



25 **Abstract**

26 This study assessed the influence of training load, exposure to match play and sleep  
27 duration on two daily wellbeing measures in youth athletes. Forty-eight youth athletes (age  
28  $17.3 \pm 0.5$  years) completed a daily wellbeing questionnaire (DWB), the Perceived  
29 Recovery Status scale (PRS), and provided details on the previous day's training loads  
30 (TL) and self-reported sleep duration (sleep) every day for 13 weeks ( $n = 2727$ ). A linear  
31 mixed model assessed the effect of TL, exposure to match play and sleep on DWB and  
32 PRS. An increase in TL had a *most likely small* effect on muscle soreness ( $d=-0.43\pm0.10$ )  
33 and PRS ( $d=-0.37\pm0.09$ ). Match play had a *likely small* additive effect on muscle soreness  
34 ( $d=-0.26\pm0.09$ ) and PRS ( $d=-0.25\pm0.08$ ). An increase in sleep had a *most likely moderate*  
35 effect on sleep quality ( $d=0.80\pm0.14$ ); a *most likely small* effect on DWB ( $d=0.45\pm0.09$ )  
36 and fatigue ( $d=0.42\pm0.11$ ); and a *likely small* effect on PRS ( $d=0.25\pm0.09$ ). All other  
37 effects were *trivial* or did not reach the pre-determined threshold for practical significance.  
38 The influence of sleep on multiple DWB subscales and the PRS suggests that practitioners  
39 should consider the recovery of an athlete alongside the training stress imposed when  
40 considering deviations in wellbeing measures.

41

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## 44 **Introduction**

45 It is essential that an optimal balance between stress and recovery is reached when  
46 constructing athletic development programmes (Rowbottom, 2000). The stress-recovery  
47 balance dictates that when a body is subjected to a stressor (e.g. training load, examination  
48 stress or social pressures), an appropriate amount of recovery time (e.g. sleep) is required  
49 to maintain equilibrium (Kellmann, 2010). In sport, failure to maintain the stress-recovery  
50 balance can result in de-training, injury, illness or overtraining (Hulin et al., 2014;  
51 Meeusen et al., 2013; Putlur et al., 2004). Consequently, it has become commonplace to  
52 monitor an athlete's stress-recovery balance using subjective daily wellbeing  
53 questionnaires (DWB; Saw, Main, & Gastin, 2015). These questionnaires, as self-report  
54 measures, are now widespread in professional adult sport due to their inexpensiveness,  
55 time efficiency and ease of analysis (Saw et al., 2015; Saw, Main, & Gastin, 2016), but are  
56 also becoming increasingly prominent at youth level (Noon, James, Clarke, Akubat, &  
57 Thake, 2015; Sawczuk, Jones, Scantlebury, & Till, 2018). However, the stress-recovery  
58 balance at youth level may vary in response to training stressors as athletes attempt to cope  
59 with educational (e.g. academic examinations), maturational (e.g. hormonal changes) and  
60 social (e.g. pressure to succeed, relationships and peer pressure) demands alongside their  
61 sporting endeavours (Mountjoy et al., 2008; Siesmaa, Blitvich, & Finch, 2011). In order  
62 for wellbeing questionnaires to be fit for purpose, it is important that they are responsive to  
63 the stress and recovery experienced by the athlete. In sport, the primary stressor imposed  
64 upon an athlete by the coaching staff, aimed at enhancing their athletic development, is the  
65 training stimulus, whereas the primary mechanism of recovery is sleep (Halsen, 2014a,  
66 2014b). However, whilst there is a growing body of literature considering the influence of  
67 training load on DWB (Buchheit et al., 2013; Thorpe et al., 2017), studies considering their  
68 relationship with sleep are scarce (Sawczuk et al., 2018).

69

70 The influence of training load on overall DWB scores appears surprisingly contentious  
71 given their widespread use in sport (Saw et al., 2015). Buchheit and colleagues (2013)  
72 found a DWB and all its individual subscales (i.e. measures of fatigue, muscle soreness,  
73 sleep quality, stress and mood) to be related to training load in Australian Rules football  
74 players during the pre-season phase. However, other studies in Australian Rules football  
75 players (Gallo, Cormack, Gabbett, & Lorenzen, 2017) and youth athletes (Sawczuk et al.,  
76 2018) have argued that the overall DWB score is not influenced by the previous day's  
77 workload. It is possible that the difference between these studies is due to the training loads  
78 present. Buchheit and colleagues (2013) reported a weekly training load of over 10,000  
79 AU in their study, whereas both studies reporting no change had weekly training loads of  
80 around 1,750 AU (Gallo et al., 2017; Sawczuk et al., 2018). Furthermore, only Buchheit  
81 and colleagues (2013) provided a DWB subscale analysis which showed all subscales to  
82 have a *small* association with training load. Given that very high training loads are  
83 believed to affect mood and stress prior to the onset of the overtraining syndrome  
84 (Meeusen et al., 2013) and neither Gallo and colleagues (2017) nor Sawczuk and  
85 colleagues' (2017) studies included very high training loads, it is possible that a masking  
86 effect between subscales occurred within the studies showing no relationship between  
87 DWB and training load. Therefore, fatigue, muscle soreness and sleep quality may have  
88 been affected by training load but a lack of association with other subscales could have  
89 blunted the overall response. Previous studies have shown that individual subscales such as  
90 fatigue (Thorpe et al., 2015, 2017), muscle soreness (Montgomery & Hopkins, 2013), and  
91 the PRS (Sawczuk et al., 2018) may be affected by training load and exposure to match  
92 play at training loads between 1,750 and 2,000 AU, supporting this hypothesis. However,  
93 none of these studies analysed the effect of these training loads on mood or stress

94 subscales. A study considering the effect of moderate weekly training loads (circa 2,000  
95 AU per week), including exposure to match play, on the overall DWB score and all  
96 individual subscales, and a comparison with the PRS, a standalone scale shown to be  
97 sensitive to training loads (Sawczuk et al., 2018), is therefore merited.

98

99 In order to recover from the training and match stimuli encountered by athletes, it is  
100 important that sleep is optimised (Halson, 2014b; Tuomilehto et al., 2017). Previous  
101 research has indicated that sleep can affect sporting performance (Fullagar et al., 2015;  
102 Mah, Mah, Kezirian, & Dement, 2011), risk of illness (Cohen, Doyle, Alper, Janicki-  
103 Deverts, & Turner, 2009; Prather, Janicki-Deverts, Hall, & Cohen, 2015) and wellbeing  
104 measures (Oginska & Pokorski, 2006). Despite this evidence showing the importance of  
105 sleep, previous studies have avoided the use of self-reported sleep duration as a predictor  
106 of changes in wellbeing measures due to its perceived lack of validity when compared to  
107 actigraphy measures (Lauderdale, Knutson, Yan, Liu, & Rathouz, 2008). However, it has  
108 recently become apparent that in athletic populations self-reported sleep duration is a valid  
109 measure when compared to actigraphy (Caia et al., 2017; Kölling, Endler, Ferrauti, Meyer,  
110 & Kellmann, 2016), although it maintains its systematic bias of overestimating sleep  
111 duration by around 1 hour. These new findings, alongside suggestions that perceptions of  
112 sleep quality are not always congruent with objective measures (Krystal & Edinger, 2008),  
113 provide rationale for the use of self-reported sleep duration as a predictor of changes in  
114 wellbeing. To date, the only study to have considered the influence of sleep duration on a  
115 sport specific wellbeing measure found DWB to be related to short, but not extended, sleep  
116 durations and found no relationship with the PRS (Sawczuk et al., 2018). However, the  
117 study only took place on four weekdays, which may not be representative of a youth  
118 athlete population as participants would likely have had to be at school by 8.30am on those

119 weekdays, whereas their sleep durations may not be similarly restricted at weekends.  
120 Furthermore, the inclusion of a sleep quality measure within the overall DWB score could  
121 have skewed the true relationship, but an individual subscale analysis was not provided in  
122 the study. Therefore, there is scope for a study considering all seven days, in which the  
123 influence of self-reported sleep length on DWB, its individual subscales and the PRS is  
124 considered, alongside training loads and match stress. Consequently, the aim of this study  
125 was to assess the influence of training load, exposure to match play and self-reported sleep  
126 duration on a DWB, its individual subscales (i.e. muscle soreness, fatigue, sleep quality,  
127 mood and stress) and PRS.

128

## 129 **Methods**

### 130 *Participants*

131 Forty-eight male and female adolescent team sport athletes aged 16-18 years (age  $17.3 \pm$   
132  $0.5$  years, height  $172.8 \pm 18.3$  cm, body mass  $73.6 \pm 12.8$  kg) participated in this study.  
133 Participants were recruited from a local independent school in the United Kingdom (UK),  
134 where they were members of the school's sport scholarship programme. The sports cricket  
135 ( $n=5$ ), football ( $n=10$ ), hockey ( $n=10$ ), netball ( $n=10$ ) and rugby union ( $n=13$ ) were  
136 represented by athletes competing at club/school ( $n=29$ ), professional academy ( $n=6$ ),  
137 county/regional ( $n=10$ ) and international ( $n=3$ ) standard in their respective sports. All  
138 participants were made aware of the benefits and risks of the study, and written informed  
139 consent was provided by all participants and their parents prior to the study. Ethics  
140 approval was granted by the University Ethics Committee.

141

### 142 *Procedures*

143 Participants completed an online Google Docs (Google Forms, Google, CA, USA)  
144 questionnaire before 11am , and prior to their first training session of the day on training  
145 days, every morning for a 13-week period. The questionnaire was emailed to participants  
146 at 6am every morning and on weekdays they were verbally reminded to complete it if they  
147 hadn't done so by 10.30am. The form included a DWB related to fatigue, muscle soreness,  
148 sleep quality, stress and mood (McLean, Coutts, Kelly, McGuigan, & Cormack, 2010),  
149 with each subscale rated 1-5 and totalled to an overall score out of 25; the PRS (Laurent et  
150 al., 2011); self-reported sleep length (in hours) and 24 hour training load recall. For the 24-  
151 hour training load recall, participants provided information with regards to the type,  
152 intensity and duration of each session from the previous day. Type included technical  
153 training, strength and conditioning training, personal gym and matches. All participants  
154 were scheduled to complete two technical training sessions, two strength and conditioning  
155 training sessions and one match per week as part of their school programmes, but club  
156 programmes varied widely by individual. Participants could participate in multiple session  
157 types on a single day, but every day where they participated in a match was used to  
158 calculate the additive effect of exposure to match play on wellbeing measures. The  
159 intensity of each session was rated via the Borg category ratio-10 scale (Foster et al., 2001)  
160 choosing the respective descriptor, which was converted to the appropriate rating of  
161 perceived exertion (RPE) number and multiplied by the session duration (in minutes) to  
162 provide the session-RPE (s-RPE). The sum of all s-RPE's on a single day gave the daily  
163 training load. The temporal robustness of the s-RPE method over 24 hours has previously  
164 been confirmed (Phibbs et al., 2017; Scantlebury, Till, Sawczuk, Phibbs, & Jones, 2017),  
165 and the between-day reliability (typical error as a coefficient of variation) of DWB and  
166 PRS has previously been evaluated in this cohort as 11.7% and 8.5% respectively  
167 (Sawczuk et al., 2018).

168

169 *Statistical analyses*

170 For statistical analysis, DWB and PRS scores were converted to scores out of 100.. Data  
171 were analysed using SAS University Edition (SAS Institute, Cary, NC). A linear mixed  
172 model (via Proc Mixed) was used to evaluate the influence of training load, sleep length  
173 and match stress on the dependent variables. The overall DWB score, individual DWB  
174 subscales (fatigue, muscle soreness, sleep quality, stress and mood) and PRS score were  
175 used as dependent variables. Sport (referring to the athlete's sport), week (referring to the  
176 week of the study), and day (referring to the day of the week) were added as fixed factors  
177 and provided estimated means for the wellbeing scores for each factor. Training load and  
178 sleep duration were mean centred by individual and added as time varying covariates. The  
179 additive effect of exposure to match play was calculated by a dummy covariate on any day  
180 where the participant reported they had taken part in a match. Athlete\*training load\*sleep  
181 duration was added as an unstructured random effect to allow for variation in the effect of  
182 the covariates on the dependent variables between individuals to be calculated. Due to the  
183 difficulty in obtaining correlation coefficients from mixed effects models with complicated  
184 random effects structures (Roy, 2006), the effect of the covariates was calculated by  
185 assessing a two standard deviation (2 SD) difference in the covariate. This evaluates the  
186 difference between a typically high and typically low training day/sleep duration, and  
187 'ensures congruence between Cohen's threshold magnitudes for correlations and  
188 standardized differences' (Hopkins, Marshall, Batterham, & Hanin, 2009).

189

190 Results were analysed for practical significance using magnitude-based inferences  
191 (Hopkins et al., 2009). The threshold for a change to be considered practically important  
192 (the smallest worthwhile change; SWC) was set as 0.2 x observed between participant SD,



193 based on Cohen's *d* effect size (ES) principle. Thresholds for ES were set as: 0.2 *small*; 0.6  
194 *moderate*; 1.2 *large*, 2.0 *very large*. The ES of random effects were doubled to fit the same  
195 ES criteria, as opposed to halving the thresholds (Hopkins, 2015). The probability that the  
196 magnitude of change was greater than the SWC was rated as: <0.5% *almost certainly not*;  
197 0.5-5% *very unlikely*; 5-25% *unlikely*; 25-75% *possibly*; 75-95% *likely*; 95-99.5% *very*  
198 *likely*; >99.5% *most likely* (Hopkins et al., 2009). In those situations where the likelihood  
199 of the magnitude of change was classified as *most likely* greater than the SWC and the ES  
200 was greater than 0.6 (i.e. *moderate*), the magnitude-based inference given is compared  
201 against the *moderate* effect size rather than the SWC. Effect sizes are reported ES;  $\pm$  90%  
202 confidence intervals for normally distributed fixed effects and ES; lower 90% confidence  
203 interval, upper 90% confidence interval for chi square distributed random effects.

204

## 205 **Results**

206 2727 complete data points were analysed for this study at a median response rate of 54/91  
207 completions per person. Overall, 2181 training sessions, 292 matches and 991 rest days  
208 were included. The mean daily training load was  $250 \pm 317$  AU and the mean sleep length  
209 was  $7.7 \pm 1.5$  hours. A 2 SD difference in training load equated to  $556 \pm 208$  AU, whereas  
210 the difference for sleep length was  $2.6 \pm 1.3$  hours.

211

212 Figure 1 depicts the influence of training load, exposure to match play and sleep duration  
213 on DWB, its individual subscales and PRS. There was *trivial* between-participant variation  
214 in the effect of training load on DWB ( $d = 0.18$ ; 0.09, 0.56) and *moderate* between-  
215 participant variation in its effect on PRS ( $d = 0.56$ ; 0.31, 1.42). Between-participant  
216 variation for the effect of training load on individual subscales ranged from *small* to  
217 *moderate* ( $d = 0.22$  to 0.80). Sleep duration showed *moderate* variation between

218 participants in its effect on DWB ( $d = 0.66$ ; 0.42, 1.21) and PRS ( $d = 0.64$ ; 0.38, 1.35).  
219 Variation in the response to sleep duration ranged from *small* to *large* for the individual  
220 DWB subscales ( $d = 0.33$  to 1.61).

221

222 \*\* INSERT FIGURE 1 HERE \*\*

223

## 224 **Discussion**

225 The aim of this study was to assess the influence of training load, exposure to match play  
226 and sleep duration on DWB, its individual subscales and PRS. The findings show that  
227 training load had a *small* negative effect on muscle soreness and PRS, and that this  
228 negative effect was enhanced by a *small* additive effect of exposure to match play on both  
229 measures. The influence of training load and match play exposure on all other wellbeing  
230 measures was *trivial*. Sleep duration had a *moderate* positive relationship with sleep  
231 quality and a *small* positive influence on DWB, fatigue and PRS, but no relationship with  
232 muscle soreness, mood or stress.

233

### 234 *Training load and match stress*

235 The *small* negative influence of training load and match play exposure on muscle soreness  
236 is consistent with Montgomery and Hopkins' (2013) similar findings using the s-RPE  
237 method in Australian Rules football players. However, the overall DWB score showed no  
238 relationship with training load, conflicting with research in adult Australian Rules football  
239 players (Buchheit et al., 2013), but confirming previous findings in youth athletes  
240 (Sawczuk et al., 2018). It is possible that these differences can be attributed to a masking  
241 effect caused by a lack of responsiveness to training load and match play exposure of other  
242 variables within the questionnaire (e.g. mood and stress), as suggested by a recent

243 systematic review (Saw et al., 2016). It has previously been suggested that academic and  
244 maturational stressors may hold greater importance than training stressors in this age group  
245 (Mountjoy et al., 2008; Sawczuk et al., 2018). Our study cannot add to that hypothesis, but  
246 can confirm that the moderate training loads and match stress used in this study have very  
247 little direct effect on the mood and stress of youth athletes as measured by this DWB. It is  
248 possible that at very high training loads mood and stress measures would be affected,  
249 particularly if occurring as a precursor to the overtraining syndrome (Meeusen et al.,  
250 2013), but further research is required to confirm this relationship. However,  
251 as overtraining only occurs in only 7% of elite youth footballers (Brink, Visscher, Coutts,  
252 & Lemmink, 2012), it may be difficult to confirm this hypothesis using a group mean  
253 effect as presented here rather than the individual response to training. The lack of  
254 relationship with training load does not mean that the mood and stress subscales should  
255 immediately be removed from DWB questionnaires though. Mood has previously shown  
256 associations with injuries in female collegiate soccer players (Watson, Brickson, Brooks, &  
257 Dunn, 2016) and stress can impair the recovery process for up to 96 hours (Stults-  
258 Kolehmainen, Bartholomew, & Sinha, 2014), suggesting that there is value in  
259 understanding these aspects of an athlete's wellbeing when considering alterations to their  
260 training programmes.

261

262 In addition to the *small* negative association with muscle soreness, training load and match  
263 play exposure showed a *small* negative relationship with PRS, but not with the fatigue  
264 subscale of DWB. In line with the super compensation curve dictating that following a  
265 training stimulus, an athlete will experience a period of fatigue (Bompa & Haff, 2009), it  
266 was expected that both scales would be responsive to training load and exposure to match  
267 play. The lack of association between training load and the fatigue subscale is therefore

268 surprising, but the *small* negative relationship between training load and PRS does agree  
269 with previous findings in this youth athlete cohort (Sawczuk et al., 2018). It is possible that  
270 the difference in the relationships shown is due to the weightings used (fatigue measure as  
271 a category scale vs PRS as a category-ratio scale), but it could also be due to the anchoring  
272 words employed by the scales. Although the terminology used between the scales is very  
273 similar, the PRS, via its terms "*very poorly recovered/extremely tired*" to "*very well*  
274 *recovered/highly energetic*", possibly places a greater balance on how *recovered* an athlete  
275 feels, whereas the fatigue scale, via its terms "*very fresh*" to "*always tired*", appears to  
276 consider how *tired* an athlete is. It is possible that the participants in this study related the  
277 term *recovery* to training load and *fatigue* to perceptions of sleep, which may explain the  
278 difference in results between the two scales and could also explain why the fatigue scale is  
279 much more responsive to sleep duration than the PRS in this population. Alternatively, it is  
280 possible that the difference in the two measures is due to the impact training load has on  
281 the sleep durations of the individuals. Our study did not consider the interaction between  
282 the two measures, but it is likely that those participants who had higher training loads due  
283 to evening club training sessions slept less than those who did not due to increased travel  
284 time or the need to catch up with academic work. It is therefore possible that their  
285 perceptions of fatigue could have been caused by the impact of the previous day's training  
286 load on their sleep duration rather than the sleep duration itself.

287

### 288 *Sleep duration*

289 Self-reported sleep duration had a *moderate* positive relationship with sleep quality and a  
290 *small* positive influence on DWB, fatigue and PRS. These relationships, with four out of  
291 the seven variables measured show the importance of sleep as a predictor of changes in  
292 sport specific wellbeing questionnaires and highlight this as an under-researched area. The

293 *moderate* positive relationship between sleep duration and sleep quality is unsurprising in  
294 its presence as both are subjective measures surrounding sleep, but its size is perhaps  
295 smaller than could have been predicted. Indeed, a 2 SD reduction in sleep length (2.6  
296 hours) resulted in only a 0.55 unit change in the sleep quality subscale. A possible reason  
297 for this could be the difficulty in defining good sleep quality between individuals,  
298 compared to sleep duration, which can be estimated as an arbitrary duration. For example,  
299 for some individuals good sleep quality may occur with a long sleep duration, which would  
300 provide a good correlation between the two variables, whereas for others it may be based  
301 on how many times they wake (consciously or subconsciously) during the night, which  
302 may have little relationship with the sleep duration they reported (Krystal & Edinger,  
303 2008). This is supported by the relationship between self-reported sleep duration and  
304 actigraphy based total sleep time being very large ( $r = 0.85$ ), whereas the relationship  
305 between subjective sleep quality and sleep efficiency was only small ( $r = 0.22-0.28$ ) in a  
306 recent validation study (Caia et al., 2017). However, the *moderate* relationship between the  
307 two variables indicates that they do not provide the same information so, given sleep  
308 quality has shown relationships with the other wellbeing measures within DWB (Pilcher,  
309 Ginter, & Sadowsky, 1997), there is scope for its consideration as a predictor of changes in  
310 DWB, rather than as part of the measure.

311

312 The only previous study to consider the influence of sleep duration on sport specific  
313 wellbeing questionnaires, such as DWB and PRS, occurred in youth athletes (Sawczuk et  
314 al., 2018). The authors found low sleep durations in particular to have a negative influence  
315 on DWB, but that PRS had no meaningful relationship with sleep duration. Our study is  
316 unable to provide further support for the theory that low sleep durations have a greater  
317 impact on DWB than high sleep durations, but does show that a practically meaningful

318 linear relationship can be derived between sleep duration and both DWB and PRS. The  
319 relationship between sleep duration and the total score of both measures suggests that it is  
320 more important to consider the recovery of youth athletes than any single individual  
321 stressor, such as training load, if changes in wellbeing are the main aim of the monitoring  
322 process. It remains to be seen whether lack of recovery or excessive training stressors are  
323 predictive of adverse outcomes or athletic performance when both are measured together.  
324 For example, previous studies have shown spikes in training loads (Putlur et al., 2004) and  
325 low sleep durations (Cohen et al., 2009; Prather et al., 2015) to be associated with illness  
326 risk, but no study has yet considered these variables together, in which situation one of the  
327 training stress imposed or the recovery experienced may be more important than the other.  
328

329 The *small* relationship between sleep duration and fatigue was expected given previous  
330 research (Oginska & Pokorski, 2006). However, the lack of relationship with mood and  
331 stress is less congruent with previous research (Oginska & Pokorski, 2006). It has been  
332 shown that sleep quality can also affect these variables (Pilcher et al., 1997) so it would be  
333 interesting to assess whether quality of sleep is a better predictor of these measures in a  
334 sport specific wellbeing questionnaire. The lack of relationship between sleep duration and  
335 muscle soreness can probably be attributed to the 24-72 hour time scale of increasing  
336 delayed onset muscle soreness (Cheung, Hume, & Maxwell, 2003). Our study only  
337 considered the previous day's sleep duration, which may have limited restorative  
338 capabilities over the expected three day cycle, whereas if we had considered the total sleep  
339 duration over three days, a relationship may have been found.

340

341 *Limitations*

342 Although our results add to the literature, particularly through the sample size which is  
343 much greater than the previous literature (Buchheit et al., 2013; Thorpe et al., 2017) and  
344 the advanced statistical methods used, they are not without their limitations. The first of  
345 these is the use of several different sports within the study. Although this increases the  
346 ecological validity of the study, it also increases the chance that meaningful effects in one  
347 sport (e.g. football) may be lost by the trivial effect of another (e.g. cricket). Unfortunately,  
348 participant numbers prevented us from breaking the analysis down into sports to confirm  
349 this theory. This is also shown statistically by the *small to large* between participant  
350 variation in the effect of the predictors on DWB, its individual subscales and PRS. Such  
351 variation is indicative of an inconsistent response to predictors (possibly between sports as  
352 well as individuals) and ensures that it is difficult to use the mean effect in practice as  
353 some athletes will respond considerably better or worse to variations in each predictor. To  
354 that end, a move towards considering individualised responses may be more appropriate  
355 when datasets allow (Bartlett, O'Connor, Pitchford, Torres-Ronda, & Robertson, 2017;  
356 Thornton, Delaney, Duthie, & Dascombe, 2017). Furthermore, the use of self-report  
357 measures can be criticised. Although the use of daily wellbeing questionnaires is time and  
358 cost efficient in both collection and analysis, they are open to cognitive (e.g. lack of  
359 understanding) and conscious (e.g. responding with the answer the athlete believes is  
360 correct rather than how they feel) bias (Saw et al., 2015). The use of the 24 hour s-RPE  
361 method for total daily training load can also be criticised. In this study, the time and cost  
362 effectiveness of the s-RPE method was important given the resources available, however it  
363 is not the gold standard of training load measurement. Although the use of s-RPE provides  
364 an understanding of how hard an athlete believes they have worked over a day, it does not  
365 consider objective markers such as GPS, accelerometer or total resistance volume  
366 measures which may provide a more accurate depiction of the total workload produced and

367 have been linked to injury incidence with much more accuracy (Hulin, Gabbett, Lawson,  
368 Caputi, & Sampson, 2016; Williams, West, Cross, & Stokes, 2016). The use of a daily s-  
369 RPE total also cannot be extrapolated to dose-response changes in fitness unlike other  
370 internal load measures, such as heart rate monitoring (Taylor et al., 2018). Self-reported  
371 sleep duration has also been criticised in the past as previous studies have shown it can be  
372 overestimated by as much as 1-1.5 hours (Caia et al., 2017; Kölling et al., 2016;  
373 Lauderdale et al., 2008), suggesting actigraphy may be a more appropriate measure.  
374 However, to date there is no research specifically proving that objective measures more  
375 accurately influence perceptions of wellbeing than subjective measures. It is therefore  
376 possible that perceptions of sleep are more important than actual sleep characteristics when  
377 considering the perceptive wellbeing response.

378

### 379 **Conclusions**

380 In conclusion, our results show that it is important to consider the recovery of an athlete as  
381 well as the training stress they encounter when considering changes in wellbeing measures.  
382 In our study, DWB was shown to be responsive to sleep duration, but not training load.  
383 However, the individual subscale of muscle soreness was related to training load  
384 suggesting that a masking effect may have occurred with the overall score. This does not  
385 mean that the subscales not showing a relationship with training load are not valuable  
386 because they were, with the exception of the mood and stress subscales, related to the  
387 recovery the athlete encountered (measured by sleep duration) and may still be important,  
388 either alone or as part of the overall DWB score, for the detection of future adverse events  
389 such as injury, illness or overtraining. The PRS on the other hand was related to both the  
390 training stressors imposed (training load and additive match play exposure) and the  
391 recovery encountered (sleep duration), suggesting that as a single measure to monitor the



392 athletes response to a training programme it may be superior to DWB and its individual  
393 subscales. However, like DWB, its relationship with "true" outcome events such as injury,  
394 illness and overtraining is yet to be elucidated.

395

396 **Disclosure of interest**

397 The authors report no conflict of interest.

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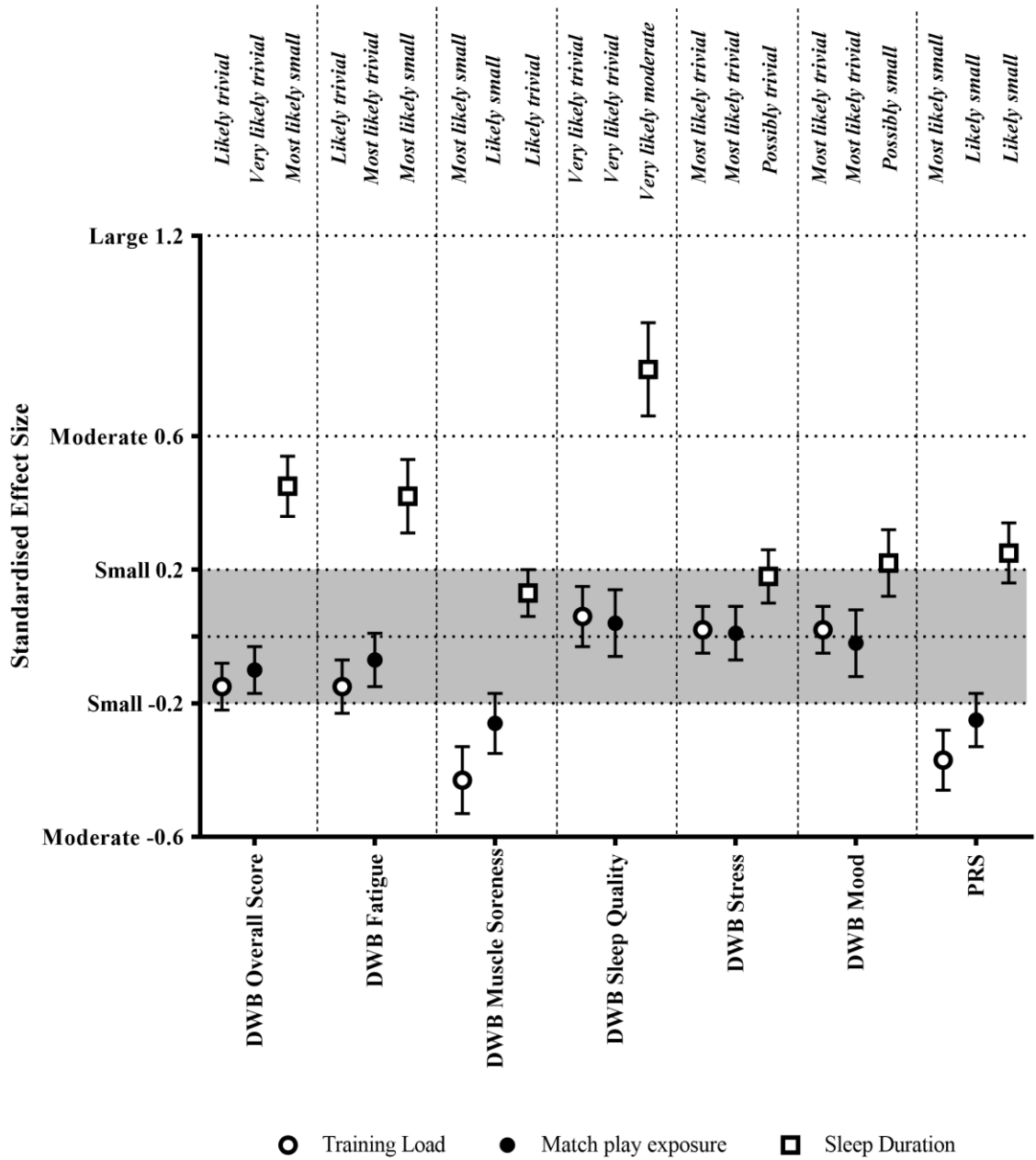
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560 **Figure 1:** The influence of training load, exposure to match play and sleep duration on the  
 561 overall DWB score, its individual subscales and PRS. Data are presented as effect size with  
 562 90% confidence intervals, shaded area denotes smallest worthwhile change.



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