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Grigoroglou, Christos, Munford, Luke, Webb, Roger T et al. (4 more authors) (2018) The association between a primary care pay-for performance scheme and suicide in England: a spatial cohort study. British Journal of Psychiatry. pp. 600-608. ISSN 1472-1465

https://doi.org/10.1192/bjp.2018.143

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The association between a primary care pay-forperformance scheme and suicide in England: a spatial cohort study.

Journal:	BJPsych
Manuscript ID	BJPsych-18-0056.R2
Manuscript Type:	Paper
Date Submitted by the Author:	09-Jun-2018
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Keywords:	Mental Health, Suicide, Quality and Outcomes Framework, financial incentivisation, pay-for-performance
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The association between a national primary care pay-for-performance scheme and suicide rates in England: a spatial cohort study.

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Word count: 4,249 words 3 Figures 2 tables

Abstract

Background: Pay-for-performance policies aim to improve population health by incentivizing improvements in quality of care.

Aim: To assess the relationship between general practice performance on severe mental illness (SMI) and depression indicators under a national incentivisation scheme and suicide risk in England for the period 2006–2014.

Methods: Longitudinal spatial analysis for 32,844 small-area geographical units (Lower Super Output Area), using population-structure adjusted numbers of suicide as the outcome variable. Negative binomial models were fitted to investigate the relationship between spatially estimated recorded quality of care and suicide risk at the LSOA level. Incidence rate ratios (IRRs) were adjusted for deprivation, social fragmentation, prevalence of depression and SMI as well as other 2011 Census variables.

Results: No association was found between practice performance on the mental health indicators and suicide incidence in practice localities (IRR 1·000; 95%CI 0·998, 1·002). IRRs indicated elevated suicide risks linked with area-level social fragmentation (1·030; 95%CI 1·027, 1·034), deprivation (1·013; 95%CI 1·012, 1014) and rurality (1·059; 95%CI 1·027, 1·092).

Conclusion: Primary care has an important role to play in suicide prevention, but we did not observe a link between practices' higher reported quality of care on incentivised mental health activities and lower suicide rates in the local population. It is likely that effective suicide prevention needs a more concerted, multi-agency approach. Better training in suicide prevention for General Practitioners (GPs) is also essential. These findings pertain to the UK but have relevance to other countries considering similar programmes.

Keywords: Mental health, suicide, Quality and Outcomes Framework, financial incentivisation, payfor-performance.

Declaration of Interest:

The authors have no conflicts of interest.

INTRODUCTION

Suicide is a major public health issue in many other countries worldwide, with close to 800,000 people per annum taking their own lives globally. (1) In the UK more than 6000 people die by suicide each year, with tens of thousands more attempting to take their own lives. (2) Primary care physicians have an important potential role in preventing suicide as they are the most common type of healthcare contact for people who experience distress or mental health problems. (3) Most people who die by suicide have contact with a primary care physician shortly beforehand (4-6) with the percentage of people seeing their general practitioners (GPs) prior to death varying from 46% to 91%. (4-8) Furthermore, primary care physicians' failure to screen patients for depression may contribute to under-diagnosis and under-treatment, as well as failure to detect patients experiencing suicidal ideation. Therefore, improving physician recognition of depression and assessment of a person with suicidal thoughts or behaviour are integral components of suicide prevention. (9, 10) Focusing on initiatives to improve the quality of care in primary care can have a big impact on improving the recognition and management of chronic conditions such as depression and severe mental illness (SMI), and can improve the detection of suicidal ideation, (11) especially when combined with coordinated care between GPs and mental health specialists across the healthcare system. However, the impact of primary care on population outcomes is limited by the wide variations that exist in the quality of care of providers. (12) This has led policymakers to link quality of care to provider remuneration in an attempt to reduce variability and to improve quality across health systems. (13)

The Quality and Outcomes Framework (QOF), one of the most costly and comprehensive pay-forperformance schemes of its kind, was introduced in 2004 in the United Kingdom (UK) to reward high quality primary care. (14) The scheme linked up to 25% of GP income to performance on quality indicators across four domains: clinical, organisational, patient experience and additional services. Varying payments were awarded to practices mainly for the ongoing management of chronic conditions, including depression and SMI, with adjustments made for practice list size and disease prevalence. The scheme was periodically reviewed and quality indicators were modified or removed, while new indicators were occasionally introduced for existing or new conditions. (15) In the first three years of its implementation, the QOF yielded improved results for intermediate outcomes and incentivised processes of care but the rate of improvement was not sustained, (16, 17) and performance in later years was no higher than levels projected from pre-intervention trends. (16, 17) In terms of improved patient outcomes, the evidence is contradictory. (18-20) For some conditions there is evidence of lower rates of emergency hospital admission, (18, 21) and reductions in inequalities of delivery of care (22) but there has been no clear improvement in mortality rates in the UK for physical conditions covered by the QOF. (20) Previous research has also found no association between practice-level performance on QOF indicators and all-cause mortality in the locality. (19)

In terms of mental health, the QOF includes indicators for patients with Serious Mental Illness (SMI), requiring practices to monitor treatment, record an up-to-date care plan and to conduct an annual review incorporating physical checks and health promotion advice. For patients with SMI, the

introduction of the QOF was associated with increases in the numbers of service interactions and the number of reported comorbidities. (23) From 2006, the QOF also includes depression indicators requiring practices to assess severity using validated assessment instruments, or indicators requiring practices to review patients within a month of diagnosis and make a biopsychosocial assessment based on 16 'themes' - with one of these being suicidal ideation. These SMI and depression indicators are used to manage the chronic nature of these conditions but also to identify individuals at risk. The complete list and descriptions of all depression and SMI indicators with information on the years they were active included are provided in the online appendix. However, existing evidence on the effectiveness of QOF mental health indicators is limited.

Suicide is an important outcome measure for assessing the quality of mental health services. Effective suicide prevention strategies need to identify individual-level risk factors but also factors associated with the social environment in which an individual lives, (24) as recent evidence indicates marked spatial variability in suicide incidence in England. (25) It is suggested that area of residence can have an impact on levels of social support, but can also increase the likelihood of developing a mental illness, both of which increase suicide risk. (26) Ecological studies have shown that area-level effects such as deprivation and social fragmentation are associated with the geographical distribution of suicide. (27-29)

Most studies investigating determinants of suicide have examined specialist mental health services and suicide (30) or suicide in local and/or regional settings. (31) However, no published studies to date have examined the relationship between performance on QOF mental health indicators and suicide. The purpose of this study was to examine for the first time the relationship between incentivisation of mental health management in primary care and suicide rates. To achieve this aim we applied spatial analysis techniques to assess the association between practice performance on the mental health domain of the QOF and suicide rates for the localities that the practices serve, whilst controlling for relevant small-area population characteristics.

METHODS

Data sources

Several data sources were accessed to extract information on numbers of suicides, population size estimates, rurality, social fragmentation, deprivation, quality of care, disease prevalence and spatial coordinates at the Lower Super Output Area (LSOA) level. LSOAs are small geographical areas containing approximately 1500 residents. From the Office for National Statistics (ONS) we obtained data on case counts for suicides and open verdicts (32) per LSOA of residence by year of death, broken down by gender and 5-year age bands , for the period 1st January 2006 to 31st December 2014. Our population at risk was people aged 20 years or older as we wished to exclude children and adolescents from our analysis because of differences in the remit of mental health services and

depression management provided to kids, adolescents and adults. (33) Classification of deaths for the study used the International Classification of Diseases, 10th edition (ICD-10) and included all deaths with a final underlying cause recorded as intentional self-harm, injury/poisoning of undetermined intent and sequelae of intentional self-harm, or injury/poisoning of undetermined intent, which follows conventional practice for government suicide statistics in the UK. We included both suicide and open verdicts in our outcome measure and we refer to this measure as suicides throughout the manuscript. Population counts by gender and age group, as well as information on ethnicity were obtained from the 2011 national Census, at the 2011 LSOA.

Area-level deprivation, as measured by the Index of Multiple Deprivation 2007, was available for LSOAs as at 2011. The IMD measures relative levels of deprivation for all the 32,844 LSOAs in England on a continuous scale of deprivation where most of the indicators are based on 2005 statistics. Moreover, we obtained data on rural/urban classification and we constructed a measure of social fragmentation based on Congdon's index (derived from 2011 Census data on private renting, single person households, single adults and population turnover). The index represents neighbourhood-level conditions which affect social resources such as social cohesion and social capital. (27) Datasets on prevalence of SMI and depression, practice list size and practice-level mental health quality of care data were obtained from NHS Digital. As 80% of LSOAs do not contain a general practice we used spatial analysis techniques to estimate healthcare data at the LSOA level from the practice-level data (a detailed description of this methodology is provided in the supplementary appendix).

The number of Primary Care practice 'hubs' (an LSOA with one or more practices) varied between 8372 in the third year of the QOF (2006-07) and 7778 in the final year (2014-15) including for each year over 99% of all patients registered with primary care in England. (34) From 2006-2007 onwards, all relevant information was available; i.e. number of practices, total list size, and prevalence rates for those over 18 years old in the two relevant QOF clinical domains: SMI and depression.

To better capture quality of care in the mental health domain, as measured by the QOF, we calculated population achievement (17, 19), defined as $PAmh = \frac{\sum N_i}{\sum (Di + Ei)}$

where N_i , Di and Ei denote the numerator, denominator and exceptions for QOF indicator *i*, respectively. Population achievement was calculated, separately within each financial year, across the full set of depression and SMI outcome and measurement QOF indicators in the mental health clinical domain.

Statistical analysis

We fitted a negative binomial model with standard errors clustered at the LSOA level for global between-area variability and/or local variability in neighbouring areas to derive and map age and sex adjusted area estimates of suicide. Negative binomial models report Incidence Rate Ratios (IRR) and

are suitable for analysing count data when there is presence of over-dispersion, as was the case in our data.

We used one set of negative binomial regression models to investigate the relationship between mental health quality of care, as measured by the QOF, and suicide incidence over time. The dependent variable was annual number of suicides between 2006 and 2014, and we investigated the relationship with the mental health quality of care measure, in the same or up to a lag of three years. The same sets of covariates were used for all models. We also conducted several sensitivity analyses under different assumptions. There was an excess of zero event counts in the study dataset; thus, more than 99% of data points (LSOAs by age group and sex, over time) did not include a single suicide. Therefore, we investigated how a zero inflated negative binomial model fitted the data. This model may intuitively seem to provide a better fit to the data, but the standard negative binomial model is more appropriate from a theoretical perspective, as every population group is 'at risk' of experiencing suicide. As an additional sensitivity analysis, we aggregated suicide numbers and population estimates in four age groups (20-24, 25-49, 50-64, 65+), according to ONS classification, to address any issues of collinearity between age groups. The economic recession of 2008 occurred in the middle of our study period; thus, to isolate the impact of the QOF on suicide risk from the effects of the economic recession, we ran a sensitivity analysis for the period 2009-2014. Furthermore, conventional practice in the UK is to include open verdicts when investigating and reporting suicide rates. However, to facilitate international comparison, we analysed open verdicts and definite suicide cases separately. Finally, we investigated the extent to which spatial autocorrelation was present in the data (i.e. level of suicide risk in one locality affecting risk in an adjacent locality) by calculating Moran's I, which is a measure of spatial autocorrelation.

Stata v14.1 was used for the principal data management and analyses. For both sets of analyses, we used the *nbreg command* with the exposure option followed by the *IRR* specification while we used clustered standard errors for our unit of analyses (i.e. 32,844 LSOAs) to take into account the correlation of observations within each LSOA.

RESULTS

Figure 1 represents the variability in incidence of suicide across regions over the study period. The North East and North West regions had the highest overall suicide rates whereas the London region had the lowest overall suicide rates across the period of study. There were 38,511 recorded suicides in England during calendar years 2006 to 2014, inclusive. Figure 2 represents the spatial distribution of suicide over the whole time period (aggregated), at the Clinical Commissioning Group (CCG) level, a middle-level organisation structure of the National Health Service (NHS). Figure 3 represents the spatial variability of quality of care in the mental health domain across LSOAs at the census year i.e. 2011. Descriptive statistics on suicide numbers, population size estimates, number of practices, census information, deprivation, prevalence of depression and SMI, and mental health quality of care across regions (Strategic Health Authorities - SHAs) for 2011 are reported in Table 1. The highest prevalence for both depression and SMI was observed in the North East region whereas the

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lowest prevalence for depression and SMI was found in the London and South Central regions respectively.

We found no association between mental health quality of care and suicide in our main model (IRR 1.000; 95%CI 0.998, 1.002). There was also no association after lagging the effects of mental health quality of care for up to three years. This finding was consistent in all sensitivity analyses, including the analysis limiting the study period to 2009-2014 (post 2008 recession). In the sensitivity analysis that investigated definite suicide deaths, results were similar (IRR 1.000; 95%CI 0.998, 1.002). All results from the sensitivity analyses are reported in the supplementary appendix.

No evidence of heterogeneity in the association of interest was found, by health region or deprivation score. Median QOF mental health quality of care across regions varied between 80.4% (East Midlands) to 83.0% (North East). As expected, variability was much greater at the LSOA level, partly due to the smaller denominators. For example, in the first analysis year (2006-07) performance varied from 12.9% to 99.2% (10th and 90th centiles were 71.3% and 94.7% respectively), in the middle year (2009-10) performance varied from 20.6% to 94.5% (10th and 90th centiles were 75.7% and 88.8% respectively) and finally for the last year (2014-15) it varied from 28.3% to 94.2% (10th and 90th centiles were 68.6% and 86.7% respectively).

Table 2 presents IRRs for the coefficients from the negative binomial regression models. Across all age groups, the highest risk for suicide was observed for the 40-44 group (1.865; 95%CI 1.775, 1.960) while a bimodal distribution was observed with peaks in the mid years and those over 85 years (1.537; 95%CI 1.425, 1.657) thus indicating a non-linear relationship of suicide and age. Female sex (0.295; 95%CI 0.288, 0.302) and social fragmentation (1.050; 95%CI 1.044, 1.056) were also strong predictors of suicide across all models, followed by deprivation (1.015; 95%CI 1.014, 1.016). Rurality (1.035; 95%CI 1.004, 1.066) was also a strong predictor of suicide risk, where inhabitants in rural areas were found to be more likely to die by suicide.

DISCUSSION

After adjusting for key social and demographic characteristics, we found no association between primary practice performance on incentivized mental health indicators and suicide rates in practice localities. Although depression and SMI are common among people who die by suicide, some people may not experience these symptoms whilst others may not be diagnosed by their GPs. (35) Not all people with mental illness see their GP in the year before death occurs (4) and not all people with mental illness will be assessed or treated. (5) Nevertheless, GP based efforts can have positive results in the reduction of suicide rates sometimes associated with an increase in antidepressant prescription. (36) One of the aims of incentivisation was to achieve better quality of care for people diagnosed with mental illnesses and, although suicide is rare, we might have expected to have observed small effects for a successful programme in that respect. For a dataset of this size, statistical significance is a secondary concern, as we should have been able to identify even a tiny effect - but there appeared to be none. Moreover, QOF mental health 'quality of care' scores do not inform as to care received in settings outside general practice. For example, a 28% of people who die

by suicide in the UK have had contact with mental health services in the previous 12 months (37), however, as this study focuses on aspects of mental health care in general practice we did not account for the important role of mental health and emergency services in preventing suicide. If local mental health services had been consistently beneficial and primary care services poor or vice versa, then the effect of the quality of primary care would be mitigated by the unknown effect of the quality of specialist mental healthcare in the same locality. Finally, even if high quality of mental health provision in primary care, as incentivised through the QOF, has beneficial impacts for patients overall this may not be apparent for important but relatively low frequency markers of population health such as death by suicide.

Our regression models have also indicated that area-level socioeconomic factors, such as social fragmentation and rurality, rather than variations in the quality of primary care, are more strongly associated with suicide risk. We identified a strong relationship between levels of deprivation and suicide from 2006 to 2014. One additional percentage point in population deprivation is associated with a 1.5% (n=61) increase in suicides nationally in one year. A change in IMD score from the 25th centile (9.72) to the median (17.23) would correspond to an 11.2% increase in suicide rates (n=458). The extent of the relationship between deprivation and suicide is also highly consistent across the sensitivity analyses reported in the supplementary appendix. Social fragmentation was a strong predictor of suicide, where one additional percentage point in social fragmentation is linked to 5.3% increase in suicide (n=215) nationally in one year. These findings agree with previous evidence on the effects of deprivation (27) and social fragmentation (27, 28) and also agree with previous evidence which showed that social fragmentation has a stronger association with suicide risk than deprivation. (28)

Additionally, we observed a moderate association between the prevalence of SMI and suicide across all models and all sensitivity analyses but small or no associations between prevalence of depression and risk of suicide. As this result may seem unusual we performed a univariate analysis for the associations between prevalence of SMI and depression with risk of suicide. The results from the univariate analysis indicate that prevalence of SMI has a strong effect on suicide (1·476; 95%CI 1·361, 1·601) while prevalence of depression has a very small but statistically significant effect (1·009; 95%CI 1·006, 1·012). The fact that unresolved depression register is weakly associated with suicide, despite the well documented association between depression and suicide, may imply that the condition is not captured well in primary care. Misidentifications and missed cases in depression diagnosis are common in primary care. (38) Moreover, it is suggested that GP's who work in socioeconomically deprived areas identify the ongoing management of depression as a burden and they may be reluctant to investigate psychosocial issues with their patients due to structural factors and workload. (39)

This analysis was conducted at the population level and used a novel spatial estimation approach to assess and quantify the relationship between mental health quality of care and suicide rates for the whole primary care-registered population of England. This analysis allows for careful control of population characteristics at a low-level geographical area (LSOA), which is not possible with analyses at the general practice-level. We avoided standardising our outcome variable as standardised death by suicide ratios are imprecise at this low geographical (LSOA) level. Instead we adjusted our analysis for population structure by including all the relevant information on age and

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sex. This level of analysis was deemed to be optimal as aggregating data at higher geographical levels would introduce heterogeneity and confounding across small-area populations.

However, our study has a number of limitations. First, the possibility of ecological fallacy cannot be ruled out as the associations we observed at small area level may not pertain to individual patients. Second, in essence we equated performance on the QOF mental health domain to quality of care for mental health, which is not necessarily the case. However, in the absence of other quantifiable metrics, this has become an accepted practice, although the QOF fails to capture other aspects of consultations that may be even more important. Third, the IMD is by definition an aggregate measure of income, employment, health, education, crime and environmental deprivation in a locality. Thus, it is always collinear with measures of poor health and income, and these were necessarily excluded from our analysis. Fourth, we could not access population measures of ethnicity for all age groups and we included this variable as '% White' at LSOA level. Fifth, we were unable to attribute practice performance to LSOAs for the first two years of the scheme (2004/2005 and 2005/2006), the period during which practices exhibited the largest improvements in performance, (22) due to data on mental health quality of care being unavailable. However, there was sufficient variation in quality of care at the LSOA level from 2006/7 to discern any association with suicide risk should it exist. Sixth, any effects that the scheme might have had on suicide may have been delayed and the time window between mental health quality of care and suicide (i.e. 3 year lag model) could be too narrow. Nonetheless, analysis with longer lag periods is subject to methodological issues, whilst it seems clinically implausible that the indicators could have long-term preventive effects without short- or medium-term effects. Seventh, some degree of spatial autocorrelation was identified, but considering the very low value of the measure (Moran's I = 0.0049), we did not account for spatial autocorrelation in the model. Even if spatial clustering was present, we would expect it to be driven by socio-economic factors we controlled for, and not by suicide cases in neighbouring areas which would warrant the use of spatial regression techniques. Eighth, multimorbidity and psychosocial risk factors for suicide (e.g. domestic abuse or substance misuse) may occur concomitantly with depression especially in deprived areas, (40) however, it is likely that these cases may not be captured adequately in the QOF (15) and our measure of guality of mental healthcare as some of them will remain unrecorded. Ninth, recording of depression in English primary care has been described as being problematic, as some GPs prefer to record symptoms rather than formal psychiatric diagnostic labels, (41) but we believe this is unlikely to have affected the validity of our findings in relation to the primary research question. Finally, as death by suicide is used as an underlying indicator in the health and deprivation domain of the IMD, we used the IMD 2007 which was calculated from 2005 data on suicide deaths. Nevertheless, the various IMD versions have been documented to be highly correlated and thus the use of an old version of IMD should impose no implications to our findings.

Our study adds to a growing body of evidence on the potential impact of the QOF on patient outcomes, and is the first study to assess whether quality of primary care, measured as a proxy via metrics generated from a national incentive scheme, is linked with subsequent suicide risk at the population level. Our findings indicate that the scheme appears to have had no impact on suicide risk through the treatment and monitoring of depression and SMI, although population benefits in

terms of improved quality of life and reductions in unnatural mortality risk may accrue in the longer term.

Suicide is a complex phenomenon and many factors may influence a person's decision to take their own life. In the absence of indicators related to suicidality and suicidal thoughts and acts, recognition of patients at risk of suicide depends entirely on the GP's ability to identify and interpret relevant signs. This may be extremely difficult without sufficient training as suicidal thoughts may be communicated differently across people at risk. This adds to concerns from GPs regarding the limited training that is provided to them regarding their role in preventing suicide. (42) Furthermore, the first few weeks following discharge from psychiatric inpatient care are characterised by increased suicide risk for patients (37) and GPs stress the need to improve communication and information sharing between services, particularly in the cases of hospital discharge and transition between services. (5) It is vital to establish a collaborative model between GPs, mental health services, community care and hospitals, where improved communication and information sharing will be priorities. However, suicide prevention has never had a higher profile in the UK than at present. The Refreshed Suicide Prevention Strategy, the Health Select Committee Inquiry into Suicide Prevention and the NHS 5-year Forward Review provide a target of a 10% reduction in the national suicide rate. It is clear that effective suicide prevention requires a more concerted, multiagency approach, including better training in suicide prevention for primary care clinicians.

It is widely recognised that primary care has an important role to play in suicide prevention, but we found no association between practices' reported quality of care on mental health activities incentivised in the QOF and suicide rates in the local population. The QOF was a major national experiment, dwarfing in scale all previous attempts to incentivise high quality care, but doubts on its impact and value mean that the UK is now retreating from using incentives in primary care, led by Scotland's decision to withdraw from the QOF in 2017. Other countries continue to use incentives to improve population health, however, hoping to draw on the experiences of earlier schemes to develop more effective frameworks. If these attempts are to succeed, attention needs to be focussed on those activities that contribute to improved patient outcomes, such as reductions in risk of dying prematurely from suicide, and those that will more accurately capture higher value care. More broadly, a better understanding is needed as to how resources can best be allocated to achieve higher quality of services that will lead to genuine improvements in population health outcomes.

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Acknowledgments

We would like to thank the Office of National Statistics and NHS Digital (formerly the Health and Social Care Information Centre) for the wealth of information they have collected and systematically organised, which made this study possible.

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Funding

This study is funded by the National Institute for Health Research School for Primary Care Research (NIHR SPCR), through CG's PhD. This report is independent research by the National Institute for Health Research. LM acknowledges financial support from the MRC Skills Development Fellowship (MR/N015126/1). The views expressed in this publication are those of the authors and necessarily those of the NHS, the National Institute for Health Research, the Department of Health, or the MRC.

Authorship & contributorship

EK and CG designed the study. CG extracted the data from all sources and performed the analyses. CG and wrote the manuscript. LM, RW, NK, TD, DA, EK critically revised the manuscript. CG is the guarantor of this work and, as such, had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Transparency declaration

CG affirms that this manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained.

Data sharing

The data used in this study are freely available and the authors are happy to share in an organised and cleaned final dataset.

Ethical Approval

No ethical approval was needed as no patients were involved and all data were publicly available except for the ONS Mortality Data.

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to per period

Strategic Health Authority*										
Aggregates	North	North	Yorkshire	East	West	East	London	South	South	South
88 8	East	West	Humber	Midlands	Midlands	England		West	Central	East
						8		Coast		
Population size	1.998.1 71	5.365.75 5	4.011.205	3.461.393	4.217.657	4.465.027	6.195.926	4.107.71 4	3.158.651	3.415.10 8
Suicide counts+	265	654	448	356	418	470	571	499	324	374
No. of practices	404	1248	787	626	956	785	1473	623	502	718
Census	Medians a	across 2011 L	SOAs	•		•				
Information [‡]										
IMD 2007 (mean	28.1	27.7	25.0	20.0	24.8	15.3	26.0	17.7	12.7	14.9
across LSOAs)	-									
P10	8.2	7.1	6.9	5.5	7.0	4.5	8.8	5.7	3.5	3.8
P25	13.7	11.7	10.9	8.8	11.3	7.4	14.5	9.2	5.7	6.9
P75	39.7	41.0	36.8	27.5	35.9	20.1	36.0	22.6	18.9	19.8
P90	53.2	57.6	52.1	41.5	50.5	30.2	45.6	33.4	28.6	29.7
Social	-0.3	-0.2	-0.7	-0.2	-0.4	-0.2	0.7	0.2	0.1	0.0
Fragmentation										
(mean across										
LSOAS)	1.9	1.7	1.6	1.6	1.7	1.5	1.4	1.2	1.4	1.5
D25	-1 0	-1 /	-1 0	-1.1	-1 /	-1 5	-1 4	-1 3	-1 4	-1 5
P75	-1.3	-1.5	-1-1	-1.1	-1.3	-1.1	-0"/	-0.9	-1.0	-1.1
P00	1.6	1.7	2.0	2.0	-0 0	0.2	2.5	-0 0	2.2	2.
00E Information: r	1 0	1 /	2.0	2.0	1 2	10	55	24	22	2
List Size		8222	8026	8803	8006	9100	7632	0063	10376	0075
OOF information (s	7410	mated): medi	ans across 2011 L	SOAs	8090	9100	7032	9005	10370	9075
Severe Mental		0.8	0.7	0.7	0.7	0.7	0.9	0.7	0.6	0.7
Illness %	0.0	0.0	0 /		0 /	0 /	0 /	0,	0.0	0,
P10	0.6	0.6	0.5	0.5	0.5	0.5	0.6	0.4	0.4	0.4
P25	0.7	0.7	0.6	0.5	0.6	0.5	0.7	0.5	0.5	0.5
P75	0.9	1.0	0.9	0.8	0.8	0.8	1.1	0.8	0.8	0.8
P90	1.0	1.1	1.0	1.0	1.0	0.9	1.3	1.0	0.9	1.0
Depression %	11.9	10.6	9.5	10.0	9.0	9.1	6.2	10.1	9.2	8.9
P10	8.7	7.0	5.5	6.5	5.9	5.5	4.0	7.1	6.0	6.1
P25	10.2	8.8	7.2	7.9	7.5	7.1	5.0	8.5	7.5	7.5
P75	14.8	12.3	11.4	11.7	10.8	11.0	7.7	11.9	11.1	10.7
P90	16.9	13.8	13.0	13.5	12.6	12.8	9.0	13.5	12.3	12.5
% Mental Health	83.0	81.6	81.2	80.4	82.7	80.7	81.80	82.3	82.0	80.6
Quality of Care										
						6				
P10	78.6	76.2	75.2	74.8	77.0	74.37	76.3	75.5	75.8	74.0
P25	80.8	79.0	78.5	77.7	80.1	78.16	79.2	79.4	79.1	77.4
P75	84.9	84.2	83.8	83.2	84.5	83.48	84.1	84.7	84.2	83.4
P90	86.2	86.0	85.6	85.4	86.0	85.58	85.7	86.4	85.9	85.5
Other Information A	Across LSOA	As	·		•	-			•	
Rural (%)	17.5	9.8	16.5	25.4	14.8	28.1	0.1	20.1	19.7	30.2
QOF Information: n List Size QOF information (s Severe Mental Illness % P10 P25 P75 P90 Depression % P10 P25 P75 P90 % Mental Health Quality of Care P10 P25 P75 P90 % Mental Health Quality of Care P10 P25 P75 P90 Other Information / Rural (%)	nedians acro 9410 patially estin 0.8 0.6 0.7 0.9 1.0 11.9 8.7 10.2 14.8 16.9 83.0 78.6 80.8 84.9 86.2 Across LSO/ 17.5	0.8 0.6 0.7 1.0 1.1 10.6 7.0 8.8 12.3 13.8 81.6 76.2 79.0 84.2 86.0 As 9.8	100 8926 ans across 2011 L 0.7 0.5 0.6 0.9 1.0 9.5 5.5 7.2 11.4 13.0 81.2 75.2 78.5 83.8 85.6 16.5	8893 SOAs 0.7 0.5 0.8 1.0 10.0 6.5 7.9 11.7 13.5 80.4 74.8 77.7 83.2 85.4 25.4	8096 0.7 0.5 0.6 0.8 1.0 9.0 5.9 7.5 10.8 12.6 82.7 77.0 80.1 84.5 86.0 14.8	9100 0.7 0.5 0.5 0.8 0.9 9.1 5.5 7.1 11.0 12.8 80.7 74.37 78.16 83.48 85.58 28.1	7632 0·9 0·6 0·7 1·1 1·3 6·2 4·0 5·0 7·7 9·0 81·80 76·3 79·2 84·1 85·7 0·1	9063 0.7 0.4 0.5 0.8 1.0 10.1 7.1 8.5 11.9 13.5 82.3 75.5 79.4 84.7 86.4 20.1	$ \begin{array}{c} 10376 \\ \hline 0.6 \\ \hline 0.4 \\ 0.5 \\ 0.9 \\ 9.2 \\ \hline 6.0 \\ 7.5 \\ 11.1 \\ 12.3 \\ 82.0 \\ \hline 75.8 \\ 79.1 \\ 84.2 \\ 85.9 \\ \hline 19.7 \\ \hline 19.7 \\ \hline $	9075 0.7 0.4 0.5 0.8 1.0 8.9 6.1 7.5 10.7 12.5 80.6 74.0 77.4 83.4 85.5 30.2

Table 1: Area and population characteristics by strategic health authority*++□

*Strategic health authorities were the highest geographical organisational level for NHS England during the study period. This structure was abolished in 2013.

⁺Suicides and open verdicts in 2011.

‡Data available from the 2011 census.

□ Quality and Outcomes Framework (QOF) year7 (2010-2011).

* Mental Health Quality of Care is measured as % achievement of the population in the relevant QOF indicators.

P10, P25, P75, P90 are the respective percentiles for each variable.

8	U	Binomial Model)	. <i>V</i>	
Year:	Lag=0	Lag=1	Lag=2	Lag=3
% Mental Health Quality of Care *	1.000 (0.998, 1.002),	1.000 (0.998, 1.003),	1.000 (0.998, 1.003),	1.000 (0.997, 1.003),
	<0.537 (0.000)	<0.569 (0.001)	<0.971 (0.001)	<0.715 (0.001)
Female	0·295 (0·288, 0·302),	0·293 (0·286, 0·301)	0·293 (0·285, 0·301)	0·293 (0·284, 0·301)
	<0·001 (0·003)	<0·001 (0·003)	<0·001 (0·003)	<0·001 (0·004)
Index of Multiple	1.013 (1.012, 1.014),	1.013 (1.012, 1.014),	1.013 (1.012, 1.014),	1.013 (1.012, 1.014),
Deprivation 2007	<0.001 (0.0003)	<0.001 (0.0004)	<0.001 (0.0004)	<0.001 (0.0004)
Rural (v urban)	1.059 (1.027, 1.092), <0.001 (0.016)	$\begin{array}{c} 1.065 \ (1.032, \ 1.099), \\ < 0.001 \ (0.016) \end{array}$	1.065 (1.030, 1.102), <0.001 (0.018)	1.053 (1.016, 1.091), <0.005 (0.019)
Index of Social Frag.	1.030 (1.027, 1.034), <0.001 (0.001)	$\begin{array}{c} 1 \cdot 031 \ (1 \cdot 027, \ 1 \cdot 035), \\ < 0 \cdot 001 \ (0 \cdot 003) \end{array}$	1.031 (1.027, 1.035), <0.001 (0.003)	1.031 (1.027, 1.035), <0.001 (0.002)
Ethnicity (%White)	1.007 (1.006, 1.008),	1.007 (1.006, 1.008),	1.007 (1.006, 1.008),	1.008 (1.007, 1.009),
	<0.001 (0.000)	<0.001 (0.0004)	<0.001 (0.0004)	<0.001 (0.000)
Prevalence of SMI	1.054 (1.004, 1.108),	1.049 (0.995, 1.105),	1.038 (0.983, 1.096),	1.040 (0.982, 1.101),
	<0.033 (0.026)	<0.071 (0.028)	<0.169 (0.028)	<0.173 (0.030)
Prevalence of Depression	1.000 (0.996, 1.004),	0·999 (0·995, 1·003),	0·999 (0·995, 1·003),	0.998 (0.994, 1.003),
	<0.803 (0.019)	<0·881 (0·019)	<0·760 (0·021)	<0.648 (0.022)
Age (20-24)	Reference Age Group	Reference Age Group	Reference Age Group	Reference Age Group
Age (25-29)	1.162 (1.101, 1.226),	1 · 147 (1 · 084, 1 · 214)	1 · 143 (1 · 077, 1 · 214)	1·146 (1·074, 1·223)
	<0.001 (0.031)	<0 · 001 (0 · 033)	<0 · 001 (0 · 035)	<0·001 (0·038)
Age (30-34)	1·360 (1·290, 1·434)	1·339 (1·267, 1·416)	1·300 (1·225, 1·380)	1·281 (1·201, 1·366)
	<0·001 (0·036)	<0·001 (0·038)	<0·001 (0·039)	<0·001 (0·042)
Age (35-39)	1·701 (1·616, 1·790)	1.681 (1.593, 1.774)	1.657 (1.564, 1.754)	1.660 (1.560, 1.767)
	<0·001 (0·043)	<0.001 (0.046)	<0.001 (0.048)	<0.001 (0.052)
Age (40-44)	1.876 (1.785, 1.972),	1.880 (1.784, 1.981)	1.883 (1.781, 1.990)	1·901 (1·790, 2·019)
	<0.001 (0.047)	<0.001 (0.050)	<0.001 (0.052)	<0·001 (0·058)
Age (45-49)	1.846 (1.755, 1.942),	1.839 (1.744, 1.939)	1.852 (1.750, 1.959)	1.891 (1.780, 2.009)
	<0.001 (0.047)	<0.001 (0.049)	<0.001 (0.053)	<0.001 (0.058)
Age (50-54)	1.841 (1.749, 1.939), <0.001 (0.048)	1.835 (1.738, 1.937) <0.001 (0.050)	1.843 (1.741, 1.952) <0.001 (0.053)	$\begin{array}{c} 1.882 \ (1.769, 2.001) \\ < 0.001 \ (0.059) \end{array}$
Age (55-59)	1.621 (1.536, 1.711),	1.604 (1.515, 1.698)	1.607 (1.513, 1.707)	1.631 (1.527, 1.741)
	<0.001 (0.044)	<0.001 (0.046)	<0.001 (0.049)	<0.001 (0.054)
Age (60-64)	1·273 (1·202, 1·349),	1·255 (1·181, 1·333)	1·241 (1·164, 1·322)	1·259 (1·176, 1·349)
	<0·001 (0·036)	<0·001 (0·038)	<0·001 (0·040)	<0·001 (0·044)
Age (65-69)	1.026 (0.963, 1.093),	1.008 (0.943, 1.078)	0.992 (0.924, 1.065)	0·994 (0·920, 1·073)
	<0.420 (0.032)	<0.803 (0.034)	<0.840 (0.035)	<0·878 (0·038)
Age (70-74)	0.999 (0.933, 1.070),	0.988 (0.919, 1.063)	0.997 (0.923, 1.076)	0.995 (0.916, 1.081)
	<0.993 (0.034)	<0.762 (0.036)	<0.939 (0.038)	<0.919 (0.041)
Age (75-79)	1.070 (0.996, 1.149),	1.051 (0.974, 1.133)	1.037 (0.957, 1.123)	1.053 (0.965, 1.148)
	<0.062 (0.038)	<0.197 (0.040)	<0.372 (0.042)	<0.242 (0.046)
Age (80-84)	1·351 (1·254, 1·455),	1·326 (1·226, 1·435)	1·284 (1·180, 1·396)	1·271 (1·160, 1·392)
	<0·001 (0·050)	<0·001 (0·052)	<0·001 (0·055)	<0·001 (0·059)
Age (85plus)	$\begin{array}{c} 1\cdot 551 \ (1\cdot 438, 1\cdot 673) \\ < 0\cdot 001 \ (0\cdot 059) \end{array}$	1.518 (1.402, 1.635) <0.001 (0.061)	1.505 (1.383, 1.638) <0.001 (0.065)	1.529 (1.396, 1.674) <0.001 (0.070)
2006	Reference Year	Reference Year	Reference Year	Reference Year
2007	0·940 (0·891, 0·992), <0·025 (0·025)	-	-	-
2008	0.999 (0.946, 1.054), <0.971 (0.027)	1.060 (1.014, 1.108), <0.010 (0.23)	-	-
2009	$\begin{array}{c} 1 \cdot 012 \ (0 \cdot 953, \ 1 \cdot 074), \\ < 0 \cdot 691 \ (0 \cdot 030) \end{array}$	$\begin{array}{c} 1.076 \ (1.027, \ 1.126), \\ < 0.002 \ (0.025) \end{array}$	1.076 (0.966, 1.057), <0.627 (0.022)	-
2010	0·961 (0·903, 1·022),	1.025 (0.978, 1.073),	1.025 (0.919, 1.008),	0·949 (0·908, 0·993),
	<0·210 (0·030)	<0.291 (0.024)	<0.108 (0.023)	<0·023 (0·021)
2011	1.018 (0.956, 1.084), <0.571 (0.032)	$ \begin{array}{c} 1.085 \ (1.034, \ 1.138), \\ < 0.001 \ (0.025) \end{array} $	$ \begin{array}{c c} 1.085 & (0.977, 1.071), \\ < 0.317 & (0.025) \end{array} $	1.009 (0.966, 1.054), <0.663 (0.022)

Table 2 - Regression Analysis set 1: Effect of Mental Health QOF Quality of Care on suicide over time (Negative Binomial Model)

Binomial Model)						
Year:	Lag=0	Lag=1	Lag=2	Lag=3		
2012	1·014 (0·966, 1·065), <0·556 (0·025)	$\begin{array}{c} 1 \cdot 080 \ (1 \cdot 031, \ 1 \cdot 130), \\ < 0 \cdot 001 \ (0 \cdot 024) \end{array}$	1.080 (0.969, 1.062), <0.526 (0.023)	0.999 (0.951, 1.049), <0.990 (0.025)		
2013	1.056 (1.002, 1.112),	1·127 (1·070, 1·184)	1·127 (1·009, 1·113)	1.045 (0.993, 1.100)		
	<0.039 (0.027)	<0·001 (0·027)	<0·020 (0·026)	<0.088 (0.027)		
2014	1.076 (1.020, 1.134),	1·158 (1·101, 1·215)	1·158 (1·040, 1·148)	1·074 (1·019, 1·131)		
	<0.006 (0.029)	<0·001 (0·028)	<0·001 (0·027)	<0·007 (0·028)		
Model intercept	0.00004 (0.00003, 0.00005),	0·00003 (0·00002, 0·00004),	0.00003 (0.00002, 0.00004),	0.00003 (0.00002, 0.00004),		
	<0.001 (0.00001)	<0·001 (0·00001)	<0.001 (0.00001)	<0.001 (0.00001)		

Table 2 - Repression Analysis set 1: Effect of Mental Health OOF Quality of Care on suicide over time (Negative

95% Confidence Intervals are in brackets, Results are reported as incidence rate ratios (IRR) followed by Pvalues and Standard Errors in parentheses.

*Mental Health Quality of Care is measured as % achievement of the population in the relevant QOF indicators.

BJPsych

Figure 1



Liez

Figure 1: Suicide rates, per 100,000 population, by English Regions over time (2006-2014).

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Figure 2: Suicide rates per 100,000 population in England aggregated across all years (2006 – 2014) at the CCG level.



Figure 3: Measure of Mental Health Quality of Care in England for specific mental health indicators at the 2011 LSOAs. (2011 Census Year).

The association between a national primary care pay-for-performance scheme and suicide rates in England: a spatial cohort study.

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Word count: 4,249 words 3 Figures 2 tables

Abstract

Background: Pay-for-performance policies aim to improve population health by incentivizing improvements in quality of care.

Aim: To assess the relationship between general practice performance on severe mental illness (SMI) and depression indicators under a national incentivisation scheme and suicide risk in England for the period 2006–2014.

Methods: Longitudinal spatial analysis for 32,844 small-area geographical units (Lower Super Output Area), using population-structure adjusted numbers of suicide as the outcome variable. Negative binomial models were fitted to investigate the relationship between spatially estimated recorded quality of care and suicide risk at the LSOA level. Incidence rate ratios (IRRs) were adjusted for deprivation, social fragmentation, prevalence of depression and SMI as well as other 2011 Census variables.

Results: No association was found between practice performance on the mental health indicators and suicide incidence in practice localities (IRR 1.000; 95%CI 0.998, 1.002). IRRs indicated elevated suicide risks linked with area-level social fragmentation (1.030; 95%CI 1.027, 1.034), deprivation (1.013; 95%CI 1.012, 1014) and rurality (1.059; 95%CI 1.027, 1.092).

Conclusion: Primary care has an important role to play in suicide prevention, but we did not observe a link between practices' higher reported quality of care on incentivised mental health activities and lower suicide rates in the local population. It is likely that effective suicide prevention needs a more concerted, multi-agency approach. Better training in suicide prevention for General Practitioners (GPs) is also essential. These findings pertain to the UK but have relevance to other countries considering similar programmes.

Keywords: Mental health, suicide, Quality and Outcomes Framework, financial incentivisation, payfor-performance.

Declaration of Interest:

The authors have no conflicts of interest.

INTRODUCTION

Suicide is a major public health issue in many other countries worldwide, with close to 800,000 people per annum taking their own lives globally. (1) In the UK more than 6000 people die by suicide each year, with tens of thousands more attempting to take their own lives. (2) Primary care physicians have an important potential role in preventing suicide as they are the most common type of healthcare contact for people who experience distress or mental health problems. (3) Most people who die by suicide have contact with a primary care physician shortly beforehand (4-6) with the percentage of people seeing their general practitioners (GPs) prior to death varying from 46% to 91%. (4-8) Furthermore, primary care physicians' failure to screen patients for depression may contribute to under-diagnosis and under-treatment, as well as failure to detect patients experiencing suicidal ideation. Therefore, improving physician recognition of depression and assessment of a person with suicidal thoughts or behaviour are integral components of suicide prevention. (9, 10) Focusing on initiatives to improve the quality of care in primary care can have a big impact on improving the recognition and management of chronic conditions such as depression and severe mental illness (SMI), and can improve the detection of suicidal ideation, (11) especially when combined with coordinated care between GPs and mental health specialists across the healthcare system. However, the impact of primary care on population outcomes is limited by the wide variations that exist in the quality of care of providers. (12) This has led policymakers to link quality of care to provider remuneration in an attempt to reduce variability and to improve quality across health systems. (13)

The Quality and Outcomes Framework (QOF), one of the most costly and comprehensive pay-forperformance schemes of its kind, was introduced in 2004 in the United Kingdom (UK) to reward high quality primary care. (14) The scheme linked up to 25% of GP income to performance on quality indicators across four domains: clinical, organisational, patient experience and additional services. Varying payments were awarded to practices mainly for the ongoing management of chronic conditions, including depression and SMI, with adjustments made for practice list size and disease prevalence. The scheme was periodically reviewed and quality indicators were modified or removed, while new indicators were occasionally introduced for existing or new conditions. (15) In the first three years of its implementation, the QOF yielded improved results for intermediate outcomes and incentivised processes of care but the rate of improvement was not sustained, (16, 17) and performance in later years was no higher than levels projected from pre-intervention trends. (16, 17) In terms of improved patient outcomes, the evidence is contradictory. (18-20) For some conditions there is evidence of lower rates of emergency hospital admission, (18, 21) and reductions in inequalities of delivery of care (22) but there has been no clear improvement in mortality rates in the UK for physical conditions covered by the QOF. (20) Previous research has also found no association between practice-level performance on QOF indicators and all-cause mortality in the locality. (19)

In terms of mental health, the QOF includes indicators for patients with Serious Mental Illness (SMI), requiring practices to monitor treatment, record an up-to-date care plan and to conduct an annual review incorporating physical checks and health promotion advice. For patients with SMI, the

introduction of the QOF was associated with increases in the numbers of service interactions and the number of reported comorbidities. (23) From 2006, the QOF also includes depression indicators requiring practices to assess severity using validated assessment instruments, or indicators requiring practices to review patients within a month of diagnosis and make a biopsychosocial assessment based on 16 'themes' - with one of these being suicidal ideation. These SMI and depression indicators are used to manage the chronic nature of these conditions but also to identify individuals at risk. The complete list and descriptions of all depression and SMI indicators with information on the years they were active included are provided in the online appendix. However, existing evidence on the effectiveness of QOF mental health indicators is limited.

Suicide is an important outcome measure for assessing the quality of mental health services. Effective suicide prevention strategies need to identify individual-level risk factors but also factors associated with the social environment in which an individual lives, (24) as recent evidence indicates marked spatial variability in suicide incidence in England. (25) It is suggested that area of residence can have an impact on levels of social support, but can also increase the likelihood of developing a mental illness, both of which increase suicide risk. (26) Ecological studies have shown that area-level effects such as deprivation and social fragmentation are associated with the geographical distribution of suicide. (27-29)

Most studies investigating determinants of suicide have examined specialist mental health services and suicide (30) or suicide in local and/or regional settings. (31) However, no published studies to date have examined the relationship between performance on QOF mental health indicators and suicide. The purpose of this study was to examine for the first time the relationship between incentivisation of mental health management in primary care and suicide rates. To achieve this aim we applied spatial analysis techniques to assess the association between practice performance on the mental health domain of the QOF and suicide rates for the localities that the practices serve, whilst controlling for relevant small-area population characteristics.

METHODS

Data sources

Several data sources were accessed to extract information on numbers of suicides, population size estimates, rurality, social fragmentation, deprivation, quality of care, disease prevalence and spatial coordinates at the Lower Super Output Area (LSOA) level. LSOAs are small geographical areas containing approximately 1500 residents. From the Office for National Statistics (ONS) we obtained data on case counts for suicides and open verdicts (32) per LSOA of residence by year of death, broken down by gender and 5-year age bands , for the period 1st January 2006 to 31st December 2014. Our population at risk was people aged 20 years or older as we wished to exclude children and adolescents from our analysis because of differences in the remit of mental health services and

depression management provided to kids, adolescents and adults. (33) Classification of deaths for the study used the International Classification of Diseases, 10th edition (ICD-10) and included all deaths with a final underlying cause recorded as intentional self-harm, injury/poisoning of undetermined intent and sequelae of intentional self-harm, or injury/poisoning of undetermined intent, which follows conventional practice for government suicide statistics in the UK. We included both suicide and open verdicts in our outcome measure and we refer to this measure as suicides throughout the manuscript. Population counts by gender and age group, as well as information on ethnicity were obtained from the 2011 national Census, at the 2011 LSOA.

Area-level deprivation, as measured by the Index of Multiple Deprivation 2007, was available for LSOAs as at 2011. The IMD measures relative levels of deprivation for all the 32,844 LSOAs in England on a continuous scale of deprivation where most of the indicators are based on 2005 statistics. Moreover, we obtained data on rural/urban classification and we constructed a measure of social fragmentation based on Congdon's index (derived from 2011 Census data on private renting, single person households, single adults and population turnover). The index represents neighbourhood-level conditions which affect social resources such as social cohesion and social capital. (27) Datasets on prevalence of SMI and depression, practice list size and practice-level mental health quality of care data were obtained from NHS Digital. As 80% of LSOAs do not contain a general practice we used spatial analysis techniques to estimate healthcare data at the LSOA level from the practice-level data (a detailed description of this methodology is provided in the supplementary appendix).

The number of Primary Care practice 'hubs' (an LSOA with one or more practices) varied between 8372 in the third year of the QOF (2006-07) and 7778 in the final year (2014-15) including for each year over 99% of all patients registered with primary care in England. (34) From 2006-2007 onwards, all relevant information was available; i.e. number of practices, total list size, and prevalence rates for those over 18 years old in the two relevant QOF clinical domains: SMI and depression.

To better capture quality of care in the mental health domain, as measured by the QOF, we calculated population achievement (17, 19), defined as $PAmh = \frac{\sum N_i}{\sum (Di + Ei)}$

where N_i , Di and Ei denote the numerator, denominator and exceptions for QOF indicator *i*, respectively. Population achievement was calculated, separately within each financial year, across the full set of depression and SMI outcome and measurement QOF indicators in the mental health clinical domain.

Statistical analysis

We fitted a negative binomial model with standard errors clustered at the LSOA level for global between-area variability and/or local variability in neighbouring areas to derive and map age and sex adjusted area estimates of suicide. Negative binomial models report Incidence Rate Ratios (IRR) and

are suitable for analysing count data when there is presence of over-dispersion, as was the case in our data.

We used one set of negative binomial regression models to investigate the relationship between mental health quality of care, as measured by the QOF, and suicide incidence over time. The dependent variable was annual number of suicides between 2006 and 2014, and we investigated the relationship with the mental health quality of care measure, in the same or up to a lag of three years. The same sets of covariates were used for all models. We also conducted several sensitivity analyses under different assumptions. There was an excess of zero event counts in the study dataset; thus, more than 99% of data points (LSOAs by age group and sex, over time) did not include a single suicide. Therefore, we investigated how a zero inflated negative binomial model fitted the data. This model may intuitively seem to provide a better fit to the data, but the standard negative binomial model is more appropriate from a theoretical perspective, as every population group is 'at risk' of experiencing suicide. As an additional sensitivity analysis, we aggregated suicide numbers and population estimates in four age groups (20-24, 25-49, 50-64, 65+), according to ONS classification, to address any issues of collinearity between age groups. The economic recession of 2008 occurred in the middle of our study period; thus, to isolate the impact of the QOF on suicide risk from the effects of the economic recession, we ran a sensitivity analysis for the period 2009-2014. Furthermore, conventional practice in the UK is to include open verdicts when investigating and reporting suicide rates. However, to facilitate international comparison, we analysed open verdicts and definite suicide cases separately. Finally, we investigated the extent to which spatial autocorrelation was present in the data (i.e. level of suicide risk in one locality affecting risk in an adjacent locality) by calculating Moran's I, which is a measure of spatial autocorrelation.

Stata v14.1 was used for the principal data management and analyses. For both sets of analyses, we used the *nbreg command* with the exposure option followed by the *IRR* specification while we used clustered standard errors for our unit of analyses (i.e. 32,844 LSOAs) to take into account the correlation of observations within each LSOA.

RESULTS

Figure 1 represents the variability in incidence of suicide across regions over the study period. The North East and North West regions had the highest overall suicide rates whereas the London region had the lowest overall suicide rates across the period of study. There were 38,511 recorded suicides in England during calendar years 2006 to 2014, inclusive. Figure 2 represents the spatial distribution of suicide over the whole time period (aggregated), at the Clinical Commissioning Group (CCG) level, a middle-level organisation structure of the National Health Service (NHS). Figure 3 represents the spatial variability of quality of care in the mental health domain across LSOAs at the census year i.e. 2011. Descriptive statistics on suicide numbers, population size estimates, number of practices, census information, deprivation, prevalence of depression and SMI, and mental health quality of care across regions (Strategic Health Authorities - SHAs) for 2011 are reported in Table 1. The highest prevalence for both depression and SMI was observed in the North East region whereas the

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lowest prevalence for depression and SMI was found in the London and South Central regions respectively.

We found no association between mental health quality of care and suicide in our main model (IRR 1.000; 95%CI 0.998, 1.002). There was also no association after lagging the effects of mental health quality of care for up to three years. This finding was consistent in all sensitivity analyses, including the analysis limiting the study period to 2009-2014 (post 2008 recession). In the sensitivity analysis that investigated definite suicide deaths, results were similar (IRR 1.000; 95%CI 0.998, 1.002). All results from the sensitivity analyses are reported in the supplementary appendix.

No evidence of heterogeneity in the association of interest was found, by health region or deprivation score. Median QOF mental health quality of care across regions varied between 80.4% (East Midlands) to 83.0% (North East). As expected, variability was much greater at the LSOA level, partly due to the smaller denominators. For example, in the first analysis year (2006-07) performance varied from 12.9% to 99.2% (10th and 90th centiles were 71.3% and 94.7% respectively), in the middle year (2009-10) performance varied from 20.6% to 94.5% (10th and 90th centiles were 75.7% and 88.8% respectively) and finally for the last year (2014-15) it varied from 28.3% to 94.2% (10th and 90th centiles were 68.6% and 86.7% respectively).

Table 2 presents IRRs for the coefficients from the negative binomial regression models. Across all age groups, the highest risk for suicide was observed for the 40-44 group (1.865; 95%CI 1.775, 1.960) while a bimodal distribution was observed with peaks in the mid years and those over 85 years (1.537; 95%CI 1.425, 1.657) thus indicating a non-linear relationship of suicide and age. Female sex (0.295; 95%CI 0.288, 0.302) and social fragmentation (1.050; 95%CI 1.044, 1.056) were also strong predictors of suicide across all models, followed by deprivation (1.015; 95%CI 1.014, 1.016). Rurality (1.035; 95%CI 1.004, 1.066) was also a strong predictor of suicide risk, where inhabitants in rural areas were found to be more likely to die by suicide.

DISCUSSION

After adjusting for key social and demographic characteristics, we found no association between primary practice performance on incentivized mental health indicators and suicide rates in practice localities. Although depression and SMI are common among people who die by suicide, some people may not experience these symptoms whilst others may not be diagnosed by their GPs. (35) Not all people with mental illness see their GP in the year before death occurs (4) and not all people with mental illness will be assessed or treated. (5) Nevertheless, GP based efforts can have positive results in the reduction of suicide rates sometimes associated with an increase in antidepressant prescription. (36) One of the aims of incentivisation was to achieve better quality of care for people diagnosed with mental illnesses and, although suicide is rare, we might have expected to have observed small effects for a successful programme in that respect. For a dataset of this size, statistical significance is a secondary concern, as we should have been able to identify even a tiny effect - but there appeared to be none. Moreover, QOF mental health 'quality of care' scores do not inform as to care received in settings outside general practice. For example, a 28% of people who die

by suicide in the UK have had contact with mental health services in the previous 12 months (37), however, as this study focuses on aspects of mental health care in general practice we did not account for the important role of mental health and emergency services in preventing suicide. If local mental health services had been consistently beneficial and primary care services poor or vice versa, then the effect of the quality of primary care would be mitigated by the unknown effect of the quality of specialist mental healthcare in the same locality. Finally, even if high quality of mental health provision in primary care, as incentivised through the QOF, has beneficial impacts for patients overall this may not be apparent for important but relatively low frequency markers of population health such as death by suicide.

Our regression models have also indicated that area-level socioeconomic factors, such as social fragmentation and rurality, rather than variations in the quality of primary care, are more strongly associated with suicide risk. We identified a strong relationship between levels of deprivation and suicide from 2006 to 2014. One additional percentage point in population deprivation is associated with a 1.5% (n=61) increase in suicides nationally in one year. A change in IMD score from the 25th centile (9.72) to the median (17.23) would correspond to an 11.2% increase in suicide rates (n=458). The extent of the relationship between deprivation and suicide is also highly consistent across the sensitivity analyses reported in the supplementary appendix. Social fragmentation was a strong predictor of suicide, where one additional percentage point in social fragmentation is linked to 5.3% increase in suicide (n=215) nationally in one year. These findings agree with previous evidence on the effects of deprivation (27) and social fragmentation (27, 28) and also agree with previous evidence which showed that social fragmentation has a stronger association with suicide risk than deprivation. (28)

Additionally, we observed a moderate association between the prevalence of SMI and suicide across all models and all sensitivity analyses but small or no associations between prevalence of depression and risk of suicide. As this result may seem unusual we performed a univariate analysis for the associations between prevalence of SMI and depression with risk of suicide. The results from the univariate analysis indicate that prevalence of SMI has a strong effect on suicide (1·476; 95%CI 1·361, 1·601) while prevalence of depression has a very small but statistically significant effect (1·009; 95%CI 1·006, 1·012). The fact that unresolved depression register is weakly associated with suicide, despite the well documented association between depression and suicide, may imply that the condition is not captured well in primary care. Misidentifications and missed cases in depression diagnosis are common in primary care. (38) Moreover, it is suggested that GP's who work in socioeconomically deprived areas identify the ongoing management of depression as a burden and they may be reluctant to investigate psychosocial issues with their patients due to structural factors and workload. (39)

This analysis was conducted at the population level and used a novel spatial estimation approach to assess and quantify the relationship between mental health quality of care and suicide rates for the whole primary care-registered population of England. This analysis allows for careful control of population characteristics at a low-level geographical area (LSOA), which is not possible with analyses at the general practice-level. We avoided standardising our outcome variable as standardised death by suicide ratios are imprecise at this low geographical (LSOA) level. Instead we adjusted our analysis for population structure by including all the relevant information on age and

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sex. This level of analysis was deemed to be optimal as aggregating data at higher geographical levels would introduce heterogeneity and confounding across small-area populations.

However, our study has a number of limitations. First, the possibility of ecological fallacy cannot be ruled out as the associations we observed at small area level may not pertain to individual patients. Second, in essence we equated performance on the QOF mental health domain to quality of care for mental health, which is not necessarily the case. However, in the absence of other quantifiable metrics, this has become an accepted practice, although the QOF fails to capture other aspects of consultations that may be even more important. Third, the IMD is by definition an aggregate measure of income, employment, health, education, crime and environmental deprivation in a locality. Thus, it is always collinear with measures of poor health and income, and these were necessarily excluded from our analysis. Fourth, we could not access population measures of ethnicity for all age groups and we included this variable as '% White' at LSOA level. Fifth, we were unable to attribute practice performance to LSOAs for the first two years of the scheme (2004/2005 and 2005/2006), the period during which practices exhibited the largest improvements in performance, (22) due to data on mental health quality of care being unavailable. However, there was sufficient variation in quality of care at the LSOA level from 2006/7 to discern any association with suicide risk should it exist. Sixth, any effects that the scheme might have had on suicide may have been delayed and the time window between mental health quality of care and suicide (i.e. 3 year lag model) could be too narrow. Nonetheless, analysis with longer lag periods is subject to methodological issues, whilst it seems clinically implausible that the indicators could have long-term preventive effects without short- or medium-term effects. Seventh, some degree of spatial autocorrelation was identified, but considering the very low value of the measure (Moran's I = 0.0049), we did not account for spatial autocorrelation in the model. Even if spatial clustering was present, we would expect it to be driven by socio-economic factors we controlled for, and not by suicide cases in neighbouring areas which would warrant the use of spatial regression techniques. Eighth, multimorbidity and psychosocial risk factors for suicide (e.g. domestic abuse or substance misuse) may occur concomitantly with depression especially in deprived areas, (40) however, it is likely that these cases may not be captured adequately in the QOF (15) and our measure of guality of mental healthcare as some of them will remain unrecorded. Ninth, recording of depression in English primary care has been described as being problematic, as some GPs prefer to record symptoms rather than formal psychiatric diagnostic labels, (41) but we believe this is unlikely to have affected the validity of our findings in relation to the primary research question. Finally, as death by suicide is used as an underlying indicator in the health and deprivation domain of the IMD, we used the IMD 2007 which was calculated from 2005 data on suicide deaths. Nevertheless, the various IMD versions have been documented to be highly correlated and thus the use of an old version of IMD should impose no implications to our findings.

Our study adds to a growing body of evidence on the potential impact of the QOF on patient outcomes, and is the first study to assess whether quality of primary care, measured as a proxy via metrics generated from a national incentive scheme, is linked with subsequent suicide risk at the population level. Our findings indicate that the scheme appears to have had no impact on suicide risk through the treatment and monitoring of depression and SMI, although population benefits in

terms of improved quality of life and reductions in unnatural mortality risk may accrue in the longer term.

Suicide is a complex phenomenon and many factors may influence a person's decision to take their own life. In the absence of indicators related to suicidality and suicidal thoughts and acts, recognition of patients at risk of suicide depends entirely on the GP's ability to identify and interpret relevant signs. This may be extremely difficult without sufficient training as suicidal thoughts may be communicated differently across people at risk. This adds to concerns from GPs regarding the limited training that is provided to them regarding their role in preventing suicide. (42) Furthermore, the first few weeks following discharge from psychiatric inpatient care are characterised by increased suicide risk for patients (37) and GPs stress the need to improve communication and information sharing between services, particularly in the cases of hospital discharge and transition between services. (5) It is vital to establish a collaborative model between GPs, mental health services, community care and hospitals, where improved communication and information sharing will be priorities. However, suicide prevention has never had a higher profile in the UK than at present. The Refreshed Suicide Prevention Strategy, the Health Select Committee Inquiry into Suicide Prevention and the NHS 5-year Forward Review provide a target of a 10% reduction in the national suicide rate. It is clear that effective suicide prevention requires a more concerted, multiagency approach, including better training in suicide prevention for primary care clinicians.

It is widely recognised that primary care has an important role to play in suicide prevention, but we found no association between practices' reported quality of care on mental health activities incentivised in the QOF and suicide rates in the local population. The QOF was a major national experiment, dwarfing in scale all previous attempts to incentivise high quality care, but doubts on its impact and value mean that the UK is now retreating from using incentives in primary care, led by Scotland's decision to withdraw from the QOF in 2017. Other countries continue to use incentives to improve population health, however, hoping to draw on the experiences of earlier schemes to develop more effective frameworks. If these attempts are to succeed, attention needs to be focussed on those activities that contribute to improved patient outcomes, such as reductions in risk of dying prematurely from suicide, and those that will more accurately capture higher value care. More broadly, a better understanding is needed as to how resources can best be allocated to achieve higher quality of services that will lead to genuine improvements in population health outcomes.

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Acknowledgments

We would like to thank the Office of National Statistics and NHS Digital (formerly the Health and Social Care Information Centre) for the wealth of information they have collected and systematically organised, which made this study possible.

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Funding

This study is funded by the National Institute for Health Research School for Primary Care Research (NIHR SPCR), through CG's PhD. This report is independent research by the National Institute for Health Research. LM acknowledges financial support from the MRC Skills Development Fellowship (MR/N015126/1). The views expressed in this publication are those of the authors and necessarily those of the NHS, the National Institute for Health Research, the Department of Health, or the MRC.

Authorship & contributorship

EK and CG designed the study. CG extracted the data from all sources and performed the analyses. CG and wrote the manuscript. LM, RW, NK, TD, DA, EK critically revised the manuscript. CG is the guarantor of this work and, as such, had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Transparency declaration

CG affirms that this manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained.

Data sharing

The data used in this study are freely available and the authors are happy to share in an organised and cleaned final dataset.

Ethical Approval

No ethical approval was needed as no patients were involved and all data were publicly available except for the ONS Mortality Data.

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to per period

Strategic Health Authority*										
Aggregates	North	North	Yorkshire	East	West	East	London	South	South	South
00 0	East	West	Humber	Midlands	Midlands	England		West	Central	East
						8		Coast		
Population size	1.998.1 71	5.365.75 5	4.011.205	3.461.393	4.217.657	4.465.027	6.195.926	4.107.71 4	3.158.651	3.415.10 8
Suicide counts+	265	654	448	356	418	470	571	499	324	374
No. of practices	404	1248	787	626	956	785	1473	623	502	718
Census	Medians a	across 2011 L	SOAs	•	1					
Information [‡]										
IMD 2007 (mean	28.1	27.7	25.0	20.0	24.8	15.3	26.0	17.7	12.7	14.9
across LSOAs)										
P10	8.2	7.1	6.9	5.5	7.0	4.5	8.8	5.7	3.5	3.8
P25	13.7	11.7	10.9	8.8	11.3	7.4	14.5	9.2	5.7	6.9
P75	39.7	41.0	36.8	27.5	35.9	20.1	36.0	22.6	18.9	19.8
P90	53.2	57.6	52.1	41.5	50.5	30.2	45.6	33.4	28.6	29.7
Social	-0.3	-0.2	-0.7	-0.2	-0.4	-0.2	0.7	0.2	0.1	0.0
Fragmentation										
(mean across										
LSOAs)										
P10	-1.8	-1.7	-1.6	-1.6	-1.7	-1.5	-1.4	-1.3	-1.4	-1.5
P25	-1.3	-1.3	-1.1	-1.1	-1.3	-1.1	-0.7	-0.9	-1.0	-1.1
P75	0.0	0.2	0.3	0.3	-0.0	0.2	1.9	-0.6	0.5	0.5
P90	1.6	1.7	2.0	2.0	1.2	1.6	3.5	2.4	2.2	2.
QOF Information: m	nedians acro	ss practice hu	ıbs				-			
List Size	9410	8322	8926	8893	8096	9100	7632	9063	10376	9075
QOF information (sp	patially estin	nated): media	ans across 2011 L	SOAs 🗆						
Severe Mental	0.8	0.8	0.7	0.7	0.7	0.7	0.9	0.7	0.6	0.7
Illness %										
P10	0.6	0.6	0.5	0.5	0.5	0.5	0.6	0.4	0.4	0.4
P25	0.7	0.7	0.6	0.5	0.6	0.5	0.7	0.5	0.5	0.5
P75	0.9	1.0	0.9	0.8	0.8	0.8	1.1	0.8	0.8	0.8
P90	1.0	1.1	1.0	1.0	1.0	0.9	1.3	1.0	0.9	1.0
Depression %	11.9	10.6	9.5	10.0	9.0	9.1	6.2	10.1	9.2	8.9
P10	8.7	7.0	5.5	6.5	5.9	5.5	4.0	7.1	6.0	6.1
P25	10.2	8.8	7.2	7.9	7.5	7.1	5.0	8.5	7.5	7.5
P75	14.8	12.3	11.4	11.7	10.8	11.0	7.7	11.9	11.1	10.7
P90	16.9	13.8	13.0	13.5	12.6	12.8	9.0	13.5	12.3	12.5
% Mental Health	83.0	81.6	81.2	80.4	82.7	80.7	81.80	82.3	82.0	80.6
Quality of Care										
						e				
P10	78.6	76.2	75.2	74.8	77.0	74.37	76.3	75.5	75.8	74.0
P25	80.8	79.0	78.5	77.7	80.1	78.16	79.2	79.4	79.1	77.4
P75	84.9	84.2	83.8	83.2	84.5	83.48	84.1	84.7	84.2	83.4
P90	86.2	86.0	85.6	85.4	86.0	85.58	85.7	86.4	85.9	85.5
Other Information A	cross LSOA	As	1	1	1	1	1	1	<u>I</u>	1
Rural (%)	17.5	9.8	16.5	25.4	14.8	28.1	0.1	20.1	19.7	30.2
			1	1	1			1		1

Table 1: Area and population characteristics by strategic health authority*++□

*Strategic health authorities were the highest geographical organisational level for NHS England during the study period. This structure was abolished in 2013.

⁺Suicides and open verdicts in 2011.

‡Data available from the 2011 census.

□ Quality and Outcomes Framework (QOF) year7 (2010-2011).

* Mental Health Quality of Care is measured as % achievement of the population in the relevant QOF indicators.

P10, P25, P75, P90 are the respective percentiles for each variable.

8	v	Binomial Model)	U U	
Year:	Lag=0	Lag=1	Lag=2	Lag=3
% Mental Health Quality of Care *	1.000 (0.998, 1.002),	1.000 (0.998, 1.003),	1.000 (0.998, 1.003),	1.000 (0.997, 1.003),
	<0.537 (0.000)	<0.569 (0.001)	<0.971 (0.001)	<0.715 (0.001)
Female	0·295 (0·288, 0·302),	0·293 (0·286, 0·301)	0.293 (0.285, 0.301)	0·293 (0·284, 0·301)
	<0·001 (0·003)	<0·001 (0·003)	<0.001 (0.003)	<0·001 (0·004)
Index of Multiple	1.013 (1.012, 1.014),	1·013 (1·012, 1·014),	1.013 (1.012, 1.014),	1.013 (1.012, 1.014),
Deprivation 2007	<0.001 (0.0003)	<0·001 (0·0004)	<0.001 (0.0004)	<0.001 (0.0004)
Rural (v urban)	$\begin{array}{c} 1.059 \ (1.027, \ 1.092), \\ < 0.001 \ (0.016) \end{array}$	1.065 (1.032, 1.099), <0.001 (0.016)	1.065 (1.030, 1.102), <0.001 (0.018)	1.053 (1.016, 1.091), <0.005 (0.019)
Index of Social Frag.	1.030 (1.027, 1.034),	1.031 (1.027, 1.035),	1.031 (1.027, 1.035),	1.031 (1.027, 1.035),
	<0.001 (0.001)	<0.001 (0.003)	<0.001 (0.003)	<0.001 (0.002)
Ethnicity (%White)	1.007 (1.006, 1.008),	1.007 (1.006, 1.008),	1.007 (1.006, 1.008),	1.008 (1.007, 1.009),
	<0.001 (0.000)	<0.001 (0.0004)	<0.001 (0.0004)	<0.001 (0.000)
Prevalence of SMI	1.054 (1.004, 1.108),	1.049 (0.995, 1.105),	1.038 (0.983, 1.096),	1.040 (0.982, 1.101),
	<0.033 (0.026)	<0.071 (0.028)	<0.169 (0.028)	<0.173 (0.030)
Prevalence of Depression	1.000 (0.996, 1.004),	0.999 (0.995, 1.003),	0·999 (0·995, 1·003),	0.998 (0.994, 1.003),
	<0.803 (0.019)	<0.881 (0.019)	<0·760 (0·021)	<0.648 (0.022)
Age (20-24)	Reference Age Group	Reference Age Group	Reference Age Group	Reference Age Group
Age (25-29)	1.162 (1.101, 1.226), <0.001 (0.031)	1·147 (1·084, 1·214) <0·001 (0·033)	1·143 (1·077, 1·214) <0·001 (0·035)	$\begin{array}{c} 1 \cdot 146 \ (1 \cdot 074, \ 1 \cdot 223) \\ < 0 \cdot 001 \ (0 \cdot 038) \end{array}$
Age (30-34)	1·360 (1·290, 1·434)	1·339 (1·267, 1·416)	1·300 (1·225, 1·380)	1·281 (1·201, 1·366)
	<0·001 (0·036)	<0·001 (0·038)	<0·001 (0·039)	<0·001 (0·042)
Age (35-39)	1.701 (1.616, 1.790)	1.681 (1.593, 1.774)	1.657 (1.564, 1.754)	1.660 (1.560, 1.767)
	<0.001 (0.043)	<0.001 (0.046)	<0.001 (0.048)	<0.001 (0.052)
Age (40-44)	1.876 (1.785, 1.972),	1.880 (1.784, 1.981)	1.883 (1.781, 1.990)	1·901 (1·790, 2·019)
	<0.001 (0.047)	<0.001 (0.050)	<0.001 (0.052)	<0·001 (0·058)
Age (45-49)	1.846 (1.755, 1.942),	1.839 (1.744, 1.939)	1.852 (1.750, 1.959)	1.891 (1.780, 2.009)
	<0.001 (0.047)	<0.001 (0.049)	<0.001 (0.053)	<0.001 (0.058)
Age (50-54)	1.841 (1.749, 1.939),	1.835 (1.738, 1.937)	1.843 (1.741, 1.952)	1.882 (1.769, 2.001)
	<0.001 (0.048)	<0.001 (0.050)	<0.001 (0.053)	<0.001 (0.059)
Age (55-59)	1.621 (1.536, 1.711),	1.604 (1.515, 1.698)	1.607 (1.513, 1.707)	1.631 (1.527, 1.741)
	<0.001 (0.044)	<0.001 (0.046)	<0.001 (0.049)	<0.001 (0.054)
Age (60-64)	1·273 (1·202, 1·349),	1·255 (1·181, 1·333)	1·241 (1·164, 1·322)	1·259 (1·176, 1·349)
	<0·001 (0·036)	<0·001 (0·038)	<0·001 (0·040)	<0·001 (0·044)
Age (65-69)	1.026 (0.963, 1.093),	1.008 (0.943, 1.078)	0.992 (0.924, 1.065)	0·994 (0·920, 1·073)
	<0.420 (0.032)	<0.803 (0.034)	<0.840 (0.035)	<0·878 (0·038)
Age (70-74)	0.999 (0.933, 1.070),	0.988 (0.919, 1.063)	0.997 (0.923, 1.076)	0.995 (0.916, 1.081)
	<0.993 (0.034)	<0.762 (0.036)	<0.939 (0.038)	<0.919 (0.041)
Age (75-79)	1.070 (0.996, 1.149),	1·051 (0·974, 1·133)	1.037 (0.957, 1.123)	1.053 (0.965, 1.148)
	<0.062 (0.038)	<0·197 (0·040)	<0.372 (0.042)	<0.242 (0.046)
Age (80-84)	1·351 (1·254, 1·455),	1·326 (1·226, 1·435)	1·284 (1·180, 1·396)	1·271 (1·160, 1·392)
	<0·001 (0·050)	<0·001 (0·052)	<0·001 (0·055)	<0·001 (0·059)
Age (85plus)	1.551 (1.438, 1.673) <0.001 (0.059)	1.518 (1.402, 1.635) <0.001 (0.061)	1.505 (1.383, 1.638) <0.001 (0.065)	$\begin{array}{c} 1\cdot 529\ (1\cdot 396,\ 1\cdot 674) \\ <0\cdot 001\ (0\cdot 070) \end{array}$
2006	Reference Year	Reference Year	Reference Year	Reference Year
2007	0.940 (0.891, 0.992), <0.025 (0.025)	-	-	-
2008	0·999 (0·946, 1·054), <0·971 (0·027)	$\begin{array}{c} 1.060 \ (1.014, \ 1.108), \\ < 0.010 \ (0.23) \end{array}$	-	-
2009	1.012 (0.953, 1.074), <0.691 (0.030)	1.076 (1.027, 1.126), <0.002 (0.025)	1.076 (0.966, 1.057), <0.627 (0.022)	-
2010	0·961 (0·903, 1·022),	1.025 (0.978, 1.073),	1.025 (0.919, 1.008),	0·949 (0·908, 0·993),
	<0·210 (0·030)	<0.291 (0.024)	<0.108 (0.023)	<0·023 (0·021)
2011	1.018 (0.956, 1.084),	1.085 (1.034, 1.138),	1.085 (0.977, 1.071),	1.009 (0.966, 1.054),
	<0.571 (0.032)	<0.001 (0.025)	<0.317 (0.025)	<0.663 (0.022)

Table 2 - Regression Analysis set 1: Effect of Mental Health QOF Quality of Care on suicide over time (Negative Binomial Model)

Binomial Model)						
Year:	Lag=0	Lag=1	Lag=2	Lag=3		
2012	1.014 (0.966, 1.065),	1.080 (1.031, 1.130),	1.080 (0.969, 1.062),	0.999 (0.951, 1.049),		
	<0.556 (0.025)	<0.001 (0.024)	<0.526 (0.023)	<0.990 (0.025)		
2013	1·056 (1·002, 1·112),	1·127 (1·070, 1·184)	1·127 (1·009, 1·113)	1.045 (0.993, 1.100)		
	<0·039 (0·027)	<0·001 (0·027)	<0·020 (0·026)	<0.088 (0.027)		
2014	1·076 (1·020, 1·134),	1·158 (1·101, 1·215)	1·158 (1·040, 1·148)	1.074 (1.019, 1.131)		
	<0·006 (0·029)	<0·001 (0·028)	<0·001 (0·027)	<0.007 (0.028)		
Model intercept	0.00004 (0.00003, 0.00005),	0·00003 (0·00002, 0·00004),	0.00003 (0.00002, 0.00004),	0.00003 (0.00002, 0.00004),		
	<0.001 (0.00001)	<0·001 (0·00001)	<0.001 (0.00001)	<0.001 (0.00001)		

Table 2 - Regression Analysis set 1: Effect of Mental Health OOF Quality of Care on suicide over time (Negative

95% Confidence Intervals are in brackets, Results are reported as incidence rate ratios (IRR) followed by Pvalues and Standard Errors in parentheses.

*Mental Health Quality of Care is measured as % achievement of the population in the relevant QOF indicators.

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Figure 1



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Figure 1: Suicide rates, per 100,000 population, by English Regions over time (2006-2014).

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Figure 2: Suicide rates per 100,000 population in England aggregated across all years (2006 – 2014) at the CCG level.



Figure 3: Measure of Mental Health Quality of Care in England for specific mental health indicators at the 2011 LSOAs. (2011 Census Year).