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**DESIGN AND PERFORMANCE OF A
POLARITY COINCIDENCE DETECTOR**

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ABSTRACT

A synchronous detector based on the polarity coincidence principle has been constructed and tested. The theory of operation is discussed with special emphasis on the influence of correlation between samples and the influence of error sources. Details of the construction are presented and test results are discussed. The present instrument works in the range of 20 c/sec to 20 Kc/sec and is capable of measuring signal-to-noise ratios as low as 10^{-4} . It has been in constant laboratory use for several months and has proven to be a very reliable tool for the synchronous measurement of small ac signals.

Introduction:

Synchronous detectors have received widespread application in various branches of experimental physics for the measurement of small signals in the presence of noise. Among the various schemes the Polarity Coincidence Detector (PCD) occupies a unique place because several outstanding features make it particularly suited for the measurement of very small signal-to-noise ratios.

Despite its simplicity and reliability, it has not yet found its way into laboratories as a standard research tool. It is the purpose of this report to summarize and extend the theory of operation of a PCD, and discuss the design and performance of a working model.

Historical Background:

Polarity coincidence schemes for signal detection have been proposed by several authors. Faran and Hills (1) calculated the cross correlation function of two waveforms after discarding all amplitude information. They constructed a detector where coincidences between two waveforms were produced after hard limiting. The output of the coincidence circuit was averaged by passing it through a low pass filter.

Melton and Karr (2) proposed a PCD-scheme where the waveform under consideration is sampled at appropriate time intervals and polarity coincidences are recorded. They restrict their consideration to the use of uncorrelated samples which is an unnecessary restriction. It leads to an underestimation of the statistical error.

Wolff, Thomas, and Williams (3) have shown that the polarity coincidence detector is nonparametric with respect to the "false alarm rate."

Bitzer (4) applied the polarity coincidence scheme in the space domain to phase quantization in connection with radar antennae.

In the following, the principle of operation of a PCD will be presented in a form suitable for its use as a synchronous detector for general laboratory purposes.

Theory of Operation:

A signal of known frequency may be embedded in noise. In the PCD-scheme the waveform of signal + noise is sampled with short pulses at the signal frequency. Coincidences are recorded whenever the sampled waveform and the sampling pulse have the same polarity (+ or -). With pure noise present, on the average half the samples will result in a polarity coincidence if we sample, i.e., with positive pulses.

In the presence of a signal, slightly more or less coincidences will be recorded according to the phase relationship between signal and sampling pulses.

In order to arrive at a quantitative description let us assume a signal $u\sqrt{2} \cdot \sin(\omega t + \varphi)$ in the presence of narrow band Gaussian noise,

$$x(t) = x_s \sin \omega t + x_c \cos \omega t \quad (1)$$

where x_s and x_c are Gaussian random variables distributed according to

$$p(x_s) = \frac{1}{\sqrt{2\pi} \cdot \sigma} \exp \left(-\frac{x_s^2}{2\sigma^2} \right)$$

$$p(x_c) = \frac{1}{\sqrt{2\pi} \cdot \sigma} \exp \left(-\frac{x_c^2}{2\sigma^2} \right) . \quad (2)$$

The sum of noise + signal can be written

$$\begin{aligned} x(t) &= (x_s + u\sqrt{2}\cdot\cos\varphi)\sin\omega t + (x_c + u\sqrt{2}\sin\varphi)\cdot\cos\omega t \\ &= X_s \sin\omega t + X_c \cdot \cos\omega t \end{aligned}$$

X_s and X_c are also distributed normally, and their joint distribution is

$$p(X_s, X_c) = \frac{1}{2\pi\sigma^2} \exp\left(-\frac{X_s^2 + X_c^2}{2\sigma^2}\right)$$

We now sample in phase with X_s and X_c , respectively. A coincidence is recorded whenever the waveform exceeds a certain threshold x_+ (for positive coincidences). The fraction of sampling pulses that results in an in-phase coincidence is given by

$$\begin{aligned} \frac{n(0)}{n_0} &= \frac{1}{2\pi\sigma^2} \int_{-\infty}^{+\infty} \exp\left(-\frac{(X_c - u\sqrt{2}\sin\varphi)^2}{2\sigma^2}\right) dX_c \int_{+x_+}^{\infty} \exp\left(-\frac{(X_s - u\sqrt{2}\cos\varphi)^2}{2\sigma^2}\right) dX_s \\ &= \frac{1}{2} \left[1 - \operatorname{erf}\left(\frac{x_+}{\sigma\sqrt{2}} - \frac{u\cos\varphi}{\sigma}\right) \right] \end{aligned} \quad (3)$$

In an analogous way we find for sampling 90° out-of-phase

$$\frac{n(90)}{n_0} = \frac{1}{2} \left[1 - \operatorname{erf}\left(\frac{x_+'}{\sigma\sqrt{2}} - \frac{u\sin\varphi}{\sigma}\right) \right]$$

n_0 is the total number of samples in a given time,

x_+ and x_+' are usually different because different channels are involved.

For small values of the argument (small signal-to-noise ratios and hard limiting) the error function can be expanded, $\operatorname{erf}(x) \approx \frac{2x}{\sqrt{\pi}}$, and we obtain

$$\frac{n(0)}{n_0} = \frac{1}{2} \left[1 - \frac{2}{\sqrt{\pi}} \left(\frac{x_+}{\sigma\sqrt{2}} - \frac{u}{\sigma} \cos\varphi \right) \right] \quad (4)$$

$$\frac{n(90)}{n_0} = \frac{1}{2} \left[1 - \frac{2}{\sqrt{\pi}} \left(\frac{x_+}{\sigma\sqrt{2}} - \frac{u}{\sigma} \sin\varphi \right) \right].$$

Let $\frac{n_N}{n_0}$ be the relative count of the noise alone. We then obtain the signal part by subtracting $\frac{n_N}{n_0}$ from $\frac{n}{n_0}$

$$\frac{1}{\sqrt{\pi}} \frac{u}{\sigma} \cos\varphi = \frac{n(0) - n_N(0)}{n_0} \quad (5)$$

$$\frac{1}{\sqrt{\pi}} \frac{u}{\sigma} \sin\varphi = \frac{n(90) - n_N(90)}{n_0}$$

$$\frac{u}{\sqrt{\pi}\sigma} = \frac{1}{n_0} \sqrt{[n(0) - n_N(0)]^2 + [n(90) - n_N(90)]^2} \quad (6)$$

We notice that the measurement does not yield the signal itself but the signal-to-noise ratio. This has the great advantage that fluctuations in the noise level affect the measurement only slightly. The signal can still be measured as accurately as the noise is known, independent of how small the signal-to-noise ratio may actually be.

The response of the PCD is linear for small enough signal-to-noise ratios. The signal phase does not have to be known if the data are processed in the described way.

Statistical Error of the Measurement:

If all samples were statistically independent, the mean error of n counts is $\Delta n = \sqrt{n}$.

Due to the finite bandwidth of the noise, the samples are not statistically independent and the error will be larger than \sqrt{n} .

The variance of the sample means is given by (5)

$$\sigma_M^2 = \frac{\sigma^2}{n} + \frac{2}{n} \sum_{k=1}^{n-1} \left(1 - \frac{k}{n}\right) R(kt_0). \quad (7)$$

σ^2 is the variance of the process being sampled. $R(\tau)$ is its correlation function, N is the total number of counts. t_0 is the time between samples.

$\frac{\sigma_M}{\sigma} = \frac{|\Delta n|}{n}$ is the relative error of the measurement. We can write this

$$\left| \frac{\Delta n}{n} \right|^2 = \frac{1}{n} + \frac{2}{n} \sum_{k=1}^{n-1} \left(1 - \frac{k}{n}\right) \rho(kt_0) \quad (8)$$

with $\rho(\tau) = \frac{R(\tau)}{R(0)}$ the normalized correlation function.

The correlation function for clipped signal + noise has been derived by Davenport (6) and Galejs (7). If V is the clipping level we have

$$R(\tau) = \frac{V^2}{\pi} \sum_{\ell=0}^{\infty} \sum_{k=0}^{\infty} \frac{1}{k!} [\sigma_N^2 \rho_N(\tau)]^k h_{\ell k}^2 e_{\ell} \cos \ell \omega_0 \tau. \quad (9)$$

σ_N^2 is the variance of the noise before clipping.

$\rho_N(\tau) = \frac{R_N(\tau)}{\sigma^2}$ is the normalized correlation function of the noise

before clipping.

$$e_{\ell} = \begin{cases} 1 & \text{for } \ell = 0 \\ 2 & \text{for } \ell > 0 \end{cases}$$

$$h_{\ell k} = \frac{\pi b^{-\frac{1}{2}k} \xi^{-\frac{1}{2}\ell}}{\ell! \Gamma\left(\frac{2-k-\ell}{2}\right)} F\left(\frac{k+\ell}{2}\right); \ell+1; -\xi$$

F is the confluent hypergeometric function,

$$b = a + \frac{1}{2}\sigma_N^2 \quad (\text{for hard limiting } a \rightarrow 0),$$

$$\xi = \frac{u^2}{\sigma_N^2}.$$

For small signal-to-noise ratios $\xi \ll 1$, and we may expand the hypergeometric function

$$F \approx 1 - \frac{k+\ell}{2(\ell+1)} \xi.$$

Retaining only linear terms in ξ we obtain

$\ell = 0$:

$$R_0 = V^2 \sum_{k=\text{odd}} \frac{1}{k!} \left[\frac{\sigma^2 \rho_N(\tau)}{b} \right]^k \frac{(1-k\xi)}{\Gamma^2\left(1-\frac{k}{2}\right)}$$

$$\text{with } \Gamma\left(1-\frac{1}{2}k\right) = \sqrt{\pi} \frac{(-2)^s}{1 \cdot 3 \cdot 5 \cdots (2s-1)}; \quad k = 2s+1$$

$$\begin{aligned} R_0 &= 2 \frac{V^2}{\pi} \left[\left(\rho_N + \frac{1}{2} \frac{\rho_N^3}{3} + \frac{1}{2} \cdot \frac{3}{4} \cdot \frac{\rho_N^5}{5} + \cdots \right) \right. \\ &\quad \left. - \xi \left(\rho_N + \frac{1}{2} \rho_N^3 + \frac{1}{2} \cdot \frac{3}{4} \cdot \rho_N^5 + \cdots \right) \right] \\ &= \frac{2V^2}{\pi} \left[\arcsin \rho_N - \xi \frac{\rho_N}{\sqrt{1-\rho_N^2}} \right] \end{aligned}$$

$\ell = 1$:

$$R_1 = 2V^2 \sum_{k=\text{even}} \frac{1}{k!} \frac{2^k \rho_N^k \xi}{\Gamma^2\left(\frac{1}{2} - \frac{k}{2}\right)} \cos \omega_0 \tau$$

$$\Gamma\left(\frac{1}{2} - \frac{k}{2}\right) = \sqrt{\pi} \frac{(-2)^s}{1 \cdot 3 \cdot 5 \cdots (2s-1)}; \quad k = 2s$$

$$\begin{aligned}
 R_1 &= \frac{2V^2}{\pi} \xi \left[1 + \frac{1}{2} \rho_N^2 + \frac{1}{2} \cdot \frac{3}{4} \cdot \rho_N^4 + \dots \right] \cos \omega_0 \tau \\
 &= \frac{2V^2}{\pi} \frac{\xi}{\sqrt{1-\rho_N^2}} \cos \omega_0 \tau.
 \end{aligned}$$

Up to linear terms in ξ we obtain finally

$$R(\tau) = \frac{2V^2}{\pi} \left[\arcsin \rho_N + \xi \frac{\cos \omega_0 \tau - \rho_N}{\sqrt{1-\rho_N^2}} \right]. \quad (10)$$

For a concrete example let us consider two different spectral densities of the narrow band noise. $\Delta\omega$ is defined to be the total amplitude bandwidth of the spectral density.

1. A lorentzian

$$S(\omega) = \frac{S_0}{1 + 12 \frac{(\omega - \omega_0)^2}{\Delta\omega^2}} \quad (11)$$

2. A gaussian

$$S(\omega) = S_0 \exp \left[-8 \cdot \ln 2 \frac{(\omega - \omega_0)^2}{\Delta\omega^2} \right].$$

1. Lorentzian

$$S_+(\omega) = \frac{S_0}{1 + 12 \frac{(\omega - \omega_0)^2}{\Delta\omega^2}} \quad \text{for } \omega > 0$$

$$S_-(\omega) = \frac{S_0}{1 + 12 \frac{(\omega + \omega_0)^2}{\Delta\omega^2}} \quad \text{for } \omega < 0$$

$$R_N(\tau) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} S(\omega) e^{i\omega\tau} d\omega$$

We write approximately

$$R_N(\tau) = \frac{1}{2\pi} \left\{ \int_{-\infty}^{+\infty} S_+(\omega) e^{i\omega\tau} d\omega + \int_{-\infty}^{+\infty} S_-(\omega) e^{i\omega\tau} d\omega \right\} = R_+ + R_- .$$

The integrals are readily evaluated by contour integration

$$\rho_N(\tau) = e^{-\frac{\Delta\omega\tau}{2\sqrt{3}}} \cdot \cos \omega_0 \tau. \quad (12)$$

We obtain

$$\rho(\tau) = \frac{2}{\pi} \arcsin \left[e^{-\frac{\Delta\omega \cdot \tau}{2\sqrt{3}}} \cdot \cos \omega_0 \tau \right]$$

with $\tau = kt_0$ and $t_0 = \frac{2\pi}{\omega_0}$ we have

$$\left(\frac{\Delta n}{n}\right)^2 = \frac{1}{n} \left\{ 1 + \frac{4}{\pi} \sum_{k=1}^{n-1} \arcsin \left[e^{-\frac{\pi k \alpha}{\sqrt{3}}} \cos 2\pi k \right] \right\}$$

with $\alpha = \frac{\Delta\omega}{\omega_0}$. If α is not too small the sum converges rapidly. We define a factor β by $\Delta n = \beta \sqrt{n}$ as the factor by which the mean error is increased over the uncorrelated error. In Figure 1, β is plotted as a function of α .

2. Gaussian

$$S_+ = S_0 \exp \left[-8 \ln 2 \frac{(\omega - \omega_0)^2}{\Delta\omega^2} \right] \quad \text{for } \omega > 0 \quad (13)$$

$$S_- = S_0 \exp \left[-8 \ln 2 \frac{(\omega + \omega_0)^2}{\Delta\omega^2} \right] \quad \text{for } \omega < 0$$

$$\rho_N(\tau) = e^{-\frac{\tau^2 \Delta\omega^2}{32 \ln 2}} \cdot \cos \omega_0 \tau. \quad (14)$$

We obtain for $\frac{\Delta n}{n}$ in an analogous way

$$\left(\frac{\Delta n}{n}\right)^2 = \frac{1}{n} \left\{ 1 + \frac{4}{\pi} \sum_{k=1}^{n-1} \arcsin \left[e^{-\frac{\pi^2 k^2 \alpha^2}{8 \ln 2}} \cos 2\pi k \right] \right\} .$$

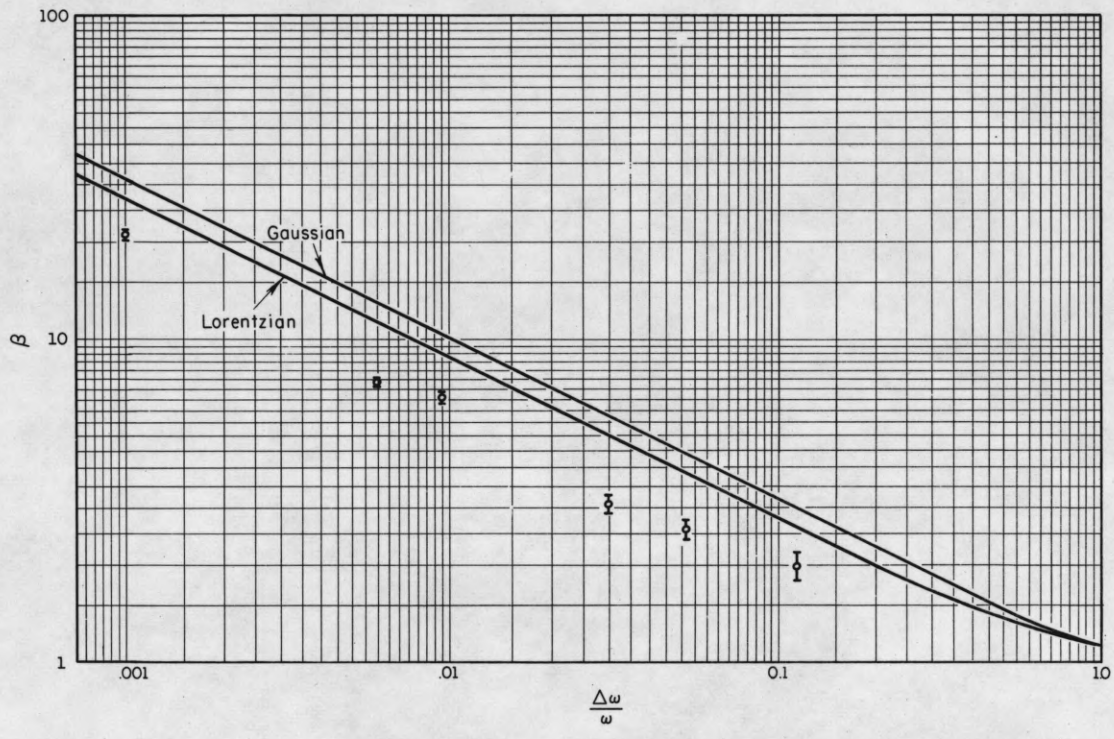


Figure 1. Statistical error as a function of bandwidth.

This sum converges very rapidly. Again, β is plotted in Figure 1.

Output Signal-to-Noise Ratio

The measurement yields the signal-to-noise ratio according to

$$\frac{u}{\sigma} = \frac{\sqrt{\pi}}{n_0} \sqrt{(n_S - n_N)_{0}^2 + (n_S - n_N)_{90}^2} . \quad (15)$$

The error of this measurement (r.m.s. deviation) for small signal-to-noise ratios is given by

$$\Delta \frac{u}{\sigma} = \sqrt{\pi} \frac{\beta}{n_0} \sqrt{n_S + n_N}$$

where β is the increase of the statistical error over the uncorrelated error

$\Delta n = \beta \sqrt{n}$. $\frac{u}{\Delta u}$ can be considered as the output amplitude signal-to-noise ratio or with $S = u^2$; $N = \sigma^2$

$$\left(\frac{S}{N}\right)_{\text{out}} = \frac{(n_S - n_N)_{0}^2 + (n_S - n_N)_{90}^2}{\beta^2 (n_S + n_N)}$$

for small signal-to-noise ratios $n_S \approx \frac{1}{2}n_0$; $n_N \approx \frac{1}{2}n_0$

$$(n_S - n_N)_{0}^2 + (n_S - n_N)_{90}^2 = \frac{n_0^2}{\pi} \left(\frac{S}{N}\right)_{\text{in}}$$

hence

$$\left(\frac{S}{N}\right)_{\text{out}} = \frac{n_0}{\pi \beta^2} \left(\frac{S}{N}\right)_{\text{in}} . \quad (16)$$

We consider first the signal to be embedded in wide-band noise. The samples taken in the PCD will then be completely uncorrelated ($\beta = 1$) and the result of the measurement will be ^{*)}

$$\left(\frac{S}{N}\right)_{\text{out}} = \frac{n_0}{\pi} \left(\frac{S}{N}\right)_{\text{in}} . \quad (17)$$

^{*}This is only approximately true, because (15) is derived under the assumption that we are dealing with narrow-band noise.

We now pass signal + noise through a filter with a system function $H(i\omega)$ which reduces the variance of the input noise, but in consequence introduces correlation between adjacent samples.

$$\left(\frac{S}{N}\right)_{\text{out}} = \frac{n_o}{\pi\beta^2} \left(\frac{S}{N}\right)_F$$

where $\left(\frac{S}{N}\right)_F$ is the signal-to-noise ratio at the output of the filter. The filter does not act on the signal but only on the noise.

Let $G_{\text{in}}(\omega)$ be the spectral density at the input, $G_o(\omega)$ the spectral density at the output, then

$$G_o(\omega) = |H(i\omega)|^2 G_{\text{in}}(\omega)$$

$$N = \frac{1}{2\pi} \int_{-\infty}^{+\infty} G(\omega) d\omega$$

hence

$$\left(\frac{S}{N}\right)_F = \frac{S}{\frac{1}{2\pi} \int_{-\infty}^{+\infty} |H(i\omega)|^2 G_{\text{in}}(\omega) d\omega}$$

or since

$$N_{\text{in}} = \frac{1}{2\pi} \int_{-\infty}^{+\infty} G_{\text{in}}(\omega) d\omega$$

$$\left(\frac{S}{N}\right)_F = \left(\frac{S}{N}\right)_{\text{in}} \frac{\int_{-\infty}^{+\infty} G_{\text{in}}(\omega) d\omega}{\int_{-\infty}^{+\infty} |H(i\omega)|^2 G_{\text{in}}(\omega) d\omega}$$

$$\left(\frac{S}{N}\right)_{\text{out}} = \frac{n_o}{\pi\beta^2} \frac{\int_{-\infty}^{+\infty} G_{\text{in}}(\omega) d\omega}{\int_{-\infty}^{+\infty} |H(i\omega)|^2 G_{\text{in}}(\omega) d\omega} \left(\frac{S}{N}\right)_{\text{in}}$$

Let us consider white input noise up to a frequency ω_N

$$G(\omega) = G_o \quad \text{for } \omega \leq \omega_N$$

$$= 0 \quad \text{for } \omega > \omega_N$$

and a system function $H(i\omega) = 1$ for $\omega_0 - \frac{1}{2}\Delta\omega < \omega < \omega_0 + \frac{1}{2}\Delta\omega$
 $= 0$ for outside.

This results in

$$\left(\frac{S}{N}\right)_{\text{out}} = \frac{n_0}{\pi\beta^2} \frac{\omega_N}{\Delta\omega} \left(\frac{S}{N}\right)_{\text{in}} \quad (18)$$

for small values of $\frac{\Delta\omega}{\omega_0}$

$$\beta \sim \sqrt{\frac{\omega_0}{\Delta\omega}}.$$

We notice that there is an advantage to filter out frequencies $> \omega_0$ before coincidence detection. Nothing, however, is gained by filtering out low frequencies, except near zero. A band-pass filter of width $\sim \omega_0$ is equally as effective as a narrow band-pass filter.

The necessity for introducing a filter before coincidence detection arises because the signal is sampled at a low rate. The effect of sampling can be illustrated more easily if the coincidence detector is replaced by a linear integrator. In the linear integrator the amplitude of the signal at the instant of sampling are summed as compared to the polarity coincidence detector where the amplitudes are converted to ± 1 before summing. The effect of summing the sampled signal is the same as transmitting the original signal through a comb filter. The center frequencies of the pass bands are separated from the neighboring bands by ω_0 , where ω_0 is the sampling rate. The width of each pass band is approximately the reciprocal of the total integration time. Since the desired part of the signal is monochromatic at frequency ω_0 , it is desirable to filter out all of the pass bands except the band centered at ω_0 . This can be accomplished by using a filter of width ω_0 .

Experimental Part:

The technical realization of the polarity coincidence principle presents no essential problems. Figure 2 shows the block diagram of an operational model that is presently in laboratory use for several months. The use of a synchronous detector in an experiment usually proceeds in the following way. A generalized carrier (which can be an rf-signal, an electron beam, a molecular beam, and so forth) interacts with the system under investigation. If the carrier is modulated with a certain frequency any change in the system that results from the interaction will exhibit the same frequency component. This change is detected and transformed into an ac signal which is cross-correlated with the modulation frequency.

In the case of the polarity coincidence detector, the signal is passed through a limiting amplifier. In the present model the limiting amplifier consists of three wideband feedback loops with an individual gain of 50 and a rise time of 0.1 microsecond. Limiting is accomplished by a pair of biased diodes after each stage. The output is a square wave of 40 V_{pp} amplitude.

The reference signal is split into two channels 90° out of phase, where it is converted into sampling pulses of 0.5 microsecond duration. Sampling pulses and output of the limiting amplifier are applied to the two control grids of 6BN6 tube. The output of the coincidence tubes triggers multi-vibrators that generate standard pulses for the counters. In order to cancel out fluctuations in bias level or noise level, signal + noise is counted for a given time determined by a pre-set count of the master counter. The modulation is switched off and the noise alone is recorded for the same time on a different set of counters. The noise + signal and the noise counts alternate typically every 10 seconds.

The instrument operates in the range 20 c/sec-20 kc/sec. The performance is demonstrated by the following set of diagrams.

Figure 3 shows a plot of the output signal, defined as $\sqrt{(n_S - n_N)_0^2 + (n_S - n_N)_{90}^2}$ in arbitrary units versus the input amplitude signal-to-noise ratio. The signal frequency was 10 kc/sec. We notice the deviation from linearity for signal-to-noise ratios larger than 0.2.

The response of the individual channels to a phase shift between signal and reference signal is demonstrated by Figure 4, where the counting rates of the individual channels are plotted versus the phase shift between signal and reference signal. The signal-to-noise ratio was 0.05, frequency 10 kc. The measured values fit smoothly a sin and cos dependence.

In order to check the increase of the error with decreasing bandwidth of the noise, series of 50 individual readings of the counting rate were taken. The computed statistical error of a single reading together with its error is plotted in Figure 1 versus the relative bandwidth of the noise. Although the error is smaller than predicted for a Lorentzian or Gaussian spectral density, the qualitative behavior is confirmed.

The appendix contains the complete circuit diagrams of the instrument.

Error Sources:

a) Fluctuations of threshold:

The main uncertainty is introduced in the first limiting stage of the amplifier. We have

$$\frac{n(0)}{n_0} = \frac{1}{2} \left[1 - \frac{2x_+}{\sqrt{2\pi} \cdot \sigma} + \frac{2u}{\sigma\sqrt{\pi}} \cos\varphi \right]$$

$$\frac{\Delta n(0)}{n_0} = \frac{\Delta x_+}{\sigma\sqrt{2\pi}}$$

σ is the variance of the noise after amplification by the first stage.

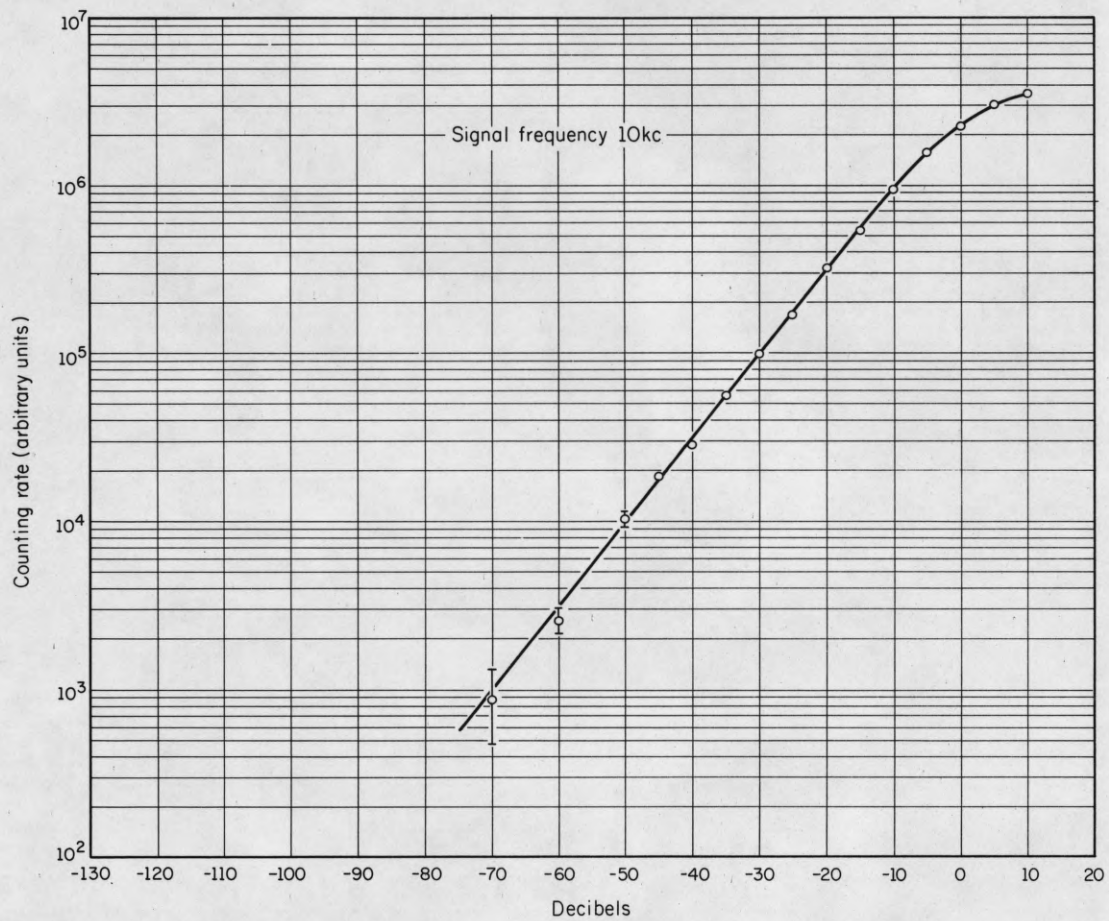


Figure 3. Response of polarity coincidence detector to input signal-to-noise ratio. (0 db corresponds to S/N=1)

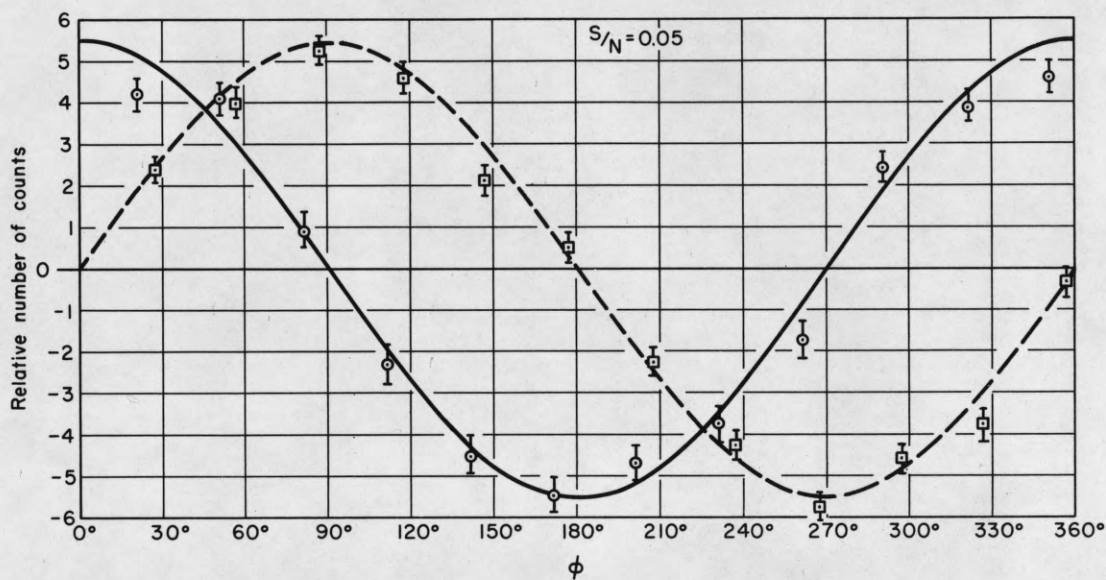


Figure 4. Counting rate as a function of phase shift between signal and reference signal.

The error of the measurement can then be written

$$\Delta \frac{u}{\sigma\sqrt{\pi}} = \sqrt{\frac{\beta^2}{n_0} + \frac{\Delta x_+^2 + \Delta x_+'^2}{2\pi\sigma^2}} . \quad (19)$$

The conditions for the influence of the threshold fluctuation to be small compared to the statistical error can then be written

$$\frac{(\Delta x)^2}{\pi\sigma^2} < \frac{\beta^2}{n_0} .$$

σ is limited by the dynamical range of the amplifier. In a typical example one may have $\sigma = 100V$

$$\Delta x = 0.1V$$

$\beta = \sqrt{2}$. This requires $n_0 < 6 \times 10^6$, and the minimum detectable signal-to-noise ratio would be 5×10^{-5} .

b) Phase Fluctuations:

When the signal is passed through narrow band filters the possibility of phase fluctuations must be considered. We have

$$\Delta n(0) = -n_0 \frac{u}{\sigma\sqrt{\pi}} \cos\varphi \cdot \Delta\varphi$$

$$\Delta n(90) = n_0 \frac{u}{\sigma\sqrt{\pi}} \sin\varphi \cdot \Delta\varphi .$$

This contributes to the error according to

$$\Delta \frac{u}{\sigma\sqrt{\pi}} = \sqrt{\frac{\beta^2}{n_0} + \frac{u^2}{\pi\sigma^2} \Delta\varphi^2} . \quad (20)$$

In the limiting case $n_0 \rightarrow \infty$ we have $\frac{\Delta u}{u} = \Delta\varphi$. The influence of phase fluctuations is negligible in most situations.

c) Noise Fluctuations:

A fluctuation in the noise level changes the counting rate by

$$\frac{\Delta n_S(0)}{n_0} = \frac{1}{\sqrt{\pi}} \left(\frac{x_+}{\sqrt{2}} - u \cos\varphi \right) \frac{\Delta\sigma}{\sigma^2} .$$

An analogous expression holds for $\Delta n_S(90)$.

The error in the signal-to-noise ratio becomes

$$\Delta \left(\frac{u}{\sigma} \right) = \sqrt{\pi} \sqrt{\frac{\beta^2}{n_0} + \frac{\Delta\sigma^2}{n_0^2 u^2 \sigma^2} \left\{ (n_S - n_N)_{00}^2 \left[\frac{x_+^2}{2} + \left(\frac{x_+}{\sqrt{2}} - u \cos\varphi \right)^2 \right] \right.} \\ \left. + (n_S - n_N)_{90}^2 \left[\frac{x_+^2}{2} + \left(\frac{x_+}{\sqrt{2}} - u \sin\varphi \right)^2 \right] \right\}} .$$

In two limiting cases this simplifies considerably

$$1) \quad u \ll x_+ \quad x_+ \approx x_+' = x$$

$$\Delta \frac{u}{\sigma} = \sqrt{\pi} \sqrt{\frac{\beta^2}{n_0} + \frac{1}{\pi} \frac{x^2}{\sigma^2} \left(\frac{\Delta\sigma}{\sigma} \right)^2} . \quad (21)$$

The minimum detectable signal-to-noise ratio then becomes

$$\left(\frac{u}{\sigma} \right)_{\min} = \frac{x}{\sigma} \cdot \frac{\Delta\sigma}{\sigma} .$$

$$2) \quad u \gg x_+$$

$$\Delta \frac{u}{\sigma} = \sqrt{\pi} \sqrt{\frac{\beta^2}{n_0} + \frac{1}{\pi} \frac{u^2}{\sigma^2} \left(\frac{\Delta\sigma}{\sigma} \right)^2 (\cos^4\varphi + \sin^4\varphi)} . \quad (22)$$

The minimum relative error becomes

$$\frac{\Delta u}{u} = \frac{\Delta\sigma}{\sigma} \sqrt{\cos^4\varphi + \sin^4\varphi} .$$

$\sqrt{\cos^4\varphi + \sin^4\varphi}$ varies between 1 and $\frac{1}{\sqrt{2}}$.

d) Linearity:

In order to get an estimate of the deviation of the measurement from linearity we expand the error function in the expression for $\frac{n_S(0)}{n_o}$.

$$\frac{n_S(0)}{n_o} \approx \frac{1}{2} \left[1 - \frac{2}{\sqrt{\pi}} \left(\frac{x}{\sigma\sqrt{2}} - \frac{u}{\sigma} \cos\varphi \right) - \frac{2}{3\sqrt{\pi}} \left(\frac{x}{\sigma\sqrt{2}} - \frac{u}{\sigma} \cos\varphi \right)^3 + \dots \right]$$

$$\frac{n_N(0)}{n_o} \approx \frac{1}{2} \left[1 - \frac{2}{\sqrt{\pi}} \frac{x}{\sigma\sqrt{2}} - \frac{2}{3\sqrt{\pi}} \frac{x^3}{\sigma^3 2^{3/2}} - \dots \right]$$

$$\frac{n_S - n_N}{n_o} = \frac{u}{\sqrt{\pi} \sigma} \left[1 + \left(\frac{u}{\sqrt{\pi} \sigma} \right)^2 + \dots \right].$$

For a signal-to-noise ratio of 0.1, i.e., the deviation from linearity is less than one per cent.

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