the coefficients are scaled for the unit of measurement, the results employing Freedom

House data (columns 2 and 3) confirm the column (1)'s estimates.

Results systematically present more beneficial effects of democracy when we use as left-hand-side variable the sector's employment growth. In column (8) the effect of an increase of a unit point on the Polity IV scale of democracy for a sector that operates with the most advanced technology (i.e. distance from the frontier is 0) has the effect of increasing the growth rate of value added of 3.9 percent. At the mean distance to frontier the effect of a similar increase in democracy on the growth rate of employment is -0.7 percent.

Considering the 10-year growth rates for both the even years and the odd years samples we maintain the finding of a negative interaction term and a positive coefficient on the level of democracy in the majority of the specifications, albeit with some exceptions that we tend to attribute to the demanding nature of our three-way fixed-effects specification. The result confirm a positive role of democracy especially when close to the technological frontier. Consider column (1) of the odd-years panel. The estimated effects are consistent with the 5-years growth rates but not statistically significant. Again the results on employment growth rates are stronger.

Overall these results, including many other specification tests which we performed and are available upon request<sup>12</sup>, suggest that our intuition of a differential effects of democracy in different sectors as a function of distance form the frontier is plausible. However, as the last part of Table 3 shows, some degree of sensitivity of the results to certain specification or measurement issues remains evident.

### 3.4 Entry and Democracy

The previous section presents evidence of differential effects of political institutions on sectorial performance, depending on the specific characteristics of each industry. In this section we make some progress in the investigation of a channel, especially focusing on entry barriers.

In the model of Section 2 we discussed the role of freedom of entry and we assumed a correlation between freedom of entry and political freedom. In Appendix Table A2 we report evidence that democracies do 12 Among the additional tests we performed are: change of sampling frequency (yearly panel results; modified dates for the computation of the sample changes); specification changes (including quadratic terms of the democracy index); exclusion of country or sector fixed effects.

indeed have lower barriers and cost of entry by employing the Djankov et al. (2002) cross-country data on regulation of entry. Indeed, a table similar to A2 is also present in the original Djankov et al. (2002) paper and here it is reported only for completeness of the argument. The left-hand-side variable is a measure of the cost of entry given by the number of bureaucratic procedures needed for a firm to enter the market (number of procedures necessary to start up a business), available from the Doing Business project by the World Bank. All measures of democratic development (constraints on the executives, political and civil rights, et similar) present a negative correlation with entry costs, even controlling for income levels. Notice that in addition to the Polity IV autocracy index and Freedom House political and civil rights indexes we also report results employing all the additional measures of democracy and political liberties that Djankov et al. (2002) employ, such as the Polity IV sub-indexes of executive de facto independence; constraints on executive power; effectiveness of legislature; government effectiveness; competition in the legislative nominating process.

Perotti and Volpin (2006) study in detail the cross national determinants of investor protection and entry. They present convincing evidence that "countries with more accountable political institutions have better investor protection and lower entry costs". Note that, as they point out, poor investor protection is indirectly a barrier for entry since new investors face higher risks in entering a new market. Perotti and Volpin (2006) also use POLITY IV as a measure of political accountability but the experiment with several others finding consistently a strong correlation between democracy and freedom of entry.

A direct way of relating our approach to past literature would be to consider the differential effect of democracy and entry using as a proxy for the probability of entry of a competitor the number of procedures needed to start up a business in the country as reported by Djankov et. al. (2002). We could check if, controlling for a differential effect of entry, the role of democracy and the significance of the interaction term change. Given the cross-sectional nature of the cost of entry variables only the interaction term between entry cost and distance can be added to the specification of Table 3.

Although promising in principle, this approach did not yield any statistically relevant result in our attempts. In tables available from the authors we considered this approach, with and without controls for GDP levels. The estimates of the coefficients on democracy and its interaction with distance seemed to be unaffected by introducing this particular measure of entry costs interacted with distance. Considering the

5-year growth rate sample and the 10-year rates for both even and odd dates presents a noisy picture at best: the marginal effects of democracy at different distances become somewhat smaller but the proxy for entry is always insignificant and often with the wrong sign. This is also true if we add further interactions on property rights protection, which we explored as well. The results on democracy and the interaction with distance remain generally significant and with a positive and a negative sign respectively. These results taken together suggest either that the Polity IV and Freedom House variables capture something more general than the number of procedures regarding the possibility of entry and competition or that a purely cross-sectional measure of cost of entry as the one proposed by Djankov et. al. (2002) does not present sufficient variation to allow our three-way panel setup (industry-country-time, which is indeed fairly demanding in this respect) to pick up any significant relationship with growth.

In order to provide a reasonable amount of variation in the data in Table 4 we replace the number of procedures with the effective level of entry in the sector, measured by the 5-year or 10-year growth rates in the number of establishments. Notice that the growth rate in the number of establishment is a realized (ex post) measure of entry and one could argue about whether or not is the relevant covariate to influence the pattern of innovation discussed in the model, as opposed to ex ante proxies. Clearly a point in favor of our approach is the scarcity of alternatives in terms of time-varying proxies of entry.

The time-varying entry variable allows to control for both the level of entry and its interaction with distance to frontier. However, the UNIDO data do not include information on the number of establishments before the eighties, therefore limiting the sample period for the analysis. In addition, casual inspection of the original data indicates a somewhat more noisy nature of the establishment figures relative to employment and value added figures, with somewhat large (abnormal) jumps (which we decide not to censor in the analysis).

The effect of democracy is now reduced substantially. If we interpret effective entry as a precise proxy for entry costs, this finding tends to support our interpretation of the data. However, only for the 10-year sample the marginal effects of entry depend on the distance to frontier in a direction in line with the model's prediction. Entry in levels has a positive and significant effect in all regressions and the interaction with distance is instead negative. Statistically significant coefficients on the interaction between entry and distance are present for both sectorial output and value added. The effect of democracy is generally negative

at the mean distance to frontier (0.16), but the coefficients on democracy and its interaction with distance are significant only in the case of political rights. This suggests that indeed an entry channel may explain why democracy is conducive to growth in sectors closer to the technological frontier rather than other less advanced sectors, at least over relatively long periods of time (10 years).

## 4 Conclusions

This paper investigates a different approach and some novel empirical evidence on the relationship between democracy, level of development, and economic growth by employing disaggregated data on industrial sectors' growth rates. We have argued (and shown empirically with some degree of success) that democratic institutions favor growth in sectors of the economy that are particularly advanced in terms of value added per worker, or in our terminology, close to the world technological frontier. Our interpretation is that it is in sectors close to the technological frontier that democracy is more beneficial, possibly through fostering entry, competition, and innovation, which are relatively more important for growth in those sectors. Thus our analysis introduces a technological motive for political freedom survival, a dimension that has not received much attention in the wide empirical literature on democratization. It also suggests that the demand for democracy should be higher in richer countries, that is in which more sectors are closer to the technological frontier.

In a general sense this paper contributes to the strand of the politico-economic literature focusing on the differential effects of political institutions on economic outcomes. The same institutional features may have quite different effects on different components of the economy.

Natural next steps in this research program would be, first, to dig further into the various channels whereby democracy fosters growth in more advanced sectors, and, second, to analyze the process by which the economic demand for democracy may or may not translate into a real transition to democracy.

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Table 1: Aggregate manufacturing growth and democracy

		Table 1: A	Aggregate ma	ınufacturing	growth and	democracy			
	F.E.	F.E.	F.E.	F.E.	F.E.	F.E.	F.E.	F.E.	F.E.
	country	country	country	country	country	country	country	country	country
	year	year	year	year	year	year	year	year	year
				5-year	5-year	5-year	5-year	5-year	5-year
	5-year	5-year	5-year	Value	Value	Value	•	Employme	•
1975-80-85-90-95-00	Output	Output	Output	added	added	added	nt growth	nt growth	nt growth
	growth rate	growth rate	growth rate			growth rate	_	rate	rate
				Siowai iace	growth rate	Stowariace	Tate	Tate	Tate
L5. Dem. polityIV	0.001			-0.003			-0.009		
	[0.006]			[0.005]			[0.005]*		
L5. Polit. rights		0.023			-0.014			0.003	
		[0.024]			[0.023]			[0.016]	
L5. Civil liberty			0.029			-0.036			-0.008
			[0.031]			[0.033]			[0.022]
Observations	627	591	591	405	362	362	438	411	411
Number of countries	133	148	148	106	111	111	116	130	130
R-squared	0.08	0.08	0.08	0.34	0.26	0.27	0.12	0.06	0.06
				10-year	10-year	10-year	10-year	10-year	10-year
	10-year	10-year	10-year	Value	Value	Value	•	Employme	•
1975-1985-1995	Output	Output	Output	added	added	added	nt growth	nt growth	nt growth
	growth rate	growth rate	growth rate			growth rate	_	rate	rate
				grower race	810 11 11110	810 11 1410	1400	1410	1400
L10. Dem. polityIV	-0.003			0.007			0.001		
	[0.007]			[0.013]			[0.012]		
L10. Polit. rights		0.05			0.002			0.102	
		[0.038]			[0.081]			[0.055]*	
L10. Civil liberty			0.05			0.057			-0.046
			[0.051]			[0.065]			[0.077]
Observations	277	226	226	181	142	142	190	157	157
Number of countries	111	121	121	82	82	82	87	92	92
R-squared	0.13	0.14	0.14	0.37	0.29	0.3	0.09	0.03	0.02
				10-year	10-year	10-year	10-year	10-year	10-year
	10-year	10-year	10-year	Value	Value	Value			Employme
1980-1990-2000	Output	Output	Output	added	added	added	nt growth	nt growth	nt growth
	growth rate	growth rate	growth rate			growth rate	_	rate	rate
				_	g,				
L10. Dem. polityIV	0.02			-0.009			-0.022		
	[0.010]*			[0.008]			[0.008]***		
L10. Polit. rights		0.087			-0.028			-0.07	
		[0.051]*			[0.053]			[0.046]	
L10. Civil liberty			0.098			-0.04			-0.084
			[0.089]			[0.053]			[0.042]**
Observations	290	217	217	184	121	121	194	133	133
Number of countries	111	119	119	82	77	77	89	85	85
R-squared	0.04	0.05	0.04	0.45	0.01	0.01	0.27	0.15	0.15
Clustered standard erro	ors in bracket	S							

Clustered standard errors in brackets

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Table 2: Growth by sector and democracy

			-	sector and	•				
	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC
	country	country	country	country	country	country	country	country	country
	year	year	year	year	year	year	year	year	year
	5-year	5-year	5-year	5-year	5-year	5-year	_	_	_
	Output	Output	Output	Value	Value	Value	5-year	5-year	5-year
	growth	growth	growth	added growth	added growth	added growth		nt growth	Employme nt growth
1975-80-85-90-95-00	rate	rate	rate	rate	rate	rate	rate	rate	rate
L5. Dem. polityIV	0.002			0.001	1410	1410	-0.008	1400	1400
- F - 7	[0.002]			[0.002]			[0.001]***		
L5. Polit. rights	. ,	-0.004		. ,	-0.006		. ,	-0.012	
O		[0.009]			[0.009]			[0.006]*	
L5. Civil liberty			-0.015			-0.015			-0.012
			[0.010]			[0.011]			[0.007]*
L5. Real GDP per capita	-0.321	-0.518	-0.505	-0.155	-0.397	-0.384	-0.141	-0.178	-0.17
	[0.048]***	[0.057]***	[0.058]***	[0.049]***	[0.059]***	[0.060]***	[0.031]***	[0.038]***	[0.038]***
Observations	9250	8234	8234	9412	8078	8078	10097	8932	8932
Number of countries	2579	2650	2650	2408	2447	2447	2619	2779	2779
R-squared	0.3	0.26	0.26	0.26	0.22	0.22	0.08	0.05	0.05
	10-year	10-year	10-year	10-year	10-year	10-year			
	Output	Output	Output	Value	Value	Value	10-year	10-year	10-year
	growth	growth	growth	added	added	added			Employme
1975-1985-1995	rate	rate	rate	growth rate	growth rate	growth rate	nt growth rate	nt growth rate	nt growth rate
L10. Dem. polityIV	0.003			-0.002	Tate	Tate	0.007	Tate	Tate
210. Deni. ponty1 v	[0.006]			[0.006]			[0.004]*		
L10. Polit. rights	[*****]	-0.001		[~.~~]	-0.043		[,1]	0.017	
O		[0.031]			[0.031]			[0.019]	
L10. Civil liberty			0.104			0.135			0.008
			[0.031]***			[0.030]***			[0.022]
L10. Real GDP per capita	-0.239	-0.142	-0.163	-0.04	0.064	0.04	-0.428	-0.356	-0.35
	[0.119]**	[0.182]	[0.184]	[0.113]	[0.171]	[0.173]	[0.085]***	[0.129]***	[0.129]***
Observations	4103	3217	3217	4164	3183	3183	4386	3486	3486
Number of countries	1963	1950	1950	1941	1867	1867	2038	2032	2032
R-squared	0.19	0.07	0.07	0.17	0.07	0.08	0.1	0.01	0.01
	10-year	10-year	10-year	10-year	10-year	10-year	10-year	10-year	10-year
	Output	Output	Output	Value added	Value added	Value added	Employme	Employme	Employme
	growth	growth	growth	growth	growth	growth	nt growth	nt growth	nt growth
1980-1990-2000	rate	rate	rate	rate	rate	rate	rate	rate	rate
L10. Dem. polityIV	-0.001			-0.006			-0.013		
210. 20m ponty1	[0.003]			[0.003]*			[0.003]***		
L10. Polit. rights	. ,	-0.006		. ,	-0.012		. ,	-0.044	
O		[0.020]			[0.021]			[0.016]***	
L10. Civil liberty		-	0.01		•	-0.028		-	-0.049
			[0.024]			[0.026]			[0.018]***
L10. Real GDP per capita	-0.715	-1.044	-1.059	-0.592	-0.984	-0.958	-0.379	-0.203	-0.165
	[0.096]***	[0.159]***	[0.160]***	[0.101]***	[0.166]***	[0.166]***	[0.072]***	[0.103]**	[0.106]
Observations	4060	2661	2661	4294	2682	2682	4485	2874	2874
Number of countries	1873	1685	1685	1962	1693	1693	2087	1811	1811
R-squared Clustered standard errors in br	0.45	0.1	0.1	0.37	0.09	0.09	0.17	0.03	0.03

Clustered standard errors in brackets

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Table 3: Differential effect of democracy depending on distance to frontier

Table	e 3: Differen					e to frontier	:		
	F.E. SIC	F.E. SIC	F.E. SIC		F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC
	country	country	country	country	country	country	country	country	country
	year	year	year	year	year	year	year	year	year
	5-year	5-year	5-year				5-year	5-year	5-year
	Output	Output	Output		5-year VA	-	_	Emp.	Emp.
	growth	growth	growth	growth	growth	growth	growth	growth	growth
1975-80-85-90-95-00	rate	rate	rate	rate	rate	rate	rate	rate	rate
L5. Distance to frontier in VA/Emp	4.283	5.939	6.044	5.06	6.307	6.262	0.873	1.845	1.626
	[1.259]***	[1.588]***	[1.588]***	[1.308]***	[1.707]***	[1.700]***	[0.952]	[1.142]	[1.159]
L5. Dist. to front. x dem. polityIV	0.008			-0.022			-0.056		
	[0.024]			[0.023]			[0.018]***		
L5. Dem. polityIV	0.001			0.005			0.004		
15 D: 6 D: 11	[0.004]	0.054		[0.004]	0.400		[0.003]	0.00	
L5. Dist. to front. x Polit. rights		-0.076			-0.138			-0.29	
		[0.076]			[0.080]*			[0.052]***	
L5. Polit. rights		0.006			0.01			0.039	
		[0.014]	0.120		[0.014]	0.24		[0.010]***	0.240
L5. Dist. to front. x Civil liberty			-0.128			-0.21			-0.249
15.0: 319			[0.087]			[0.093]**			[0.063]***
L5. Civil liberty			0.003			0.019			0.04
15 División 16DD	0.402	0.260	[0.015]	0.062	0.047	[0.016]	0.455	0.4	[0.011]***
L5. Dist. to front. x real GDP p.c.	-0.192	-0.269	-0.256	-0.063	-0.047	-0.007	-0.155	-0.1	-0.104
15 D 10DD	[0.147]	[0.179]	[0.175]	[0.153]	[0.192]	[0.190]	[0.110]	[0.130]	[0.128]
L5. Real GDP p.c.	-0.148	-0.298	-0.29	0.087	-0.11	-0.117	-0.218	-0.223	-0.237
01	[0.061]**	[0.070]***	[0.070]***	[0.059]	[0.069]	[0.069]*	[0.044]***	[0.051]***	[0.050]***
Observations	8110	7449	7449	8427	7649	7649	8607	7893	7893
Number of industry-country R-squared	2205 0.34	2295 0.31	2295 0.31	2250 0.33	2318 0.3	2318 0.3	2257 0.09	2368 0.07	2368 0.06
K-squared									
	10-year Output	10-year Output	10-year Output	10-year VA	10-year VA	10-year VA	10-year	10-year	10-year
	growth	growth	growth	growth	growth	growth	Emp. growth	Emp. growth	Emp. growth
1975-1985-1995	rate	rate	rate	rate	rate	rate	rate	rate	rate
L10. Distance to frontier in VA/Emp	13.985	15.496	16.527	10.689	11.224	12.987	1.274	-3.171	-1.676
E10. Distance to Honder in VII, Emp	[2.278]***		[4.212]***	[2.503]***		[5.384]**	[1.713]	[2.997]	[3.181]
L10. Dist. to front. x dem. polityIV	-0.04	[5.7 15]	[1.212]	-0.066	[5.250]	[5.501]	-0.035	[2.777]	[5.101]
Ero. Dist. to front. x defin. pontyr v	[0.057]			[0.053]			[0.044]		
L10. Dem. polityIV	0.01			0.012			0.017		
E10. Dem. ponty1 v	[0.010]			[0.009]			[0.007]**		
L10. Dist. to front. x Polit. rights	[0.010]	-0.58		[0.007]	-0.447		[0.007]	-0.469	
		[0.216]***			[0.223]**			[0.157]***	
L10. Polit. rights		0.105			0.046			0.095	
		[0.040]***			[0.042]			[0.027]***	
L10. Dist. to front. x Civil liberty		[0.0.0]	-0.387		[0.0]	-0.334		[~.~=.]	-0.576
,			[0.266]			[0.266]			[0.198]***
L10. Civil liberty			0.157			0.172			0.123
			[0.040]***			[0.041]***			[0.029]***
I 10 Dist to front wroal CDD a s		-1.073	-1.318	-0.655	-0.362	-0.654	-0.309	0.529	0.395
LIU, DISI, IO HOIH, X ICALVITATI DA	-1.319					[0.580]	[0.209]	[0.339]	[0.335]
L10. Dist. to front. x real GDP p.c.	-1.319 [0.277]***		[0.441]***	[0.2961**]	10.000			[~~~]	[]
-	[0.277]***	[0.450]**	[0.441]*** 0.219	[0.296]** 0.305	[0.600] 0.399	-	-0.422	-0.523	-0.5
L10. Real GDP p.c.	[0.277]*** 0.183	[0.450]** 0.218	0.219	0.305	0.399	0.413	-0.422 [0.104]***	-0.523 [0.151]***	-0.5 [0.151]***
L10. Real GDP p.c.	[0.277]*** 0.183 [0.147]	[0.450]** 0.218 [0.214]	0.219 [0.216]	0.305 [0.139]**	0.399 [0.226]*	0.413 [0.225]*	[0.104]***	[0.151]***	[0.151]***
L10. Real GDP p.c. Observations	[0.277]*** 0.183 [0.147] 3803	[0.450]** 0.218 [0.214] 3045	0.219 [0.216] 3045	0.305 [0.139]** 3900	0.399 [0.226]* 3114	0.413 [0.225]* 3114	[0.104]*** 3999	[0.151]*** 3238	[0.151]*** 3238
L10. Real GDP p.c.	[0.277]*** 0.183 [0.147]	[0.450]** 0.218 [0.214]	0.219 [0.216]	0.305 [0.139]**	0.399 [0.226]*	0.413 [0.225]*	[0.104]***	[0.151]***	[0.151]***

Clustered standard errors in brackets. \* significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Table 3: Differential effect of democracy depending on distance to frontier (cont.)

	F.E. SIC								
	country								
	year								
	10-year								
	Output	Output	Output	VA	VA	VA	Emp.	Emp.	Emp.
	growth								
1980-1990-2000	rate								
L10. Distance to frontier in VA/Emp	6.672	11.612	11.296	4.619	13.327	13.666	8.025	15.233	15.414
	[2.682]**	[4.183]***	[4.117]***	[2.955]	[4.260]***	[4.234]***	[2.196]***	[3.648]***	[3.612]***
L10. Dist. to front. x dem. polityIV	0.158			0.147			0.062		
	[0.047]***			[0.046]***			[0.039]		
L10. Dem. polityIV	-0.027			-0.033			-0.021		
	[0.006]***			[0.007]***			[0.005]***		
L10. Dist. to front. x Polit. rights		-0.676			-0.763			-0.683	
_		[0.201]***			[0.194]***			[0.161]***	
L10. Polit. rights		0.064			0.067			0.004	
		[0.023]***			[0.024]***			[0.022]	
L10. Dist. to front. x Civil liberty			-0.697			-0.852			-0.841
·			[0.214]***			[0.214]***			[0.166]***
L10. Civil liberty			0.115			0.1			0.023
•			[0.033]***			[0.037]***			[0.026]
L10. Dist. to front. x real GDP p.c.	-0.483	-0.444	-0.411	-0.121	-0.427	-0.414	-1.083	-1.263	-1.197
•	[0.306]	[0.428]	[0.414]	[0.334]	[0.451]	[0.432]	[0.246]***	[0.395]***	[0.378]***
L10. Real GDP p.c.	-0.542	-0.89	-1.071	-0.511	-0.83	-0.959	-0.254	0.151	0.099
•	[0.127]***	[0.179]***	[0.173]***	[0.140]***	[0.180]***	[0.172]***	[0.111]**	[0.142]	[0.141]
Observations	3546	2499	2499	3662	2546	2546	3729	2638	2638
Number of industry-country	1737	1599	1599	1784	1634	1634	1782	1619	1619
R-squared	0.47	0.22	0.22	0.4	0.25	0.25	0.18	0.08	0.09

Clustered standard errors in brackets

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Table 4.	Growth	Democracy.	and Actual	entry

	F.E. SIC								
	country								
	year								
	Output	Output	Output	Value	Value	Value	5-year	5-year	5-year
	growth	growth	growth	added	added	added	Employm	Employm	Employm
1985-90, 1990-95, 1995-2000	rate	rate	rate	growth	growth	growth	ent	ent	ent
L5. Distance to frontier in VA/Emp	-16.168	-10.385	-0.197	-15.656	-11.037	0.732	-0.187	2.102	3.607
	[9.651]*	[7.660]	[8.171]	[10.027]	[9.887]	[9.496]	[7.299]	[5.304]	[5.394]
L5. Dist. to front. x 5-yr No. Esta.growth	-0.296	-0.372	-0.448	0.236	0.13	0.163	-0.192	-0.381	-0.33
	[0.549]	[0.514]	[0.526]	[0.450]	[0.462]	[0.451]	[0.465]	[0.453]	[0.480]
L5. 5-year No. Establ. growth rate	-0.058	-0.066	-0.055	-0.143	-0.165	-0.176	-0.055	-0.047	-0.046
	[0.081]	[0.078]	[0.080]	[0.045]***	[0.045]***	[0.046]***	[0.053]	[0.051]	[0.054]
L5. Dist. to front. x dem. polityIV	-0.368			-0.36			-0.208		
	[0.205]*			[0.200]*			[0.150]		
L5. Dem. polityIV	0.105			0.137			0		
	[0.033]***			[0.030]***			[0.028]		
L5. Dist. to front. x Polit. rights		0.497			0.611			0.062	
		[0.283]*			[0.274]**			[0.149]	
L5. Polit. rights		-0.111			-0.119			0.05	
		[0.070]			[0.066]*			[0.047]	
L5. Dist. to front. x Civil liberty			-0.453			-0.394			-0.206
			[0.340]			[0.368]			[0.254]
L5. Civil liberty			-0.074			-0.1			0.022
			[0.048]			[0.049]**			[0.032]
L5. Dist. to front. x Real GDP p.c.	2.577	1.749	1.086	2.83	2.102	1.288	0.103	-0.24	-0.266
	[1.115]**	[0.867]**	[0.913]	[1.164]**	[1.088]*	[1.044]	[0.883]	[0.620]	[0.624]
L5. Real GDP p.c.	-1.673	-1.478	-1.731	-1.791	-1.777	-2.01	-0.864	-0.575	-0.624
	[0.347]***	[0.224]***	[0.217]***	[0.309]***	[0.225]***	[0.219]***	[0.201]***	[0.158]***	[0.164]***
Observations	1694	1789	1789	1742	1835	1835	1819	1910	1910
Number of industry-country	1117	1161	1161	1145	1187	1187	1184	1228	1228
R-squared	0.31	0.33	0.34	0.37	0.4	0.41	0.12	0.1	0.1

Clustered standard errors in brackets

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

	Table 4: G	rowth, Den	nocracy, an	d Actual er	ntry (Cont.)				
	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC	F.E. SIC
	country	country	country	country	country	country	country	country	country
	year	year	year	year	year	year	year	year	year
	10-year Output growth	10-year Output growth	10-year Output growth	10-year Value added growth	10-year Value added growth	10-year Value added growth	10-year Employm ent growth	10-year Employm ent growth	10-year Employm ent growth
1985-95, 1990-2000	rate	rate	rate	rate	rate	rate	rate	rate	rate
L10. Distance to frontier in VA/Emp	-8.102	-7.676	-7.387	-10.077	-8.546	-13.865	1.111	0.285	-6.124
•	[9.748]	[9.374]	[10.137]	[9.889]	[8.825]	[9.523]	[5.762]	[4.884]	[5.726]
L10. Dist. to front. x 10-yr No. Esta. growth	-1.916	-1.985	-2.332	-1.939	-1.857	-1.983	-0.217	-0.56	-0.612
_	[0.646]***	[0.639]***	[0.652]***	[0.689]***	[0.638]***	[0.664]***	[0.616]	[0.631]	[0.649]
L10. 10-yr No. Establ. growth rate	0.691	0.637	0.695	0.565	0.521	0.555	0.486	0.486	0.505
	[0.104]***	[0.112]***	[0.113]***	[0.119]***	[0.117]***	[0.119]***	[0.119]***	[0.128]***	[0.127]***
L10. Dist. to front. x dem. polityIV	-0.277			-0.071			-0.022		
	[0.140]**			[0.124]			[0.082]		
L10. Dem. polityIV	0.029			0.014			-0.009		
	[0.021]			[0.019]			[0.015]		
L10. Dist. to front. x Polit. rights		-1.335			-1.611			-0.283	
		[0.466]***			[0.328]***			[0.198]	
L10. Polit. rights		0.288			0.31			0.152	
		[0.096]***			[0.056]***			[0.036]***	
L10. Dist. to front. x Civil liberty			-0.639			-1.148			-0.53
			[0.393]			[0.362]***			[0.248]**
L10. Civil liberty			0.136			0.172			0.153
			[0.062]**			[0.054]***			[0.034]***
L10. Dist. to front. x Real GDP p.c.	1.524	2.025	1.525	2.079	2.733	2.994	-0.334	-0.005	0.748
	[1.131]	[1.269]	[1.276]	[1.135]*	[1.068]**	[1.158]***	[0.653]	[0.538]	[0.686]
L10. Real GDP p.c.	-1.478	-1.965	-1.994	-1.671	-1.763	-1.963	-0.562	-0.757	-0.944
	[0.307]***	[0.254]***	[0.298]***	[0.267]***	[0.245]***	[0.272]***	[0.184]***	[0.155]***	[0.185]***
Observations	1718	1852	1852	1732	1838	1838	1837	1980	1980
Number of industry-country	1186	1267	1267	1170	1225	1225	1220	1314	1314

Clustered standard errors in brackets

R-squared

0.55

0.49

0.47

0.57

0.52

0.5

0.5

0.45

0.44

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%

Table A1: Summary Statistics by Year

Year	1975	1975	1975	Table A1: S	ummary Sta 1975	1980	ar 1980	1980	1980	1980
Variable	Obs	Mean	St.dev.	Min	Max	Obs	Mean	St.dev.	Min	Max
	220		Sudov.			2 20		200011		
logY	2249	19.16	2.75	8.15	26.73	2348	19.60	2.73	9.04	26.76
logVA	2272	18.09	2.63	10.26	25.66	2269	18.52	2.63	10.50	25.97
logEmp	2494	8.45	2.43	0.00	15.79	2576	8.63	2.41	0.00	15.79
logEsta	0					0				
dlogY	2176	0.04	0.32	-3.54	4.25	2184	0.06	0.34	-3.58	7.79
O										
dlogVA	2238	0.03	0.33	-3.18	4.13	2133	0.04	0.35	-3.19	7.96
dlogEmp	2428	0.04	0.26	-2.85	3.76	2469	0.02	0.26	-2.84	6.58
dlogEsta	0					0				
d5logY	1869	0.61	0.58	-2.35	5.39	2054	0.41	0.57	-3.48	4.44
d5logVA	2074	0.57	0.60	-2.70	4.48	2101	0.39	0.60	-3.21	4.15
d5logEmp	2191	0.26	0.47	-1.61	3.24	2342	0.16	0.50	-3.67	6.92
d5logEsta	0					0				
d10logY	1429	0.93	0.78	-2.42	5.59	1812	0.98	0.81	-3.49	6.99
d10logVA	1475	0.90	0.77	-1.67	5.71	1971	0.91	0.84	-3.64	6.51
d10logEmp	1584	0.44	0.63	-2.45	4.38	2143	0.41	0.70	-3.64	6.45
d10logEsta	0					0				
distfront	2171	0.15	0.08	0.00	0.45	2159	0.15	0.08	0.00	0.49
db_proc_99	1911	10.19	4.51	2.00	21.00	1969	10.22	4.45	2.00	21.00
polityIV	2837	-0.91	7.83	-10.00	10.00	2836	-0.33	7.98	-10.00	10.00
pr	3168	3.64	2.24	1.00	7.00	3262	3.96	2.20	1.00	7.00
cl	3168	3.98	1.94	1.00	7.00	3262	3.95	1.96	1.00	7.00
Year	1985	1985	1985	1985	1985	1990	1990	1990	1990	1990
Variable	Obs	Mean	St.dev.	Min	Max	Obs	Mean	St.dev.	Min	Max
10 oV	2249	19.51	2.69	9.31	26.85	2286	19.73	2.79	10.20	26.88
logY					25.94					25.96
logVA	2067	18.51	2.54	10.34		2124	18.89	2.62	9.29	
logEmp	2387	8.73	2.44	0.00	15.80	2441	8.88	2.31	0.00	16.11
logEsta	2069 2179	4.37 -0.02	2.24 0.32	0.00 -2.50	11.02 4.78	2202 2048	4.38 0.00	2.25 0.43	0.00 -5.83	10.97 3.49
dlogY	21/9	-0.02	0.32	-2.30	4.70	2046	0.00	0.43	-3.63	3.49
dlogVA	1992	-0.04	0.36	-2.40	4.86	1941	0.03	0.43	-6.08	4.46
dlogEmp	2309	0.01	0.21	-3.89	3.82	2159	0.03	0.43	-3.21	2.99
dlogEsta	1892	0.03	0.21	-1.61	1.72	1912	0.02	0.42	-4.62	3.76
d5logY	2008	-0.26	0.58	-4.94	4.30	1787	0.32	0.63	-7.91	4.47
d5logVA	1868	-0.20	0.63	-5.15	4.52	1701	0.34	0.66	-7.12	3.35
4310g v 11	1000	0.51	0.03	J.1J	1.74	1/01	U.JT	0.00	1.14	5.55
d5logEmp	2181	-0.01	0.43	-3.60	4.91	1881	0.09	0.45	-6.38	3.93
d5logEsta	0	0.01	J. 13	5.00	1	1576	0.23	0.43	-4.31	4.27
d10logY	1875	0.15	0.77	-4.11	4.68	1736	0.07	0.76	-4.62	7.27
d10logVA	1835	0.13	0.79	-2.78	3.85	1717	0.05	0.82	-6.09	7.39
d10logEmp	2086	0.15	0.65	-4.22	5.15	1871	0.07	0.61	-4.97	7.76
aronogramp	2000	0.13	0.03	1.44	5.15	10/1	0.07	0.01	1.77	1.10
d10logEsta	0					0				
distfront	2100	0.15	0.08	0.00	0.47	1953	0.16	0.09	0.00	0.53
db_proc_99	1989	10.20	4.43	2.00	21.00	2133	10.19	4.33	2.00	21.00
polityIV	2861	0.37	7.93	-10.00	10.00	2806	2.45	7.41	-10.00	10.00
L OTTO 1		0.57		10.00	-0.00	_000	10	,	10.00	10.00

pr	3287	4.12	2.23	1.00	7.00	3151	4.46	2.19	1.00	7.00
cl	3287	3.90	2.05	1.00	7.00	3151	4.53	1.86	1.00	7.00

Table A1: Summary Statistics by Year (Cont.)

				le A1: Sumn			-			
Year	1995	1995	1995	1995	1995	2000	2000	2000	2000	2000
Variable	Obs	Mean	St.dev.	Min	Max	Obs	Mean	St.dev.	Min	Max
logY	2289	19.31	2.99	8.35	27.00	1649	19.71	3.05	6.68	27.12
logVA	1982	18.69	2.85	8.00	26.04	1405	19.16	2.71	8.70	26.20
logEmp	2408	8.68	2.39	0.00	15.77	1780	8.90	2.38	0.00	15.41
logEsta	2225	4.40	2.50	0.00	11.38	1695	5.05	2.44	0.00	11.47
dlogY	2171	0.08	0.49	-7.69	6.70	1579	0.03	0.31	-2.31	2.31
dlogVA	1877	0.07	0.51	-6.60	5.68	1312	-0.01	0.27	-1.92	1.73
dlogEmp	2280	0.00	0.34	-4.12	4.20	1674	0.00	0.20	-2.39	1.78
dlogEsta	2109	0.10	0.48	-2.69	3.84	1621	0.05	0.36	-3.76	2.68
d5logY	1709	0.06	0.91	-5.92	5.09	1495	-0.07	0.75	-7.10	6.02
d5logVA	1654	0.04	0.80	-6.47	4.08	1246	-0.07	0.63	-4.64	3.94
d5logEmp	1944	-0.11	0.68	-5.89	4.30	1613	-0.07	0.50	-4.69	2.78
d5logEsta	1673	0.33	0.97	-4.23	4.87	1396	0.15	0.70	-4.22	5.46
d10logY	1546	0.43	1.03	-10.86	5.31	1233	0.02	1.11	-5.82	5.68
d10logVA	1486	0.47	0.97	-6.89	4.67	1126	0.05	0.83	-4.87	5.29
d10logEmp	1677	0.12	0.73	-6.20	4.33	1417	-0.14	0.78	-5.37	3.43
d10logEsta	1328	0.50	1.13	-3.57	5.01	1289	0.53	1.11	-4.25	5.23
distfront	1969	0.18	0.10	0.00	0.82	1410	0.16	0.10	0.00	0.85
db_proc_99	2194	10.30	4.34	2.00	21.00	1954	10.17	4.39	2.00	21.00
polityIV	3241	3.79	6.54	-10.00	10.00	2542	4.96	6.07	-10.00	10.00
pr	3576	4.74	2.06	1.00	7.00	2729	5.11	2.06	1.00	7.00
cl	3576	4.48	1.79	1.00	7.00	2729	4.87	1.64	1.00	7.00

	Procedure	Procedure	Procedure	Procedure	Procedure	Procedure	Procedure	Procedure
	s No.	s No.	s No.	s No.	s No.	s No.	s No.	s No.
GDP per capita, log	-0.544	-0.687	-0.111	-0.367	-0.757	-0.853	-1.094	-0.783
	[0.337]	[0.367]*	[0.382]	[0.511]	[0.320]**	[0.338]**	[0.440]**	[0.383]**
executive de facto independence	-0.853							
	[0.270]***							
constraints on executive power		-0.658						
		[0.290]**						
effectiveness of legislature			-2.553					
			[0.661]***					
government effectiveness				-1.88				
				[0.933]**				
competition in the legisl. nominating process					-2.051			
					[0.963]**			
autocracy						0.322		
						[0.185]*		
political rights average 1972-99							-0.655	
							[2.263]	
civil rights average 1972-99							. ,	-4.171
								[2.362]*
Observations	84	84	73	85	73	84	84	84
R-squared	0.26	0.24	0.3	0.25	0.22	0.22	0.2	0.23

Robust standard errors in brackets

<sup>\*</sup> significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%