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# Figures of merit for present and future dark energy probes

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We compare current and forecasted constraints on dynamical dark energy models from Type Ia supernovae and the cosmic microwave background using figures of merit based on the volume of the allowed dark energy parameter space. For a two-parameter dark energy equation of state that varies linearly with the scale factor, and assuming a flat universe, the area of the error ellipse can be reduced by a factor of  $\sim 10$  relative to current constraints by future space-based supernova data and CMB measurements from the Planck satellite. If the dark energy equation of state is described by a more general basis of principal components, the expected improvement in volume-based figures of merit is much greater. While the forecasted precision for any single parameter is only a factor of 2–5 smaller than current uncertainties, the constraints on dark energy models bounded by  $-1 \le w \le 1$  improve for approximately 6 independent dark energy parameters resulting in a reduction of the total allowed volume of principal component parameter space by a factor of  $\sim 100$ . Typical quintessence models can be adequately described by just 2–3 of these parameters even given the precision of future data, leading to a more modest but still significant improvement. In addition to advances in supernova and CMB data, percent-level measurement of absolute distance and/or the expansion rate is required to ensure that dark energy constraints remain robust to variations in spatial curvature.

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### I. INTRODUCTION

In the absence of physically compelling models for dark energy, figures of merit (FoMs) are a useful tool that encapsulate the constraining power of cosmological data. FoMs combine various constraints on the expansion history of the universe (e.g. the distance-redshift relation) in a single number, or at most a handful of numbers. These numbers serve as simple and quantifiable metrics by which to evaluate the accuracy of constraints on dark energy parameters from current and proposed experiments [1–3]. The simplest schemes adopt fixed functional forms for the evolution of the dark energy equation of state and define the FoM as the inverse of the allowed parameter volume [4]. One widely used version is the two-dimensional  $w_0 - w_a$  parametrization [1], but other higher-dimensional versions have also been considered [4–6].

Any simple parametrization of the expansion history risks biasing the FoM in favor of or against certain types of data by choosing a fixed functional form [2]. To avoid this problem, one can use more complicated schemes that parametrize the whole functional freedom in the dark energy equation of state evolution and separate the expansion history and growth of structure information. For example, uncorrelated modes of piecewise-constant discretizations of the equation of state that are local in redshift [7–15] or constructed from principal components (PCs) [7,16–21] have been employed to characterize both current and future data. In particular, the inverse parameter volume

of the PC amplitudes, defined separately for each experiment, has been advocated as a FoM [3].

Besides avoiding biasing results towards a particular functional form, a more model-independent FoM has the advantage of being able to identify improvements in dark energy constraints that might be missed by FoMs with fewer parameters. On the other hand, not all improvements in a multidimensional PC FoM reflect improvements in constraining the space of known dark energy models [22,23]. For example, stronger constraints may exclude regions of the parameter space not occupied by typical models.

In this paper, we study the FoMs defined both with the commonly used  $w_0 - w_a$  parametrization and with PCs based on forecasts for Type Ia supernova (SN) and cosmic microwave background (CMB) data. Previous studies of PC-based FoMs have generally relied on the Fisher matrix approximation, and the implementation and utility of these FoMs for real data has not been addressed. Here we define straightforward generalizations of PC FoMs and apply them to both actual data from current measurements and forecasts for future data. While we still construct the PCs using the Fisher matrix approach, we then treat the PC amplitudes as free parameters and constrain them using full Markov chain Monte Carlo (MCMC) likelihood explorations with current or future data. This methodology follows that employed in previous papers where we studied generalized predictions of classes of dark energy models based on forecasts [17] and current data [18]. The new element here is the application of these methods to the study of dark energy FoMs.

The standard approach for constructing the principal component-based FoM is to use the PCs specific to the experiment and the cosmological probe(s) considered; see e.g. Ref. [3]. However, this approach makes it difficult to directly compare PC-based FoMs for different experiments and to assess improvements in specific regions of parameter space. To facilitate such direct comparison between FoMs for current and future data, we choose instead to fix the set of PCs based on a specific projection of future data rather than computing separate PCs for future and current data.

This paper is organized as follows. In Sec. II, we describe the current and forecasted SN and CMB data sets as well as additional priors from baryon acoustic oscillations (BAO) and measurements of the Hubble constant. We also briefly review the MCMC analysis methods used to infer dark energy constraints. In Sec. III, we compare FoMs from the current and future SN and CMB constraints for both the commonly used  $w_0 - w_a$  model (Sec. III A) and a more general PC-based parametrization (Sec. III B). We summarize and discuss these results in Sec. IV.

## II. METHODOLOGY

In this section, we review the current data sets and assumptions about future experiments for forecasts that we use in this paper. We refer the reader to Refs. [17] (hereafter MHH) and [18] for more details concerning the forecasts and current data sets, respectively. All forecasts in this paper assume that the data originate from a flat cosmological constant (w=-1) model with present matter fraction  $\Omega_m=0.24$  and Hubble constant  $H_0=73~{\rm km\,s^{-1}\,Mpc^{-1}}$ .

## A. Current SN and CMB data

The Type Ia SN sample we use is the Union compilation [24]. These SN observations measure relative distances,  $D(z_1)/D(z_2)$ , over a range of redshifts spanning 0.015  $\leq z \leq$  1.551, with most SNe at  $z \leq$  1 (see Fig. 1). We include SN constraints using the likelihood code for the Union data

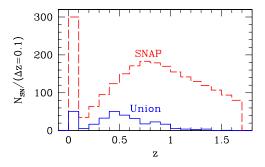


FIG. 1 (color online). Redshift distributions of Type Ia supernovae in the Union compilation (solid blue) and assumed for SNAP (dashed red), including the anticipated low-redshift sample from other surveys at z < 0.1.

sets [25], which includes estimated systematic errors in the covariance matrix [24].

For the current CMB data, we use the 5-year data release from the WMAP satellite [26–28] employing the likelihood code available at the LAMBDA Web site [29]. Unlike the CMB priors used for the forecasts below, the likelihood used here contains the full information from the CMB angular power spectra, except for the small effects of gravitational lensing of the CMB that add little to current dark energy constraints from WMAP. We compute the CMB angular power spectra using the code CAMB [30,31] modified with the parametrized post-Friedmann dark energy module [32,33] to include models with general dark energy equation of state evolution where w(z) may cross w = -1.

## **B. SN and CMB forecasts**

For our SN forecasts, we take the expected redshift distribution for the SuperNova/Acceleration Probe (SNAP) [34,35] plus a low-z sample of 300 SNe at 0.03 < z < 0.1. The SNAP magnitude errors include both statistical and systematic components:

$$\sigma_{\alpha}^{2} = \left(\frac{\Delta z}{\Delta z_{\text{sub}}}\right) \left[\frac{0.15^{2}}{N_{\alpha}} + 0.02^{2} \left(\frac{1 + z_{\alpha}}{2.7}\right)^{2}\right], \tag{1}$$

where  $N_{\alpha}$ , shown in Fig. 1, is the number of SNe in each bin of width  $\Delta z$  ( $\Delta z = 0.1$  except for the statistical uncertainties in the low-z SN bin, for which  $\Delta z = 0.1 - z_{\min}^{\rm SN} = 0.07$ ), and  $\Delta z_{\rm sub}$  is the width of the sub-bins used to smooth the distribution of SNe in redshift. We use 500 sub-bins up to  $z_{\rm max} = 1.7$ . The second term on the right-hand side of Eq. (1) models a systematic floor that increases linearly with z up to a maximum at  $z_{\rm max}$  of 0.02 mag per  $\Delta z = 0.1$  bin [36].

For the Planck CMB forecasts, we use a  $2 \times 2$  covariance matrix corresponding to the inner error ellipses in Fig. 2,

$$\mathbf{C}^{\text{CMB}} = \begin{pmatrix} (0.0018)^2 & -(0.0014)^2 \\ -(0.0014)^2 & (0.0011)^2 \end{pmatrix}, \tag{2}$$

with parameters

$$\boldsymbol{\theta}^{\text{CMB}} = \{\ln(D(z_*)/\text{Mpc}), \Omega_m h^2\}. \tag{3}$$

Here  $D(z_*)$  is the comoving angular diameter distance to recombination.

In the Planck forecasts we ignore additional CMB information about dark energy such as the integrated Sachs-Wolfe effect and gravitational lensing of the CMB. Planck data are expected to obtain limits on the fraction of early dark energy of  $\sigma(\Omega_{\rm DE}(z_*)) \approx 0.004$  [37,38]; however, since these limits may depend on the modeling of early dark energy, we include in our forecasts only a conservative prior of  $\sigma(\Omega_{\rm DE}(z_*)) = 0.025$  [with fiducial value  $\Omega_{\rm DE}(z_*) \sim 10^{-9}$ ] which approximates the current constraint from WMAP [39].

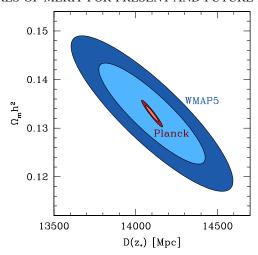


FIG. 2 (color online). Approximate constraints on  $D(z_*)$  and  $\Omega_m h^2$  from WMAP 5-year data (outer contours, blue shading) and forecasts for Planck (inner contours, red shading), showing 68% C.L. (light shading) and 95% C.L. (dark shading) contours.

## C. Additional priors

For both the current and forecasted constraints we add additional priors from recent measurements of baryon acoustic oscillations and the Hubble constant. The BAO constraint we use is based on the measurement of the correlation function of Sloan Digital Sky Survey Luminous Red Galaxies (LRGs) [40], which determines the distance and expansion rate at  $z_{\rm BAO} \approx 0.35$  through the combination  $D_V(z) \equiv [zD^2(z)/H(z)]^{1/3}$ . We implement this constraint by taking the volume average of this quantity,  $\langle D_V \rangle$ , over the LRG redshifts, 0.16 < z < 0.47, and comparing with the value of  $A \equiv \langle D_V \rangle \sqrt{\Omega_m h^2}/z_{\rm BAO}$  given in Ref. [40],  $A = 0.472 \pm 0.017$  (taking the scalar spectral tilt to be  $n_s = 0.96$ ). Using more recent BAO constraints, e.g. from Ref. [41], has only a small effect on the current constraints with SN and CMB data [18].

We include the recent Hubble constant ( $H_0$ ) constraint from the SHOES team [42], based on SN distances at 0.023 < z < 0.1 that are linked to a maser-determined absolute distance using Cepheids observed in both the maser galaxy and nearby galaxies hosting Type Ia SNe. The SHOES measurement determines the absolute distance to a mean SN redshift of  $z_h = 0.04$ , which effectively corresponds to a constraint on  $H_0$  for models with relatively smooth dark energy evolution in the recent past (cf. [43]) such that  $\lim_{z\to 0} D(z) = cz/H_0$ . We implement this constraint as a measurement of  $D(z_h) = cz_h/(74.2 \pm 3.6 \text{ km s}^{-1} \text{ Mpc}^{-1})$ .

## D. MCMC methodology

Given the current or forecasted data, we use MCMC likelihood analysis (e.g. see [44–46]) to determine dark energy parameter constraints and figures of merit for both the simple  $w_0 - w_a$  models and general PC

parametrization. From the likelihood  $\mathcal{L}(\mathbf{x}|\boldsymbol{\theta})$  of the data  $\mathbf{x}$  given each proposed parameter set  $\boldsymbol{\theta}$ , Bayes' Theorem tells us the posterior probability distribution of the parameter set given the data

$$\mathcal{P}(\boldsymbol{\theta}|\mathbf{x}) = \frac{\mathcal{L}(\mathbf{x}|\boldsymbol{\theta})\mathcal{P}(\boldsymbol{\theta})}{\int d\boldsymbol{\theta} \mathcal{L}(\mathbf{x}|\boldsymbol{\theta})\mathcal{P}(\boldsymbol{\theta})},\tag{4}$$

where  $\mathcal{P}(\boldsymbol{\theta})$  is the prior probability density. The MCMC algorithm generates random draws from the posterior distribution. We test convergence of the samples to a stationary distribution that approximates  $\mathcal{P}(\boldsymbol{\theta}|\mathbf{x})$  by applying a conservative Gelman-Rubin criterion [47] of  $R-1 \lesssim 0.01$  across a minimum of four chains for each model class. We use the code COSMOMC [48,49] for the analysis of current data and an independent MCMC code for forecasts.

#### III. FIGURE OF MERIT COMPARISONS

# A. $w_0 - w_a$ figure of merit

We first consider the two-parameter model for the dark energy equation of state [50,51]

$$w(z) = w_0 + w_a \frac{z}{1+z}. (5)$$

The FoM for this model defined by the Dark Energy Task Force [1,4] is the inverse of the area of the 95.4% C.L. region  $A_{95}$  in the  $w_0-w_a$  plane. Figure 3 shows these regions for the current and forecasted data with and without marginalization of spatial curvature  $\Omega_{\rm K}$ .

For a Gaussian error distribution,  $A_{95}$  is proportional to the square root of the determinant of the 2D covariance matrix  $\mathbb{C}$  for  $w_0$  and  $w_a$ . Since the constant of proportionality used in practice for the FoM varies widely in the literature (e.g., see [52]), we simply define

FoM 
$$^{(w_0 - w_a)} \equiv (\det \mathbf{C})^{-1/2} \approx \frac{6.17\pi}{A_{95}}$$
. (6)

The approximate equality in Eq. (6) becomes exact for a Gaussian posterior distribution. Although the posterior in  $w_0$  and  $w_a$  is not perfectly Gaussian, the FoM computed using detC in Eq. (6) remains a good approximation to the area-based FoM. The difference between the two is  $\sim 10\%$  in the worst case (current data with  $\Omega_{\rm K} \neq 0$ ) and  $\lesssim 2\%$  in the other cases. Values of the detC version of FoM<sup>( $w_0 - w_a$ )</sup> are given in Table I. The FoMs for current data are consistent with those found in previous studies of  $w_0 - w_a$  constraints from similar data sets (e.g., [5,53–55]).

<sup>&</sup>lt;sup>1</sup>The constant of proportionality can be derived by considering a 2D circular Gaussian distribution with width  $\sigma$ . Defining  $\alpha \equiv A_{95}/(\pi\sigma^2)$  and integrating the distribution over the area  $A_{95}$  yields  $1 - e^{-\alpha/2}$ ; setting this equal to 0.954, we obtain  $\alpha \approx 6.17$ .

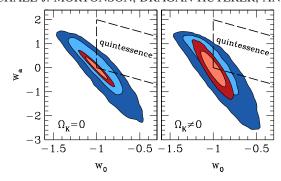


FIG. 3 (color online). Constraints on  $w_0$  and  $w_a$ , assuming a flat universe (left panel) or marginalizing over spatial curvature (right panel). The regions of 68% C.L. (light shading) and 95% C.L. (dark shading) are shown for both current Union + WMAP constraints (outer contours, blue shading) and SNAP + Planck forecasts (inner contours, red shading). Dashed lines mark the boundary of the quintessence prior,  $-1 \le w(z) \le 1$ .

While analyses of  $w_0 - w_a$  models typically allow w(z)to cross -1, it is useful to also consider a more restricted class of models that satisfy the quintessence bound  $-1 \le$  $w \le 1$  in light of the PC description below. Imposing this quintessence prior greatly restricts the allowed parameter space as shown in Fig. 3. Assuming a flat (nonflat) universe, the detC FoM is a factor of ~8 (6.5) larger than without the quintessence prior for current data, and a factor of  $\sim$ 13 (9) larger for forecasts. Thus the addition of the quintessence prior increases the ratio of future to current FoM values by about 60% (35%) relative to the ratios in Table I. Note that the effect of the quintessence prior on the FoM for forecasts depends on the choice of the true model from which the data are drawn; for example, had we chosen fiducial  $w_0$  and  $w_a$  values that lie as far within the quintessence prior boundaries as the current data allow, the forecasted area allowed within the priors would be greater and the improvement in the FoM relative to current constraints would be smaller.

Even in this case where the posterior distribution with the quintessence prior is far from Gaussian, det C still approximates the area-based FoM through Eq. (6) reasonably well. Although values of the det C FoM are smaller by 35%–40% than the area FoM, the ratio of future to current FoM values is nearly unchanged. We shall see that this type of agreement for quintessence models carries over to the more general PC-based FoM in the next section.

TABLE I. Figures of merit for  $w_0 - w_a$  models and 10-PC quintessence models without early dark energy.

	$FoM^{(w_0-w_a)}$		FoM <sub>10</sub> <sup>(PC)</sup>	
Data	$\Omega_{\rm K} = 0$	$\Omega_{\rm K} \neq 0$	$\Omega_{\rm K} = 0$	$\Omega_{\rm K} \neq 0$
SNAP + Planck	160	46	53 000	19 000
Union + WMAP	15	11	370	260
Ratio	11	4.3	140	73

## B. Principal component figure of merit

We generalize the dark energy parametrization to allow arbitrary variations of the equation of state at  $z < z_{\rm max}$  with a basis of principal components (PCs) [7,16]. Details of the PC construction can be found in MHH; here we highlight the points of special relevance for the FoM.

We construct the PCs based on the SN and CMB forecasts from Sec. II B. The PCs are a set of orthogonal functions ordered by the precision with which they can be measured by the future SN and CMB data. Specifically, the principal component functions  $e_i(z_j)$  are eigenvectors of the SNAP + Planck covariance matrix in the space of piecewise-constant values of the equation of state in redshift bins,  $w(z_j)$ . The principal components form a basis in which an arbitrary function  $w(z_i)$  may be expressed as

$$w(z_j) = -1 + \sum_{i=1}^{N_{z,PC}} \alpha_i e_i(z_j),$$
 (7)

where  $\alpha_i$  are the PC amplitudes,  $N_{z,PC}=1+z_{\rm max}/\Delta z$  is the number of redshift bins of width  $\Delta z$ , and  $z_j=(j-1)\Delta z$ . We choose the maximum redshift for variations in w(z) to be  $z_{\rm max}=1.7$ , matching the largest redshift for the SNAP supernova data. The impact of dark energy evolution at higher redshifts is expected to be small, but perhaps non-negligible; to account for this possibility, we parametrize the equation of state at  $z>z_{\rm max}$  by a constant,  $w_{\infty}$ . Likewise, we consider models with spatial curvature,  $\Omega_{\rm K}\neq 0$ . Note that  $w_{\infty}$  and  $\Omega_{\rm K}$  are not allowed to vary from their fiducial values of -1 and 0, respectively, in the Fisher analysis used to construct the PCs.

Since the highest-variance PCs correspond to modes of w(z) to which even future data are insensitive, we truncate the sum in Eq. (7) by replacing  $N_{z,PC}$  with  $N_{max} < N_{z,PC}$ . As shown in MHH, the 10 lowest-variance PCs ( $N_{max} = 10$ ) form a basis which, for the classes of models we consider here, is sufficiently complete for the forecasts. Therefore, 10 PCs more than suffice for the current data as well. This set of basis functions is displayed in Fig. 4.

We impose priors on the PC amplitudes corresponding to the range of w(z) allowed for scalar field quintessence,  $-1 \le w \le 1$  following MHH. Our conservative implementation excludes only parameter values that must violate these bounds even when considering possible compensation from the omitted higher-variance PCs (e.g. see the lower panel of Fig. 4 at z < 0.1). This approach yields top hat priors of width

$$\Delta \alpha_i = \frac{2}{N_{z,PC}} \sum_{j=1}^{N_{z,PC}} |e_i(z_j)|, \tag{8}$$

which follows from Eq. (A10) in MHH.

In analogy to the  $w_0 - w_a$  FoM in the previous section, we base the FoM for dark energy PCs on the determinant of the covariance matrix of  $\alpha_i$  for the *n* lowest-variance PCs,

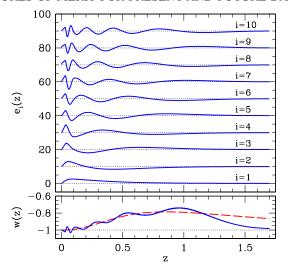


FIG. 4 (color online). Top panel: The first 10 PCs of w(z) (increasing variance from bottom to top), with 500 redshift bins between z=0 and  $z_{\rm max}=1.7$ . The PCs are offset vertically from each other for clarity. Bottom panel: An example of w(z) for a quintessence model [red dashed curve; see Eq. (12)] and its representation using the first 10 PCs in Eq. (7) (blue solid curve). Dotted lines show the w=-1 zero point for each component (top) and for the example model (bottom).

 $\det \mathbf{C}_n$ . Even without informative data, this determinant is finite:

$$\det \mathbf{C}_n^{(\text{prior})} = \prod_{i=1}^n \left(\frac{\Delta \alpha_i}{\sqrt{12}}\right)^2, \tag{9}$$

where the individual factors of  $\Delta \alpha_i / \sqrt{12}$  (listed in Table II) are the rms values of the corresponding top hat priors. These priors impact the FoM in a manner similar to those imposed in Ref. [3], where a Gaussian prior is adopted that requires the rms variation of  $1 + w_i$  for each PC mode, averaged over scale factor, to be no more than unity. However, our priors are slightly stronger since we use the quintessence bounds to impose a top hat prior on 1 + w(z) at all redshifts. In order that the FoM values do

not reflect information that comes exclusively from the quintessence prior, we follow the convention in Ref. [3] and renormalize the detC statistic to obtain

$$\operatorname{FoM}_{n}^{(\operatorname{PC})} \equiv \left(\frac{\det \mathbf{C}_{n}}{\det \mathbf{C}_{n}^{(\operatorname{prior})}}\right)^{-1/2}.$$
 (10)

Table II shows the mean and rms of each of the 10 PCs from MCMC likelihood analysis using the current data and forecasts. Note that for components 6–10 in the current data the rms is dominated by the prior, and likewise for 7–10 in the forecast. Figure 5 shows the one-dimensional posterior probability distributions of the first 6 PCs where the information from the data resides.

Figure 6 (upper panel) shows the FoM for the current data and forecasts. Note that for both data sets the FoM starts to saturate around the sixth PC as expected. The lower panel shows the ratio of FoMs of future and current data. Once again the saturation point is around the sixth PC with the total level of improvement varying from a factor of  $\sim$ 60 to  $\sim$ 140 depending mainly on whether spatial curvature is included.

This dependence of the FoM improvement on curvature is mainly due to a degeneracy between  $\Omega_{\rm K}$  and the first PC which reduces the FoM when marginalizing over  $\Omega_{\rm K}$ , particularly for the forecasts. This degeneracy is largely a consequence of our choice not to marginalize over  $\Omega_{\rm K}$  when constructing the PCs, but the leading degeneracy between dark energy and curvature is included in the first PC. There is a related difference between the flat and nonflat cases in the  $w_0-w_a$  contours in Fig. 3 and FoM $^{(w_0-w_a)}$  in Table I.

Current BAO and  $H_0$  measurements (Sec. II C) constrain  $\Omega_{\rm K}$  enough that current SN + CMB PC uncertainties are affected little by curvature, but for the SN + CMB forecasts we need ~1% measurements of an absolute distance scale to achieve a FoM improvement comparable to that in the flat case. Figure 7 shows an example of breaking the  $\Omega_{\rm K}-\alpha_1$  degeneracy using a 1%  $\langle D_V \rangle$  measurement at z=0.8–1.2 as might be achieved from a future BAO

TABLE II. Top hat prior rms  $\Delta \alpha_i / \sqrt{12}$  and mean and rms of  $\alpha_i$  from current data and forecasts, assuming flat quintessence models without early dark energy.

		Union + WMAP		SNAP + Planck	
PC i	$\Delta \alpha_i / \sqrt{12}$	$ar{lpha}_i$	$\sigma_i$	$ar{lpha}_i$	$\sigma_i$
1	0.375	0.061	0.041	0.011	0.008
2	0.421	0.087	0.132	-0.013	0.037
3	0.428	0.165	0.203	0.011	0.086
4	0.411	0.206	0.278	-0.040	0.141
5	0.450	0.184	0.278	-0.028	0.206
6	0.452	-0.069	0.394	0.053	0.284
7	0.425	-0.071	0.360	0.117	0.343
8	0.454	-0.063	0.436	0.043	0.374
9	0.461	-0.281	0.438	-0.147	0.418
10	0.463	0.026	0.448	-0.003	0.424

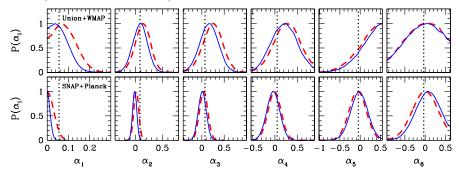


FIG. 5 (color online). Marginalized 1D posterior distributions for the first 6 PCs of flat (solid blue curves) and nonflat (dashed red curves) quintessence models without early dark energy. Marginalizing over the early dark energy parameter  $w_{\infty}$  has little effect on the distributions. Top row: current Union + WMAP data; bottom row: forecasts for SNAP + Planck assuming a realization of the data with  $\alpha_i = 0$ . Plot boundaries that cut off the distributions at nonzero probability correspond to the top hat priors on  $\alpha_i$  for quintessence models. Higher-variance PCs are mainly limited by the quintessence priors for both current constraints and forecasts (see Table II). Vertical dotted lines show the predictions of an example quintessence model from Eq. (12) (see Fig. 4).

experiment. As shown in Fig. 7 of MHH, the shift in absolute distance corresponding to the curvature degeneracy is largely independent of redshift and its elimination could also be achieved with comparable measurements at z = 0–0.1 from improved Hubble constant probes or z = 3 from high-redshift BAO.

The small gap in  $FoM_n^{(PC)}$  ratios in Fig. 6 between models with and without early dark energy is driven by the current constraint on the fifth PC; as shown in Fig. 5,  $P(\alpha_5)$  is not centered on zero and the distribution is cut off by the prior. Including early dark energy shifts this distribution further outside the prior, reducing the PC volume

FIG. 6 (color online). Top panel: PC figures of merit  $FoM_n^{(PC)}$  with forecasted uncertainties for SNAP + Planck and with measured uncertainties for Union + WMAP, normalized as in Eq. (10) to account for the quintessence prior. Bottom panel: Ratios of  $FoM_n^{(PC)}$  forecasts to current values. In both panels, point types indicate different quintessence model classes: flat (solid points) or nonflat (open points), either with (squares) or without (circles) early dark energy.

(and increasing the FoM) for current data. Forecasts for PCs are basically unaffected by early dark energy (see Fig. 6), so the overall effect of early dark energy is a slightly smaller FoM ratio when the fifth PC is included. This particular feature is not generic in the sense that it would not necessarily show up for other choices of data sets or dark energy parametrizations, but similar effects could appear in other analyses where the dark energy priors play an important role.

Since we are using the PCs constructed assuming the forecasted data rather than the current data, we could in principle have strong covariances between the PC

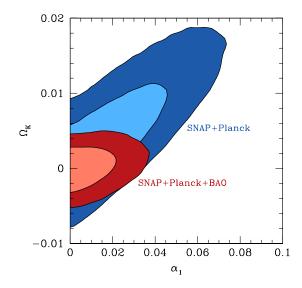


FIG. 7 (color online). Forecasts for  $\Omega_{\rm K}$  and the amplitude of the first PC,  $\alpha_1$ , showing 68% C.L. (light shading) and 95% C.L. (dark shading) regions. The large contours (blue shading) include only SNAP and Planck data as well as the additional priors from Sec. II C. The small contours (red shading) add to these data a 1% measurement of  $\langle D_V \rangle$  averaged over the redshift bin 0.8 < z < 1.2 as might be obtained from a future BAO experiment, reducing the degeneracy between curvature and dark energy.

amplitudes that would be hidden in Fig. 5. Likewise, we could have substantial differences between FoM<sub>n</sub><sup>(PC)</sup> and definitions of the PC FoM involving ratios of  $\sigma_i$  [2,3]. In practice, however, the covariances between PC amplitudes remain small in all cases. Even for current data, the difference between  $(\det \mathbf{C}_n)^{-1/2}$  and  $\prod_{i=1}^n \sigma_i^{-1}$  is  $\leq 20\%$  for flat quintessence models and  $\leq 30\%$  for nonflat models. The effect of covariances on the forecast FoM values is even smaller. Thus the 1D distributions in Fig. 5 accurately depict the current and future constraints on PC amplitudes.

A separate question is whether the use of detC gives misleading results due to non-Gaussianity of the PC posterior distributions (for example, due to distributions being cut off by the quintessence prior). To test the significance of such effects, we consider an alternate FoM analogous to the area-based FoM for  $w_0 - w_a$  models,

$$V_n^{-1} \equiv (2\sqrt{\pi})^n \int d\boldsymbol{\alpha} \mathcal{P}^2(\boldsymbol{\alpha}|\mathbf{x}), \tag{11}$$

where  $\alpha$  is the parameter subset consisting of the first n PC amplitudes. The motivation of this form is that the allowed volume in parameter space is proportional to the inverse of the average number density of MCMC samples. The number density is proportional to  $\mathcal{P}$ , and the averaging over the posterior probability gives another factor of  $\mathcal{P}$ . The normalization of Eq. (11) is chosen so that for an nD Gaussian posterior with covariance  $C_n$ ,  $V_n^{-1} = (\det C_n)^{-1/2}$ .

For the 4–5 lowest-variance PCs, we find good agreement between  $V_n^{-1}$  and  $(\det \mathbf{C}_n)^{-1/2}$  with differences of no more than ~30%. Using the  $V_n^{-1}$  FoM increases the ratio of future to current FoM values by 10%–15%. For n > 5, accurate computation of the integral in Eq. (11) becomes more difficult due to the sparsity of MCMC samples.

Improvements in  $FoM_n^{(PC)}$  do not necessarily represent significant improvements in the ability to limit the parameter space of known dark energy models, especially for the higher PCs [22,23]. In order to address such issues, one can project the predictions of a model for w(z) onto the PC space and examine whether the predictions lie in the volume excluded by forecasted constraints. As an illustrative example, we consider a quintessence model with the potential

$$V(\phi) = V_0 + \frac{1}{2}m^2\phi^2. \tag{12}$$

This model provides examples in the thawing class [56] for  $V_0 \rightarrow 0$  and can also mimic the low-redshift behavior of Albrecht-Skordis models with oscillations around an offset minimum [57]. We consider an example where  $m=7\times 10^{-33}$  eV and  $V_0/\rho_{\rm crit}=0.717$  with  $\Omega_{\rm DE}=0.733$ , h=0.69, and Hubble-drag frozen initial conditions  $\dot{\phi}_i=0$ ; the equation of state for this model is shown in the lower panel of Fig. 4. These parameters are chosen to be allowed by the current data but testable with future data.

Figure 5 compares the predictions for this model with the data constraints. Note that even for the future data only

the first two PCs are stimulated at a level that the data can constrain. Thus for these types of models  $FoM_2^{(PC)}$  is more representative of the parameter volume improvements than higher-dimensional FoMs. Taking into account the quintessence prior, the improvement from current to future data in this case is comparable to that implied by the  $w_0 - w_a$  FoM.

More generally, studies have found that up to 3 PCs are useful for distinguishing amongst different commonly used quintessence models with data sets comparable to our future forecasts [23]. On the other hand, these studies do not preclude the possibility that FoM improvements in the higher components can distinguish between other yet to be investigated classes of models.

#### IV. DISCUSSION

We have compared current and forecasted figures of merit (FoMs) for dark energy using both a simple  $w_0 - w_a$  description of the equation of state, and a more complicated but more complete principal component (PC) parametrization. By consistently using a fixed set of PCs based on future data, and by generalizing the FoM definition to include possible parameter covariance, we showed how PC FoMs can be applied to likelihood analyses of both existing data sets and forecasts.

We have also shown that the covariance-based FoMs accurately represent relative changes in the parameter volume, even in the presence of strongly non-Gaussian posterior distributions such as those caused by imposing top hat priors. Traditional variance-based PC FoMs are consistent with those that include the full covariance for the cases we have tested, but we do not expect such agreement to hold in general. For example, if our PC eigenfunctions from SN and CMB data are applied to qualitatively different data, e.g. weak lensing or BAO, the FoMs that account for the parameter covariances should be employed.

For the  $w_0 - w_a$  FoM, future space-based supernova data and CMB measurements from the Planck satellite can improve on current measurements by a factor of  $\sim$ 10. For the PC FoM, the expected improvement is much greater still. While the forecasted precision for any single principal component is only a factor of 2-5 smaller than current uncertainties, the constraints on general quintessence models bounded by  $-1 \le w \le 1$  improve for approximately 6 components resulting in a reduction of the total allowed volume of dark energy parameter space by a factor of  $\sim$ 60–140. The FoM improvement depends mainly on whether or not variations in spatial curvature are allowed, with the maximum ratio of  $\sim$ 140 requiring either fixing the curvature with a theoretical prior ( $\Omega_{\rm K}=0$ ) or combining future SN and CMB data with a measurement of absolute distance and/or the expansion rate with at least  $\sim$ 1% accuracy.

Although improvements from future SN and CMB data extend to 6 components of w(z), many commonly considered quintessence models are adequately described by the first three PCs for which the allowed volume only decreases by a factor of  $\sim$ 20–40 relative to current measurements. While figures of merit provide a convenient model-independent metric to evaluate experiments, they must be interpreted with caution when science goals include discriminating amongst specific dark energy models.

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