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A New Control Chart for Monitoring Reliability Using Sudden Death Testing Under Weibull Distribution

MUHAMMAD ASLAM^{®1}, OSAMA H. ARIF¹, AND CHI-HYUCK JUN², (Member, IEEE)

¹Department of Statistics, Faculty of Science, King Abdulaziz University, Jeddah 21551, Saudi Arabia ²Department of Industrial and Management Engineering, POSTECH, Pohang 790-784, South Korea

 $Corresponding \ Author: \ Muhammad \ Aslam \ (aslam_ravian@hotmail.com)$

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ABSTRACT In this paper, a new control chart using sudden death testing is designed by assuming that the lifetime/failure time of the product follows the Weibull distribution. The structure of the proposed chart is presented. The control chart coefficient is determined using some specified average run length for the in control process and the shifted process. Simulation study is given for the illustration purpose.

INDEX TERMS Life test, Weibull distribution, average run length.

I. INTRODUCTION

Statistical process control (SPC) has been considered a powerful tool in quality management decisions. Control chart and acceptance sampling plans are considered as integral parts of SPC. The earlier is used to monitor the manufacturing process and later one is used for the inspection of the finished product. The manufacturing process is monitoring through two control limits called lower control limit (LCL) and upper control limit (UCL). Shewhart control chart is considered to be more effective in detecting large shift in the process. A process beyond these control limits is said to be out-of-control. During the manufacturing, the process can shift from target value due to several factors such as temperature, workers and machines etc. [1] A timely indication is necessary to indicate the shift in the process to avoid the defective/non-conforming items. The control charts help in taking corrective decision for improving the quality of the product. The control charts have much applications in variety of fields including health care, nuclear engineering and epidemiology etc. see, for example [2]–[4].

The control chart can be classified into two classes such as attribute control chart and variable control chart. If the plotting statistic is computed from the attribute data, then attribute control chart is used to monitor the process. The variable control charts are used when the data is obtained from the measurement process. Both control charts have much application in the industry. Attribute chart is easy to apply but less informative than the variable control charts. Several authors designed and discussed the applications of control charts in verity of fields including for example, see [5]–[16] and [17]. Due to high reliability of product, it may not possible to wait failure of all products being tested for inspection. Recently, [18] designed attribute control chart for Weibull distribution for truncated life test. More details about the control chart using the Weibull distribution can be seen in [17], [19]–[24] and [25]. The main objective of this paper to develop a complete structure of proposed chart using the sudden death testing. We will present the necessary measures to evaluate the performance of proposed chart. A simulation study will be given for illustration purpose.

II. DESIGN OF PROPOSED CONTROL CHART

Suppose that the lifetime of a part (denoted by *X*) follows a Weibull distribution with shape parameter *m* and scale parameter λ such that the cumulative distribution function is given by

$$F(x) = 1 - \exp(-(\lambda x)^{m})$$
(1)

According to [26], sudden death testing is frequently adopted by parts manufacturers to reduce testing time. In this testing, a sample of items are distributed into g groups having r items in one group. By exploring the literature, we note that there is no work on the designing of a control chart using the sudden death testing. The operational procedure of the proposed control chart is stated as follows

Step-1: Select a random sample of size n from a lot and distribute r items into g groups so that n = rg.

TABLE 1. The values of ARL when $ARL_0 = 300$; $\lambda = 0.5$.

	r=5				r=10			
k	3.102291	3.102291	3.102289	3.10229	3.489702	3.4897	3.489701	3.489701
g	3	3	3	3	4	4	4	4
m	1	2	2.5	3	1	2	2.5	3
с				AR	L			
1	300.00	300.00	300.00	300.00	300.00	300.00	300.00	300.00
0.99	291.11	282.48	278.25	274.10	288.20	276.86	271.36	265.97
0.98	282.39	265.81	257.89	250.21	276.75	255.30	245.22	235.52
0.97	257.29	220.66	204.36	189.26	244.45	199.19	179.82	162.34
0.95	241.41	194.28	174.29	156.36	224.54	168.10	145.45	125.87
0.93	218.84	159.67	136.40	116.53	197.00	129.43	104.93	85.09
0.9	184.43	113.46	89.02	69.87	156.84	82.11	59.47	43.10
0.85	153.85	79.01	56.68	40.70	123.18	50.75	32.66	21.09
0.8	126.85	53.81	35.12	22.99	95.27	30.49	17.37	10.00
0.75	103.23	35.74	21.14	12.59	72.41	17.77	8.97	4.67
0.7	65.19	14.48	7.01	3.55	39.32	5.55	2.36	1.28
0.6	37.94	5.20	2.20	1.22	19.22	1.74	1.03	1.00
0.5	19.67	1.79	1.04	1.00	8.18	1.01	1.00	1.00
0.3	8.60	1.01	1.00	1.00	2.96	1.00	1.00	1.00
0.2	2.93	1.00	1.00	1.00	1.14	1.00	1.00	1.00
0.1	1.04	1.00	1.00	1.00	1.00	1.00	1.00	1.00

TABLE 2. The values of ARL when $ARL_0 = 370$; $\lambda = 0.5$.

	r=5				r=10				
	3.14653	3.14653	3.14653	3.5305	3.53057	3.53061	3.53055	3.14653	
k	2	2	2	5	4	2	4	7	
g	3	3	3	4	4	4	4	3	
m	1	2	2.5	3	1	2	2.5	3	
c				Α	RL				
1	370.00	370.00	370.00	370.00	370.05	370.12	370.01	370.01	
0.99	359.03	348.38	343.17	328.02	355.49	341.57	334.68	338.05	
0.98	348.27	327.82	318.05	290.46	341.36	314.96	302.42	308.58	
0.97	317.30	272.12	252.00	200.17	301.50	245.71	221.74	233.38	
0.95	297.71	239.56	214.91	155.17	276.94	207.33	179.34	192.80	
0.93	269.87	196.87	168.16	104.86	242.96	159.61	129.34	143.66	
0.9	227.42	139.86	109.71	53.06	193.41	101.22	73.25	86.09	
0.85	189.68	97.36	69.81	25.90	151.87	62.51	40.18	50.10	
0.8	156.38	66.26	43.22	12.21	117.43	37.51	21.31	28.25	
0.75	127.24	43.97	25.96	5.63	89.23	21.81	10.95	15.42	
0.7	80.31	17.74	8.52	1.41	48.40	6.72	2.77	4.25	
0.6	46.69	6.29	2.58	1.00	23.60	2.00	1.07	1.33	
0.5	24.15	2.07	1.08	1.00	9.97	1.02	1.00	1.00	
0.3	10.48	1.03	1.00	1.00	3.52	1.00	1.00	1.00	
0.2	3.48	1.00	1.00	1.00	1.23	1.00	1.00	1.00	
0.1	1.07	1.00	1.00	1.00	1.00	1.00	1.00	1.00	

Step-2: Perform sudden death testing and observe Y_i , the time to the first failure from the ith group (i = 1, 2, ...g). Calculate the quantity $v = \sum_{i=1}^{g} Y_i^m$ and transform it to $v^* = v^{1/3}$.

Step-3: Plot v^* on the control chart. Declare the process in control if LCL $\leq v^* \leq$ UCL. Declare the process out-of-control if $v^* > UCL$ or $v^* < LCL$.

The proposed control chart is based on two control limits, namely LCL and UCL. As in Jun et al. (2006), the quantity v in Step 2 follows a gamma distribution with shape parameter g and scale parameter $r\lambda^m$. In order to set up

symmetric Shewhart-type control limits, we need to transform v to a random variable having a symmetric distribution. Wilson and Hilferty [29] suggested that a gamma random variable can be transformed to an approximate normal variable through power transformation. In fact, the transformation of $v^* = v^{1/3}$ leads to an approximate normal distribution with mean

$$\mu_{v^*} = \frac{(r\lambda^m)^{1/3} \,\Gamma(g+1/3)}{\Gamma(g)} \tag{2}$$

TABLE 3. The values of ARL when $ARL_0 = 300$; $\lambda = 1$.

		1	=5		r=10			
k	3.4897	3.102294	3.102294	3.489702	3.102294	3.489701	3.489705	3.489783
g	4	3	3	4	3	4	4	4
m	1	2	2.5	3	1	2	2.5	3
с				A	RL			
1	300.00	300.01	300.01	300.00	300.01	300.00	300.01	300.13
0.99	288.20	282.48	278.26	265.98	291.11	276.86	271.37	266.08
0.98	276.75	265.82	257.90	235.53	282.39	255.30	245.22	235.62
0.97	244.45	220.67	204.36	162.34	257.29	199.20	179.83	162.41
0.95	224.54	194.28	174.29	125.87	241.41	168.10	145.46	125.92
0.93	197.00	159.67	136.40	85.09	218.84	129.43	104.93	85.12
0.9	156.84	113.46	89.02	43.10	184.44	82.11	59.47	43.12
0.85	123.18	79.02	56.68	21.09	153.85	50.75	32.66	21.09
0.8	95.26	53.81	35.13	10.00	126.86	30.49	17.37	10.00
0.75	72.41	35.74	21.14	4.67	103.23	17.77	8.97	4.67
0.7	39.32	14.48	7.01	1.28	65.19	5.55	2.36	1.28
0.6	19.22	5.20	2.20	1.00	37.94	1.74	1.03	1.00
0.5	8.18	1.79	1.04	1.00	19.67	1.01	1.00	1.00
0.3	2.96	1.01	1.00	1.00	8.60	1.00	1.00	1.00
0.2	1.14	1.00	1.00	1.00	2.93	1.00	1.00	1.00
0.1	1.00	1.00	1.00	1.00	1.04	1.00	1.00	1.00

TABLE 4. The values of ARL when $ARL_0 = 370$; $\lambda = 1$.

	r=5				r=10			
	3.53055	3.53055	3.14653	3.14653	3.14653	3.14653	3.5305	3.53057
k	3	4	4	7	4	2	5	9
g	4	4	3	3	3	3	4	4
m	1	2	2.5	3	1	2	2.5	3
c				AR	L			
1	370.01	370.01	370.00	370.01	370.00	370.00	370.00	370.06
0.99	355.45	341.46	343.18	338.05	359.03	348.38	334.67	328.07
0.98	341.32	314.87	318.05	308.58	348.27	327.82	302.41	290.50
0.97	301.47	245.64	252.01	233.38	317.30	272.12	221.74	200.20
0.95	276.91	207.27	214.91	192.80	297.71	239.56	179.33	155.20
0.93	242.93	159.56	168.16	143.66	269.87	196.87	129.34	104.88
0.9	193.39	101.19	109.71	86.09	227.42	139.86	73.25	53.07
0.85	151.85	62.49	69.81	50.10	189.69	97.36	40.18	25.90
0.8	117.42	37.49	43.22	28.25	156.38	66.26	21.31	12.21
0.75	89.22	21.81	25.96	15.42	127.24	43.97	10.95	5.63
0.7	48.39	6.72	8.52	4.25	80.31	17.74	2.77	1.41
0.6	23.60	2.00	2.58	1.33	46.69	6.29	1.07	1.00
0.5	9.97	1.02	1.08	1.00	24.15	2.07	1.00	1.00
0.3	3.52	1.00	1.00	1.00	10.48	1.03	1.00	1.00
0.2	1.23	1.00	1.00	1.00	3.48	1.00	1.00	1.00
0.1	1.00	1.00	1.00	1.00	1.07	1.00	1.00	1.00

and variance

$$\sigma_{v^*} = \frac{(r\lambda^m)^{2/3} \, \Gamma(g+2/3)}{\Gamma(g)} - (\mu_{v^*})^2 \tag{3}$$

Therefore, the control limits for the proposed control chart are given as

$$LCL = \frac{(r\lambda_0^{m})^{1/3} \Gamma(g+1/3)}{\Gamma(g)} - k \sqrt{\frac{(r\lambda_0^{m})^{2/3} \Gamma(q+2/3)}{\Gamma(q)} - \mu_{v^*}^2}$$
(4)

$$UCL = \frac{(r\lambda_0^{m})^{1/3} \Gamma(g+1/3)}{\Gamma(g)} + k \sqrt{\frac{(r\lambda_0^{m})^{2/3} \Gamma(g+2/3)}{\Gamma(g)}} - \mu_{v^{*2}}$$
(5)

where λ_0 is the scale parameter when the process is in control.

For the proposed control chart, the process is declared to be out-of-control if $v^* (LCL \text{ or } v^*) UCL$. So, the probability that the process is declared as out-of-control when the process is actually in control is given as follows:

$$P_{out}^{0} = P\{v^{*} < LCL|\lambda_{0}\} + P\{v^{*} > UCL|\lambda_{0}\}$$
(6)



TABLE 5. The values of ARL when $ARL_0 = 300$; $\lambda = 2$.

		r=5				r=10			
k		3.102293	3.10229	3.10229	3.489704	3.4897	3.4897	3.489747	3.489962
g		3	3	3	4	4	4	4	4
m		1	2	2.5	3	1	2	2.5	3
с					AR	L			
	1	300.01	300.00	300.00	300.01	300.00	300.00	300.07	300.40
	0.99	291.11	282.47	278.26	265.98	288.20	276.86	271.43	266.33
	0.98	282.39	265.81	257.89	235.53	276.75	255.30	245.27	235.84
	0.97	257.29	220.66	204.36	162.34	244.45	199.19	179.87	162.55
	0.95	241.41	194.27	174.29	125.87	224.54	168.10	145.49	126.04
	0.93	218.84	159.67	136.40	85.09	197.00	129.43	104.96	85.20
	0.9	184.43	113.46	89.02	43.10	156.84	82.11	59.48	43.16
	0.85	153.85	79.01	56.68	21.09	123.18	50.75	32.67	21.11
	0.8	126.85	53.81	35.12	10.00	95.26	30.49	17.37	10.01
	0.75	103.23	35.74	21.14	4.67	72.41	17.77	8.97	4.67
	0.7	65.19	14.48	7.01	1.28	39.32	5.55	2.36	1.28
	0.6	37.94	5.20	2.20	1.00	19.22	1.74	1.03	1.00
	0.5	19.67	1.79	1.04	1.00	8.18	1.01	1.00	1.00
	0.3	8.60	1.01	1.00	1.00	2.96	1.00	1.00	1.00
	0.2	2.93	1.00	1.00	1.00	1.14	1.00	1.00	1.00
	0.1	1.04	1.00	1.00	1.00	1.00	1.00	1.00	1.00

TABLE 6. The values of ARL when $ARL_0 = 370$; $\lambda = 2$.

	r=5				r=10			
	3.14653	3.14653	3.53055	3.14653	3.53054	3.53055	3.14662	3.53054
k	3	4	2	6	9	5	4	9
g	3	3	4	3	4	4	3	4
m	1	2	2.5	3	1	2	2.5	3
c				AI	RL .			
1	370.00	370.01	370.01	370.01	370.00	370.01	370.16	370.00
0.99	359.03	348.38	334.68	338.05	355.44	341.47	343.32	328.02
0.98	348.27	327.82	302.41	308.58	341.32	314.87	318.19	290.45
0.97	317.30	272.12	221.74	233.38	301.46	245.64	252.11	200.16
0.95	297.71	239.57	179.34	192.80	276.91	207.27	215.00	155.17
0.93	269.87	196.87	129.34	143.66	242.93	159.56	168.24	104.86
0.9	227.42	139.86	73.25	86.09	193.38	101.19	109.76	53.06
0.85	189.69	97.36	40.18	50.10	151.85	62.50	69.84	25.90
0.8	156.38	66.27	21.31	28.25	117.41	37.50	43.24	12.21
0.75	127.24	43.97	10.95	15.42	89.22	21.81	25.97	5.63
0.7	80.31	17.74	2.77	4.25	48.39	6.72	8.53	1.41
0.6	46.69	6.29	1.07	1.33	23.60	2.00	2.58	1.00
0.5	24.15	2.07	1.00	1.00	9.97	1.02	1.08	1.00
0.3	10.48	1.03	1.00	1.00	3.52	1.00	1.00	1.00
0.2	3.48	1.00	1.00	1.00	1.23	1.00	1.00	1.00
0.1	1.07	1.00	1.00	1.00	1.00	1.00	1.00	1.00

It is rewritten by

$$P_{out}^{0} = P\left\{v^{\frac{1}{3}} < LCL\right\} + P\left\{v^{\frac{1}{3}} > UCL\right\}$$
$$= P\left\{v < LCL^{3}\right\} + 1 - P\left\{v < UCL^{3}\right\}$$
(7)

Here,

$$P\left(v \leq LCL^{3}\right) = \sum\nolimits_{j=1}^{g-1} \frac{e^{-\frac{LCL^{3}}{r\lambda^{m}}} \left(LCL^{3}/r\lambda^{m}\right)^{j}}{j!}$$

and

$$P\left\{v < UCL^{3}\right\} = \sum_{j=1}^{g-1} \frac{e^{-\frac{UCL^{3}}{r\lambda^{m}}} (UCL^{3}/r\lambda^{m})^{j}}{j!}$$

Therefore, P_{out}^0 can be obtained by

$$P_{out}^{0} = 1 - \sum_{j=1}^{g-1} \frac{e^{-LCL^{3}/r\lambda_{0}^{m}} (LCL^{3}/r\lambda_{0}^{m})^{j}}{j!} + \sum_{j=1}^{g-1} \frac{e^{-UCL^{3}/r\lambda_{0}^{m}} (UCL^{3}/r\lambda_{0}^{m})^{j}}{j!}$$
(8)

The performance of a control chart is usually evaluated by the average run length (ARL). The in-control ARL is given by (9), shown at the bottom of the next page. Suppose now that the scale parameter is shifted to $\lambda_1 = c\lambda_0$. Then, the probability (P¹_{out}) of the process being declared to be out-of-control when the process is shifted is:

$$P_{out}^{1} = P\{v^{*} < LCL|\lambda_{1} = c\lambda_{0}\} + P\{v^{*} > UCL|\lambda_{1} = c\lambda_{0}\}$$
(10)

$$P_{out}^{1} = 1 - \sum_{j=1}^{g-1} \frac{e^{-LCL^{3}/r\lambda_{1}^{m}} (LCL^{3}/r\lambda_{1}^{m})^{J}}{j!} + \sum_{j=1}^{g-1} \frac{e^{-UCL^{3}/r\lambda_{1}^{m}} (UCL^{3}/r\lambda_{1}^{m})^{j}}{j!}$$
(11)

The out-of-control ARL for the shifted process is given as (12), shown at the bottom of this page. The values of ARL₁ for various values of specified parameters are presented in Tables 1-6. We fixed two values of number of testers r = 5, 10, shift c and ARLs. The values of g and m are determined. Table 1-2 are presented for various shift c, $\lambda = 0.5$ and specified ARL₀, say r_0 is 300 and 370. Tables 3-4 are presented for $\lambda = 1$ and ARL_0 is 300 and 370. Tables 5-6 are placed for $\lambda = 2$ and ARL_0 is 300 and 370.

From Tables 1-6, we note following trends in control chart parameters.

- 1) For same values of all other parameters, ARL_1 decreases as λ increases from 0.5 to 2.
- For same values of all other parameters, ARL₁ increases as *m* increases from 1 to 3.
- 3) For same values of all other parameters, ARL₁ decreases as *r* increases from 5 to 10.

The following algorithm is used to find ARL_1 , k and g.

- 1) Specify $ARL_0 = 300, 370; r = 5,10;$ and *m*.
- 2) Determine *k* such that $ARL_0 \ge ARL_0$
- 3) Find ARL₁ according to the values of c for selected k

III. APPLICATION OF PROPOSED CHART

In this section, we will discuss the application of the proposed control chart using an artificial data and a real data obtained from a bearing manufacturing company.

A. ARTIFICIAL DATA

In this section, we will show the performance of proposed control chart using simulated data. For this simulation study, we assume that $\lambda = 2$, m = 2, r = 5 and g = 3. The data Y_i 's are generated from the Weibull distribution with shape parameter m and scale parameter $\lambda r^{1/m}$ when the process is in control and the data Y_i 's are generated from the Weibull distribution with shape parameter m and scale parameter $0.7\lambda r^{1/m}$ when the process is shifted. First 20 values are generated from in-control process and next 20 observations

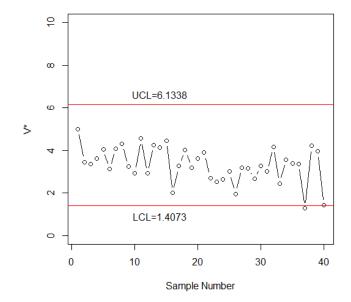


FIGURE 1. The control chart for simulated data.

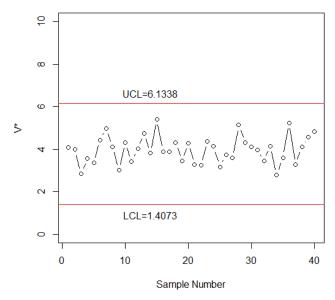


FIGURE 2. The proposed chart for real data.

from shifted process. The simulated data is shown in Table 7. The statistic v^* is computed and shown in Table 7.

The control limits from the in-control data are obtained by LCL = 1.4073 and UCL = 6.1338. The values of v^{*} are plotted on the proposed control chart in Figure 1.

$$ARL_{0} = \frac{1}{1 - \sum_{j=1}^{g-1} \frac{e^{-LCL^{3}/r\lambda_{0}^{m}}(LCL^{3}/r\lambda_{0}^{m})^{j}}{j!} + \sum_{j=1}^{g-1} \frac{e^{-UCL^{3}/r\lambda_{0}^{m}}(UCL^{3}/r\lambda_{0}^{m})^{j}}{j!}}$$
(9)

$$ARL_{1} = \frac{1}{1 - \sum_{j=1}^{g-1} \frac{e^{-LCL^{3}/r\lambda_{1}^{m}}(LCL^{3}/r\lambda_{1}^{m})^{j}}{j!} + \sum_{j=1}^{g-1} \frac{e^{-UCL^{3}/r\lambda_{1}^{m}}(UCL^{3}/r\lambda_{1}^{m})^{j}}{j!}}$$
(12)

TABLE 7. The simulated data.

0.11		Yi's		
Sr#	1	2	3	v*
1	9.521185	3.687743	4.607741	5.006441
2	1.865519	4.038575	4.614244	3.4505
3	1.431465	3.385011	4.912179	3.351233
4	1.135338	6.547314	1.65866	3.606456
5	6.492342	4.528849	1.899786	4.046747
6	4.147889	1.73047	3.241334	3.13141
7	4.139434	3.653806	6.161775	4.090692
8	7.393612	4.565883	2.144854	4.3109
9	4.143483	3.471952	2.179491	3.238757
10	1.973474	4.609762	0.445603	2.937332
11	7.670984	5.118874	3.047797	4.552246
12	3.731001	2.005349	2.601238	2.912598
13	4.195149	6.577462	3.985337	4.249623
14	3.734449	5.846805	4.743502	4.133652
15	3.998439	8.066407	2.709039	4.454577
16	1.589059	0.280453	2.327163	2.00162
17	2.86003	3.038659	4.171112	3.26518
18	2.851136	6.084647	4.495552	4.028174
19	5.203626	0.577831	2.245158	3.189692
20	3.846271	4.746208	3.07018	3.602321
21	3.783086	0.542585	6.66767	3.894403
22	1.344575	2.36778	3.527983	2.708111
23	1.257962	2.762949	2.638212	2.529075
24	1.342806	3.891105	1.257952	2.646038
25	2.40425	3.543137	2.963193	3.004244
26	0.959815	1.212899	2.25691	1.956216
27	3.858921	3.515457	2.238231	3.183357
28	2.711807	4.52549	1.901871	3.156544
29	2.395621	3.493858	0.856643	2.653331
30	1.285116	3.89909	4.208398	3.25746
31	1.329027	2.882516	4.149054	3.010697
32	4.669327	4.185677	5.753682	4.168382
33	0.914754	2.525036	2.726456	2.446665
34	1.625006	5.140723	3.974389	3.553292
35	4.665399	1.846819	3.698544	3.38703
36	2.605654	4.404121	3.469738	3.368592
37	0.293305	0.809499	1.156123	1.276079
38	2.591531	2.40597	7.915057	4.220026
39	1.560297	4.123714	6.488563	3.9481
40	0.983747	1.037736	0.921615	1.425064

From Figure 1, it can be noted that the proposed control chart detects shift at 37th sample, which is the 17th sample after the process shift. So, the proposed control chart has efficiency to detect shift in the process.

TABLE 8. The data for ball bearing.

Yi's							
Sr#	1		2	V*			
1	1 2.009385	2 7.779006	<u>3</u> 1.752351	4.074064			
2	5.960827	3.217039	4.13765	3.979077			
3	1.729362	4.127599	1.659021	2.834775			
4	3.609626	4.108306	3.892758	3.558504			
5	5.054601	1.877135		3.351524			
6	4.426042	5.617833	2.928153	4.430714			
7	6.123741	6.243373	6.859577	4.980372			
8							
	6.536124	2.37621	4.574555	4.10738			
9	1.566379	3.114121	3.858613	3.001488			
10	4.773247	2.235248	7.252549	4.315675			
11	4.62858	3.478084	2.52234	3.416615			
12	3.875296	4.26413	5.612815	4.014622			
13	6.241229	5.970449	5.67346	4.744312			
14	6.863653	1.599389	2.34034	3.806291			
15	8.844831	7.478871	4.875911	5.405425			
16	3.452253	4.972192	4.659262	3.878635			
17	5.611566	0.507887	5.083183	3.861653			
18	3.447337	3.134791	7.592218	4.297219			
19	1.695614	5.297071	3.13846	3.442151			
20	3.449106	5.70535	5.801958	4.274668			
21	2.74742	4.989276	1.635562	3.274684			
22	3.108467	4.727357	1.328097	3.232428			
23	3.457512	6.644438	5.202562	4.36504			
24	2.612608	3.234709	7.259367	4.121039			
25	0.704539	5.320732	1.704934	3.165294			
26	3.129609	4.061275	5.037083	3.724373			
27	3.484034	5.70558	1.066568	3.578621			
28	10.80973	2.10554	3.781144	5.137272			
29	2.920218	7.199859	4.328848	4.292733			
30	5.867873	2.070135	5.462089	4.092666			
31	7.426672	1.852443	1.965285	3.96743			
32	3.72919	1.474369	4.984183	3.446049			
33	6.281662	5.382718	1.688685	4.146343			
34	0.947883	4.323917	1.340808	2.776007			
35	4.090849	4.059047	3.60756	3.588891			
36	5.886786	5.752082	8.687733	5.23197			
37	4.255701	3.032014	2.759001	3.268453			
38	6.916949	2.316732	4.041306	4.112308			
39	3.300406	9.049077	1.796128	4.578929			
40	3.563152	9.971883	0.897838	4.833741			

B. REAL DATA

Now, we apply the proposed control chart to a real data obtained from a bearing manufacturing company. The similar example was given by [20] Suppose that quality assurance department of this company decides to monitor the reliability of their product using the proposed control chart with the target in-control ARL of 370. They use sudden death testing with r = 5 and g = 3. The data is well fitted to the Weibull distribution with m = 2. The first failure time from each group is noted and the data is reported in Table 8. The values of statistic v* are computed and placed in Table 8.

The values of statistics v^* are plotted on Figure 2. It can be seen that all plotted values are within the control limits which indicate that the process of ball bearing is in-control.

IV. CONCLUSIONS

A variable control chart for monitoring reliability of a product under sudden death life testing is proposed under the assumption that the lifetime of the product follows a Weibull distribution. A power transform is utilized so as to design symmetric type control limits. The in-control ARL and the ARL for a shifted process are derived. The simulation study shows that the proposed control chart has ability to detect a shift in the process. An example is also given to illustrate the use of the proposed control chart in the industry. As Weibull distribution may be also asymmetric depending on the values of the parameters, it could be very interesting in determining the control limits for the gamma distribution and in comparing the performance with asymmetric versus symmetric distributions as future research.

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REFERENCES

- D. C. Montgomery, Introduction to Statistical Quality Control. Hoboken, NJ, USA: Wiley, 2007.
- [2] P. Masson, "Quality control techniques for routine analysis with liquid chromatography in laboratories," *J. Chromatogr. A*, vol. 1158, nos. 1–2, pp. 168–173, 2007.
- [3] S.-L. Hwang *et al.*, "Application control chart concepts of designing a prealarm system in the nuclear power plant control room," *Nucl. Eng. Des.*, vol. 238, no. 12, pp. 3522–3527, 2008.
- [4] W. H. Woodall, "The use of control charts in health-care and public-health surveillance," J. Quality Technol., vol. 38, no. 2, pp. 89–104, 2006.
- [5] L. Ahmad, M. Aslam, and C.-H. Jun, "The design of a new repetitive sampling control chart based on process capability index," *Trans. Inst. Meas. Control*, vol. 38, no. 8, pp. 971–980, 2015.
- [6] H. A. Al-Oraini and M. Rahim, "Economic statistical design of x¥ control charts for systems with gamma (< i > λ < /i >, 2) in-control times," *Comput. Ind. Eng.*, vol. 43, no. 3, pp. 645–654, 2002.
- [7] A. F. Costa and M. S. De Magalhaes, "An adaptive chart for monitoring the process mean and variance," *Quality Rel. Eng. Int.*, vol. 23, no. 7, pp. 821–831, 2007.
- [8] E. K. Epprecht, A. F. B. Costa, and F. C. T. Mendes, "Adaptive control charts for attributes," *IIE Trans.*, vol. 35, no. 6, pp. 567–582, 2003.
- [9] H. Luo and Z. Wu, "Optimal np control charts with variable sample sizes or variable sampling intervals," *Econ. Quality Control*, vol. 17, no. 1, pp. 39–61, 2002.
- [10] A. A. De Araújo Rodrigues, E. K. Epprecht, and M. S. De Magalhães, "Double-sampling control charts for attributes," *J. Appl. Statist.*, vol. 38, no. 1, pp. 87–112, 2011.
- [11] Z. Wu, M. B. C. Khoo, L. Shu, and W. Jiang, "An < i > np < /i > control chart for monitoring the mean of a variable based on an attribute inspection," *Int. J. Prod. Econ.*, vol. 121, no. 1, pp. 141–147, 2009.

- [12] F. Y. Yen, K. M. B. Chong, and L. M. Ha, "Synthetic-type control charts for time-between-events monitoring," *PLoS ONE*, vol. 8, no. 6, p. e65440, 2013.
- [13] M. Aslam, M. Azam, and C.-H. Jun, "New attributes and variables control charts under repetitive sampling," *Ind. Eng. Manage. Syst.*, vol. 13, no. 1, pp. 101–106, 2014.
- [14] Z. Wu, H. Luo, and X. Zhang, "Optimal np control chart with curtailment," *Eur. J. Oper. Res.*, vol. 174, no. 3, pp. 1723–1741, 2006.
- [15] Z. Wu and Q. Wang, "An NP control chart using double inspections," J. Appl. Statist., vol. 34, no. 7, pp. 843–855, 2007.
- [16] W. Zhang, S. H. Yeo, and T. A. Spedding, "A synthetic control chart for detecting fraction nonconforming increases," *J. Quality Technol.*, vol. 33, no. 1, pp. 104–111, 2001.
- [17] M. Aslam, O. H. Arif, and C.-H. Jun, "An attribute control chart for a Weibull distribution under accelerated hybrid censoring," *PLoS ONE*, vol. 12, no. 3, p. e0173406, 2017.
- [18] M. Aslam and C.-H. Jun, "Attribute control charts for the Weibull distribution under truncated life tests," *Quality Eng.*, vol. 27, no. 3, pp. 283–288, 2015.
- [19] S. B. Akhundjanov and F. Pascual, "Moving range EWMA control charts for monitoring the Weibull shape parameter," J. Statist. Comput. Simul., vol. 85, no. 9, pp. 1864–1882, 2015.
- [20] R. M. Dickinson, D. A. O. Roberts, A. R. Driscoll, W. H. Woodall, and G. G. Vining, "CUSUM charts for monitoring the characteristic life of censored Weibull lifetimes," *J. Quality Technol.*, vol. 46, no. 4, pp. 340–358, 2014.
- [21] X. Huang and F. Pascual, "ARL-unbiased control charts with alarm and warning lines for monitoring Weibull percentiles using the first-order statistic," J. Statist. Comput. Simul., vol. 81, no. 11, pp. 1677–1696, 2011.
- [22] M. D. Nichols and W. J. Padgett, "A bootstrap control chart for Weibull percentiles," *Quality Rel. Eng. Int.*, vol. 22, no. 2, pp. 141–151, 2006.
- [23] F. Pascual, "EWMA charts for the Weibull shape parameter," J. Quality Technol., vol. 42, no. 4, pp. 400–416, 2010.
- [24] F. Pascual and D. Nguyen, "Moving range charts for monitoring the Weibull shape parameter with single observation samples," *Quality Rel. Eng. Int.*, vol. 27, no. 7, pp. 905–919, 2011.
- [25] L. Zhang and G. Chen, "EWMA charts for monitoring the mean of censored Weibull lifetimes," J. Quality Technol., vol. 36, no. 3, pp. 321–328, 2004.
- [26] C.-H. Jun, S. Balamurali, and S.-H. Lee, "Variables sampling plans for Weibull distributed lifetimes under sudden death testing," *IEEE Trans. Rel.*, vol. 55, no. 1, pp. 53–58, Mar. 2006.



MUHAMMAD ASLAM received the M.Sc. degree in statistics from GC University Lahore with the Chief Minister of the Punjab Merit Scholarship in 2004, the M.Phil. degree in statistics from GC University Lahore with the Governor of the Punjab Merit Scholarship in 2006, and the Ph.D. degree in statistics from the National College of Business Administration and Economics Lahore in 2010, under the supervision of Prof. Dr. M. Ahmad. He was a Lecturer of Statistics with

the Edge College System International from 2003 to 2006. He was also a Research Assistant with the Department of Statistics, GC University Lahore, from 2006 to 2008. Then he joined as a Lecturer with the Forman Christian College University in 2009, where he was an Assistant Professor from 2010 to 2012. He was with the Department of Statistics as an Associate Professor from 2012 to 2014. He was an Associate Professor of Statistics with the Department of Statistics, Faculty of Science, King Abdulaziz University,

Jeddah, Saudi Arabia from 2014 to 2017. He taught summer course as a Visiting Faculty of statistics with Beijing Jiaotong University, China, in 2016. He is currently a Full Professor of Statistics with the Department of Statistics, King Abdulaziz University. He has authored over 235 research papers in national and international well reputed journals, including IEEE Access, Journal of Applied Statistics, the European Journal of Operation Research, the Journal of the Operational Research Society, Applied Mathematical Modeling, the International Journal of Advanced Manufacturer Technology, Communications in Statistics, the Journal of Testing and Evaluation, and the Pakistan Journal of Statistics. His papers have been cited 1634 times with h-index 22 and i-10 index 52. He has authored one book published in Germany. He has been a HEC approved Ph.D. Supervisor since 2011. His areas of interest include reliability, decision trees, Industrial Statistics, acceptance sampling, rank set sampling and applied Statistics. He has supervised five Ph.D. theses over 25 M.Phil. theses and 3 M.Sc. theses. He is currently supervising one Ph.D. thesis and over five M.Phil. theses in statistics. He is a reviewer of over 50 well-reputed international journals. He reviewed over 75 research papers for various well reputed international journals. He is also a member of Islamic Countries Society of Statistical Sciences. He received the Meritorious Services Award in research from the National College of Business Administration and Economics Lahore in 2011. He received the Research Productivity Award in 2012 by Pakistan Council for Science and Technology. His name Listed at the Second Position among Statistician in the Directory of Productivity Scientists of Pakistan 2013. His name Listed at the First Position among Statistician in the Directory of Productivity Scientists of Pakistan 2014. He received 371 positions in the list of top 2210 profiles of Scientist of Saudi Institutions 2016. He is selected for the Innovative Academic Research and Dedicated Faculty Award 2017 by SPE, Malaysia. He received King Abdulaziz University Excellence Awards in Scientific Research for the paper entitled Aslam, M., Azam, M., Khan, N. and Jun, C.-H. (2015). A New Mixed Control Chart to Monitor the Process, the International Journal of Production Research, 53 (15), 4684-4693. He is a member of the Editorial Board of the Electronic Journal of Applied Statistical Analysis, Asian Journal of Applied Science and Technology, and the Pakistan Journal of Commence and Social Sciences.



OSAMA H. ARIF received the M.Sc. degree from Sheffield University and the Ph.D. degree from Sheffield Hallam University, U.K. He is currently a member of Staff with the Department of Statistics, Faculty of Science, King Abdulaziz University, Saudi Arabia. He has over 15 years' experience in research in statistics and quality fields. He has authored several papers in various international statistical and quality journals. He has appointed as the Director General of TQM Program with KAU

from 2005 to 2009. In addition, he has been an Executive Manager of Quality and Excellence Consultant, Think Tank, KAU, since 2007, and a Strategic Plan Consultant with Jouf University.



CHI-HYUCK JUN (M'91) was born in Seoul, South Korea, in 1954. He received the B.S. degree in mineral and petroleum engineering from Seoul National University in 1977, the M.S. degree in industrial engineering from KAIST in 1979, and the Ph.D. degree in operations research from University of California at Berkeley, in 1986. Since 1987, he has been with the Department of Industrial and Management Engineering, POSTECH, and he is currently a Professor, and the Department

Head. He is interested in reliability and quality analysis, and data mining techniques. He is a member of INFORMS, and ASQ.