

# Measurements of flavour-dependent fragmentation functions in $Z^0 \rightarrow q\bar{q}$ events

OPAL Collaboration

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**Abstract.** Fragmentation functions for charged particles in  $Z^0 \rightarrow q\bar{q}$  events have been measured for bottom (b), charm (c) and light (uds) quarks as well as for all flavours together. The results are based on data recorded between 1990 and 1995 using the OPAL detector at LEP. Event samples with different flavour compositions were formed using reconstructed  $D^{*\pm}$  mesons and secondary vertices. The  $\xi_p = \ln(1/x_p)$  distributions and the position of their maxima  $\xi_0$  are also presented separately for uds, c and b quark events. The fragmentation function for b quarks is significantly softer than for uds quarks.

## 1 Introduction

Experimental measurements of the inclusive momentum distribution of charged particles in  $e^+e^-$  collisions provide important insight into the process of how quarks turn into hadrons. This distribution is commonly normalised to the total hadronic cross-section  $\sigma^{\text{tot}}$  and presented as a function of the scaled momenta  $x_p = 2p_h/\sqrt{s}$  of the charged hadrons, where  $\sqrt{s}$  is the centre-of-mass energy. In this form, the spectrum is usually referred to as the fragmentation function and can be obtained experimentally from the total number of hadronic final states,  $N_{\text{event}}$ , and the number of charged particles in each  $x_p$  bin,  $N_{\text{track}}(x_p)$ :

$$F(x_p) = \frac{1}{\sigma^{\text{tot}}} \frac{d\sigma^h}{dx_p} = \frac{1}{N_{\text{event}}} \frac{N_{\text{track}}(x_p)}{\Delta x_p}. \quad (1)$$

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The charged particle momentum spectrum can also be studied as the distribution of  $\xi_p = \ln(1/x_p)$ . The  $\xi_p$  distribution emphasises the low momentum component and the  $x_p$  distribution the high momentum component of the momentum spectrum.

In the naive quark parton model, the scaled momentum distribution is expected to be independent of the centre-of-mass energy. A violation of this scaling is expected due to gluon radiation in the final state. Experimentally, scaling violation in fragmentation functions had indeed been observed by combining measurements at different centre-of-mass energies and could be used to determine  $\alpha_s$  [1]. The position of the maximum of the  $\xi_p$  distribution,  $\xi_0$ , has been studied in the past in various experiments (see for example [2] and references therein). The energy dependence of the position of the maximum provides an important test of the quantum chromodynamics (QCD) prediction for the emission of soft gluons [3].

In events with a heavy primary quark, the possibility of cascade decays of bottom or charm hadrons results in more particles sharing the same energy than in light quark events and a softer momentum spectrum can be

expected. Since the flavour composition of the primary quarks in  $e^+e^- \rightarrow q\bar{q}$  is predicted by the electroweak theory to change with centre-of-mass energy, this flavour dependence of the momentum spectra affects the energy dependence of the  $x_p$  and  $\xi_p$  distributions of the inclusive event sample. To correct for this contribution, in [4] not only the inclusive fragmentation function was measured but also fragmentation functions in event samples with different flavour compositions were studied. In [5], measurements of fragmentation functions in samples with different flavour composition were used to extract flavour-dependent fragmentation functions for events with primary light (uds), charm (c) or bottom (b) quarks.

Here we present a measurement of flavour-dependent fragmentation functions, based on the methods developed for the OPAL measurement of charged particle multiplicities in uds, c and b quark events [6,7]. Events were divided into two hemispheres by a plane perpendicular to the thrust axis. Secondary vertices and reconstructed  $D^{*\pm}$  mesons were used to tag hemispheres to create samples of events with different quark flavour mixtures (Sect. 2). To reduce the biases induced by the tagging, the measurement of the fragmentation functions was based on the momentum spectrum of charged particles in the event hemisphere opposite to the tag. Corrections for hemisphere correlations and for distortions due to detector effects are described in Sect. 3. The flavour-dependent  $x_p$  and  $\xi_p$  distributions were obtained from a simultaneous fit to the momentum spectra of the different hemisphere samples (Sect. 4). For the first time, the measured position of the maximum of the  $\xi_p$  distribution,  $\xi_0$ , is presented separately for uds, c and b events. A measurement of the inclusive distribution of all five flavours was also performed, based on the track momentum spectrum of all events, i.e. without considering any flavour tagging.

## 2 Selection and event tagging

A complete description of the OPAL detector can be found elsewhere [8]. This analysis relied on the precise reconstruction of charged particles in the central detector, consisting of a silicon microvertex detector, a vertex drift chamber, a large volume jet chamber and chambers measuring the  $z$ -coordinate<sup>1</sup> of tracks as they leave the jet chamber.

This analysis used data recorded with the OPAL detector in the years 1990 to 1995 at centre-of-mass energies around 91.2 GeV comprising an integrated luminosity of about 177 pb<sup>-1</sup>.  $Z^0$  decays were selected using the criteria described in [9]. To ensure that most charged particles were well contained in the detector, the polar angle of the thrust axis was required to satisfy  $|\cos\theta_{\text{thrust}}| < 0.8$ . To reduce systematic errors in the application of the sec-

<sup>1</sup> The OPAL coordinate system is defined with positive  $z$  along the electron beam direction and with positive  $x$  pointing towards the centre of the LEP ring. The polar angle  $\theta$  is defined relative to the  $+z$  axis and the azimuthal angle  $\phi$  relative to the  $+x$  axis

ondary vertex tag, it was only based on a homogeneous data sample taken in the year 1994, representing an integrated luminosity of about 34 pb<sup>-1</sup>. The full integrated luminosity was used in the case of the  $D^{*\pm}$  meson tag.

Charged tracks used in the measurement of the fragmentation function were required to have a measured momentum in the  $x$ - $y$  plane,  $p_t$ , of at least 0.150 GeV/ $c$  and to satisfy  $|d_0| < 0.5$  cm, where  $d_0$  is the distance of closest approach to the origin in the  $x$ - $y$  plane.

Simulated hadronic  $Z^0$  decays were generated with the Jetset 7.4 Monte Carlo program [10] tuned to OPAL data [11]. The events were passed through a detailed simulation of the OPAL detector [12] and processed using the same reconstruction and selection algorithms as the data.

### 2.1 Secondary vertex tag

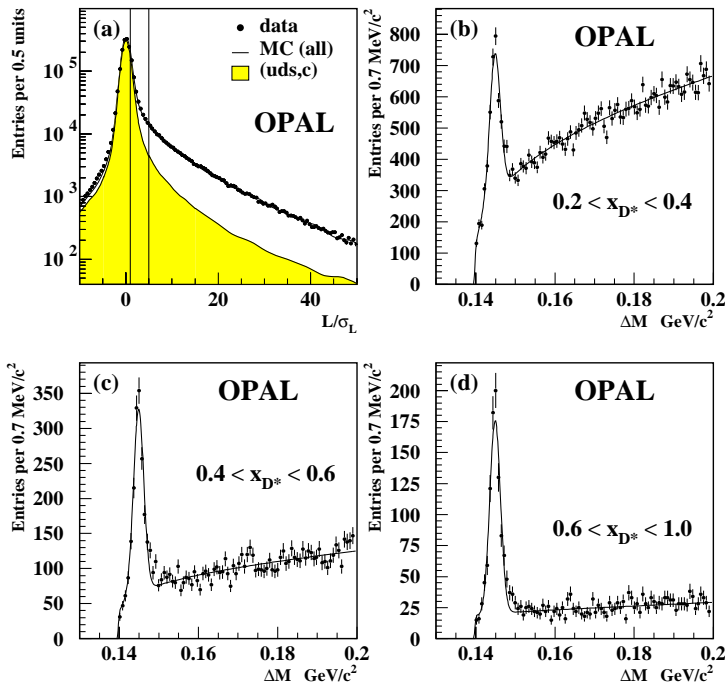
Samples with varying purity of b quark events were selected by reconstructing secondary vertices, following the procedure described in [6]. Events were divided into two hemispheres by the plane perpendicular to the thrust axis and comprising the interaction point. Jets were reconstructed by combining charged tracks and electromagnetic clusters not associated to tracks, using the scaled invariant mass algorithm described in [13] with the JADE-E0 recombination scheme and the invariant mass cut-off being set to 7 GeV/ $c^2$ . A vertex fit was then attempted in the highest energy jet in each hemisphere separately. Each track used in these vertex fits was required to have at least one hit in the silicon microvertex detector. All such tracks in the jet were fitted to a common vertex point in the  $x$ - $y$  plane and the track with the largest contribution to the  $\chi^2$  was removed if this contribution was greater than four. The remaining tracks were then refitted until either all tracks contributed less than four to the  $\chi^2$  or there were fewer than four remaining tracks. For each successfully reconstructed secondary vertex, the projected decay length  $L$  in the  $x$ - $y$  plane with respect to the primary vertex was calculated, where the primary vertex was reconstructed from all tracks in the event together with a constraint to the average beamspot position as in [14].

The decay length significance, i.e. the decay length divided by its uncertainty  $L/\sigma_L$  was used to obtain three event samples  $k$  of varying b flavour purity. According to the simulation, the b purities  $f_k^b$  in these samples vary from 11% to 90% (see Table 1). Figure 1(a) shows the distribution of the decay length significance in the data and in the Monte Carlo simulation.

### 2.2 $D^{*\pm}$ meson tag

Event samples with an enriched c quark contribution were obtained by reconstructing  $D^{*\pm}$  meson candidates.  $D^{*\pm}$  candidates were selected via the decay<sup>2</sup>  $D^{*+} \rightarrow K^- \pi^+ \pi^+$  closely following the procedure described in [7]:

<sup>2</sup> Throughout this paper, charge conjugate particles and decay modes are always implied



**Fig. 1.** **a** The decay length significance distribution in the data (symbols) and Monte Carlo (solid curve) simulations. The contribution from uds and from c quarks in the Monte Carlo distribution has been shaded. The boundaries of the three decay length significance bins used in this analysis:  $-10 < L/\sigma_L < 1$ ,  $1 < L/\sigma_L < 5$  and  $5 < L/\sigma_L < 50$  are indicated by vertical lines. **b** to **d** The distribution of the mass difference of the  $D^{*\pm}$  candidate and  $D^0$  candidate in the three different  $x_{D^*}$  bins. The symbols show the data while the solid lines are the results of the fits described in the text

**Table 1.** Number of tagged hemispheres in data and the flavour composition derived from the Monte Carlo simulation in three decay length significance regions

	$-10.0 < L/\sigma_L < 1.0$	$1.0 < L/\sigma_L < 5.0$	$5.0 < L/\sigma_L < 50.0$
Number of hemispheres	940 275	268 500	117 665
uds quark fraction $f_k^{\text{uds}}$	0.71	0.39	0.04
c quark fraction $f_k^c$	0.18	0.23	0.06
b quark fraction $f_k^b$	0.11	0.38	0.90

**Table 2.** Number of  $D^{*\pm}$  candidates, the fitted background fraction and the flavour composition of events with a genuine  $D^{*\pm}$  as taken from [15] in three  $x_{D^*}$  regions

	$0.2 < x_{D^*} < 0.4$	$0.4 < x_{D^*} < 0.6$	$0.6 < x_{D^*} < 1.0$
Number of $D^{*\pm}$ candidates	5109	1951	985
combinatorial background fraction $f_k^{\text{BG}}$	0.57	0.36	0.24
c quark fraction $\mathcal{P}_k^c$	$0.22 \pm 0.06$	$0.50 \pm 0.06$	$0.90 \pm 0.04$
b quark fraction $\mathcal{P}_k^b$	$0.78 \pm 0.06$	$0.50 \pm 0.06$	$0.10 \pm 0.04$

- A subset of tracks was selected with  $p_t > 0.250 \text{ GeV}/c$  and  $|d_0| < 0.5 \text{ cm}$ .
- Candidates of  $D^0 \rightarrow K^- \pi^+$  decays were selected by taking all combinations of two oppositely charged tracks, with one of them assumed to be a pion and the other assumed to be a kaon.  $D^{*\pm}$  candidates were selected by combining  $D^0$  candidates with a third track. This ‘slow pion’ track was required to have the same charge as the track presumed to be the pion in the  $D^0$  decay.
- The probability that the measured rate of energy loss,  $dE/dx$ , for the kaon candidate track was consistent with that expected for a real kaon was required to be greater than 10%.
- At least two of the three tracks were required to have either  $z$ -chamber hits or a polar angle measurement

derived from the point at which the track has left the jet chamber.

- The invariant mass of the  $D^0$  candidate was required to be between  $1.790 \text{ GeV}/c^2$  and  $1.940 \text{ GeV}/c^2$  and the mass difference between the  $D^0$  and the  $D^{*\pm}$  candidate,  $\Delta M$ , was required to be between  $0.142 \text{ GeV}/c^2$  and  $0.149 \text{ GeV}/c^2$ .
- Making use of the fact that real  $D^0$  mesons decay isotropically in their rest frames whereas the combinatorial background is peaked in the forward and the backward direction, the following cuts were applied:  $|\cos \theta^*| < 0.8$  for  $x_{D^*} < 0.5$  and  $|\cos \theta^*| < 0.9$  for  $x_{D^*} > 0.5$ , where  $\theta^*$  is the angle between the kaon in the  $D^0$  rest frame and the direction of the  $D^0$  in the laboratory frame and  $x_{D^*}$  is the scaled energy of the  $D^{*\pm}$ , i.e.  $x_{D^*} = 2E_{D^{*\pm}}/\sqrt{s}$ .

To provide samples with differing charm purity, the data were divided into three  $x_{D^*}$  regions. To evaluate their flavour composition,  $\Delta M$  distributions obtained without the cut on  $\Delta M$  were fitted with a Gaussian for the signal and a function  $A \exp(-B\Delta M)(\Delta M/m_\pi - 1)^C$  for the background [15]. The signals, together with the fitted functions are shown in Fig. 1(b)–(d). These fits were used to determine the fraction of background,  $f_k^{\text{BG}}$  in each  $D^{*\pm}$  sample in the signal  $\Delta M$  region. The results of these fits are summarised in Table 2.

The selected samples of  $D^{*\pm}$  candidates have three components: genuine  $D^{*\pm}$  mesons from b quark decays, genuine  $D^{*\pm}$  mesons from c quark decays and combinatorial background. No other sources of  $D^{*\pm}$  candidates were considered since Monte Carlo simulations predicts that only 0.3% of  $D^{*\pm}$  mesons with  $x_{D^*} > 0.2$  are produced via gluon splitting in light quark events [15]. To evaluate the effect of the contribution from fake  $D^{*\pm}$ , a side-band sample was selected by requiring that the two pions of the  $D^{*\pm}$  candidates had opposite charge and that  $0.150 \text{ GeV}/c^2 < \Delta M < 0.170 \text{ GeV}/c^2$ . Once this contribution was taken into account, the flavour composition of the  $D^{*\pm}$  samples was taken as the fractions  $\mathcal{P}_k^c$  and  $\mathcal{P}_k^b$  of genuine  $D^{*\pm}$  mesons originating from a primary c quark and b quark as measured in [15]. While in [15] the fractions  $\mathcal{P}_k^c$  and  $\mathcal{P}_k^b$  were derived for  $D^{*\pm}$  candidates after corrections for detector efficiency and acceptance were made, they were applied in this analysis to uncorrected data. No modifications were made since tests with Monte Carlo simulated events showed no significant flavour dependence of these corrections.

### 3 Corrections

The track momentum distributions in the hemispheres opposite to the secondary vertex or  $D^{*\pm}$  tag were measured. If secondary vertices or  $D^{*\pm}$  mesons were reconstructed in both hemispheres of an event then the momentum distributions opposite each of them were used. If, however, a secondary vertex and a  $D^{*\pm}$  were reconstructed in the same hemisphere, then only the  $D^{*\pm}$  information was used. Since there are many more secondary vertices reconstructed than  $D^{*\pm}$  mesons, this introduces negligible bias in the secondary vertex samples.

Six track momentum distributions (label  $k$ ) were obtained, corresponding to the three decay length regions and the three  $x_{D^*}$  regions. To obtain fragmentation functions from these distributions, three sets of corrections were applied. Firstly, a correction was made to take into account track momentum resolution and reconstruction efficiency. Secondly, the effects due to the event selection and the correlation between hemispheres were accounted for. In addition, the measured track momentum spectra in the  $D^{*\pm}$  tag samples were corrected for the contribution of fake  $D^{*\pm}$  mesons. The different corrections are described in the following.

After this procedure, the track momentum distributions are defined as the momentum distributions of all

promptly produced stable charged particles and those produced in the decays of particles with lifetimes shorter than  $3 \times 10^{-10}$  s, corrected for initial state radiation. This means that charged decay products from  $K_s^0$ , hyperons and weakly decaying b and c flavoured hadrons are included in the definition, regardless of how far away from the interaction point the decay actually occurred.

#### 3.1 Track momentum resolution and efficiency

The number  $N_{j,k}^{\text{observed}}$  of tracks in a tag sample was measured in 22 different  $x_p$  bins  $j$ .<sup>3</sup> The corrected distribution for a given tag sample is

$$N_{i,k}^{\text{corrected}} = \sum_j \sum_q \frac{\mathcal{M}_{ij}^q}{\epsilon_i^q} (w_{j,k}^q N_{j,k}^{\text{observed}}). \quad (2)$$

Here,  $\mathcal{M}_{ij}^q$  is the probability that a track measured in  $x_p$  bin  $j$  originates from a true  $x_p$  bin  $i$ . This correction was applied to account for the migration of the tracks between different bins due to the track momentum resolution. The reconstruction efficiency for tracks belonging to a true  $x_p$  bin  $i$  is accounted for by factors  $\epsilon_i^q$ . Differences in the slope of the  $x_p$  spectrum between uds, c and b quark events lead to flavour-dependent migration effects and to a flavour-dependent efficiency. Consequently, the matrices  $\mathcal{M}_{ij}^q$  and  $\epsilon_i^q$  are flavour dependent and have to be applied to the fraction  $w_{j,k}^q$  of observed tracks created in a  $q = \text{uds, c or b}$  event.

These weights  $w_{j,k}^q$  are the normalised products of the flavour-dependent fragmentation function  $F_j^q$  in an  $x_p$  bin  $j$  and the fraction  $f_k^q$  of events of a primary quark  $q$  in the considered tag sample:

$$w_{j,k}^q = \frac{f_k^q F_j^q}{\sum_{q'} f_k^{q'} F_j^{q'}}. \quad (3)$$

The applied weights and the obtained fragmentation functions are strongly correlated. This was taken into account in an iterative procedure, whereby the result of the measurement was used to re-calculate  $w_{j,k}^q$  and to repeat the correction procedure until the results were stable. Initial values for the weighting factors were taken from Monte Carlo simulations, but alternative values were also tried to confirm that the results did not depend on the choice of the exact initial values.

The values for  $\mathcal{M}_{ij}^q$  and  $\epsilon_i^q$  were obtained from Monte Carlo simulations. The diagonal elements  $\mathcal{M}_{ii}^q$  of the matrix, i.e. the probabilities that tracks are measured in their true  $x_p$  bins, are around 80% in most bins but become smaller than 50% for very high  $x_p$  values in c and b flavoured events. Values for the efficiency are typically

<sup>3</sup> For the measurement of the  $\xi_p$  distribution, a different binning with 29  $\xi_p$  bins was used. Apart from the binning, there were no differences between the analysis of the  $x_p$  and the  $\xi_p$  distribution, so the measurement of the  $\xi_p$  distribution is not explicitly described in the following sections

around  $\epsilon_i^q \approx 90\%$  with the exception of the lowest  $x_p$  bin where the efficiency is about 50%. The efficiency shows only a weak flavour dependence, the values differ for different flavours by less than 5%.

The corrected number of tracks  $N_{i,k}^{\text{corrected}}$  in each tag sample was divided by the corresponding number of tagged hemispheres  $N_k^{\text{hemi}}$  to form a fragmentation function for each tagged sample:

$$\mathcal{F}_{i,k} = \frac{N_{i,k}^{\text{corrected}}}{N_k^{\text{hemi}}}. \quad (4)$$

### 3.2 Flavour tagging and hemisphere correlations

About 80% of the  $D^{*\pm}$  mesons with  $x_{D^*} > 0.6$  are reconstructed in the jet with the highest energy. Likewise, secondary vertices with large values for the decay length significance  $L/\sigma_L$  are more likely to be found in high energy jets. The hemisphere containing the highest energy jet also tends to have a higher charged particle multiplicity and a harder track momentum spectrum than the opposite hemisphere. Indeed, the fragmentation function measured in the hemisphere opposite to a tag in the highest energy jet was seen to be 10% higher for low  $x_p$  values and up to 60% lower for high  $x_p$  values than that measured in the hemisphere opposite to a tag that is not in the highest energy jet. Consequently, the measured fragmentation functions in samples with high values for  $L/\sigma_L$  or  $x_{D^*}$  would be too soft since in an unbiased event sample only 50% of the analysed hemispheres would be opposite to the highest energy jet. To correct for this effect, the whole analysis was performed separately for the case where the tag-hemisphere contains the highest energy jet and where this is not the case. The unweighted average of the two results was taken at the end.

The dependence of the track momentum spectrum on the actual value of the decay length significance and on the  $D^{*\pm}$  energy was also considered. Requiring a high value for  $L/\sigma_L$  or  $x_{D^*}$  reduces the phase space for gluon bremsstrahlung and thus introduces a kinematic correlation between the hemispheres. The effect is flavour dependent, becoming more important for higher values of  $x_p$  and is more pronounced for the  $D^{*\pm}$  tag than for the secondary vertex tag. Besides this kinematical effect, correlations also occur due to geometrical effects: in a typical two-jet event, the jets are back to back, thus pointing into geometrically opposite parts of the detector. This introduces a hemisphere correlation if the detector response is not uniform. In addition to the kinematical and the geometrical correlations, the difference in the fragmentation functions in tagged events and in unselected events had to be taken into account.

All these effects were accounted for by applying correction factors for each tag sample, flavour and  $x_p$  bin:

$$\mathcal{T}_{i,k}^q = \frac{F_{i,k}^q(\text{generated})}{F_i^q(\text{generated})}, \quad (5)$$

i.e. the ratio of the generated fragmentation functions in tagged events and in events where the tag has not been

applied. These correction factors  $\mathcal{T}$  lead to a 10% correction for high  $x_p$  values opposite a hemisphere tagged by a secondary vertex and up to a 50% correction for high  $x_p$  values opposite a  $D^{*\pm}$  tagged hemisphere. Technically, these factors are applied as corrections to the purities in the fit procedure, thus taking into account the *a priori* unknown flavour composition of the tracks in a specific  $x_p$  bin.

### 3.3 Background subtraction in the $D^{*\pm}$ samples

The measured fragmentation functions in the  $D^{*\pm}$  signal samples  $\mathcal{F}_{i,k}^{\text{signal}}$  and the side-band samples  $\mathcal{F}_{i,k}^{\text{SB}}$  were used to determine the fragmentation functions for samples that contain genuine  $D^{*\pm}$  mesons:

$$\mathcal{F}_{i,k}^{D^{*\pm}} = \frac{1}{(1 - f_k^{\text{BG}})} \left( \mathcal{F}_{i,k}^{\text{signal}} - f_k^{\text{BG}} c_{i,k} \mathcal{F}_{i,k}^{\text{SB}} \right). \quad (6)$$

The background fractions  $f_k^{\text{BG}}$  derived from fits to the  $\Delta M$  distribution are listed in Table 2. To take into account differences of the hemisphere correlations for events in the signal and those in the side-band region, correction factors  $c_{i,k} = \mathcal{T}_{i,k}^{\text{signal}} / \mathcal{T}_{i,k}^{\text{SB}}$  were applied to the fragmentation functions of the side-band samples, i.e. they were multiplied by the ratio of the correction factors  $\mathcal{T}$  for the flavour mix of the signal and the side-band sample as predicted by the simulation.

## 4 Fits

A simultaneous fit was performed on the fragmentation functions  $\mathcal{F}_{i,k}^{D^{*\pm}}$  and  $\mathcal{F}_{i,k}^{\text{vtx}}$  of the three  $D^{*\pm}$  and the three secondary vertex tagged samples to extract the flavour-dependent fragmentation functions  $F^{\text{uds}}$ ,  $F^c$  and  $F^b$ . The fragmentation functions obtained from samples tagged by  $D^{*\pm}$  decays  $\mathcal{F}_{i,k}^{D^{*\pm}}$ , corrected for detector effects and for the combinatorial background for each of the three  $x_{D^*}$  regions  $k$  and each of the 22  $x_p$  bins  $i$  were described in the fit by

$$\mathcal{F}_{i,k}^{D^{*\pm}} = (\mathcal{P}_k^c \mathcal{T}_{i,k}^{\text{q=c}}) F_i^c + (\mathcal{P}_k^b \mathcal{T}_{i,k}^{\text{q=b}}) F_i^b, \quad (7)$$

where the purities  $\mathcal{P}_k^c$ ,  $\mathcal{P}_k^b$  are given in Table 2 and the correction factors  $\mathcal{T}_{i,k}^q$  are defined in (5). Naively, the fragmentation function corrected for detector effects for each of the three samples tagged by secondary vertices and each of the 22  $x_p$  bins could be described by

$$\mathcal{F}_{i,k}^{\text{vtx}} = (f_k^{\text{uds}} \mathcal{T}_{i,k}^{\text{q=uds}}) F_i^{\text{uds}} + (f_k^c \mathcal{T}_{i,k}^{\text{q=c}}) F_i^c + (f_k^b \mathcal{T}_{i,k}^{\text{q=b}}) F_i^b. \quad (8)$$

However, the fraction of c events  $f_k^c$  in all three vertex tagged samples is small (see Table 1), hence the relative uncertainty on these fractions large. If (8) were used, the vertex tagged samples would dominate the fit results due

**Table 3.** Fragmentation functions of uds, c and b events. The first error is statistical, the second systematic

$x_p$	$(1/\sigma_{\text{tot}})(d\sigma^h/dx_p)$					
	uds events		c events		b events	
0.00–0.01	388.0 ± 5.0	± 9.0	413.0 ± 19.0	± 18.0	416.0 ± 1.0	± 8.0
0.01–0.02	390.0 ± 5.0	± 10.0	381.0 ± 17.0	± 11.0	447.0 ± 1.0	± 8.0
0.02–0.03	241.0 ± 4.0	± 7.0	287.0 ± 13.0	± 8.0	300.0 ± 1.0	± 7.0
0.03–0.04	176.0 ± 3.0	± 5.0	178.0 ± 11.0	± 6.0	215.0 ± 1.0	± 5.0
0.04–0.05	122.6 ± 2.7	± 3.9	159.0 ± 10.0	± 5.0	160.7 ± 0.6	± 4.1
0.05–0.06	95.7 ± 2.2	± 2.9	116.0 ± 8.0	± 4.0	126.1 ± 0.5	± 3.4
0.06–0.07	79.3 ± 1.9	± 2.3	79.0 ± 7.0	± 3.0	101.4 ± 0.4	± 2.7
0.07–0.08	65.0 ± 1.6	± 1.7	61.0 ± 6.0	± 2.0	81.9 ± 0.4	± 2.2
0.08–0.09	53.3 ± 1.6	± 1.3	59.0 ± 6.0	± 2.0	68.9 ± 0.4	± 1.9
0.09–0.10	43.3 ± 1.5	± 1.0	53.0 ± 5.0	± 2.0	57.1 ± 0.3	± 1.6
0.10–0.12	35.1 ± 0.9	± 0.7	41.9 ± 3.2	± 1.5	44.0 ± 0.2	± 1.3
0.12–0.14	27.7 ± 0.7	± 0.4	27.6 ± 2.6	± 1.2	30.9 ± 0.2	± 1.0
0.14–0.16	21.2 ± 0.7	± 0.4	23.8 ± 2.4	± 1.0	22.5 ± 0.1	± 0.8
0.16–0.18	17.1 ± 0.6	± 0.3	17.6 ± 2.0	± 0.8	16.8 ± 0.1	± 0.6
0.18–0.20	13.3 ± 0.6	± 0.3	16.5 ± 1.9	± 0.7	12.3 ± 0.1	± 0.5
0.20–0.25	9.86 ± 0.26	± 0.30	9.8 ± 0.9	± 0.5	7.82 ± 0.05	± 0.40
0.25–0.30	6.30 ± 0.19	± 0.25	5.5 ± 0.7	± 0.3	4.16 ± 0.04	± 0.29
0.30–0.40	3.42 ± 0.09	± 0.17	2.49 ± 0.31	± 0.19	1.84 ± 0.02	± 0.18
0.40–0.50	1.50 ± 0.05	± 0.10	0.95 ± 0.19	± 0.26	0.65 ± 0.01	± 0.10
0.50–0.60	0.668 ± 0.033	± 0.048	0.36 ± 0.11	± 0.10	0.210 ± 0.006	± 0.052
0.60–0.80	0.241 ± 0.008	± 0.024	0.014 ± 0.041	± 0.015	0.038 ± 0.001	± 0.020
0.80–1.00	0.031 ± 0.007	± 0.007	0.003 ± 0.046	± 0.012	0.0040 ± 0.0005	± 0.0035

to their larger statistical weight as compared to the  $D^{*\pm}$  tagged samples. To ensure that the  $D^{*\pm}$  samples are used to obtain the charm fragmentation function, the flavour fractions  $f_k^{\text{uds}}$  and  $f_k^c$  were replaced by  $(1 - f_k^b)R_{\text{uds}}/(R_{\text{uds}} + R_c)$  and  $(1 - f_k^b)R_c/(R_{\text{uds}} + R_c)$  where  $R_{\text{uds}}$  and  $R_c$  are the standard model values for the branching fractions  $R_q = \Gamma(Z^0 \rightarrow q\bar{q})/\Gamma(Z^0 \rightarrow \text{hadrons})$ . The decay length dependence of the ratio of uds to c events was accounted for by correction factors  $d_{i,k}$  to the measured fragmentation function. The fragmentation function of the secondary vertex tagged samples was then described by

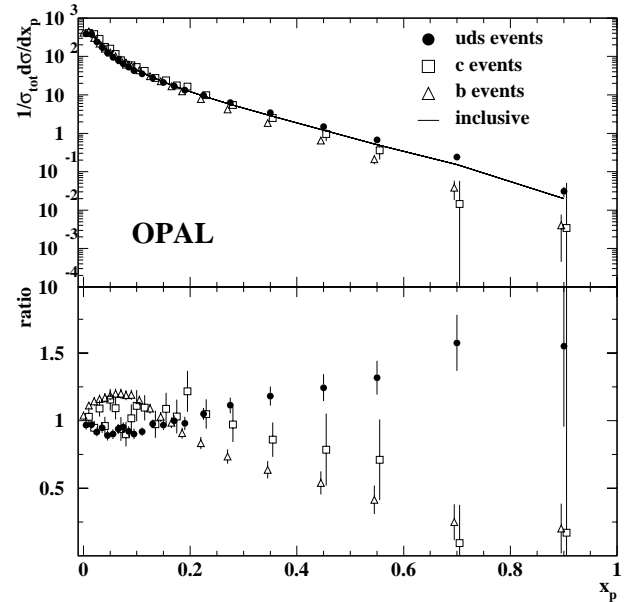
$$d_{i,k}\mathcal{F}_{i,k}^{\text{vtx}} = (1 - (f_k^b\mathcal{T}_{i,k}^{\text{q=b}}))F_i^{\text{udsc}} + (f_k^b\mathcal{T}_{i,k}^{\text{q=b}})F_i^b, \quad (9)$$

where

$$F_i^{\text{udsc}} = \frac{R_{\text{uds}}F_i^{\text{uds}} + R_cF_i^c}{R_{\text{uds}} + R_c}. \quad (10)$$

The correction factors  $d_{i,k}$  were derived from the momentum spectrum in Monte Carlo events with the ratio of uds to c events taken to be the same in all vertex samples, divided by the unmodified momentum spectrum with a variable ratio of uds to c events. These corrections are of the order of 1% for most of the  $x_p$  bins except for the highest  $x_p$  bins where they exceed 10%.

A simultaneous fit was performed to extract  $F_i^{\text{uds}}$ ,  $F_i^c$  and  $F_i^b$ . In fact, the secondary vertex data using (9) essentially fixes  $F_i^{\text{udsc}}$  and  $F_i^b$  and then the  $D^{*\pm}$  data provide  $F_i^c$  through (7), allowing (10) to give  $F_i^{\text{uds}}$ . The fit was based on the track momentum spectrum of the hemisphere opposite the tag. Therefore, the results had to be multiplied by a factor of two to obtain the full event fragmentation functions as they are shown in Fig. 2 and in



**Fig. 2.** The upper plot shows the measured fragmentation functions for uds events (filled symbols), c events (open squares) and b events (open triangles) as well as the inclusive fragmentation function (solid line). The lower plot shows the ratio of the flavour-dependent fragmentation functions to the inclusive fragmentation function. The error bars include statistical and systematic uncertainties. The systematic uncertainties are correlated between bins as well as between flavours



**Table 4.** Inclusive fragmentation function. The first error is statistical, the second systematic

$x_p$	$(1/\sigma_{\text{tot}})(d\sigma^h/dx_p)$	
0.00–0.01	401.2 ± 0.3	± 7.4
0.01–0.02	401.6 ± 0.2	± 4.9
0.02–0.03	262.8 ± 0.2	± 3.9
0.03–0.04	185.5 ± 0.2	± 2.9
0.04–0.05	137.5 ± 0.1	± 2.1
0.05–0.06	106.2 ± 0.1	± 1.6
0.06–0.07	84.5 ± 0.1	± 1.2
0.07–0.08	68.3 ± 0.1	± 0.9
0.08–0.09	57.9 ± 0.1	± 0.7
0.09–0.10	47.9 ± 0.1	± 0.6
0.10–0.12	38.19 ± 0.05	± 0.42
0.12–0.14	28.30 ± 0.04	± 0.28
0.14–0.16	21.88 ± 0.03	± 0.19
0.16–0.18	17.10 ± 0.03	± 0.16
0.18–0.20	13.54 ± 0.03	± 0.14
0.20–0.25	9.37 ± 0.01	± 0.11
0.25–0.30	5.66 ± 0.01	± 0.08
0.30–0.40	2.89 ± 0.01	± 0.05
0.40–0.50	1.208 ± 0.004	± 0.036
0.50–0.60	0.506 ± 0.002	± 0.018
0.60–0.80	0.153 ± 0.001	± 0.012
0.80–1.00	0.0199 ± 0.0003	± 0.0044

Table 3. The mean values of these distributions and their statistical uncertainties are:

$$\begin{aligned}\langle x_p \rangle^{\text{uds}} &= 0.0630 \pm 0.0003 \\ \langle x_p \rangle^{\text{c}} &= 0.0576 \pm 0.0012 \\ \langle x_p \rangle^{\text{b}} &= 0.0529 \pm 0.0001.\end{aligned}$$

The results can be compared with results for the inclusive fragmentation function which were obtained from the track momentum spectrum of all events without considering any flavour tagging. These results are shown in Table 4 and the mean value of the distribution was found to be:

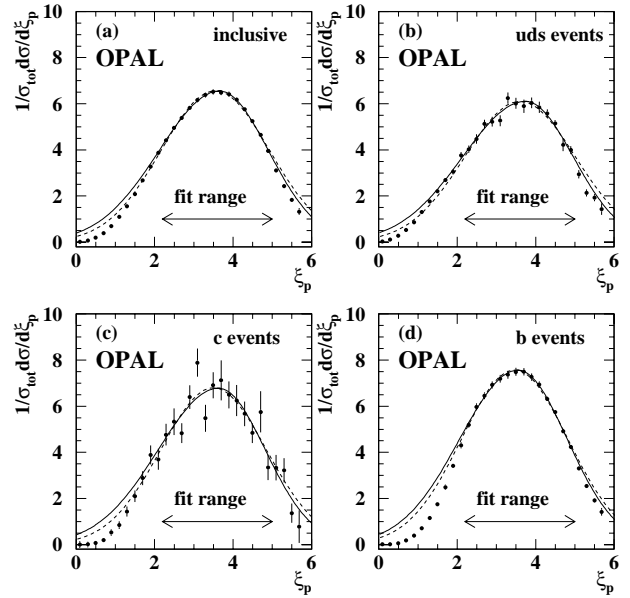
$$\langle x_p \rangle^{\text{incl}} = 0.05938 \pm 0.00002.$$

The results for the  $\xi_p$  distribution are shown in Fig. 3 and Table 5. To determine the positions of the maxima,  $\xi_0$ , skewed Gaussians, i.e. combinations of two Gaussians with different widths to the left and to the right from the centre were fitted to these distribution as motivated by the next-to-leading-log (NLLA) approximation [3]. Following the procedure in [17], the fit was performed in the region  $2.2 < \xi_p < 5.0$ . The results for the positions of the maxima with their statistical uncertainties are:

$$\begin{aligned}\xi_0^{\text{uds}} &= 3.74 \pm 0.06 \\ \xi_0^{\text{c}} &= 3.63 \pm 0.16 \\ \xi_0^{\text{b}} &= 3.55 \pm 0.01.\end{aligned}$$

Again, these flavour-dependent results can be compared with the results of the inclusive  $\xi_p$  distribution as obtained from all events without any flavour tagging. The results are shown in Table 6, the position of the maximum was determined to be:

$$\xi_0^{\text{incl}} = 3.656 \pm 0.003.$$

**Fig. 3.**  $\xi_p = \ln(1/x_p)$  distribution for **a** all events, **b** uds events, **c** c events and **d** b events. The solid lines show the results of the skewed Gaussian fitted to the distributions in the indicated fit range and the dashed lines show the results of a normal Gaussian fit. The error bars include statistical and systematic uncertainties

## 5 Systematic errors and cross-checks

The systematic uncertainties affecting the above results are due to the following sources: (i) uncertainties of the purities of the samples tagged by secondary vertices and (ii) by  $D^{*\pm}$  decays, (iii) the  $D^{*\pm}$  reconstruction, (iv) the hemisphere correlation, (v) uncertainties inherent to the correction procedure and (vi) the track and event selection. These were estimated from the difference between the central value and the result of the repeated analysis after a cut, a purity or the correction procedure was modified. In each case, the largest deviation was taken as the systematic error.

- (i) The uncertainties on the purities in the secondary vertex event samples were estimated using the published results from the measurement of the charged multiplicity in b, c and uds events [7]. There, variations of the measured multiplicity due to the uncertainties in the b lifetime, the fragmentation of b and c quark events, the production rates and the mixture of b hadrons produced as well as the decay multiplicities were studied and can be used to derive the uncertainties on the purities. The secondary vertex sample purities were then varied in the range of their uncertainty and the resulting differences of the results for the fragmentation function were taken as the systematic error due to this source.
- (ii) The purities  $\mathcal{P}_k^{\text{c}}$  and  $\mathcal{P}_k^{\text{b}}$  of the  $D^{*\pm}$  tag bins were taken from [15]. They were modified within their systematic errors to obtain the contribution to the systematic uncertainty on the fragmentation function.

**Table 5.**  $\xi_p = \ln(1/x_p)$  distribution of uds, c and b events. The first error is statistical, the second systematic

$\xi_p$	$(1/\sigma_{\text{tot}})(d\sigma^h/d\xi_p)$		
	uds events	c events	b events
0.0–0.2	0.024±0.006±0.006	0.002±0.147±0.005	0.0034±0.0001±0.0027
0.2–0.4	0.114±0.003±0.011	0.005±0.123±0.008	0.014 ± 0.001 ± 0.010
0.4–0.6	0.277±0.009±0.025	0.07 ± 0.03 ± 0.06	0.066 ± 0.002 ± 0.023
0.6–0.8	0.529±0.016±0.032	0.20 ± 0.06 ± 0.07	0.188 ± 0.004 ± 0.034
0.8–1.0	0.86 ± 0.02 ± 0.05	0.52 ± 0.08 ± 0.11	0.40 ± 0.01 ± 0.05
1.0–1.2	1.31 ± 0.03 ± 0.06	0.84 ± 0.12 ± 0.12	0.71 ± 0.01 ± 0.07
1.2–1.4	1.76 ± 0.05 ± 0.07	1.43 ± 0.16 ± 0.15	1.15 ± 0.01 ± 0.07
1.4–1.6	2.22 ± 0.06 ± 0.06	2.10 ± 0.20 ± 0.14	1.74 ± 0.01 ± 0.09
1.6–1.8	2.70 ± 0.07 ± 0.06	2.88 ± 0.24 ± 0.18	2.49 ± 0.02 ± 0.10
1.8–2.0	3.06 ± 0.08 ± 0.09	3.89 ± 0.29 ± 0.29	3.41 ± 0.02 ± 0.08
2.0–2.2	3.76 ± 0.09 ± 0.11	3.70 ± 0.31 ± 0.34	4.30 ± 0.02 ± 0.13
2.2–2.4	4.03 ± 0.10 ± 0.13	4.77 ± 0.35 ± 0.31	5.19 ± 0.02 ± 0.11
2.4–2.6	4.48 ± 0.10 ± 0.18	5.32 ± 0.37 ± 0.45	5.98 ± 0.02 ± 0.15
2.6–2.8	5.12 ± 0.11 ± 0.16	4.83 ± 0.37 ± 0.21	6.46 ± 0.03 ± 0.15
2.8–3.0	5.22 ± 0.12 ± 0.17	6.39 ± 0.41 ± 0.31	6.92 ± 0.03 ± 0.15
3.0–3.2	5.26 ± 0.13 ± 0.19	7.89 ± 0.48 ± 0.37	7.19 ± 0.03 ± 0.16
3.2–3.4	6.24 ± 0.12 ± 0.21	5.48 ± 0.43 ± 0.40	7.37 ± 0.03 ± 0.17
3.4–3.6	6.02 ± 0.12 ± 0.20	6.90 ± 0.44 ± 0.36	7.49 ± 0.03 ± 0.17
3.6–3.8	5.89 ± 0.13 ± 0.26	7.12 ± 0.45 ± 0.74	7.49 ± 0.03 ± 0.16
3.8–4.0	6.04 ± 0.12 ± 0.20	6.50 ± 0.43 ± 0.40	7.26 ± 0.03 ± 0.16
4.0–4.2	5.85 ± 0.13 ± 0.20	6.24 ± 0.46 ± 0.49	6.93 ± 0.03 ± 0.14
4.2–4.4	5.58 ± 0.11 ± 0.14	5.67 ± 0.40 ± 0.35	6.32 ± 0.03 ± 0.09
4.4–4.6	5.15 ± 0.11 ± 0.09	4.84 ± 0.40 ± 0.16	5.76 ± 0.03 ± 0.08
4.6–4.8	4.21 ± 0.12 ± 0.24	5.76 ± 0.41 ± 0.80	4.91 ± 0.02 ± 0.05
4.8–5.0	3.99 ± 0.10 ± 0.14	3.35 ± 0.36 ± 0.43	4.24 ± 0.02 ± 0.07
5.0–5.2	2.94 ± 0.10 ± 0.15	3.34 ± 0.33 ± 0.45	3.30 ± 0.02 ± 0.06
5.2–5.4	2.14 ± 0.10 ± 0.12	3.22 ± 0.34 ± 0.40	2.54 ± 0.02 ± 0.06
5.4–5.6	1.93 ± 0.08 ± 0.13	1.36 ± 0.28 ± 0.32	1.92 ± 0.02 ± 0.09
5.6–5.8	1.43 ± 0.09 ± 0.23	0.78 ± 0.32 ± 0.62	1.42 ± 0.03 ± 0.16

- (iii) To study the impact of the details of the  $D^{*\pm}$  candidate selection, the analysis was repeated with four sets of modified selection criteria. The  $M_{D_0^{\text{cand}}}$  mass window was increased to  $1.765 \text{ GeV}/c^2 < M_{D_0^{\text{cand}}} < 1.965 \text{ GeV}/c^2$ ; the  $\Delta M$  window was increased to  $0.141 \text{ GeV}/c^2 < \Delta M < 0.150 \text{ GeV}/c^2$ ; only one of the three tracks was required to have a  $z$ -chamber or jet chamber end-point  $z$  measurement and the cuts based on  $dE/dx$  were removed. The last modification lead to a reduction of the signal-to-noise ratio of more than 25%.
- (iv) The correlation between hemispheres caused by the kinematics of gluon radiation is corrected for by the factors  $\mathcal{T}(5)$ . A good description of the energy spectrum and the angular distribution of jets in the Monte Carlo simulation is important for a reliable prediction of this effect. To estimate the effects of small discrepancies between data and simulation in these distributions, the analysis was repeated applying weights to Monte Carlo events so that the energy distribution of the most energetic jet and the distribution of the angle between the two most energetic jets in data and Monte Carlo simulation agreed. Most of the resulting event weights had values between 0.95 and 1.05. The difference of the results with and without weighting was taken as the systematic uncertainty.
- (v) To estimate systematic uncertainties due to the correction for track momentum resolution and efficiency (Sect. 3.1) and corrections applied in the fitting procedure (Sect. 4), the following two modifications to the correction procedure were tested. First, the weighting factors  $w_{j,k}^q$  as defined in (3) were not re-calculated in an iterative procedure but were based on the initial values from the Monte Carlo simulation. Secondly, the correction factors  $d_{i,k}$  in (9) were omitted, i.e. the secondary vertex samples were not corrected for the variation of the uds to c flavour fraction.
- (vi) To account for imperfections in the tracking detector simulation, results were obtained in six different ways with modified event and track selections and variations of track quantities in the Monte Carlo simulation. The cut on the angle of the thrust axis  $|\cos\theta_{\text{thrust}}| > 0.8$  was removed; instead of accepting all tracks, a cut on  $|\cos\theta_{\text{track}}| < 0.7$  was applied; tracks were rejected if their  $z$ -coordinate at the point of closest approach to the event origin was larger than  $|z_0| > 10 \text{ cm}$ ; tracks were rejected if their momentum was smaller than  $0.250 \text{ GeV}/c$ ; the track momenta in simulated events were modified by an additional smearing fac-

**Table 6.** Inclusive  $\xi_p = \ln(1/x_p)$  distribution. The first error is statistical, the second systematic

$\xi_p$	$1/\sigma_{\text{tot}} \cdot d\sigma^h/d\xi_p$
0.0–0.2	$0.0153 \pm 0.0003 \pm 0.0035$
0.2–0.4	$0.071 \pm 0.001 \pm 0.006$
0.4–0.6	$0.191 \pm 0.001 \pm 0.008$
0.6–0.8	$0.392 \pm 0.001 \pm 0.011$
0.8–1.0	$0.693 \pm 0.002 \pm 0.018$
1.0–1.2	$1.087 \pm 0.002 \pm 0.021$
1.2–1.4	$1.556 \pm 0.003 \pm 0.025$
1.4–1.6	$2.082 \pm 0.003 \pm 0.028$
1.6–1.8	$2.674 \pm 0.004 \pm 0.028$
1.8–2.0	$3.272 \pm 0.004 \pm 0.032$
2.0–2.2	$3.866 \pm 0.005 \pm 0.036$
2.2–2.4	$4.403 \pm 0.005 \pm 0.044$
2.4–2.6	$4.96 \pm 0.01 \pm 0.06$
2.6–2.8	$5.39 \pm 0.01 \pm 0.07$
2.8–3.0	$5.82 \pm 0.01 \pm 0.08$
3.0–3.2	$6.15 \pm 0.01 \pm 0.10$
3.2–3.4	$6.38 \pm 0.01 \pm 0.11$
3.4–3.6	$6.52 \pm 0.01 \pm 0.12$
3.6–3.8	$6.48 \pm 0.01 \pm 0.10$
3.8–4.0	$6.41 \pm 0.01 \pm 0.12$
4.0–4.2	$6.17 \pm 0.01 \pm 0.09$
4.2–4.4	$5.75 \pm 0.01 \pm 0.06$
4.4–4.6	$5.24 \pm 0.01 \pm 0.04$
4.6–4.8	$4.65 \pm 0.01 \pm 0.04$
4.8–5.0	$3.95 \pm 0.01 \pm 0.05$
5.0–5.2	$3.110 \pm 0.005 \pm 0.051$
5.2–5.4	$2.423 \pm 0.004 \pm 0.055$
5.4–5.6	$1.835 \pm 0.004 \pm 0.081$
5.6–5.8	$1.33 \pm 0.01 \pm 0.15$

tor leading to a degradation of the momentum resolution in Monte Carlo events of 10%; the simulated track momenta were shifted by 1%. Since these effects are flavour independent, the uncertainty due to the track and event selection has been set to be the same for the inclusive and the flavour-dependent fragmentation functions.

The systematic uncertainties from the above groups of effects were added in quadrature and are shown for the flavour-dependent distributions in the last column of Tables 3 and 5. The result for the inclusive fragmentation function was obtained without flavour tagging and consequently does not depend on the tagging efficiency, purity or hemisphere correlation, so only the last two groups of effects contributed to the systematic uncertainty shown in the last column of Tables 4 and 6. For most of the  $x_p$  range, the relative systematic error is below 5% for uds and b events and below 10% for c events. For very high momenta ( $x_p > 0.5$ ), the systematic uncertainty becomes larger. Note that in uds and c events, the systematic and statistical errors are roughly equal. Detailed results are shown in Table 7 for a typical low momentum  $x_p$  bin ( $0.05 < x_p < 0.06$ ) and a typical high momentum  $x_p$  bin ( $0.3 < x_p < 0.4$ ). The systematic uncertainties are strongly correlated between adjacent bins and between different flavours. Since the results were obtained in a com-

bined fit, the variations of the results for uds and c events in the systematic studies were in general anti-correlated. Also, the systematic errors for uds and c events combined were in most cases anti-correlated with those obtained for b events.

In Table 8, the systematic uncertainties on the measurement of the mean value  $\langle x_p \rangle$  of the  $x_p$  distributions are shown. Further systematic checks were done for the determination of the position of the maximum of the  $\xi_p$  distributions,  $\xi_0$ . Instead of evaluating the position of the maximum using a skewed Gaussian fit, a normal Gaussian fit as motivated by LLA [18] was used. Furthermore, the fit range was modified and the skewed Gaussian was fitted to the measured  $\xi_p$  distribution in the regions  $2.0 < \xi_p < 5.2$  and  $2.4 < \xi_p < 4.8$ . The uncertainty obtained from this test has been added in quadrature to the previous six contributions as listed in Table 9.

To cross-check the results, the tagging methods were modified. An alternative b tag was applied, based on impact parameter information rather than on decay length information: the third largest impact parameter of a track was taken as the tag quantity. The impact parameter distribution of tracks from a decay are independent of the energy of the decaying particle while the decay length of a particle is proportional to its energy. Hence, this simple alternative method is less affected by kinematical correlations due to gluon radiation but was affected by other systematic effects which are not found in the standard method. The result obtained with the alternative b tagging method is consistent with the central values within the assigned systematic error. Another cross-check was performed using an alternative background treatment in the  $D^{*\pm}$  tagged event samples. Instead of taking the  $D^{*\pm}$  purities from [15] and subtracting the effect of the combinatorial background with the help of a fit to the  $D^{*\pm}$  signal, purities and background were taken from Monte Carlo. Again, the results obtained with this alternative method and the central values agreed within the systematic errors.

The model dependence of the matrix  $\mathcal{M}_{ij}^q$  was tested by repeating the calculation of the matrix in each new iteration, taking into account the difference between the measured and the generated track momentum spectra. The difference between the results obtained with the matrix  $\mathcal{M}_{ij}^q$  derived from this iterative procedure and the central values was found to be small for most of the  $x_p$  range. Only in the highest  $x_p$  bin, a deviation of around 15% has been observed. However, this effect can be absorbed partially by the statistical uncertainty and the other systematic errors, and the remaining part is small compared to the total error.

Effects from the binning in the tag variable have also been cross-checked. Instead of using three secondary vertex bins and three  $D^{*\pm}$  bins, alternative results were obtained using two and four secondary vertex bins and likewise two and four  $D^{*\pm}$  bins. The deviations from the central value were in all cases smaller than the estimated systematic error.

**Table 7.** Relative systematic and statistical uncertainties in percent on the results for  $0.05 < x_p < 0.06$  ( $0.3 < x_p < 0.4$ )

Effect	Inclusive		uds		c		b	
(i) Purities of sec. vertex samples	—	(—)	2.2	(4.0)	0.9	(1.5)	1.4	(6.5)
(ii) Purities of $D^{*\pm}$ samples	—	(—)	0.2	(0.3)	0.5	(1.0)	< 0.1	(< 0.1)
(iii) $D^{*\pm}$ selection	—	(—)	1.2	(1.8)	2.6	(4.2)	< 0.1	(0.1)
(iv) Hemisphere correlation	—	(—)	0.2	(0.9)	0.4	(0.6)	0.3	(0.1)
(v) Correction procedure	<0.1	(<0.1)	0.8	(1.7)	1.0	(3.3)	1.7	(7.6)
(vi) Track and event selection	1.5	(1.9)	1.5	(1.9)	1.5	(1.9)	1.5	(1.9)
Total systematic uncertainty	1.5	(1.9)	3.0	(5.2)	3.4	(6.0)	2.6	(10.2)
Total statistical uncertainty	0.1	(0.2)	2.3	(2.6)	6.7	(12.5)	0.4	(1.0)

**Table 8.** Relative systematic and statistical uncertainties in percent on the results for the mean value of the  $x_p$  distribution ( $x_p$ )

Effect	inclusive	uds	c	b
(i) Purities of sec. vertex samples	—	1.4	0.8	1.3
(ii) Purities of $D^{*\pm}$ samples	—	0.1	0.3	< 0.1
(iii) $D^{*\pm}$ selection	—	0.6	2.1	< 0.1
(iv) Hemisphere correlation	-	< 0.1	0.1	0.2
(v) Correction procedure	0.04	0.1	1.3	1.7
(vi) Track and event selection	0.96	1.0	1.0	1.0
Total systematic uncertainty	0.96	1.8	2.8	2.4
Total statistical uncertainty	0.03	0.6	2.0	0.1

The whole analysis procedure has also been tested globally with simulated events. It was shown that the generated fragmentation function and the result of the unfolding procedure agree within the statistical uncertainty.

By integrating the fragmentation functions, the charged multiplicity in uds, c and b events can be obtained. The results,

$$\begin{aligned}
n^{\text{uds}} &= 20.25 \pm 0.11 \pm 0.37 \\
n^c &= 21.55 \pm 0.37 \pm 0.64 \\
n^b &= 23.16 \pm 0.02 \pm 0.45 \\
n^{\text{incl}} &= 21.16 \pm 0.01 \pm 0.21,
\end{aligned}$$

are in good agreement with results of direct measurements of the charged multiplicities [6, 7, 23, 24].

The average of the three flavour-dependent fragmentation functions can be formed and weighted with the standard model branching fractions  $R_{\text{uds}}$ ,  $R_c$  and  $R_b$ . This combined fragmentation function can be compared with the results for the inclusive fragmentation function and with previously published OPAL results [23]. All three results show good agreement with each other.

## 6 Results

The results for the flavour-dependent fragmentation functions for uds, c and b events as well as the inclusive fragmentation function are shown in Fig. 2 and in Tables 3 and 4. The mean values of these distributions are:

**Table 9.** Relative systematic and statistical uncertainties in percent on the results for the position of the maximum of the  $\xi_p$  distribution,  $\xi_0$ 

Effect	inclusive	uds	c	b
(i) Purities of sec. vertex samples	—	0.6	0.1	0.2
(ii) Purities of $D^{*\pm}$ samples	—	< 0.1	0.2	< 0.1
(iii) $D^{*\pm}$ selection	—	0.5	2.0	< 0.1
(iv) Hemisphere correlation	—	0.4	0.2	< 0.1
(v) Correction procedure	0.9	0.6	0.3	0.5
(vi) Track and event selection	0.6	0.6	0.6	0.6
(vii) Fit range and fit type	3.0	5.3	8.4	1.9
Total systematic uncertainty	3.1	5.4	8.7	2.1
Total statistical uncertainty	0.1	1.5	4.1	0.3

$$\begin{aligned}
\langle x_p \rangle^{\text{uds}} &= 0.0630 \pm 0.0003 \pm 0.0011 \\
\langle x_p \rangle^c &= 0.0576 \pm 0.0012 \pm 0.0016 \\
\langle x_p \rangle^b &= 0.0529 \pm 0.0001 \pm 0.0013 \\
\langle x_p \rangle^{\text{incl}} &= 0.05938 \pm 0.00002 \pm 0.00057.
\end{aligned}$$

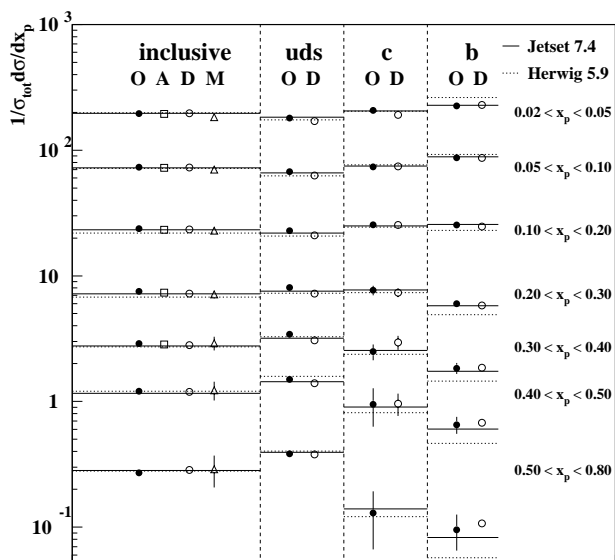
The light quark fragmentation function is found to be harder than the b quark fragmentation function as expected due to the cascade decays of b hadrons in b quark events with more particles sharing the energy. This observation is also consistent with the results of comparisons of gluon, uds and b jets [21, 22].

In Fig. 3 and in Tables 5 and 6, the results are presented for the  $\xi_p = \ln(1/x_p)$  distribution which emphasises the lower momenta of the spectrum. Skewed Gaussians were fitted to these distributions to obtain the position of their maxima:

$$\begin{aligned}
\xi_0^{\text{uds}} &= 3.74 \pm 0.06 \pm 0.21 \\
\xi_0^c &= 3.63 \pm 0.16 \pm 0.31 \\
\xi_0^b &= 3.55 \pm 0.01 \pm 0.07 \\
\xi_0^{\text{incl}} &= 3.656 \pm 0.003 \pm 0.115.
\end{aligned}$$

The result for the inclusive distribution is in good agreement with previous results [19, 20], whereas the position of the maxima of the flavour-dependent  $\xi_p$  distribution is reported here for the first time. Part of the systematic uncertainties cancel when the ratio of the flavour-dependent results to the inclusive result is taken:

$$\begin{aligned}
\xi_0^{\text{uds}}/\xi_0^{\text{incl}} &= 1.023 \pm 0.017 \pm 0.028 \\
\xi_0^c/\xi_0^{\text{incl}} &= 0.993 \pm 0.044 \pm 0.082 \\
\xi_0^b/\xi_0^{\text{incl}} &= 0.971 \pm 0.003 \pm 0.022.
\end{aligned}$$



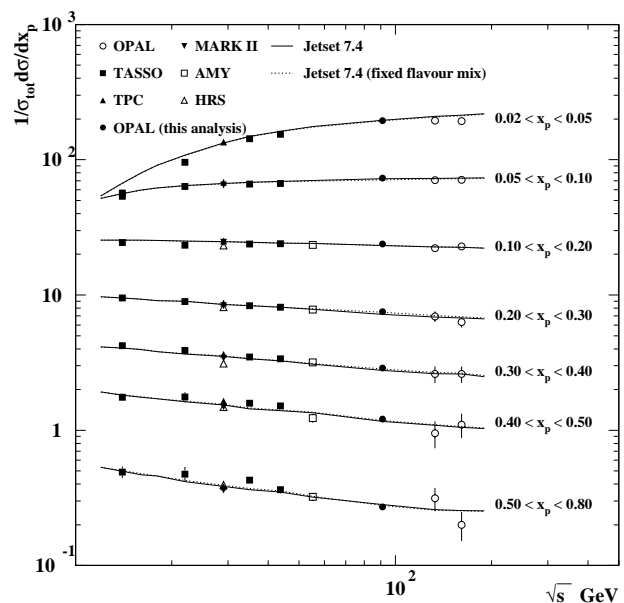
**Fig. 4.** Comparison of the results for the inclusive fragmentation function for this analysis (O) with results from ALEPH (A), DELPHI (D) and MARK II (M) at  $\sqrt{s} = m_{Z^0}$  [20,27] and of the flavour-dependent fragmentation function with the results from DELPHI [5]. The error bars include statistical and systematic uncertainties. The Jetset 7.4 predictions for the fragmentation function are shown as full horizontal lines and the Herwig 5.9 predictions as dotted horizontal lines

Another indication for a flavour dependence of the  $\xi_p$  distribution is given by the differences of the shape of the distributions in Fig. 3.

In Fig. 4, the total and the flavour-dependent fragmentation functions are compared with results from other experiments at the same centre-of-mass energy. There is good agreement with [5] where the  $D^{*+} \rightarrow K^- \pi^+ \pi^+$  decay was used as well, but a different b tagging method and a different correction procedure was applied. Also the quoted systematic uncertainty is similar in size with the exception of the high momentum region, where the uncertainty in [5] is smaller than in this paper. However, a direct comparison of the systematic errors is difficult since the error sources dominating the systematic uncertainty in the high momentum region in this paper (Table 7) are not explicitly considered in [5]. Also in Fig. 4, the results are shown to be consistent with the Jetset 7.4 expectation while the Herwig 5.9 Monte Carlo program [25]<sup>4</sup> fails to describe fully the b fragmentation function.

The measurement of the total fragmentation function in comparison to measurements at lower energies [26] and at centre-of-mass energies between 130 GeV and 161 GeV [2,17] is shown in Fig. 5. Apart from the lowest  $x_p$  region, there is good agreement with the Jetset 7.4 prediction (solid line) despite the fact that the parameters used

<sup>4</sup> The parameter set used was the same as in [21] for Herwig 5.8, except for the value of the cluster mass cutoff CLMAX which has been increased from 3.40 GeV/ $c^2$  to 3.75 GeV/ $c^2$ . Alternatively, studies were done with the default parameter set of the Herwig Monte Carlo program, but in this case, even the inclusive fragmentation function fails to describe the data



**Fig. 5.** Comparison of the results for the inclusive fragmentation function with results at different lower [26] and higher centre-of-mass energies [2,17]. The error bars include statistical and systematic uncertainties. The solid lines show the Jetset 7.4 prediction, assuming the centre-of-mass energy dependence of the flavour composition as predicted by the electroweak theory. The dotted lines show the Jetset 7.4 prediction assuming for all energies the same flavour mix as at  $\sqrt{s} = 18$  GeV. The dotted line is almost entirely hidden behind the full line and even at  $\sqrt{s} = 91.2$  GeV, only a negligible difference between the two curves can be seen because the effect of an increased b contribution is compensated largely by the effect of a decreased c contribution

in Jetset 7.4 were optimised to describe data at the  $Z^0$  resonance. In the highest  $x_p$  bin, the difference between the total fragmentation function measured at  $\sqrt{s} = 14.0$  GeV and at  $\sqrt{s} = 161$  GeV is of the same order of magnitude as the difference between uds and b fragmentation functions at  $\sqrt{s} = 91.2$  GeV. Observing the good agreement between data and the Jetset 7.4 prediction of the  $\sqrt{s}$  dependence of the total fragmentation function and of the flavour-dependent fragmentation functions at  $\sqrt{s} = 91.2$  GeV, we can use Jetset to estimate the effect of the change of the flavour mix to the apparent scaling violation. In Fig. 5, the Jetset 7.4 prediction is shown when the flavour mix at all centre-of-mass energies was forced to be the same as the flavour mix at  $\sqrt{s} = 18$  GeV. Although this fixed flavour mix is very different from that at the  $Z^0$  peak and the flavour-dependent fragmentation functions differ significantly, the changes on the total fragmentation function at  $\sqrt{s} = 91.2$  GeV are of the order of only two to four percent. This is due to a cancellation of the effect of an increased b contribution and a decreased c contribution at centre-of-mass energies close to the  $Z^0$  resonance.

Comparing the positions of the maxima,  $\xi_0$ , of the  $\xi_p$  distribution, the flavour dependence is less pronounced than for the fragmentation function at high  $x_p$ . The values for  $\xi_0$  at  $\sqrt{s} = 14.0$  GeV and  $\sqrt{s} = 161$  GeV differ by almost a factor of two, while the difference between the

flavour-dependent results at  $\sqrt{s} = 91.2$  GeV is one order of magnitude smaller. Like in the case of the fragmentation functions, an increased b contribution and a simultaneously decreased c contribution leads to cancellations when comparing low energy measurements and results at the  $Z^0$  peak. This is expected on the basis of simulations where the  $\xi_0$  obtained for the flavour mixture at  $\sqrt{s} = 91.2$  GeV and at  $\sqrt{s} = 18.0$  GeV differ by less than one percent.

## 7 Conclusions

Flavour-dependent fragmentation functions in  $Z^0 \rightarrow q\bar{q}$  events have been measured separately for bottom, charm and light (uds) quarks and as well as for all flavours together. These measurements are based on OPAL data recorded between 1990 and 1995. Event samples with different flavour compositions were formed using reconstructed  $D^{*\pm}$  mesons and secondary vertices in jets. The charged particle momentum spectrum has been studied in the event hemisphere opposite to the tag. A simultaneous fit was performed to extract the flavour-dependent  $x_p$  distribution as well as the flavour-dependent  $\xi_p$  distribution.

The fragmentation function for b quarks is significantly softer than for uds quarks. The fragmentation functions are well described by the Jetset 7.4 Monte Carlo program while Herwig 5.9 fails to describe fully the b fragmentation function.

For the first time, flavour-dependent  $\xi_p$  distributions have been studied. The flavour dependence of the position of the maximum has been determined and was found to be small compared with the differences of this value at different centre-of-mass energies.

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