

1 Measuring  $\beta$ -diversity by remote sensing: a  
2 challenge for biodiversity monitoring

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## 51 **Abstract**

52 Biodiversity includes multiscalar and multitemporal structures and  
53 processes, with different levels of functional organization, from genetic  
54 to ecosystemic levels. One of the mostly used methods to infer bio-  
55 diversity is based on taxonomic approaches and community ecology  
56 theories. However, gathering extensive data in the field is difficult due  
57 to logistic problems, overall when aiming at modelling biodiversity  
58 changes in space and time, which assumes statistically sound sam-  
59 pling schemes. In this view, [airborne](#) or satellite remote sensing allow  
60 to gather information over wide areas in a reasonable time.

61 Most of the biodiversity maps obtained from remote sensing have  
62 been based on the inference of species richness by regression analy-  
63 sis. On the contrary, estimating compositional turnover ( $\beta$ -diversity)

64 might add crucial information related to relative abundance of dif-  
65 ferent species instead of just richness. Presently, few studies have  
66 addressed the measurement of species compositional turnover from  
67 space.

68 Extending on previous work, in this manuscript we propose novel  
69 techniques to measure  $\beta$ -diversity from [airborne](#) or satellite remote  
70 sensing, mainly based on: i) multivariate statistical analysis, ii) the  
71 spectral species concept, iii) self-organizing feature maps, iv) multi-  
72 dimensional distance matrices, and the v) Rao's Q diversity. Each of  
73 these measures allow to solve one or several issues related to turnover  
74 measurement. This manuscript is the first methodological example  
75 encompassing (and enhancing) most of the available methods for es-  
76 timating  $\beta$ -diversity from remotely sensed imagery and potentially  
77 relate them to species diversity in the field.

78 *Keywords:*  $\beta$ -diversity, Kohonen self-organising feature maps, Rao's Q  
79 diversity index, remote sensing, satellite imagery, Sparse Generalized Dis-  
80 similarity Model, spectral species concept.

## 81 1 Introduction

82 Biodiversity cannot be fully investigated without considering the spatial com-  
83 ponent of its variation. In fact, it is known that the dispersal of species over  
84 wide areas is driven by spatial constraints directly related to the distance  
85 among sites. A negative exponential dispersal kernel is usually adopted to  
86 mathematically describe the occupancy of new sites by species, as:

$$F = \sum_{K=1}^N e^{-\frac{d_{ik}}{a}} \quad (1)$$

87 where  $d_{ik}$  = distance between two locations  $i$  and  $k$  and  $a$  is a parameter  
88 regulating the dispersal from localized areas (low values of  $a$ ) to widespread  
89 ones (high values of  $a$ , Meentemeyer et al. (2008)).

90 In this sense, distance acquires a significant role in ecology to estimate bio-  
91 diversity change. Hence, spatially explicit methods have been acknowledged  
92 in ecology for providing robust estimates of diversity at different hierarchical  
93 levels: from individuals (Tyre et al., 2001), to populations (Vernesi et al.,  
94 2012), to communities (Rocchini et al., 2005).

95 When dealing with spatial explicit methods, remote sensing images rep-  
96 resent a powerful tool, overall when coupling information on compositional  
97 properties of the landscape with its structure (Figure 1). Remote sensing has

98 widely been used for conservation practices including very different types of  
99 data such as nighlights data (Mazor et al., 2013), Land Surface Temperature  
100 estimated from MODIS data (Metz et al., 2014), spectral indices (Gillespie,  
101 2005).

102 Most of the remote sensing applications for biodiversity estimate have  
103 relied on the estimate of local diversity hotspots, considering land use diver-  
104 sity (Wegmann et al., 2017) or continuous spatial variability of the spectral  
105 signal (Rocchini et al., 2010). This is mainly grounded on the assumption  
106 that a higher landscape heterogeneity is strictly related to a higher amount  
107 of species occupying different niches. However, given two sites  $s_1$  and  $s_2$ ,  
108 the final diversity is not only related to the species / spectral richness of  $s_1$   
109 and  $s_2$ , but overall to the amount of shared species / spectral values. In  
110 other terms the lower their intersection  $s_1 \cap s_2$ , the higher will be the total  
111 diversity, while a low total diversity will be reached when  $s_1 \cap s_2 = s_1 \cup s_2$ .  
112 Such intersection has been widely studied in ecology, [after the development](#)  
113 [of  \$\beta\$ -diversity theory](#) (Whittaker, 1960).

114 Tuomisto et al. (2003) demonstrated the power of substituting distance  
115 in Eq. 1 by spectral distance to directly account for the distance among sites  
116 in an environmental space, instead of a merely spatial one. However, while  
117 spectral distance examples exist when measuring the  $\beta$ -diversity among pairs  
118 of sites (e.g. Rocchini et al. (2015)), few studies have tested the possibility of  
119 measuring  $\beta$ -diversity over wide areas considering several sites at the same  
120 time (however see Alahuhta et al. (2017); Harris et al. (2015)). This is  
121 overall true considering the development of remote sensing tools for diversity  
122 estimate in which the concept of  $\beta$ -diversity is still pioneering.

123 The aim of this paper is to present the most novel methods to measure  
124  $\beta$ -diversity from remotely sensed imagery based on the the most recently  
125 published ecological models. In particular we will deal with: i) multivariate  
126 statistical techniques, ii) the applicability of the spectral species concept,  
127 iii) multidimensional distance matrices, iv) metrics coupling abundance and  
128 distance-based measures.

129 This manuscript is the first methodological example encompassing (and  
130 enhancing) most of the available methods for estimating  $\beta$ -diversity from  
131 remotely sensed imagery and potentially relate them to species diversity in  
132 the field.

## 133 **2 Multivariate statistical analysis for species** 134 **diversity estimate from remote sensing**

135 Univariate statistics have been used to directly find relations between spectral  
136 and species diversity. However, the amount of variability explained by single  
137 bands / vegetation indices versus species diversity is generally relatively low,  
138 due to the fact that different aspects related to the complexity of habitats  
139 might act in shaping diversity, from disturbance and land use at local scales  
140 to climate and element fluxes at global scales.

141 Ordination techniques are designed to quantitatively describe multivari-  
142 ate gradual transitions in the species composition of sampled sites. Measuring  
143 the distance between two sampling sites in the multi-dimensional ordination  
144 space is a good proxy of the change in species composition. When this mea-  
145 sure is related to the geographical distance between the considered sites, the  
146 beta diversity at this particular scale can be assessed.

147 Of the various available ordination techniques, Detrended Correspondence  
148 Analysis (DCA, Hill and Gauch (1980)) is particularly suitable for  
149 such analyses. The axes (i.e. gradients) of the DCA ordination space are  
150 scaled in standard deviation (SD) units, where a distance of 4 SD is related  
151 to a full species turnover. This enables a versatile analysis that easily reveals  
152 whether two sampled sites still have species in common.

153 Several studies have mapped the ordination space using remote sensing  
154 data (e.g., Schmidtlein and Sassin (2004); Schmidtlein et al. (2007); Feil-  
155 hauer et al. (2009, 2011, 2014); Gu et al. (2015); Harris et al. (2015); Leitao  
156 et al. (2015); Neumann et al. (2015)). For this purpose, the axes scores of  
157 the sampled sites are regressed against the corresponding canopy reflectance  
158 values extracted from air- or spaceborne image data. The resulting multi-  
159 variate regression models, one per ordination axis and most often generated  
160 with machine learning regression techniques, are subsequently applied on the  
161 image data for a spatial prediction of ordination scores. Each pixel of the  
162 image data is assigned to a specific position in the ordination space that in-  
163 dicates its species composition. The resulting gradient maps are a powerful  
164 tool for analyses of beta diversity across different spatial scales (Feilhauer et  
165 al., 2009; Hernandez-Stefanoni et al., 2012).

166 A simple analysis of the variability of the DCA scores in a defined pixel  
167 neighborhood (i.e. a moving window) results in a efficient beta diversity  
168 assessment. The spatial scale of this assessment can be varied by either re-  
169 sampling the gradient map to a coarser spatial resolution (i.e. pixel size) or  
170 by changing the kernel size of the considered pixel neighborhood. Such tech-  
171 niques has been further developed e.g. for spatial conservation prioritization

172 programmes such as zonation (Moilanen et al., 2005, 2009).

173 Figure 2 shows an example of a DCA-based assessment of beta diversity  
174 on a very local scale (10 m) following the approach described in Feilhauer et  
175 al. (2009). The analyzed landscape is a mosaic of raised bogs, fens, transition  
176 mires and *Molinia* meadows. For a detailed description of the data and site  
177 please refer to Feilhauer et al. (2014, 2016).

178 Analyses like this require two different data sets: (1) a sample of field  
179 data that is representative for the vegetation in the studied area and is used  
180 to generate the ordination space; (2) image data with a sufficient spectral  
181 resolution to discriminate the vegetation types within the ordination space  
182 and with a spatial resolution that is in line with the sampling design of the  
183 field data (Feilhauer et al., 2013).

184 Using these data, the continuous spatial variability of the spectral signal  
185 in the image pixels is translated into a spatially continuous measure of species  
186 composition. The advantages of this approach are obvious: since the diversity  
187 analyses are conducted in the floristic gradient space, the resulting measures  
188 resemble field studies and are thus easier to interpret than spectral proxies  
189 and closer to the point of view of many end-users. [Furthermore](#), the analysis  
190 of ordination scores in defined pixel neighborhoods is not restricted to a  
191 single spatial scale but offers the opportunity to implement assessments of  
192 beta diversity on multiple scales.

### 193 **3 The spectral species concept**

194 The spectral species concept has been proposed by Féret and Asner (2014a)  
195 to map both  $\alpha$  and  $\beta$  component of the biodiversity using a unique frame-  
196 work. It is rooted in the convergence between two other concepts, the spec-  
197 tral variation hypothesis (SVH) proposed by Palmer et al. (2002), and the  
198 plant optical types proposed by Ustin and Gamon (2010), sustained by the  
199 technological advances in the domain of high spatial resolution imaging spec-  
200 troscopy. The SVH states that the spatial variability in the remotely sensed  
201 signal, that is the spectral heterogeneity, is related to environmental hetero-  
202 geneity and could therefore be used as a powerful proxy of species diversity.  
203 SVH has been tested in different situations (Rocchini et al., 2010) and con-  
204 clusions show that the performances of this approach are very dependent on  
205 several factors, including the instrumental characteristics (spectral, spatial  
206 and temporal resolution), the type of vegetation investigated, and the metrics  
207 derived from remotely sensed information to estimate spectral heterogeneity.  
208 Plant optical types refer to the capacity of sensors to measure signal aggreg-  
209 ating information about vegetation structure, phenology, biochemistry and

210 physiology. Therefore, this concept is also tightly linked to the performances  
211 of the sensor and finds particular echo with the increasing use of high spa-  
212 tial resolution imaging spectroscopy for the estimation and identification of  
213 multiple vegetation properties.

214 The details provided by high spatial resolution imaging spectroscopy are  
215 sufficient to perform analyzes of plant optical traits at the individual tree  
216 scale in order to differentiate tree species, obtain information about leaf chem-  
217 ical traits and estimate the  $\alpha$  component of biodiversity (Asner et al., 2008,  
218 2015; Chadwick and Asner , 2016; Clark et al., 2005; Clark and Roberts ,  
219 2012; Féret and Asner, 2013; Vaglio Laurin et al., 2014). These results il-  
220 lustrate that spectral information can be related to taxonomic or functional  
221 information of the vegetation, which supports the SVH under the hypothesis  
222 that the metrics used to compute spectral heterogeneity and a given com-  
223 ponent of vegetation diversity are properly defined. However these applica-  
224 tions are currently limited by the important amount of field data required  
225 to train regression or classification models, which is also directly linked to  
226 their low generalization ability in time and space. Unsupervised approaches  
227 then appear as valuable alternatives for the analysis of ecosystem heterogene-  
228 ity (Baldeck and Asner , 2013; Baldeck et al., 2014; Feilhauer et al., 2011;  
229 Baldeck and Asner , 2013; Féret and Asner, 2014b), as ecological indicators  
230 of  $\alpha$  and  $\beta$  diversity at landscape scale usually require one or several levels  
231 of abstraction beyond the correct taxonomic identification (Tuomisto et al.,  
232 2006).

233 Clustering (properly pre-processed) spectral information should result in  
234 pixels from the same species naturally grouping together rather than dis-  
235 tributing randomly among clusters, Féret and Asner (2014a) proposed a  
236 grouping method aiming at assigning labels to pixels based on multiple clus-  
237 tering of spectroscopic data acquired at landscape scale. These pixels labeled  
238 with a set of so-called spectral species can then be used straightforwardly in  
239 order to compute various diversity metrics such as Shannon index for  $\alpha$  diver-  
240 sity, and Bray-Curtis dissimilarity for  $\beta$  diversity. The pre-processing stage  
241 is divided into several stages. After masking all non-vegetated pixels, a nor-  
242 malization based on continuum removal is applied to each pixel and over the  
243 full spectral domain, then a principal component analysis is performed on  
244 the continuum removed spectral data. The normalization allows reducing  
245 effects due to changes in illumination, canopy geometry and other factors  
246 unrelated to vegetation, while enhancing the signal corresponding to veg-  
247 etation. The components including individual-specific information are the  
248 components of interest. They can be identified after visual inspection or au-  
249 tomated routines, if initial data show sufficient signal to noise ratio. Once  
250 a limited number of components have been selected, k-means clustering is

251 then applied to a certain number of subsets, and for each of these subsets,  
252 centroids are computed and each pixel in the image is labeled based on the  
253 closest centroid. The repetition of clustering based on various subsets of the  
254 image tends to minimize the risk of assigning centroids to irrelevant groups  
255 of pixels. Experimental results showed that the averaging of diversity indices  
256 computed from multiple centroid maps can be seen as an analogous to signal  
257 averaging, which consists in increasing signal to noise ratio by replicating  
258 measurements. For each repetition, the closest centroid corresponds to the  
259 spectral species, and for each spatial unit of a given size, the spectral species  
260 distribution is derived in order to compute any diversity metric requiring  
261 either information at the local scale, or comparison of information across  
262 spatially distant plots.

263 The concepts of spectral species and spectral species distribution have  
264 been tested successfully on a limited number of situations and types of ecosys-  
265 tems (see (Rocchini et al., 2016) for a review, and (Lausch et al., 2016) for  
266 an application to similar concepts). As an example, Féret and Asner (2014a)  
267 showed ability to properly estimate landscape heterogeneity at moderate spa-  
268 tial scale, up to few dozen square kilometers over tropical forests, based on  
269 high spatial resolution imaging spectroscopy (Figure 3). A generic parame-  
270 terization of the method showed robust performances for  $\alpha$  diversity mapping  
271 across space and time, but mapping  $\beta$  diversity across large spatial scales us-  
272 ing images acquired during different airborne campaign remains challenging,  
273 which leads to unsolved problem when considering operational regional map-  
274 ping. In the perspective of global monitoring of biodiversity, and based on the  
275 unprecedented remote sensing capacity allowed by the Copernicus program,  
276 including the Sentinel-2 multispectral satellites, several other challenges are  
277 foreseen and currently investigated. The influence of decreased spatial and  
278 spectral resolution on the ability to properly differentiate ecologically mean-  
279 ingful spectral species across landscapes and over regions will need to be  
280 investigated. The application of this concept beyond tropical forests and  
281 savanna ecosystems should also be investigated, as it may not hold when  
282 applied on moderately diverse ecosystems or systems with individuals with  
283 lower than metric dimensions.

## 284 4 Self organizing feature maps

285 The Kohonen self-organising feature map (SOFM, Kohonen (1982)) is a neu-  
286 ral network that may be used to undertake unsupervised clustering of data.  
287 Critically, the input to a SOFM can be a large multi-variate data set such as  
288 may be acquired on species from quadrat based field surveys and summarise



289 the data in a low, typically two, dimensional output (Figure 4). In this out-  
290 put space the data for individual quadrats are topologically ordered – with  
291 sites that are similar close together while those of highly different species  
292 composition more distant. Because the data sites in the output space are ar-  
293 ranged by relative similarity the output space may also be used to aggregate  
294 or classify a data set. As such the SOFM is attractive as a non-parametric  
295 clustering analysis and as a means to undertake an ordination (Chon et al.,  
296 1996).

297 A SOFM is, unlike some of the approaches used commonly in community  
298 ecology, not constrained by assumptions such as those relating the statistical  
299 distribution of the data used. The SOFM uses unsupervised learning to pro-  
300 duce a topologically ordered output space in which the samples are arranged  
301 spatially in relation to their relative similarity in species composition. The  
302 SOFM thus performs a non-parametric ordination analysis (Foody, 1999).  
303 The production of a classification by a SOFM comprises two main stages  
304 (Giraudel and Lek, 2001). An iterative analysis, in which time-decaying pa-  
305 rameters that control network learning and the size of local neighbourhoods  
306 located around output units, is used. For this, the user must specify a num-  
307 ber of key parameters such as the size and shape of the network, number  
308 of iterations of the algorithm, the learning rate and its rate of decline and  
309 a neighbourhood parameter. The need for such parameters can add some  
310 uncertainty to the analysis. While there are no formal rules to follow in the  
311 design of a SOFM there are recommendations for the determination of SOFM  
312 parameter settings (Giraudel and Lek, 2001). A further concern is that as  
313 an unsupervised classifier the classes defined may not always be the most  
314 useful for an investigation. In addition, the nature of the analysis means the  
315 direction of the gradients cannot be controlled (Fritzke, 1995) but the anal-  
316 ysis performs well in comparison to popular ordination techniques such as  
317 PCA and DCA (Foody and Cutler, 2003). The SOFM may also use a variety  
318 of different data types such as presence/absence, abundance or importance  
319 values and can solve complex non-linear problems (Giraudel and Lek, 2001).

## 320 **5 Multidimensional distance matrices: GDMs** 321 **and SGDMs**

322 One of the most widespread methods for assessing  $\alpha$ -diversity is using distance  
323 matrices (Legendre et al., 2005). Indeed, early work by Whittaker (1960) sug-  
324 gested that  $\beta$ -diversity could be quantified by dissimilarity matrices among  
325 (pairs of) sites. Furthermore, the Mantel test (Mantel and Valand, 2017),

326 designed to estimate the association between two independent dissimilarity  
327 matrices, has been widely used to correlate a community composition dissim-  
328 ilarity matrix with an environment dissimilarity one, thus providing useful  
329 insights into community composition and turnover (Legendre et al., 2005;  
330 Tahvanainen et al., 2011).

331 Generalized Dissimilarity Modelling (GDM; Ferrier (2007) can be con-  
332 sidered as an extension of the Mantel test, which is able to accommodate  
333 multidimensional environmental data, to be compared with the composi-  
334 tional data. GDMs also allow for the prediction of compositional turnover  
335 as well as for, e.g. environmental classification constrained to the compo-  
336 sitional dissimilarity (Ferrier, 2007; Leathwick et al., 2011). In GDM, the  
337 compositional dissimilarities between all pairs of samples are modelled as a  
338 function of their respective environmental distances. This is done through a  
339 linear combination of monotonic I-spline basis functions, under the assump-  
340 tion that increasing environmental dissimilarity (e.g. along a gradient) can  
341 only result in increasing compositional dissimilarity. This method is thus well  
342 suited for measuring and mapping  $\beta$ -diversity, and is thus becoming widely  
343 used in conservation science and macroecology, and recently been subject to  
344 several developments as we describe below.

345 One such development is the phylogenetic GDM (phylo-GDM; Rosauer  
346 et al. (2014)), which incorporates phylogenetic dissimilarities into GDM and  
347 allows for analysing and predicting phylogenetic  $\beta$ -diversity, thus linking  
348 ecological and evolutionary processes. This method can provide novel in-  
349 sights into the mechanisms underlying current patterns of biological diversity  
350 (Graham et al., 2008). Another recent development of GDM is the multi-  
351 site GDM (MS-GDM; Latombe et al. (2017)), which extends GDMs from  
352 pairwise to multi-site dissimilarity modelling. In such paper, the authors  
353 tested MS-GDM by means of both constrained (monotonical) additive mod-  
354 els and I-splines, although with no conclusive results relating to the best  
355 method overall. They concluded, however, that when applying MS-GDM to  
356 a high number of samples, they could better explain the drivers of species  
357 turnover. Also, an important development of GDM is the Bayesian bootstrap  
358 GDM (BBGDM; Woolley et al. (2017)) designed to characterize uncertainty  
359 in generalized dissimilarity models. This approach allows better represent-  
360 ing the underlying uncertainty in the data, by estimating the variance in  
361 parameters based on the available data.

362 Finally, an implementation of GDM, which was created particularly for  
363 dealing with high-dimensional (and potentially high-collinear) remote sensing  
364 data as input in GDM is the Sparse Generalized Dissimilarity Model (SGDM,  
365 Figure 5, Leitao et al. (2015)). This method is a two-stage approach that  
366 consists of initially reducing the environmental space (e.g. reflectance data)

367 by means of a Sparse Canonical Correlation Analysis (SCCA, Figure 5; Wit-  
 368 ten et al. (2013)), and then fitting the resulting components with a GDM  
 369 model. The SCCA is a form of penalized canonical correlation analysis based  
 370 on the L1 (lasso) penalty function, and is thus designed to deal with high-  
 371 dimensional data. The two algorithms are coupled in a way that the SCCA  
 372 penalization is selected through a heuristic grid search manner, in order to  
 373 minimize the cross-validate root mean square error in the dissimilarities pre-  
 374 dicted by the GDM. In this procedure, the high-dimensional environmental  
 375 data (such as coming from time series of multispectral or hyperspectral data)  
 376 are subject to a supervised ordination approach that reduces their dimen-  
 377 sion while capturing the axes of variation that most correlate to those of  
 378 the community compositional matrix. SGDM has been successfully used for  
 379 modelling and mapping the compositional turnover of both animal and plant  
 380 species, using several different sources of remote sensing (and auxiliary) data  
 381 (Leitao et al., 2015; Leitão et al., 2017).

## 382 **6 Rao's Q diversity**

383 Most of the previously shown metrics are based on the distance among pixel  
 384 values in a multidimensional spectral space. None of them considers the  
 385 relative abundance of such pixel values in a neighbourhood.

386 By contrast, abundance-based metrics such as the Shannon entropy could  
 387 output similar results despite a variable distance among pixel values. As an  
 388 example, consider a 3x3 matrix of remotely sensed data:

$$\begin{pmatrix} x_{11} & x_{12} & x_{13} \\ x_{21} & x_{22} & x_{23} \\ x_{d1} & x_{d2} & x_{d3} \end{pmatrix} \quad (2)$$

389 composed by the following values:

$$\begin{pmatrix} 10 & 13 & 15 \\ 18 & 20 & 23 \\ 19 & 21 & 22 \end{pmatrix} \quad (3)$$

390 then consider a different matrix:

$$\begin{pmatrix} 10 & 121 & 227 \\ 1 & 40 & 251 \\ 7 & 100 & 149 \end{pmatrix} \quad (4)$$

391 From a Shannon's entropy perspective, such matrices are equal in terms of  
 392 heterogeneity. The Shannon's entropy is indeed based on the relative abun-  
 393 dance (and richness) of a sample, and its value is 2.197 for both the matrices.

394 This value, equalling the natural logarithm of the number of classes (pixel  
395 values), is also Shannon’s maximum theoretical value given a 3x3 matrix,  
396 due to the lack of identical numbers in the matrices. This example explicitly  
397 shows that accounting for the distance among values and their relative abun-  
398 dance is crucial to discriminate among areas in terms of measured (modeled)  
399 heterogeneity.

400 One of the metrics accounting for both the abundance and the pairwise  
401 spectral distance among pixels is the Rao’s Q diversity index, as:

$$Q = \sum \sum d_{ij} \times p_i \times p_j \quad (5)$$

402 where  $d_{ij}$  = spectral distance among pixels  $i$  and  $j$  and  $p$  = proportion of  
403 occupied area.

404 Hence, Rao’s Q is capable to discriminate among the ecological diversity  
405 of matrices 3 and 4, turning out to be 4.59 and 90.70, respectively. Appendix  
406 1 provide an example spreadsheet to perform the calculation while the com-  
407 plete R code is stored in the GitHub repository  
408 <https://github.com/mattmar/spectralrao>.

409 We decided to make use of a case study to highlight the importance of  
410 considering the distance among pixel values in remote sense ecological appli-  
411 cation. The performance of Rao’s Q index in describing landscape diversity  
412 was tested in a complex agro-forestry landscape located in southern Portu-  
413 gal. A test site with an area of about  $10 \times 10 \text{ km}^2$  (centroid located at  $38^\circ$   
414  $39' 10.74''$  N;  $8^\circ 12' 52.30''$  W) was selected to conduct the analysis. In this  
415 area, a savanna-like ecosystem called montado occupies about 40% of the test  
416 site, followed by traditional olive groves, pastures, vineyards, and irrigated  
417 monocultures (e.g. corn fields). Montado is spatially characterized by the  
418 variability of its tree density (e.g. Godinho et al. (2016)), and the gradient  
419 between low and high tree density over space can lead to different structural  
420 heterogeneity and habitat diversity.

421 Within the test site, polyculture under small farming context (e.g. veg-  
422 etable gardens, orchards, and cereal crops) is an important feature of this  
423 landscape by generating a high compositional and configurational spatial  
424 heterogeneity (Figure 6). The main goal in using this case study is to demon-  
425 strate the potential and effectiveness of the Rao’s Q index in producing ac-  
426 curately remote-sensing based maps of spatial diversity over such complex  
427 landscape. For this study, a cloud-free Sentinel-2A (S2A) image acquired  
428 on 8 of August 2016 was used to compute the NDVI at a 10 meters spatial  
429 resolution. The S2A image download, as well as the atmospheric correction  
430 (DOS method) were performed using the Semi-Automatic Classification plu-  
431 gin (SCP) implemented in the QGIS software (QGIS Development Team ,  
432 2016(@)).

433 The NDVI was used as input data for Rao's Q index computation using  
434 a window size of 3 x 3 pixels. The performance of the Rao's Q was compared  
435 to the Shannon Entropy index (Shannon's H), which is one of the simplest,  
436 and widely used, remote sensing-based diversity measures for landscape het-  
437 erogeneity assessment (Rocchini et al., 2016). To investigate whether both  
438 diversity indices differ between land cover types, one-way ANOVA tests were  
439 performed. This approach was used for analysing the degree of dissimilarity  
440 between Rao's Q and Shannon H index across two high complex land cover  
441 types; i) montado, and ii) polyculture. To do so, a sample of 60 squares with  
442 250 x 250 meters size was randomly selected over these two land cover types.  
443 Each square represents a sample of 625 S2A NDVI pixels, thus corresponding  
444 to a total of 37,500 pixels over the 60 squares. For the comparison between  
445 both indices, the coefficient of variation (CV) was calculated for each 250 x  
446 250 m squares. Regarding the Rao's Q performance, Figure 6 clearly points  
447 to the significant improvements shown by Rao's Q index compared to the  
448 Shannon H index in describing the spatial diversity. In particular, it can be  
449 seen through the Figure 6, that Rao's Q index can highlight different gra-  
450 dients of spatial diversity of montado areas, which present high tree density  
451 variability (Figure 6), and thus high spatial heterogeneity. One-way ANOVA  
452 tests revealed that both indices values were significantly different between  
453 the two land cover types (montado:  $F = 503.3$ ,  $p < 0.001$ ; polyculture:  $F =$   
454  $889.8$ ,  $p < 0.001$ ). Overall, the obtained results demonstrate the capability of  
455 Rao's Q index in producing accurate landscape diversity maps in a complex  
456 landscape such as the Mediterranean agro-forestry systems.

## 457 7 Conclusion

458 In this paper, we showed several methods based on ecological  $\beta$ -diversity,  
459 which can be investigated by remote sensing through the calculation of  
460 ecosystem heterogeneity, to estimate the spatial variability of biodiversity.  
461 When there is a wide range of heterogeneity, as an example the data include  
462 homogeneous and heterogeneous zones, no single measure might capture all  
463 the different aspects of  $\beta$ -diversity (e.g. (Baselga, 2013)). That is why we sug-  
464 gested in this manuscript multivariate and multidimensional methods (e.g.  
465 multivariate statistics and multidimensional distance matrices) based on the  
466 spectral signal and its variability over space to account for different aspects  
467 of diversity, also including distance- and abundance-based methods (e.g. the  
468 Rao's Q).

469 Biodiversity measured as species richness is often used for conservation  
470 purposes, hence the importance of avoiding an under- or over-estimate has

471 been highlighted (Chiarucci et al., 2009). Furthermore, pairwise distance-  
472 based methods might be profitably used to detect not only diversity hotspots  
473 in an area but also the variation of biodiversity over space, and potentially  
474 over time, once multitemporal sets of images are used.

475 In this paper we focused on optimising measures of  $\beta$ -diversity based on  
476 remote sensing data. Such measures might be used to regress species diversity  
477 against remotely sensed heterogeneity, based on new regression techniques  
478 which maximise the possibility of predicting the zones in a study area, or at  
479 larger spatial scales, of peculiar conservation value. As an example, shrink-  
480 age regression, recently applied in biodiversity conservation (Authier et al.,  
481 2017) could allow to directly focus on habitat modelling, which is one of the  
482 major strengths of remote sensing (Gillespie et al., 2008). Moreover, such  
483 analysis might be performed in a Bayesian framework allowing to i) model  
484 multidimensional covariates with non-stationary variation over space (Ran-  
485 dell et al., 2016), such as the bands of satellite images, and ii) model the  
486 errors in the output and their variation over space (Rocchini et al., 2017).

487 The suggested methods for  $\beta$ -diversity estimate from remote sensing are  
488 mainly based on distances, but they could be effectively translated to relative  
489 abundance-based methods. As an example Rocchini et al. (2013) introduced  
490 the possibility of applying generalized entropy theory to satellite images with  
491 one single formula representing a countinuum of diversity measures changing  
492 one parameter. One of the best examples in this framework could be the use  
493 of Hill numbers, in which diversity is expressed as:

$${}^qD = \left( \sum_{i=1}^S p_i^q \right)^{\frac{1}{1-q}} \quad (6)$$

494 where  $S$  = number of samples / pixels and  $p_i$  = relative abundance of a  
495 species / spectral value. varying the parameter  $q$ ,  ${}^qD$  varies accordingly in  
496 several diversity indices, e.g. for  $q = 0$   ${}^qD$  is the simple number of species,  
497 for  $\lim(q) = 1$   ${}^qD$  equals Shannon's entropy, etc. (Hsieh et al., 2016).

498 Furthermore, connectivity analysis might also be taken into account (Moila-  
499 nen et al., 2005, 2009). For instance, a remote sensing based connectivity  
500 network among different sites, based on  $\beta$ -diversity measures, could be ap-  
501 plied for the estimate of landscape connectivity and consequent genetic flow,  
502 as demonstrated by Vernesi et al. (2012). It has also been shown that commu-  
503 nity related biodiversity indicators are often missing from current monitoring  
504 programmes (Vihervaara et al., 2017); thus methods such as remote sensing  
505 based Rao's Q diversity applied for various ecosystems might improve other-  
506 wise challenging monitoring of biological communities.

507 With this manuscript we hope to stimulate discussion on the available  
508 methods for estimating  $\beta$ -diversity from remotely sensed imagery by propos-  
509 ing innovative techniques grounded on ecological theory.

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767 Figures

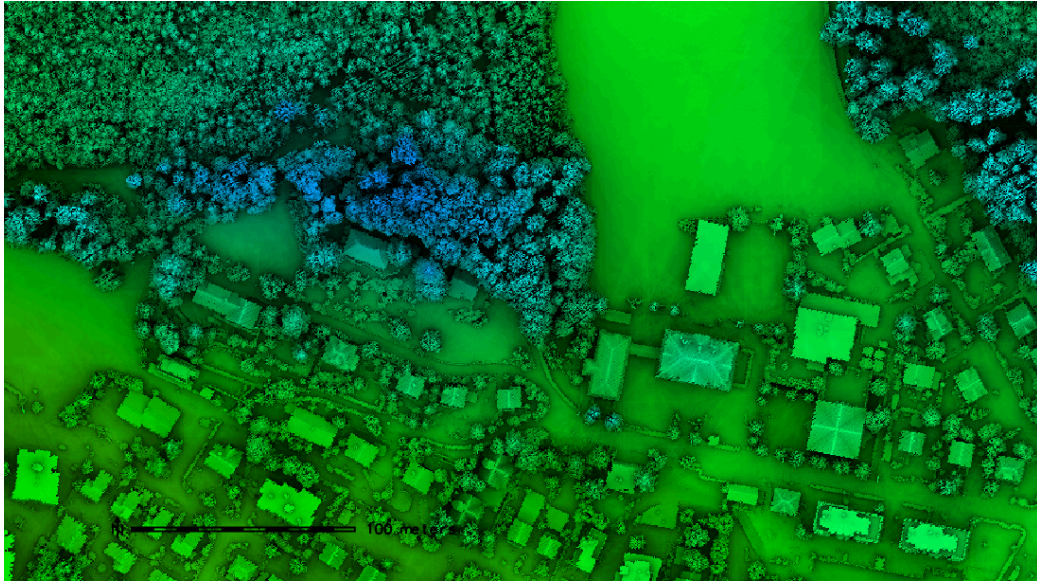


Figure 1: An example of how to couple information on compositional properties of the landscape by optical data together with structural (3D) properties by laser scanning LiDAR data.

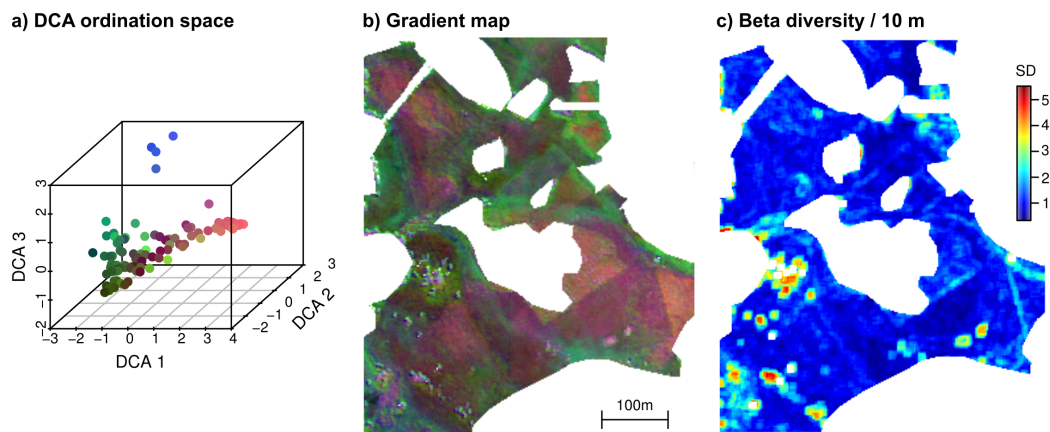


Figure 2: Beta diversity assessment with a combination of ordination techniques and remote sensing. a) Three dimensional DCA ordination space of  $n=100$  vegetation plots sampled in raised bogs, fens, transition mires and *Molinia* meadows in the alpine foothills of Southern Germany. An inter-plot distance of 4 SD corresponds to a full species turnover. b) Maps of the ordination axes resulting from a spatial prediction based on canopy reflectance. Each pixel has a predicted position in the ordination space that is indicated by its color. The color scheme corresponds to a). The map has a spatial resolution of 2 m x 2 m, which is in line with the sampled plot size. c) Cumulative change rates along the three DCA axes in a 5 x 5 pixel neighborhood. A high change rate indicates a high beta diversity.



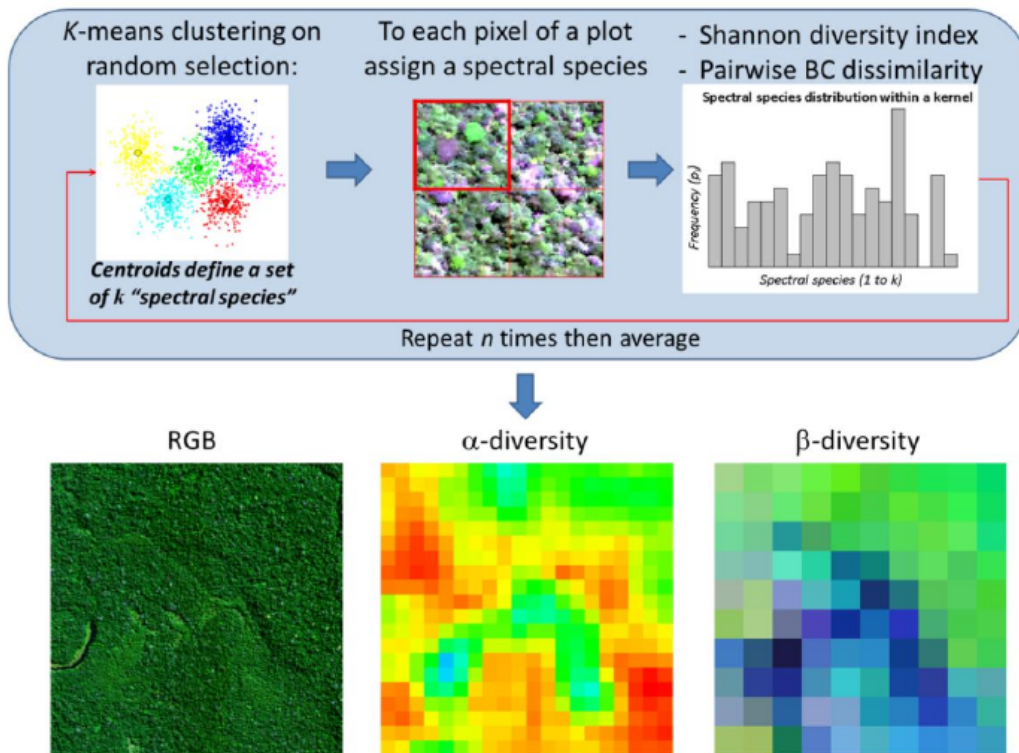


Figure 3: Spectral species can be identified in a hyper- or multi-spectral image by spatial clustering method and their distribution can be mapped. Such maps can further be used to apply local-based heterogeneity measurements ( $\alpha$ -diversity) as well as iterative distance based methods to build  $\beta$ -diversity maps. Reproduced from (Féret and Asner, 2014a).

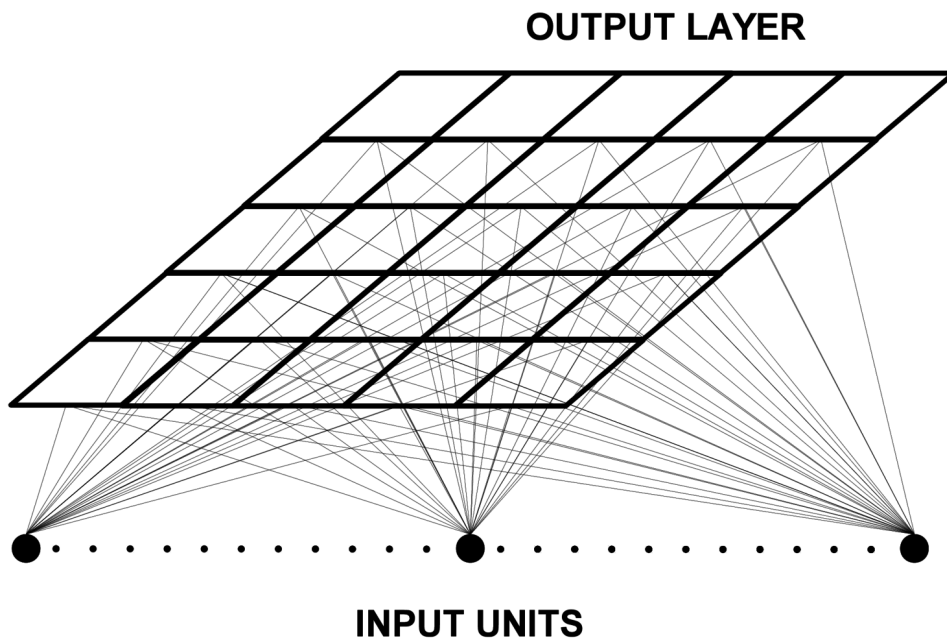


Figure 4: A self-organising feature map can be built starting from an input layer, e.g. the presence absence of a tree species or of a peculiar spectral value) which is connected to every unit in the output layer by a weighted connection. The self organising feature map uses unsupervised learning to map the location of field sites within the output space on the basis of their relative similarity in species or spectral composition. [Redrawn](#) from (Foody and Cutler, 2003).

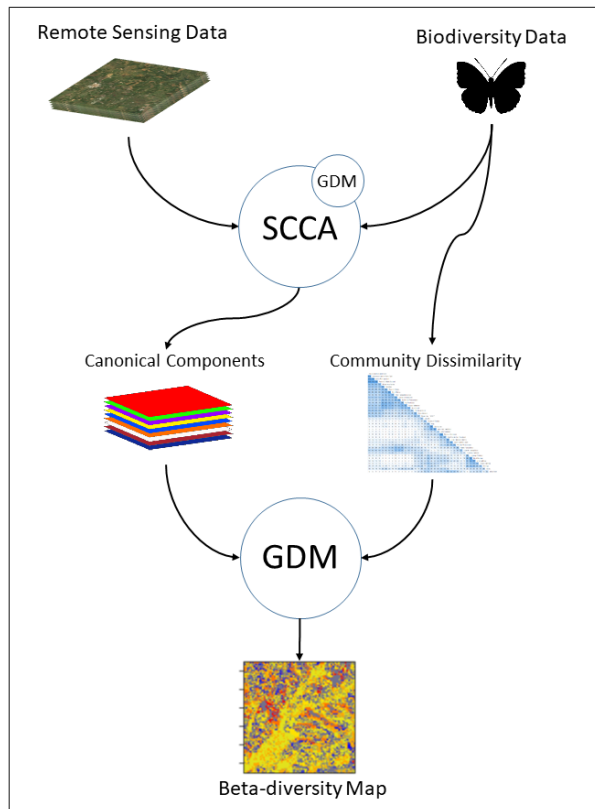


Figure 5: An example of the Sparse Generalized Dissimilarity Model (SGDM) approach. Remote sensing data and biodiversity data in the field can be coupled by Sparse Canonical Correlation Analysis to produce canonical components and a community dissimilarity matrix, which are then used to build a Generalized Dissimilarity Model to finally derived a  $\beta$ -diversity map.

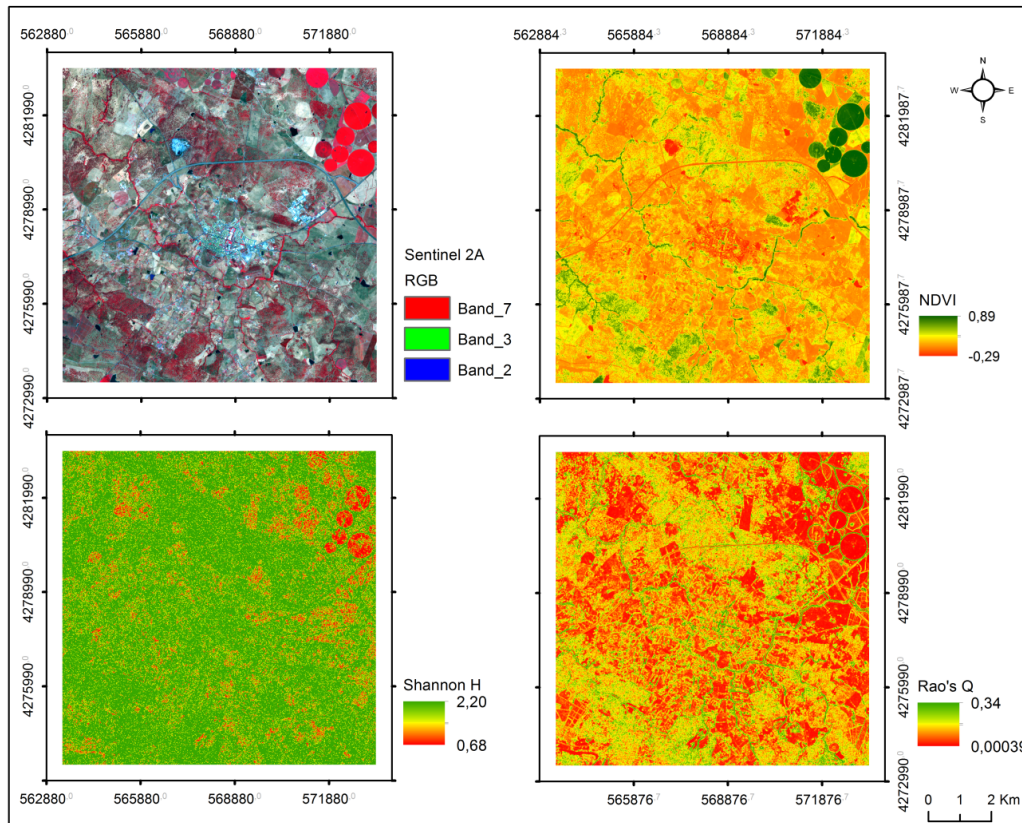


Figure 6: Upper panels: Sentinel-2A scene (8 August 2016) and derived NDVI for the agro-forestry systems test site located in southern Portugal. Lower panels: results from Shannon's H and Rao's Q indices computation. Shannon index tends to overestimate the landscape diversity when compared to the Rao's Q index.