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Evaluation of the causal effects between subjective wellbeing and cardiometabolic health: mendelian randomisation study

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ABSTRACT

OBJECTIVE

To investigate whether the association between subjective wellbeing (subjective happiness and life satisfaction) and cardiometabolic health is causal.

DESIGN

Two sample, bidirectional mendelian randomisation study.

SETTING

Genetic data taken from various cohorts comprised of the general population (mostly individuals of European ancestry, plus a small proportion of other ancestries); follow-up analysis included individuals from the United Kingdom.

PARTICIPANTS

Summary data were used from previous genome wide association studies (number of participants ranging from 83 198 to 339 224), which investigated traits related to cardiovascular or metabolic health, had the largest sample sizes, and consisted of the most similar populations while minimising sample overlap. A follow-up analysis included 337 112 individuals from the UK Biobank (54% female (n=181 363), mean age 56.87 years (standard deviation 8.00) at recruitment).

MAIN OUTCOME MEASURES

Subjective wellbeing and 11 measures of cardiometabolic health (coronary artery disease;

WHAT IS ALREADY KNOWN ON THIS TOPIC

Higher subjective wellbeing is associated with better physical health outcomes, with observational evidence highlighting effects on body mass index and cardiovascular health

Use of genetic data in a mendelian randomisation framework avoids reverse causation and residual confounding (problems associated with observational evidence), hence allowing stronger causal inference

A previous mendelian randomisation analysis has shown evidence of an effect of body mass index on subjective wellbeing, but did not adjust for sample overlap and did not look at other cardiometabolic traits

WHAT THIS STUDY ADDS

Mendelian randomisation showed strong evidence of a causal effect of higher body mass index on lower subjective wellbeing (β =-0.045 (95% confidence interval -0.084 to -0.006), P=0.02), driven by lower satisfaction with health (β =-0.035 (-0.043 to -0.027), P<0.001)

This study adds further support to the need to prevent obesity because of the downstream consequences on mental health as well as physical health

No clear evidence of a causal association between subjective wellbeing and any other measure of cardiometabolic health was seen in either direction, suggesting that previous associations might be the result of shared confounding by lifestyle factors myocardial infarction; total, high density lipoprotein, and low density lipoprotein cholesterol; diastolic and systolic blood pressure; body fat; waist to hip ratio; waist circumference; and body mass index).

RESULTS

Evidence of a causal effect of body mass index on subjective wellbeing was seen; each 1 kg/m² increase in body mass index caused a -0.045 (95% confidence interval -0.084 to -0.006, P=0.02) standard deviation reduction in subjective wellbeing. Follow-up analysis of this association in an independent sample from the UK Biobank provided strong evidence of an effect of body mass index on satisfaction with health (β =-0.035 unit decrease in health satisfaction (95% confidence interval -0.043 to -0.027) per standard deviation increase in body mass index, P<0.001). No clear evidence of a causal effect was seen between subjective wellbeing and the other cardiometabolic health measures, in either direction.

CONCLUSIONS

These results suggest that a higher body mass index is associated with a lower subjective wellbeing. A follow-up analysis confirmed this finding, suggesting that the effect in middle aged people could be driven by satisfaction with health. Body mass index is a modifiable determinant, and therefore, this study provides further motivation to tackle the obesity epidemic because of the knock-on effects of higher body mass index on subjective wellbeing.

Introduction

Subjective wellbeing is most commonly defined as a combination of life satisfaction and happiness (having high positive affect in the absence of negative affect).¹ Life satisfaction and happiness capture both the cognitive and affective components of subjective wellbeing, respectively. The importance of wellbeing is emphasised by the World Health Organization in their definition of health,² and observational evidence suggests an association between higher subjective wellbeing and better physical health or longevity,^{3 4 5} especially cardiovascular and metabolic health outcomes including cardiovascular disease,⁶ cholesterol levels,⁷ and extremes of body mass index.⁸

More frequently studied, depression has been shown to have the opposite association with cardiometabolic health, increasing the risk of coronary artery disease (especially the chance of a heart attack),⁹ altering serum cholesterol,¹⁰ and having a U shaped relation with body mass index.¹¹ A mendelian randomisation study of body mass index on multiple mental health outcomes found a consistent effect of higher body mass index on increased likelihood of depression, although the effect sizes were small.¹² This causal effect was replicated in the follow-up analysis of the most recent genome wide association study (GWAS) of depression,¹³ and was replicated using a continuous measure of depressive symptoms.¹⁴ The study also showed suggestive evidence of a causal link between body mass index and subjective wellbeing using a two sample approach and a relaxed instrument threshold. However, this study did not examine cardiovascular health or other measures of adiposity and did not adjust for sample overlap, so results could be biased towards the confounded observational effect.¹⁵

Twin analyses suggest partly distinct genetic (and environmental) causes for depression and subjective wellbeing,¹⁶ which indicates that separate analyses of associations between subjective wellbeing and depression on health outcomes might be appropriate. Observational research suggests that the association between subjective wellbeing and cardiometabolic health goes beyond the absence of negative affect states, reduced arousal, or confounding from socioeconomic position,¹⁷ and subjective wellbeing is more predictive of health outcomes than negative feelings.¹⁸ Therefore, subjective wellbeing might be a more effective target for improving cardiometabolic health outcomes than depression. From a public health perspective, it is important to understand whether increasing subjective wellbeing can increase health in later life, given that wellbeing interventions can be cost effective to administer.¹⁹

Studies suggesting a link between subjective wellbeing and cardiometabolic health are mostly observational. Owing to reverse causation and residual confounding, causal inferences are difficult to make from observational evidence.20 Mendelian randomisation uses genetic variants as instrumental variables for the exposure of interest. The inheritance of specific alleles is largely independent of genetic variants affecting other traits and of conventional disease risk factors, and associations are less prone to other biases inducing reverse causation, because genotype is unchanged over the lifetime.^{20 21} In an instrumental variable analysis, the genetic variant (Z) acts as the instrument that is related to differences in the exposure (X). If the exposure is on the causal pathway with the outcome (Y), then genetic variants that affect the exposure should be associated with the outcome (fig 1).²⁰ For example, genetic variants

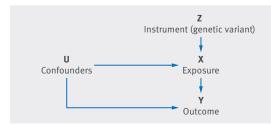


Fig 1 | Directed acyclic graph representing the basic instrumental variable analysis in mendelian randomisation

(*Z*) shown to predispose individuals to have a higher body mass index (exposure X) are associated with lower income, suggesting that increases in body mass index reduce income (outcome Y).²² Mendelian randomisation is one method available to triangulate evidence about whether a particular intervention is clinically viable.²³

The two sample mendelian randomisation method uses summary statistics from each GWAS in one analysis.²⁴ Single nucleotide polymorphisms (SNPs) reliably associated with an exposure of interest can essentially be looked up in the GWAS summary statistics of the outcome. For power, multiple genetic variants, rather than single genetic variants, are often used in two sample mendelian randomisation.¹⁵ Mendelian randomisation requires the genetic instrument to act on the outcome only via the exposure. Genetic variants may be pleiotropic (that is, when one genetic variant affects multiple phenotypes). If the SNPs chosen as instruments act on the outcome via a pathway other than through the exposure (horizontal pleiotropy), estimates can be biased.^{20 24} Use of a larger number of SNPs could protect against this bias if the pleiotropic effects balance out.¹⁵ Consistent estimates across multiple methods with different assumptions about pleiotropy also ensures that bias is less likely. Further information on the interpretation and assumptions of mendelian randomisation is presented elsewhere.²³

In this study, we used mendelian randomisation to investigate the association between subjective wellbeing and the cardiometabolic health traits of body mass index, waist to hip ratio, waist circumference, body fat, coronary artery disease, high density lipoprotein cholesterol, low density lipoprotein cholesterol, total cholesterol, systolic blood pressure, diastolic blood pressure, and myocardial infarction. We conducted two sample bidirectional analyses to establish whether subjective wellbeing affects cardiometabolic health traits, or vice versa. We extended previous research by conducting follow-up analysis in an independent sample to avoid sample overlap, adding novel mendelian randomisation methods and examining the causal relation between subjective wellbeing and various cardiometabolic health measures.

Methods

Our mendelian randomisation analysis consisted of two parts: two sample mendelian randomisation using GWAS summary data, and a follow-up analysis exploring results with sample overlap.

Two sample mendelian randomisation *Data sources*

Details of the contributing GWAS consortiums are listed in table 1. The studies were selected for investigating traits related to cardiovascular or metabolic health, having the largest sample sizes, and consisting of the most similar populations while minimising sample overlap. Percentage sample overlap is presented in supplementary table S1. Subjective wellbeing was

Subjective wellbeing Body mass index Waist to hip ratio Waist circumference	Okbay ²⁵ (2016) Locke ²⁶ (2015) Shungin ²⁷ (2015) Shungin ²⁷ (2015) Lu ²⁸ (2016)	SSGAC GIANT GIANT GIANT	298 420 339 224 210 088 232 101	100% European 95% European 100% European	Mixed [†] 53% female 56% female
Waist to hip ratio Waist circumference	Shungin ²⁷ (2015) Shungin ²⁷ (2015)	GIANT	210 088	100% European	56% female
Waist circumference	Shungin ²⁷ (2015)				
	0 (/	GIANT	232101	1000/ 5	
	1^{28} (2016)		2 2 2 1 0 1	100% European	55% female
Body fat percentage	Lu (2010)	Not available	100716	89% European	48% female
HDL cholesterol	Willer ²⁹ (2013)	GLGC	92860	100% European	Mixed [†]
LDL cholesterol	Willer ²⁹ (2013)	GLGC	83198	100% European	Mixed [†]
Total cholesterol	Willer ²⁹ (2013	GLGC	92260	100% European	Mixed [†]
Coronary artery disease	Nikpay ³⁰ (2015)	CARDIoGRAMplusC4D	Cases=60801; controls=123504	77% European	Mixed [†]
Myocardial infarction	Nikpay ³⁰ (2015)	CARDIoGRAMplusC4D	Cases=43676; controls=128199	Mixed†	Mixed [†]
Diastoli blood pressure	Wain ³¹ (2017)	Not available	150 134	100% European	60% female
Systolic blood pressure	Wain ³¹ (2017)	Not available	150 134	100% European	60% female

SSGAC=Social Science Genetics Association Consortium; GLGC=Global Lipids Genetics Consortium; GIANT=Genetic Investigation of Anthropometric Traits consortium; HDL=high density lipoprotein; LDL=low density lipoprotein.

*If not reported, percentage sex and European ancestry were calculated from contributing cohort data in the supplementary materials. All GWAS had similar sex ratios and ancestries included. The largest difference was between the consortiums for coronary artery disease and subjective wellbeing, which used 77% and 100% individuals of European ancestry, respectively. If two populations differ, two sample mendelian randomisation can still be used to test for a causal effect, but the magnitude of the effect might not be as precise.³²

tInformation on the sex ratios and ancestry proportions for the whole sample were not reported or not possible to calculate in the CARDIoGRAMplusC4D, GLGC, and SSGAC consortiums.

measured using any items relating to happiness or positive affect and overall life satisfaction. GWAS of each component were meta-analysed to capture subjective wellbeing.²⁵ For further information on the phenotype definitions and GWAS methods for all traits, see supplementary table S2. All phenotype scores were z scored apart from blood pressure.³¹

Table 1 | Description of CWAS concertiums used for each phonetum

Statistical analyses

We applied four complementary methods of two sample mendelian randomisation (inverse variance weighted method, mendelian randomisation-Egger (MR-Egger) method, weighted median method, and weighted mode based estimation), which make different assumptions about horizontal pleiotropy. A consistent effect across the four methods is less likely to be a false positive.³³ If the genetic variants have horizontally pleiotropic effects but are independent of the effects of the genetic variants on the exposure, this is known as balanced pleiotropy. If all the pleiotropic effects are biasing the estimate in the same direction (directional pleiotropy), this will bias the results (with the exception of the MR-Egger method). We used the MR-Egger intercept and MR-PRESSO (mendelian randomisation pleiotropy residual sum and outlier) to test for the presence of directional pleiotropy. Analyses were conducted using MR-Base,³⁴ a package for two sample mendelian randomisation.

Instrument identification—For all phenotypes, apart from subjective wellbeing, our genetic instruments were composed of genome wide significant SNPs (P<5×10⁻⁸) from published GWAS studies. Only three genome wide significant SNPs were available for subjective wellbeing. We tested the strength of these instruments by checking whether they predicted happiness in a large independent sample (n=242 219) from the UK Biobank. There was evidence that only one SNP (rs2075677) was associated with happiness (supplementary table S3). Therefore, we used a more liberal P value threshold of P<5×10⁻⁵ as the instrument for subjective wellbeing. We selected independent SNPs (linkage disequilibrium (LD) R²=0.001, >10000 kb) using the "clump_data" function on MR-Base. If an SNP was unavailable in the outcome GWAS summary statistics, then proxy SNPs were searched for with a minimum LD R^2 =0.8. If a SNP was palindromic, we aligned strands using minor allele frequency up to 0.3, after which 84 SNPs remained associated at P<5×10⁻⁵. This genetic instrument was checked for overlap with depression¹³ to ensure that the significant SNPs were unique to the subjective wellbeing GWAS. None of the 84 SNPs was associated with major depressive disorder at the genome wide level of significance, suggesting that, based on the current GWAS data, the instruments were independent (supplementary table S4). Power calculations are presented in supplementary table S5.

Inverse variance weighted method—We obtained the inverse variance weighted estimate by meta-analysing the SNP specific Wald estimates using multiplicative random effects. The random effects model was chosen to account for heterogeneity, also measured by Cochran's Q. The Wald estimate is the SNP outcome regression divided by the SNP exposure regression. This method assumes balanced pleiotropy.^{35 36}

MR-Egger method-This method relaxes the assumptions of mendelian randomisation and allows for directional pleiotropic effects, such that some SNPs could be acting on the outcome through a pathway other than through the exposure. The intercept is not constrained to pass through zero and provides an estimate of the directional pleiotropic effect.³⁶ MR-Egger has the lowest power of the four methods to detect a causal effect. It requires variation in the SNP effects, and therefore is most effective when more SNPs are used to create the instrument. MR-Egger is based on the INSIDE assumption (instrument strength independent of the direct effects). It requires that the SNPs' pleiotropic effects on the outcome are independent of the SNPs' association with the exposure.³⁶ The MR-Egger method is also based on the NOME assumption

(no measurement error in the SNP exposure effects),³⁶ which is evaluated by the regression dilution $I^2_{(GX)}$ statistic.³⁷ If the statistic was lower than 0.9, simulation extrapolation corrections were performed³⁷ (see supplementary table S7 for further information).

Weighted median method—This approach assumes that at least 50% of the total weight of the instrument comes from valid variants. It is more likely to give a valid causal estimate than MR-Egger or the inverse variance weighted method because it is more consistent with the true causal effect in the presence of up to 50% invalid variants.³²

Weighted mode based estimation—This method assumes that the most common causal effect is consistent with the true causal effect.³⁸ Hence, the remaining instruments could be invalid (that is, violate the assumptions of mendelian randomisation) without biasing the estimated causal effect.

MR-PRESSO—This test identifies possible bias from horizontal pleiotropy.³⁹ The test consists of three parts: the MR-PRESSO global test, which detects horizontal pleiotropy; the outlier corrected causal estimate, which corrects for the detected horizontal pleiotropy; and the MR-PRESSO distortion test, which estimates whether the causal estimate is significantly different (at P<0.05) after adjustment for outliers.³⁹ We conducted all three stages and present the outlier adjusted causal estimates when both global and distortion tests were significant.

Follow-up analysis in the UK Biobank

We followed up our results from the two sample mendelian randomisation (to overcome potential bias from sample overlap) using a mendelian randomisation analysis where participants for the exposure and outcome were from the same sample (UK Biobank), with a weighted genetic score calculated by use of estimates from GWAS data. The follow-up sample and measures are described below.

Study sample

UK Biobank is a national health resource in the United Kingdom with biological measures collected over 10 years (www.ukbiobank.ac.uk). A total of 502 647 participants aged 39-72 years were recruited from across the UK between 2006 and 2010.⁴⁰ After restricting to European ancestry and excluding related individuals, withdrawn consent, and sex mismatched individuals, the study included 337 112 participants.⁴¹ Of these individuals, the mean age was 56.87 years (standard deviation 8.00) at recruitment, and 54% (n=181 363) were female.

A subsample of participants (n=150000) were genotyped first, and this sample was selected on the basis of smoking status, to include more current smokers than would be representative of the UK population.⁴² These 150000 genotyped individuals also contributed measures of happiness to the Social Science Genetics Association Consortium (SSGAC) of subjective wellbeing.²⁵ The remaining UK Biobank participants have since been genotyped. To avoid any possible biases associated with smoking, we used

the full UK Biobank sample in the follow-up analysis presented here (n=337112). However, because of partial sample overlap with the SSGAC group, we repeated the same follow-up analysis including only individuals who did not contribute to the consortium (n=242219).

Body mass index allele score and observed body mass index

To further test the association between body mass index and wellbeing, we constructed a polygenic score for body mass index in UK Biobank. This polygenic score was constructed by extracting the 97 variants found to be associated at genome wide significance with body mass index in an independent GWAS²⁶ (see supplementary table S6 for full SNP list). Allele scores for each SNP were calculated as a sum of the number of increasing alleles weighted by the effect sizes from the GWAS summary statistics. Therefore, higher polygenic scores indicated an increased risk of higher body mass index. The allele score was standardised to mean zero and standard deviation one. The F statistic was calculated to assess instrument strength in the sample. Observed body mass index was calculated (weight (kg) divided by height (m) squared) from measurements of height and weight taken during the initial assessment centre visit.

Outcomes and confounders

We used phenotypic measurements collected at initial assessment (in 2006-10). The measures included subjective happiness, life satisfaction, and baseline demographics. Subjective happiness was assessed by a single item questionnaire measure. Responses to the question "In general, how happy are you?" were collected on a 6 point Likert scale, ranging from "Extremely unhappy" to "Extremely happy." Individuals who responded with "Do not know" or "Prefer not to answer" were coded as missing.

Life satisfaction was assessed by five single item measures relating to domains of life satisfaction. Domains were: family and relationships, work or job, health, finances, and friendships (eg, "In general, how satisfied are you with your family relationships?"). Responses were collected on a 6 point Likert scale ranging from "Extremely unhappy" to "Extremely happy." Individuals who responded with "Do not know" or "Prefer not to answer" (as well as "I am not employed" for the work/job domain) were coded as missing.

Baseline demographic measures were collected at initial assessment, including sex, age, and socioeconomic position. Socioeconomic position was measured by the Townsend deprivation index (Townsend, 1987) on the basis of participants' location in the UK (calculated from output area) and information from the most recent national census.

Statistical analysis

We first calculated observational associations using linear regression. We then did mendelian randomisation through instrumental variable regressions run in R,⁴³ using the ivreg command from the AER package. We attempted to replicate the effect of body mass index on subjective wellbeing using the body mass index polygenic score as the instrument. Age, sex, and 10 principal components of population structure were controlled for in all instrumental variable regression analyses.

Patient and public involvement

The current research was not informed by patient and public involvement because it used secondary data. However, future research following on from our findings should be guided by patient and public opinions.

Results

Two sample mendelian randomisation

Subjective wellbeing predicting cardiometabolic health outcomes

There was no clear evidence to suggest a causal effect of subjective wellbeing on any of the cardiometabolic health outcomes (fig 2). A more stringent analysis using only genome wide significant SNPs as the instrument produced a similar pattern of results (supplementary figure S1). Phenotype scores were standardised for all outcomes except blood pressure, which is represented on a different scale.

Cardiometabolic health predicting subjective wellbeing

We found evidence that higher body mass index caused lower subjective wellbeing (fig 3). The direction of effect remained consistent across all four methods. There was no clear evidence of a causal effect of any of the cardiovascular health or adiposity exposures on subjective wellbeing (fig 3). Genome wide significant SNPs ($P<5\times10^{-8}$) for each cardiometabolic health measure were used as genetic instruments, and the number of SNPs given for each analysis is shown in figure 3.

Cochran's Q and I² statistics were calculated to check for the presence of heterogeneity (dispersion of SNP effects), which can indicate pleiotropy. We found little evidence of heterogeneity for the association between body mass index and wellbeing (see supplementary table S8 for results and further information). The MR-Egger intercept suggested little evidence of directional pleiotropy (supplementary table S9, all P>0.07). The funnel plot of individual SNP effects showed a symmetrical distribution of SNP effects around the effect estimate, suggesting balanced pleiotropy (supplementary figure S2). We also conducted a leaveone-out analysis, which showed that the SNP with the largest contribution to the effect is rs1421085 located on chromosome 16 in the second intron of the FTO (fat mass and obesity associated) gene (supplementary figure S3).

Follow-up analysis in the UK Biobank

Mean body mass index in the UK Biobank follow-up sample was 27.39 (standard deviation 4.75). Means and standard deviations for the subjective wellbeing

measures are given in supplementary table S10, along with observational associations between body mass index and subjective wellbeing in the UK Biobank sample.

Association with baseline confounders

The association of the body mass index genetic score and observed body mass index with baseline confounders (age, sex, socioeconomic position, education, smoking, and alcohol consumption) were compared (supplementary figure S4). We found evidence of an association between our body mass index genetic score and socioeconomic position, educational attainment, smoking behaviour, and alcohol consumption. For educational attainment and socioeconomic position, the association was weaker for the genetic score than for the observed body mass index.

Association between polygenic score and body mass index

The genetic score was strongly associated with observed body mass index (strength of instrument F(1, 336027)=6237, R²=0.018, P<0.001).

Follow-up analysis of body mass index (exposure) on subjective wellbeing (outcome)

We found strong evidence of a causal effect of body mass index (per 1 kg/m²) on satisfaction with health $(\beta = -0.035, 95\%$ confidence interval -0.043 to -0.027, P<0.001; fig 4), which is consistent with previous estimates from an automated hypothesis free MR-PheWAS (which took the polygenic score for body mass index and looked at the association with every outcome in the UK Biobank).⁴⁴ We saw little evidence of a causal effect of body mass index on any of the other measures of subjective wellbeing, and little evidence that this effect differed in older and younger participants (although the age range in UK Biobank is narrow, with all participants over 40 years old). When individuals were split by median age, the evidence for a causal effect of body mass index on satisfaction with health remained strong in both age groups (age ≤58 years, -0.040 (95% confidence interval -0.051 to -0.029), P<0.001; age >58 years, -0.028 (-0.040 to -0.016), P<0.001). When individuals were split by sex, the evidence for a causal effect of body mass index on satisfaction with health remained strong for both sexes (women, -0.034 (-0.044 to -0.024), P<0.001; men, -0.035 (-0.048 to -0.023), P<0.001). The results remained the same in the independent sample after removal of the contributors to the SSGAC consortium (the subjective wellbeing GWAS; supplementary table S11).

Discussion

Evidence for higher body mass index reducing subjective wellbeing

In the present study, we have found evidence indicating that higher body mass index has a causal relation with lower subjective wellbeing, consistent with previous mendelian randomisation findings.¹⁴ Sensitivity

Outcome	Method	N SNPs	β (95% CI)	β (95% CI)	P value
Body mass index	Inverse variance weighted	82	_ _ _	0.01 (-0.06 to 0.08)	0.74
	MR-Egger (SIMEX)	82	<>	0.10 (-0.66 to 0.56)	0.37
	Weighted median	82		-0.04 (-0.13 to 0.05)	0.36
	Weighted mode	82		-0.09 (-0.32 to 0.14)	0.45
Waist to hip ratio	Inverse variance weighted	82		-0.06 (-0.13 to 0.01)	0.12
	MR-Egger (SIMEX)	82		-0.05 (-0.29 to 0.19)	0.68
	Weighted median	82		-0.09 (-0.19 to 0.00)	0.06
	Weighted mode	82	_	-0.13 (-0.39 to 0.12)	0.31
Waist circumference	Inverse variance weighted	82		0.00 (-0.07 to 0.07)	0.97
	MR-Egger (SIMEX)	82		0.02 (-0.20 to 0.25)	0.85
	Weighted median	82	_	-0.01 (-0.10 to 0.09)	0.90
	Weighted mode	82		0.02 (-0.21 to 0.23)	0.90
Body fat percentage	Inverse variance weighted	82		-0.05 (-0.14 to 0.04)	0.27
	MR-Egger (SIMEX)	82		-0.10 (-0.41 to 0.22)	0.55
	Weighted median	82		-0.01 (-0.13 to 0.11)	0.87
	Weighted mode	82		0.09 (-0.18 to 0.36)	0.53
HDL cholesterol	Inverse variance weighted	69		-0.03 (-0.12 to 0.06)	0.47
	MR-Egger (SIMEX)	69	← →	0.26 (-0.66 to 0.56)	0.10
	Weighted median	69		0.02 (-0.09 to 0.14)	0.69
	Weighted mode	69		0.06 (-0.21 to 0.33)	0.68
LDL cholesterol	Inverse variance weighted	69	_	0.03 (-0.06 to 0.13)	0.48
	MR-Egger (SIMEX)	69	← →	0.26 (-0.66 to 0.56)	0.14
	Weighted median	69		0.01 (-0.11 to 0.14)	0.85
	Weighted mode	69		-0.15 (-0.45 to 0.16)	0.35
Total cholesterol	Inverse variance weighted	69		0.01 (-0.08 to 0.10)	0.84
	MR-Egger (SIMEX)	69	← →	0.28 (-0.66 to 0.56)	0.09
	Weighted median	69		0.02 (-0.11 to 0.14)	0.81
	Weighted mode	69		0.01 (-0.26 to 0.27)	0.97
		-0	0.5 0 0.5		
Outcome	Method	N SNPs	Odds ratio (95% CI)	Odds ratio (95% CI)	P value
Coronary heart disease	Inverse variance weighted	82		1.04 (0.90 to 1.21)	0.62
	MR-Egger (SIMEX)	82		1.19 (0.68 to 2.08)	0.55
	Weighted median	82		1.13 (0.92 to 1.39)	0.25
	Weighted mode	82		1.26 (0.74 to 2.16)	0.40
Myocardial infarction	Inverse variance weighted	82		0.97 (0.82 to 1.15)	0.71
,	MR-Egger (SIMEX)	82		0.95 (0.52 to 1.76)	0.87
	Weighted median	82		1.04 (0.83 to 1.30)	0.75
	Weighted mode	82		1.51 (0.79 to 2.91)	0.22
	-	0	.5 1.0 1.5 2.0 2.5 3.0		
Outcome	Method	N SNPs	β (95% CI)	β (95% Cl)	P value
Diastolic blood pressure	Inverse variance weighted	81		0.32 (-0.48 to 1.12)	0.44
	MR-Egger (SIMEX)	81		1.39 (-1.63 to 4.41)	0.37
	Weighted median	81		-0.24 (-1.18 to 0.71)	0.62
	Weighted mode	81	<>	-0.25 (-2.71 to 2.20)	0.84
Systolic blood pressure	Inverse variance weighted	81		0.41 (-0.87 to 1.68)	0.54
e,stone stoba pressure	MR-Egger (SIMEX)	81		3.40 (-0.93 to 7.74)	0.13
	Weighted median	81		0.90 (-0.54 to 2.34)	0.22
	Weighted mode	81	<	0.97 (-2.57 to 4.51)	0.60
	0				
		-	2 -1 0 1 2		

Fig 2 | Two sample mendelian randomisation analysis showing the effect of subjective wellbeing on cardiometabolic health outcomes using single nucleotide polymorphisms (SNPs) significant at $P(5 \times 10^{-5})$. One unit increase of subjective wellbeing is equivalent to one standard deviation increase of the subjective wellbeing composite continuous scale. HDL=high density lipoprotein; LDL=low density lipoprotein; N SNP=number of the 84 SNPs associated with wellbeing that were available in the outcome summary statistics (SNPs might be unavailable in the outcome owing to imputation platform or not passing quality control procedures). The regression dilution $1^2_{(GX)}$ estimate was less than 90% for the subjective wellbeing instrument (see supplementary table S7 for further information); therefore, simulation extrapolation (SIMEX) correction was applied in mendelian randomisation-Egger (MR-Egger) analysis³⁷ β values are provided for continuous outcomes and odds ratios are provided for binary outcomes

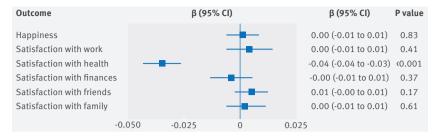
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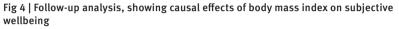
Exposure	Method	N SNPs	β (95% CI)	β (95% CI)	P value
Body mass index	Inverse variance weighted	78		-0.05 (-0.09 to -0.01)	0.02
	MR-Egger	78	←∎ ──	-0.13 (-0.22 to -0.03)	0.01
	Weighted median	78		-0.05 (-0.10 to -0.01)	0.03
	Weighted mode	78	_	-0.11 (-0.19 to -0.03)	0.01
	MR-PRESSO	75		-0.05 (-0.08 to -0.02)	0.01
Waist to hip ratio	Inverse variance weighted	36		0.01 (-0.03 to 0.05)	0.74
	MR-Egger (SIMEX)	36	← ■	0.10 (-0.23 to 0.42)	0.57
	Weighted median	36		0.01 (-0.05 to 0.06)	0.87
	Weighted mode	36		0.02 (-0.09 to 0.14)	0.69
Waist circumference	Inverse variance weighted	40		-0.03 (-0.09 to 0.03)	0.32
	MR-Egger (SIMEX)	40	_	0.23 (0.06 to 0.40)	0.01
	Weighted median	40	- - -	-0.05 (-0.11 to 0.02)	0.14
	Weighted mode	40	_	-0.02 (-0.14 to 0.10)	0.72
	MR-PRESSO	39	_ _	-0.02 (-0.07 to 0.04)	0.59
Body fat percentage	Inverse variance weighted	9		0.00 (-0.09 to 0.09)	0.97
	MR-Egger (SIMEX)	9		0.46 (-0.16 to 1.07)	0.18
	Weighted median	9	_	0.03 (-0.06 to 0.12)	0.49
	Weighted mode	9		0.04 (-0.10 to 0.19)	0.58
HDL cholesterol	Inverse variance weighted	85	÷-	0.00 (-0.02 to 0.02)	0.76
	MR-Egger	85	_ _	0.01 (-0.03 to 0.05)	0.68
	Weighted median	85	-	0.01 (-0.01 to 0.04)	0.33
	Weighted mode	85	.	0.01 (-0.02 to 0.04)	0.45
	MR-PRESSO	84	÷	0.00 (-0.02 to 0.02)	0.93
LDL cholesterol	Inverse variance weighted	69		0.01 (-0.01 to 0.03)	0.17
	MR-Egger	69	-	0.03 (0.00 to 0.06)	0.04
	Weighted median	69		0.03 (0.00 to 0.05)	0.05
	Weighted mode	69	- - -	0.03 (0.01 to 0.06)	0.02
Total cholesterol	Inverse variance weighted	80		0.01 (-0.01 to 0.03)	0.46
	MR-Egger	80		0.01 (-0.03 to 0.05)	0.48
	Weighted median	80		0.03 (-0.00 to 0.06)	0.08
	Weighted mode	80		0.03 (-0.01 to 0.06)	0.10
Coronary heart disease	Inverse variance weighted	31		0.00 (-0.02 to 0.02)	0.89
	MR-Egger	31		-0.01 (-0.05 to 0.03)	0.67
	Weighted median	31	<u> </u>	0.00 (-0.02 to 0.02)	0.99
	Weighted mode	31		0.02 (-0.03 to 0.07)	0.41
Myocardial infarction	Inverse variance weighted	17	-	-0.01 (-0.03 to 0.01)	0.44
	MR-Egger	17		-0.01 (-0.07 to 0.05)	0.74
	Weighted median	17	<u> </u>	0.00 (-0.03 to 0.03)	0.99
	Weighted mode	17		0.03 (-0.03 to 0.09)	0.31
Diastolic blood pressure	•	43		0.00 (0.00 to 0.01)	0.45
	MR-Egger (SIMEX)	43	T. T	0.01 (-0.01 to 0.02)	0.32
	Weighted median	43		0.00 (-0.01 to 0.01)	0.92
	Weighted mode	43	I	0.00 (-0.01 to 0.01)	0.89
Systolic blood pressure	-	36	I	0.00 (0.00 to 0.00)	0.67
	MR-Egger (SIMEX)	36	I	0.00 (-0.02 to 0.01)	0.95
	Weighted median	36	I	0.00 (0.00 to 0.00)	0.80
	Weighted mode	36	I	0.00 (0.00 to 0.00)	0.67
	weighten moue	50		0.00 (0.01 (0 0.01)	0.07

Fig 3 | Two sample mendelian randomisation analysis showing the effect of cardiometabolic health exposures on subjective wellbeing per unit of exposure. HDL=high density lipoprotein; LDL=low density lipoprotein; MR-Egger=mendelian randomisation-Egger method; MR-PRESSO=mendelian randomisation pleiotropy residual sum and outlier; N SNPs=number of single nucleotide polymorphisms (SNPs might be unavailable in the outcome owing to imputation platform or not passing quality control procedures). If the regression dilution $I^2_{(GX)}$ estimates were less than 90%, simulation extrapolation (SIMEX) corrections were applied (supplementary table S7 and supplementary note for further information)

analyses suggested that this association was not due to directional pleiotropy, and the finding was replicated in the UK Biobank cohort. The follow-up analysis suggested that the causal effect of body mass index on subjective wellbeing was driven by satisfaction with health, such that higher body mass index caused lower health satisfaction.

The pathway from body mass index to health satisfaction could be biological or social. Biological pathways include body mass index as a risk factor





for other negative health outcomes such as diabetes, cardiovascular illness, and cancers,⁴⁵ with randomised controlled trials and mendelian randomisation strengthening evidence of a causal effect.^{46 47 48} Therefore, the effect of body mass index on satisfaction with health seen in the current study could reflect accurate perceptions of health. Alternatively, societal influences could cause individuals to associate negative health consequences with a higher body mass index and consequently report lower health satisfaction. Subjective wellbeing and health are in a complex and dynamic system of causal pathways, and further work is needed to understand these pathways using mediation analysis.³

Improved subjective wellbeing as an outcome of reducing body mass index might serve as a motivator for behaviour change. Despite knowing the physical health consequences of having a high body mass index, obese individuals often struggle to maintain diets and lose weight.⁴⁹ Individuals value happiness very highly¹; therefore, emphasis of the potential benefits of weight loss to improve subjective wellbeing could be an unexplored motivator for encouraging weight reduction. Further work should explore the clinical use of subjective wellbeing as a desirable outcome.

In understanding this causal effect further, another important consideration might be the influence of age. Individuals recruited for the UK Biobank were middle aged or older, with an average age of 57 years. Body mass index could be an important determinant of health satisfaction in an older generation as the health implications of obesity (heart disease, diabetes, cancer) begin to emerge.⁵⁰ In younger age groups, body dissatisfaction, self esteem, and bullying might be more important mediators of the association between body mass index and wellbeing.⁵¹ Further investigation of the causal pathways in a younger sample should be explored, especially as some genetic variants for body mass index show a developmentally specific pattern of association.^{52 53}

Despite causal effects of body mass index seen on subjective wellbeing, there was no causal effect of waist to hip ratio, waist circumference, or body fat percentage. Body mass index is used as a proxy for adiposity because it is widely available, easy to collect in large samples, and other more precise measures of adiposity have not been shown to differ substantially in observational studies.¹⁴ All confidence intervals for these precise adiposity measures and body mass index overlap; therefore, we cannot rule out a small effect of these phenotypes on subjective wellbeing. Additionally, we might not have seen a clear effect of body fat percentage on subjective wellbeing because the instrument for body fat percentage was relatively weak, using only nine SNPs identified in a GWAS with a relatively small sample size. Furthermore, there are differences in waist to hip ratio and waist circumference between the sexes that we were unable to explore because the GWAS for subjective wellbeing was in a mixed sex sample. Sex differences could have been masking underlying effects of these more precise adiposity measures on subjective wellbeing. Body mass index can also vary for reasons other than adiposity. Future research should therefore explore which aspect of body mass index is driving the identified effect on subjective wellbeing, because a clear understanding of the mechanisms is important for designing interventions.¹⁵

Observational evidence suggests a non-linear association between mental health and body mass index, where extremely high and low body mass index both predict lower rates of wellbeing.^{8 11} The association between very low body mass index and depression seen in observational studies could be driven by eating disorders such as anorexia nervosa. The two disorders are commonly comorbid, with a 50% lifetime prevalence of major depression in individuals with anorexia.54 Twin studies have suggested that this comorbidity is due to shared genetic influence between anorexia and major depressive disorder.55 However, our mendelian randomisation estimators assumed a linear relation. Therefore, if individuals with low body mass index also have lower subjective wellbeing, this could lead to the effect observed in mendelian randomisation being smaller than it truly is. New methods to allow for non-linear associations in mendelian randomisation are being developed,^{56 57} but are currently too underpowered to apply here.

Evidence for subjective wellbeing influencing health outcomes

In our study, the two sample mendelian randomisation analyses showed no clear evidence of a causal effect of subjective wellbeing on cardiometabolic health outcomes. This result is consistent with a prospective analysis in over 700 000 women, which found no effect of happiness on later mortality, if baseline health was controlled for.⁵⁸ Previous observational associations³⁴ could be due to residual confounding, reverse causation, or publication bias.^{59 60} In our analysis, we saw little evidence that subjective wellbeing had a causal effect on cardiometabolic health outcomes. The genetic variants for subjective wellbeing have small effect sizes, accounting for about 0.01% of the variance, and confidence intervals were wide. Weak instrument bias in mendelian randomisation biases towards the null,⁶¹ which could explain the lack of evidence for a causal effect of subjective wellbeing. Owing to the direction of bias, false positive effects will not be induced by weak instruments (as could be the

case in traditional instrumental variable analysis⁶¹). Further analysis is needed when stronger instruments are available.

Evidence for cardiovascular health influencing subjective wellbeing

We saw no clear evidence for a causal effect of cholesterol, coronary artery disease risk, blood pressure, or myocardial infarction risk on subjective wellbeing, meaning that residual confounding is likely to be responsible for the previous observational associations. This conclusion is supported by recent evidence from a new approach called Bayesian direct multimorbidity mapping. This approach found that coronary artery disease was only associated with depression because of an association with body mass index.⁶² However, coronary artery disease and myocardial infarction are rare outcomes, and the SNPs for coronary artery disease risk used in our analysis all had small effect sizes,³⁰ resulting in limited power to detect an effect on subjective wellbeing. However, the continuous phenotype of blood pressure still showed no causal effects on subjective wellbeing, further supporting residual confounding. Overall, the evidence suggests no causal effect of cardiovascular health on subjective wellbeing.

General limitations of the study

In addition to the specific limitations of weak instruments and statistical power outlined above, a more general limitation of this study could be the influence of population structure on the genetic instrument. In large samples such as the UK Biobank cohort, it is difficult to fully remove population structure without removing true effects.⁶³ Coincident structure could confound the association between body mass index and subjective wellbeing, generating spurious signal. Although we cannot completely remove the possible influence of structure in our followup analysis, we are reasonably confident that the effect of body mass index on satisfaction with health is not spurious because we do not see the same inflation for the negative control outcomes of domain satisfaction or happiness. Furthermore, non-genetic instrumental variables give the same results as genetic instruments in the UK Biobank for educational attainment, a trait largely influenced by structure.⁶⁴

Secondly, the follow-up analysis showed evidence of an association between the body mass index polygenic score and baseline confounders. If the body mass index SNPs affect these confounders independently (horizontal pleiotropy), then the assumptions of mendelian randomisation are violated, and bias would be introduced. Alternatively, alcohol consumption and smoking could be on the causal pathway from body mass index to wellbeing (vertical pleiotropy), which does not violate the assumptions of mendelian randomisation.^{65 66 67} Given the complexity of these genetic and phenotypic relations, it is difficult to disentangle horizontal and vertical pleiotropy; however, we can be reasonably confident that our results are due to vertical pleiotropy, because the MR-Egger intercept did not show evidence of horizontal pleiotropy and MR-PRESSO gave a consistent result after removal of pleiotropic outliers.

Conclusion

Using mendelian randomisation, we found no clear evidence for a causal effect of cardiovascular health on subjective wellbeing, or of subjective wellbeing on cardiometabolic health. This lack of evidence suggests that previously reported observational associations could have resulted from residual confounding. We found strong evidence for a causal effect of increased body mass index on decreased subjective wellbeing. Follow-up analysis in UK Biobank suggested that the effect of body mass index on subjective wellbeing was driven by an adverse effect of higher body mass index on health satisfaction. Our findings add further support to the need to reduce obesity because of the downstream consequences on mental health and wellbeing. Further work is needed to understand the pathways from body mass index to subjective wellbeing and to explore how the effect of body mass index on mental health varies at different ages.

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Ethical approval: UK Biobank has received ethics approval from the UK National Health Service's National Research Ethics Service (ref 11/ NW/0382).

Data sharing: Scripts are available on GitHub at: https://github.com/ MRCIEU/Health-and-Wellbeing-MR.

The lead author affirms that this manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

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Web appendix: Supplementary materials