Global fairness of additive—increase and multiplicative—decrease with heterogeneous round—trip times

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Abstract—Consider a network with an arbitrary topology and arbitrary communication delays, in which congestion control is based on additive-increase and multiplicativedecrease. We show that the source rates tend to be distributed in order to maximize an objective function called F_A^h (" F_A^h fairness"). We derive this result under the assumption of rate proportional negative feedback and for the regime of rare negative feedback. This applies to TCP in moderately loaded networks, and to those TCP implementations that are designed to interpret multiple packet losses within one RTT as a single congestion indication and do not rely on re-transmission timeout. This result provides some insight into the distribution of rates, and hence of packet loss ratios, which can be expected in a given network with a number of competing TCP or TCP-friendly sources. We validate our findings by analyzing the parking lot scenario, and comparing with previous results [1], [2], and an extensive numerical simulation with realistic parameter settings. We apply F_A^h fairness to gain a more accurate understanding of the bias of TCP against long round trip times.

Keywords— Additive-Increase, Multiplicative-Decrease, Fairness, Best-Effort, TCP, TCP-Friendly, TCP throughput-loss formula, RTT, parking-lot, Stochastic Approximation, ODE, Lyapunov.

I. Introduction

There is a continuing interest on throughput and fairness issues of TCP [3] congestion avoidance. This interest is particularly nourished by the proliferation of real-time "stream" applications over the Internet (e.g. voice, video) for which it is required to be TCP-Friendly, i.e. to fairly coexist with already existing TCP applications.

In one of the pioneering works, Chiu and Jain [4] formulated a set of basic principles of the additive–increase and multiplicative–decrease congestion avoidance to achieve efficiency and fairness, by analyzing the simple model of a single bottleneck.

In [5] Kelly, Maulloo, and Tan showed that a large-scale network deploying some specific form of additive—

increase and multiplicative—decrease congestion avoidance tends to distribute rates according to proportional fairness. This result is commonly misinterpreted as being applicable to congestion avoidance in the Internet with TCP.

Recently, Hurley, Le Boudec, and Thiran [6] showed that in a network employing additive–increase and multiplicative–decrease, the source rates tend to be distributed in order to maximize an objective function called F_A . The authors call this " F_A fairness". This result is obtained by the limit mean ordinary differential equation (ODE) method, for a network operating in the regime of rare negative feedback. The pivotal assumption of that work is the rate proportional negative feedback, which the authors claim to be more realistic than one which depends exclusively on the overall load [5]. However, the result is restricted to the homogeneous round–trip time (RTT) case where the rates are updated synchronously.

In this paper, we extend the modeling of [6] to the heterogeneous RTT case. Our result is a generalization of F_A fairness, which we call F_A^h fairness. It gives the distribution of rates in a arbitrary network employing the additive increase and multiplicative-decrease method for congestion control, with the assumption that negative feedback is rare. We allow the round-trip times to differ from one source to another. The rates tend to maximize an objective function called F_A^h , whose parameters reflect the rate adaptation algorithm. To the best of our knowledge, this is the first general result encompassing many of the relevant system parameters, applicable to an arbitrary network topology with multiple bottlenecks. Our results allows to find a first order approximation of rate distributions; combined with a loss-throughput formula such as [2], [7], this gives a prediction of the loss rates. Extensive simulation results confirm these predictions.

The novelty of our approach is an application of the recent weak convergence results of decentralized asynchronous stochastic approximation algorithms [8]. Our

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model essentially differs from [6] in that we do not assume that rate-adaptation is performed synchronously by all sources; in contrast, we use an asynchronous model where every source updates its rate based on its own round trip time interval. Unlike the synchronous model in [4], [5] or [6], this allows us to address the case with different round trip times. But even in the case where all round trip times are equal, this gives a more accurate model. Indeed, with the synchronous model, rate adjustment is based on the most recent previous rates. In reality, the feedback received by one source at the end of one round trip time interval depends on the rates during the previous interval, shifted in time by the delay required for feedback to reach the sources. The synchronous model assumes implicitly that feedback reaches sources instantaneously. We call this assumption "stolen lag". We show with our modeling that the stolen lag assumption does not affect the distribution of average rates; by simulation, we see however that it affects the amplitude of oscillations. Note that our model explicitly considers all communication delays.

We assume in this paper that the negative feedback received by sources is rare, and is proportional to the source rate. The rare negative feedback assumption is valid in a reasonably loaded network; the proportional assumption should be true with active queue management [9] (e.g. RED [10]) applied to otherwise FIFO queues. In addition, our model assumes a single rate updating per RTT; this fits with TCP implementations designed to cope with multiple packet losses within single RTT, i.e. that treat multiple packet losses within one RTT as a single congestion signal, and avoid re-transmission timeouts.

Our model does not incorporate the effect of the variation of RTT for one given source from one feedback interval to the other. It is known that, for a network with fixed windows [11], the variation of round trip times due to queues building up has in itself a congestion avoidance effect, which is not captured by our modeling. Another limitation is that we assume the rates to be piecewise constant, i.e. to be adjusted only once per round trip time. Thus, the effect of burstiness at the timescale of the round trip time is not taken into account. In contrast, our study captures the effect of the window or rate adaptation mechanism found for example with TCP or ABR. Our results may be used as a reference fairness measure in performance evaluations of TCP–friendly rate adjustment algorithms.

In the next subsection we outline our main results.

A. Summary of the Main Results

We consider a network with multiple bottlenecks and heterogeneous round-trip times. Then, under the condition that there is no substantial queuing delay variation, and the network is operating in the regime of the rare negative feedback, the collection of rates $x \neq x_1, ..., x_{-i}, ...$) is distributed such that x maximizes the objective function

$$F_A^h(x) = \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log \frac{x_i}{r_i + \eta_i x_i},$$

subject to the constraints $\sum_{j \in \mathcal{S}} A_{l,j} x_j \leq c_l$, $\forall l \in \mathcal{L}$. In the formula, \mathcal{S} is the set of sources, \mathcal{L} the set of links, $A_{l,i}$ the routing matrix $(A_{l,i} \text{ is } 0 \text{ or } 1)$, c_l the capacity of link l, and τ_i is the RTT for flow or source i. There is one flow per source. The rate adaptation parameters are r_i (additive—increase element) and η_i (multiplicative—decrease factor); they may depend on source i.

The above result is applied to the parking—lot network topology; we obtain a closed-form for the distribution of rates. This allows us to verify the consistency of our result with existing work and with conducted simulations. We find that the results in [1] are an asymptotic case of F_A^h fairness for small additive—increase/multiplicative—decrease ratio relative to connection throughput.

We also gain a more accurate understanding of the bias of TCP against long round trip times. We point out that it is important to make the difference between a bias against long RTTs (perhaps an undesirable feature) and a bias against flows with many hops (perhaps a desired feature). We see that the bias against flows with many hops is in the nature of any rate adaptation algorithm based on additive—increase and multiplicative—decrease. In contrast, a bias against long RTTs can be attenuated with corrections such as mentioned in [1] and [12]. Finally, we also confirm throughput loss formulas, within the limitations of our modeling.

B. Outline of the Paper

The paper is organized as follows. In Section II, the main results are derived. Following the basic model definitions, feedback modeling is described in more detail. Then, asymptotic convergence results of the decentralized asynchronous stochastic approximation algorithms [8] are sketched. In the rest of the section, objective function F_A^h of the algorithm of concern is derived and analyzed. In Section III, F_A^h result is applied to the parking–lot network topology for which a closed-form rate distribution is computed, and results are verified through numerical simulation. In Section IV, the results are discussed and compared to the related previous work. Implications of the result to the Internet are addressed in Section V. In Section VI, concluding remarks are given. In Appendix A we give the main theorem of the underlying theory [8].

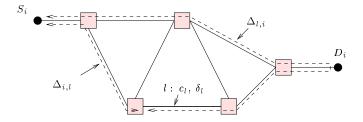


Fig. 1. An illustration of the defined delays.

II. DERIVATION OF THE MAIN RESULTS

A. Model Setup

The notation is developed as follows. Let set $\mathcal L$ contain network links. Then let c_l and δ_l be the capacity and delay of link $l \in \mathcal L$, respectively. Let set $\mathcal S$ comprise sources (flows) that are active on the given network. The routing setting we describe by routing matrix $A \notin A_{l,i}l \in \mathcal L$, $i \in \mathcal S$), such that $A_{l,i}=1$, if flow i traverses link l, and $A_{l,i}=0$, otherwise. l

Further, we define communication delays. Let $\Delta_{i,l}$ denote delay from source i to link l, and let $\Delta_{l,i}$ be the delay from link l to source i. Then, set $\Delta_{i,j} = \Delta_{j,l} + \Delta_{l,i}$. Defining τ_i as the RTT of source i, clearly, $\tau_i = \Delta_{i,l} + \Delta_{l,i}$, for all l such that $A_{l,i} > 0$. In Fig. 1, a sample network illustrates defined delays.

Let $\{\tau_{i,n}\}_{n\geq 0}$ be a non-decreasing $[0,\infty)$ -valued sequence of rate updating times of source i. Then, a number of rate updates of source i on the interval [0,t) is $N_i(t) = \sum_{n=1}^{\infty} 1_{\{\tau_{i,n} < t\}}$.

Let $\{x_{i,n}\}_{n\geq 0}$ be a $[0,\infty)$ -valued stochastic process, where $x_{i,n}$ is a rate of source i at the n-th update. Then, define a continuous time interpolation on real time as $x_i(t) = x_{i,n}$, for $t \in [\tau_{i,n}, \tau_{i,n+1})$.

For $b \geq a \geq 0$ define a σ -algebra of the form $\mathcal{F}^{i,l}_{[a,b)} = \sigma(x_{j,k}:A_{l,j},A_{l,i}>0,N_j(a-\Delta_{i,j})\leq k < N_j(b-\Delta_{i,j})).$ Finally, let $\mathcal{F}^i_{[a,b)} = \cup_{l:\ A_{l,i}>0} \mathcal{F}^{i,l}_{[a,b)}.$

An additive-increase and multiplicative-decrease algorithm has the following form

$$x_{i,n+1} = x_{i,n} + r_i(1 - I_{i,n}) - \eta_i I_{i,n} x_{i,n},$$
 (1)

where r_i and η_i are the additive–increase element and the multiplicative–decrease factor, respectively. The random sequence $\{I_{i,n}\}_{n\geq 0}$ is a negative feedback indication with values in $\{0,1\}$. We assume that the negative feedback indication $I_{i,n}$ is based on the feedback received between n-th and n+1-th rate updating. Consequently, it turns out that $I_{i,n}$ is measurable on σ -algebra $\mathcal{F}^i_{[\tau_{i,n},\tau_{i,n+1})}$.

¹We define $A_{l,i}$ on $\{0,1\}$ which can be extended to [0,1] to accommodate for instance load sharing, etc. [6]

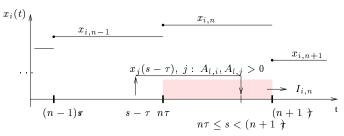


Fig. 2. Feedback modeling for the HOMRTT case.

Now let us briefly comment on the special case addressed in [6], where it is assumed that all round—trip times are equal, and the rates are updated synchronously. Following the definition of $I_{i,n}$ in the full extent, it is rather easy to see that $I_{i,n}$ is a function of $x_{j,n}, x_{j,n-1}$, and $x_{j,n-2}$, for all $j \in \mathcal{S}$ such that $A_{l,i}, A_{l,j} > 0$, depending on values of $\Delta_{j,l}$. In the related work [4]–[6], it is commonly assumed that $I_{i,n}$ is computed based on $x_{j,n}$, for all $j \in \mathcal{S}$ such that $A_{l,i}, A_{l,j} > 0$, which is indeed an unrealistic assumption.

Let us assume that all flows traversing link l have equal access delay to that link. Formally, $\Delta_{i,l} \not\equiv \Delta_{j,l}$, for all $i,j \in \mathcal{S}$, such that $A_{l,i}, A_{l,j} > 0$. Then, it follows that $I_{i,n}$ depends on $x_{j,n-1}$, for all $j \in \mathcal{S}$ such that $A_{l,i}, A_{l,j} > 0$. We refer to this assumption as a HOMRTT assumption. In addition, whenever $x_{i,n}$ is used where it should be $x_{i,n-1}$ we refer to this as a *stolen lag*. Feedback modeling is illustrated in Fig. 2.

B. Feedback Modeling

First, we introduce a notion of the link cost function $g_l(\cdot)$ $[0,\infty) \to [0,1]$, where $l \in \mathcal{L}$. At a given time t, the link cost is a function of the link load

$$f_l(t) = \sum_{i \in \mathcal{S}} A_{l,i} x_i (t - \Delta_{i,l}). \tag{2}$$

One can interpret $g_l(f_l(t))$ as a probability of marking a single packet at time t. We are concerned with negative feedback indication, $I_{i,n}$, based on feedback received within $[\tau_{i,n},\tau_{i,n+1})$. Let us partition the interval $[\tau_{i,n},\tau_{i,n+1})$ into non-overlapping intervals $[a_k,b_k)$ such that $x_j(s)$, is constant for $s\in [a_k-\Delta_{i,j},b_k-\Delta_{i,j})$, for all $j\in\mathcal{S}$ such that $A_{l,i},A_{l,j}>0$.

We define $M_{a,b}^{i,l}$ as an amount of negative feedback received by source i from link l within interval [ab], which is equal to a number of marked packets of the flow i by the link l within interval $[a-\tau_i,b-\tau_i)$.

Let $P_{a,b}^{i,l}(\cdot)$ and $P_{a,b}^{i}(\cdot)$ be conditional probabilities given $\mathcal{F}_{[a,b)}^{i,l}$ and $\mathcal{F}_{[a,b)}^{i}$, respectively. Analogously, let $E_{a,b}^{i,l}[\cdot]$ and $E_{a,b}^{i}[\cdot]$ be respective conditional expectations.

Admitting the interpretation of $g_l(\cdot)$ as a probability of marking a single packet, it is easy to see that we do have a binomial conditional probability

$$P_{a_k,b_k}^{i,l}(M_{a_k,b_k}^{i,l} = m) = \begin{pmatrix} \lceil x_{i,N_i(a_k - \tau_i)}(b_k - a_k) \rceil \\ m \end{pmatrix} \cdot g_l(f_l(a_k^{\Delta_i}))^m [1 - g_l(f_l(a_k^{\Delta_i}))]^{\lceil x_{i,N_i(a_k - \tau_i)}(b_k - a_k) \rceil - m},$$
(3)

where $a_k^{\Delta_i}$ is written in lieu of $a_k - \Delta_{l,i}$. Note that $\lceil x_{i,N_i(a_k-\tau_i)}(b_k-a_k) \rceil$ corresponds to a number of packets of flow i that are present on link l within $\lceil a_k - \tau_i, b_k - \tau_i \rceil$. Clearly, the expected amount of negative feedback is

$$E_{a_k,b_k}^{i,l}[M_{a_k,b_k}^{i,l}] = \lceil x_{i,N_i(a_k-\tau_i)}(b_k - a_k) \rceil g_l(f_l(a_k^{\Delta_i})).$$

Similarly, let $M_{a,b}^i$ be an amount of negative feedback received within [ab] by source i from all links l such that $A_{l,i} > 0$. It follows

$$P_{a_k,b_k}^i(M_{a_k,b_k}^i=0\)=\prod_{l:\;A_{l\;i}>0}P_{a_k,b_k}^{i,l}(M_{a_k,b_k}^{i,l}=0\ ,)$$

and from (3) follows

$$\begin{split} & P_{a_k,b_k}^i(M_{a_k,b_k}^i = 0) = \\ & = \prod_{l: \ A_{l,i} > 0} [1 - g_l(f_l(a_k^{\Delta_i}))]^{\lceil x_{i,N_i(a_k - \tau_i)}(b_k - a_k) \rceil}. \end{split} \tag{4}$$

By definition of $[a_k, b_k]$ we have

$$P_{\tau_{i,n},\tau_{i,n+1}}^{i}(M_{\tau_{i,n},\tau_{i,n+1}}^{i}=0) = \prod_{k} P_{a_{k},b_{k}}^{i}(M_{a_{k},b_{k}}^{i}=0)$$

Finally, $I_{i,n} \dashv$, if source i has received an indication, within $[\tau_{i,n},\tau_{i,n+1})$, that at least one packet has been marked, thus

$$P_{\tau_{i,n},\tau_{i,n+1}}^{i}(I_{i,n}=1) = P_{\tau_{i,n},\tau_{i,n+1}}^{i}(M_{\tau_{i,n},\tau_{i,n+1}}^{i} \ge 1) = 1 - P_{\tau_{i,n},\tau_{i,n+1}}^{i}(M_{\tau_{i,n},\tau_{i,n+1}}^{i} = 0).$$
(6)

Let us examine (6) for the HOMRTT case. Here we have a single partition of $[\tau_{i,n}, \tau_{i,n+1})$, hence, from (4)–(6) follows

$$P^{i}_{\tau_{i,n},\tau_{i,n+1}}(I_{i,n}=1) = 1 - \prod_{l: A_{l,i}>0} [1-g_l(f_l)]^{\lceil x_{i,n-1}\tau \rceil},$$

where $\tau_{i,n+1} - \tau_{i,n} = \tau$, for all $i \in \mathcal{S}$, $n \geq 0$, and f_l stands for $f_l(\tau_{i,n} - \Delta_{l,i}) = \sum_{j \in \mathcal{S}} A_{l,j} x_{j,n-1}$. In the limit case $g_l(\cdot) \to 0$, limited development yields $[1 - g_l(f_l)]^{\lceil x_{i,n-1}\tau \rceil} \simeq 1 - \lceil x_{i,n-1}\tau \rceil g_l(f_l)$, then replacing this in (7), and neglecting the higher order products, yield

$$P_{\tau_{i,n},\tau_{i,n+1}}^{i}(I_{i,n}=1) \succeq \sum_{l \in \mathcal{L}} A_{l,i} g_{l}(f_{l}) \lceil x_{i,n-1}\tau \rceil.$$
 (8)

Therefore, it is shown that, under the rare negative feedback assumption, (7) degenerates to the rate proportional feedback as is implicitly assumed in [6]. However, note that (8) depends on $x_{i,n-1}$ and not on $x_{i,n}$.

C. Asymptotic Convergence

Traditional theory of the stochastic approximation algorithms [13]–[14] is concerned with an algorithm of the general form

$$x_{i,n+1} = x_{i,n} + \gamma_n H_{i,n}(x_{i,n}, \xi_{i,n}) i \in \mathcal{S},$$

where $x_{i,n}$ is defined on R, $H_{i,n}(\cdot)$ $R \times R \to R$, $\xi_{i,n}: R \to R$ is a random noise, and γ_n a step size. It is assumed that components $x_{i,\cdot}$, $i \in \mathcal{S}$, are updated synchronously, facilitating the association of continuous interpolation $x_i(t)$ to discrete process $\{x_{i,n}\}_{n\geq 0}$ on the "natural" common iterate time $\{n\gamma\}_{n\geq 0}$. However, it follows that for an asynchronous updating one has to work in real time, or at least an appropriately scaled real time [8]. In general, for decreasing γ_n , in respect to n, convergence with probability one can be obtained, while for constant small $\gamma_n \equiv \gamma$, only convergence in probability can be proven (the weak convergence). In the rest of this section we briefly sketch results of [8] that are applied in our work. For a complete treatment of the underlying theory the reader is referred to [8].

Let $\{\delta au_{i,k}^{\gamma}\}_{k\geq 0}$ be a random sequence of updating intervals of $\{x_{i,k}\}_{k\geq 0}$, i $\mathfrak S$. Then denote (scaled) real updating time of $x_{i,n}$ as

$$\tau_{i,n}^{\gamma} = \gamma \sum_{k=0}^{n-1} \delta \tau_{i,k}^{\gamma}, \tag{9}$$

and a continuous interpolation on the iterate time $\tau_i^{\gamma}(t) = \tau_{i,n}^{\gamma}$, for $t \in [n\gamma, (n+1)\gamma)$. Further, let $N_i^{\gamma}(t)$ is a number of updates of $\{x_{i,k}\}_{k>0}$ before t/γ . Formally,

$$N_i^{\gamma}(t) = \gamma \sum_{n=1}^{\infty} 1_{\{\tau_{i,n}^{\gamma} < t/\gamma\}}.$$
 (10)

From the definitions, it turns out that $N_i^{\gamma}(\tau_i^{\gamma}(t)) = n\gamma$, $t \in [n\gamma, (n+1 \ \gamma), \text{ i.e. } N_i^{\gamma}(\cdot) \text{ is inverse of } \tau_i^{\gamma}(\cdot).$

Then, let $\hat{x}_i^{\gamma}(t) = x_{i,n}^{\gamma}$, $t \in [\tau_{i,n}^{\gamma}, \tau_{i,n+1}^{\gamma})$, is a continuous interpolation on scaled real time, and $\hat{x}^{\gamma}(\cdot) = (x_i^{\gamma}(\cdot), i = 1, 2, \dots, S)$, $S = |\mathcal{S}|$. From definitions of "time" processes (9) and (10) it follows $x_i^{\gamma}(t) = x_i^{\gamma}(\tau_i^{\gamma}(t))$, and $\hat{x}_i^{\gamma} = x_i^{\gamma}(N_i^{\gamma}(t))$.

Furthermore, let $\Delta_{i,j,n}^{\gamma}$ be a non-negative random variable representing a scaled (multiplied by γ) communication delay between source i and source j at the n-th rate updating of source i.

Finally, the decentralized asynchronous algorithm can be written in the form

$$x_{i,n+1}^{\gamma} = \Pi_{[a_i,b_i]} \left(x_{i,n}^{\gamma} + \gamma H_{i,n}^{\gamma} (\hat{x}_j (\tau_{i,n+1}^{\gamma} - \Delta_{i,j,n}^{\gamma})) \right) =$$

$$= x_{i,n}^{\gamma} + \gamma H_{i,n}^{\gamma} (\hat{x}_j (\tau_{i,n+1}^{\gamma} - \Delta_{i,j,n}^{\gamma})) + \gamma Z_{i,n}^{\gamma} i \in \mathcal{S},$$

$$\tag{11}$$

where $\Pi_{[a_i,b_i]}(\cdot)$ denotes projection of the argument on $[a_i,b_i]$, for the constrained x on $C=[a_1,b_1]\times [a_2,b_2]\times \ldots \times [a_S,b_S]$, and $Z_{i,n}^{\gamma}$ is a reflection term.

Let for all i, $\mathcal{F}_{i,n}^{\gamma}$ and $\mathcal{F}_{i,n}^{\gamma,+}$ be non-decreasing σ -algebras measuring the past data (including $x_{i,0}^{\gamma}, H_{j,k}^{\gamma}$, and $\delta au_{j,k}^{\gamma}, j \in \mathcal{S}$) available on $[0, au_{i,n+1}^{\gamma}]$, and $[0, au_{i,n+1}^{\gamma}]$, respectively. Then, with $P_{i,n}^{\gamma}$ and $P_{i,n}^{\gamma,+}$ denote respective conditional probabilities, and analogously $E_{i,n}^{\gamma}$ and $E_{i,n}^{\gamma,+}$ conditional expectations.

Subsequently, we have the following conditions. It is assumed that

$$\{H_{i,n}^{\gamma}, \delta \ \tau_{i,n}^{\gamma}; \gamma_{i,n} \ \},$$
 (12)

and $\Delta_{i,j,n}^{\gamma}$, $\Delta_{i,j,n}^{\gamma,+}$ are uniformly integrable. We consider the Martingale difference noise [8], for which we have

$$E_{i,n}^{\gamma}H_{i,n}^{\gamma} = h_{i,n}^{\gamma}(\hat{x}_{i}^{\gamma}(\tau_{i,n+1}^{\gamma} - \Delta_{i,j,n}^{\gamma}), j \in \mathcal{S}) + \beta_{i,n}^{\gamma},$$

where $h_{i,n}^{\gamma}(\cdot)$ are real-valued functions continuous in n and γ , $\beta_{i,n}^{\gamma}$ is asymptotically negligible noise, and $\sup_{n\leq T/\gamma}\Delta_{i,j,n}^{\gamma}\to 0$. There are real-valued functions $u_{i,n}^{\gamma}(\cdot)$ that are strictly positive $(\inf_{n,\gamma,x,\xi}u_{i,n}^{\gamma}(x\xi)>0)$ and are continuous uniformly in n and γ , and non-negative random variables $\Delta_{i,i,n}^{\gamma,+}$ such that

$$E_{i,n}^{\gamma,+} \delta \tau_{i,n+1}^{\gamma} = u_{i,n+1}^{\gamma} (\hat{x}_j^{\gamma} (\tau_{i,n+1}^{\gamma} - \Delta_{i,j,n+1}^{\gamma,+}), j \in \mathcal{S}), \tag{13}$$

where $\sup_{n\leq T/\gamma} \Delta_{i,j,n}^{\gamma,+} \to 0$, inprobability a $s\gamma \to 0$. There are continuous real-valued functions $\bar{h}_i(\cdot)$ such that for each $x\in C$, ²

$$\lim_{m,n,\gamma} \frac{1}{m} \sum_{k=n}^{n+m-1} \gamma [h_{i,k}^{\gamma}(x) - \bar{h}_i(x)] = 0 \quad . \tag{14}$$

There are continuous real-valued functions $\bar{u}_i(\cdot)$ such that for all $x \in C$,

$$\lim_{m,n,\gamma} \frac{1}{m} \sum_{k=n}^{n+m-1} E_{i,n}^{\gamma,+} [u_{i,k}^{\gamma}(x) - \bar{u}_i(x)] = 0 \quad . \tag{15}$$

Suppose

$$\lim_{m,n,\gamma} \sum_{k=n}^{n+m-1} E_{i,n}^{\gamma} \beta_{i,k}^{\gamma} = 0, \text{ inmean } .$$
 (16)

²In (14), (15), and (16), $\lim_{m,n,\gamma} \equiv \lim_{m\to\infty,n\to\infty,\gamma\to0}$, simultaneously in any way.

Finally, from the Theorem [8] (Appendix A) particularly follows that, for the unconstrained algorithm, the weak convergence subsequence \hat{x} is the limit set of ODE

$$\dot{\hat{x}}_i = \frac{\bar{h}_i(\hat{x})}{\bar{u}_i(\hat{x})} i \in \mathcal{S}. \tag{17}$$

Thus, the limit mean ODE is the same as in the synchronous case, except for an additional weight factor that takes into account frequency of the updating.

D. F_A^h Fairness

We identify $H_{i,n}^{\gamma}$ of (11) in the algorithm (1) as

$$H_{i,n}^{\gamma} = r_i - (r_i + \eta_i x_{i,n}) 1_{\{\xi_{i,n} < P_{i,n+1}^i(I_{i,n}^{\gamma} = 1)\}}, \quad (18)$$

where $\{\xi_{i,n}\}_{n\geq 0}$ is a sequence of independent random variables uniformly distributed on [0,1].³ Then,

$$E_{i,n+1}^{\gamma}H_{i,n}^{\gamma} = r_i - (r_i + \eta_i x_{i,n})P_{i,n+1}^{\gamma}(I_{i,n}^{\gamma} = 1),$$

where $P_{i,n+1}^{\gamma}(I_{i,n}^{\gamma}=1)$ is given by (6).

In the limit case, as $\gamma \to 0$, and $n \to \infty$, we neglect scaled delays $\Delta_{i,j}$, then the probability of negative feedback is

$$P_{i,n+1}^{\gamma}(I_{i,n}^{\gamma}=1) = \hat{x_i}(\tau_{i,n+1}^{\gamma,-})\delta\tau_{i,n}\sum_{l\in\mathcal{L}}A_{l,i}g_l(\hat{f_l}(\tau_{i,n+1}^{\gamma,-})).$$

For $\delta \tau_{i,n} = \tau_i$, for all n > 0, the mean vector field is

$$\bar{h}_i(\hat{x}(t)) = r_i - \hat{x}_i(t)(r_i + \eta_i \hat{x}_i(t))\tau_i \sum_{l \in \mathcal{L}} A_{l,i} g_l(\hat{f}_l(t))$$
(19)

where $\hat{f}_l(\cdot) = \sum_{j \in \mathcal{S}} A_{l,j} \hat{j}(\cdot)$.

Seemingly, $\bar{u}_i(\hat{x}) = \tau_i$, then with (19) the limit mean ODE (17) becomes

$$\dot{\hat{x}}_i = \frac{r_i}{\tau_i} - \hat{x}_i(r_i + \eta_i \hat{x}_i) \sum_{l \in \mathcal{L}} A_{l,i} g_l(\hat{f}_l).$$
 (20)

Following the same steps as in [6] we express ODE (20) as

$$\dot{\hat{x}}_i = \hat{x}_i (r_i + \eta_i \hat{x}_i) \frac{\partial J_A^h(\hat{x})}{\partial \hat{x}_i}, \tag{21}$$

where

$$J_A^h(\hat{x}) = \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log \frac{\hat{x}}{r_i + \eta_i \hat{x}} - G(\hat{x}), \qquad (22)$$

and by definition $G(\hat{x}) = \sum_{l \in \mathcal{L}} G_l(\hat{f}_l)$, where $G_l(\cdot)$ is a primitive of $g_l(\cdot)$.

³Note that instead of r_i and η_i it should be written r_i^{γ} and η_i^{γ} , where $r_i = \gamma r_i^{\gamma}$ and $\eta_i = \gamma \eta_i^{\gamma}$, but we abuse this for notation simplicity.

It is easy to see that J_A^h is strictly concave and consequently has a unique maximum over any bounded region. It turns out that J_A^h is Lyapunov for ODE and with an unique attractor, for which J_A^h is maximized.

Along the same lines as in [6], one can neglect the second term in (22). Then, it follows that the rates x are distributed such that x maximizes

$$F_A^h(x) = \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log \frac{x_i}{r_i + \eta_i x_i}$$
 (23)

subject to the constraints

$$\sum_{j \in \mathcal{S}} A_{l,j} x_j \le c_l, \ \forall l \in \mathcal{L}.$$
 (24)

D.1 F_A^h asymptotic limits

Let us write (23) in the following form

$$F_A^h(x) = \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log \frac{1}{\eta_i} - \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log \left(1 + \frac{r_i}{x_i \eta_i} \right). \tag{25}$$

Then we develop the second term to obtain

$$F_A^h(x) = \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log \frac{1}{\eta_i} - \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \sum_{n=1}^{\infty} \frac{(-1)^{n-1}}{n} \left(\frac{r_i}{\eta_i x_i}\right)^n,$$

for $x_i > r_i/\eta_i$. By the limited development for $x_i >> r_i/\eta_i$ we obtain that rates x are distributed such that x maximizes

$$F_A^{h+}(x) = -\sum_{i \in \mathcal{S}} \frac{r_i}{\tau_i \eta_i x_i},\tag{26}$$

$$F_A^{h-}(x) = \sum_{i \in \mathcal{S}} \frac{1}{\tau_i} \log x_i. \tag{27}$$

We use (26) and (27) in Section V to understand the bias of TCP against long RTTs.

III. An Example of ${\cal F}_A^h$ Fairness – parking–lot

In this section we consider a parking—lot topology depicted in Fig. 3. In this scenario we distinguish class 0 flows, traversing tandem of I links, and class i flows traversing a single link i, i = 1, 2, ..., I. Let n_i be the number of flows belonging to class i.

⁴For TCP congestion avoidance $r_i = 1/\tau_i$ and $r_i = 1/(2\tau_i)$ for not delayed ACKs and delayed ACKs, respectively. Henceforth, (26) can be safely applied to the respective cases as long as the number of packets per RTT, $x_i\tau_i >> 2$, and $x_i\tau_i >> 1$, for all $i \in \mathcal{S}$.

Feasibility constraints $n_0x_0 + n_ix_i = c_i$, i = 1, 2, ..., I, allow us to express (23) in terms of x_0 as follows

$$F_A^h(x_0) = \frac{n_0}{\tau_0} \log \frac{x_0}{r_0 + \eta_0 x_0} + \sum_{i=1}^I \frac{n_i}{\tau_i} \log \frac{c_i - n_0 x_0}{r_i n_i + \eta_i (c_i - n_0 x_0)}.$$
(28)

In the sequel, we suppose

$$c_i = c, r_i = r_I, \eta_i = \eta_I, n_i = w, i = 1, 2, \dots, I, n_0 = v.$$
(29)

Then, (28) becomes

$$F_A^h(x_0) = \frac{v}{\tau_0} \log \frac{x_0}{r_0 + \eta_0 x_0} + T_I w \log \frac{c - v x_0}{r_I n_I + \eta_I (c - v x_0)}, \quad (30)$$

where $T_I \stackrel{\Delta}{=} \sum_{i=1}^{I} \frac{1}{\tau_i}$.

Lemma 1 (parking–lot) F_A^h -fairness distribution for the parking–lot scenario with (29) is

$$x_0 = \frac{-B - \sqrt{B^2 - 4AC}}{2A},\tag{31}$$

where

$$\begin{split} A &= v^2 r_0 \eta_I - w^2 T_I r_I \eta_0 \tau_0 \\ B &= -r_0 (v(w r_I + 2c \eta_I) + w^2 T_I r_I \tau_0) \\ C &= r_0 c(w r_I + c \eta_I), \end{split}$$

for $v^2 r_0 \eta_I - w^2 r_I \eta_0 T_I \tau_0 \neq 0$, otherwise

$$x_0 = \frac{c(wr_I + c\eta_I)}{v(wr_I + 2c\eta_I) + w^2 T_I r_I \tau_0}.$$
 (32)

Then,

$$x_i = \frac{c - vx_0}{w},$$

for i = 1, 2, ..., I.

Proof: Proof is simple and is same as in [6].

Subsequently, with $x_0 > r_0/\eta_0$ and $x_i > r_I/\eta_I$, for each $i = 1, 2, \ldots, I$, the first–order approximation (26) yields⁵

$$\frac{x_0}{c} = \frac{1}{v + w\sqrt{T_I \frac{r_I \eta_0 \tau_0}{r_0 \eta_I}}},$$
 (33)

or in a simplified case $\tau_i = \tau_I$, for i = 1, 2, ..., I, so that $T_I = \frac{I}{\tau_I}$,

$$\frac{x_0}{c} = \frac{1}{v + w\sqrt{I\frac{r_I\eta_0\tau_0}{r_0\eta_I\tau_I}}}.$$
 (34)

Finally, we consider a few special cases for some reasonable parameters setup as is given in Table I. It can be observed that various fairness objectives are achieved depending on the end-system parameters and RTTs.

⁵In [6] it is referred to this case as $\lim_{c\to\infty} x_0/c$, which is encompassed in $x_i > r_i/\eta_i$, for $r_i < \infty$ and $\eta_i > 0$.

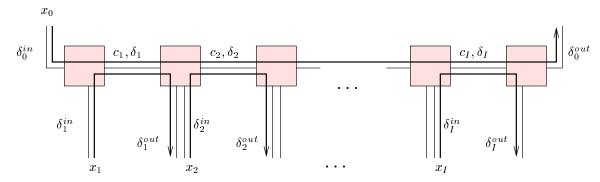


Fig. 3. Parking-lot topology.

TABLE I

Fraction of Capacity c given to a class 0 flows, for the parking–lot, with $x_i> > r_i/\eta_i$ for all $i\in \mathcal{S}$.

Fairness	x_0/c	parameters setup
F_A	$\frac{1}{v+w\sqrt{I}}$	$r_0 = r_I, \tau_0 = \tau_I; r_0 = K\tau_0, r_I = K\tau_I$
Proportional	$\frac{1}{v+wI}$	$r_0 = r_I, \tau_0 = I\tau_I$
Max-min	$\frac{1}{v+w}$	$r_0 = KI au_0, r_I = K au_I$
TCP-Reno	$\frac{1}{v+wI\sqrt{I}}$	$r_0 = 1/\tau_0, r_I = 1/\tau_I, \tau_0 = I\tau_I$

Remarks: Line $1-F_A$ -like result [6] is achieved for homogeneous RTTs, and additive-increase parameters proportional to the RTTs. Line 2 – Proportional fairness result holds for the HET-RTT network and equal values of r_i and η_i for all sources. Line 3 – Max-min fairness can be obtained by indicated parameters setting. Line 4 – Indicates rate distribution for TCP Reno and the HETRTT network (def. Section III-A), which complies to the other results mentioned in Section III-A.

A. Simulation Results

We perform an extensive numerical simulation study of the parking–lot scenario shown in Fig. 3. Let δ_i^{in} and δ_i^{out} be access delays of class i flow, as depicted in Fig. 3. Note that the definition of the parking–lot encompasses both a single bottleneck with different RTTs (I=1), and a multiple bottleneck case, as is considered separately in [1]. To facilitate computation, all delays and rate updating intervals are set as multiples of $\varepsilon>0$. All events corresponding to the rate updating are aligned in respect to the sequence of ε slots. We use the same form of the link cost function as is suggested in [6]

$$g_{l}(f_{l}) = \begin{cases} 0.f & l < 0 \\ \left(\frac{f_{l}/c - d}{1 - d}\right)^{p} d & \leq f_{l} \leq 1 \\ 1f & l > 1 \end{cases},$$

where c is a link capacity, $d \in [0, 1]$, and p > 0.

We consider all combinations of the simulation parameters that are set as follows: $I = \{2,5\}$, $v,w = \{1,2,6,12\}$, $c = \{250,625\}$, $d = \{0,0.5,1\}$, and $p = \{0,0.5,1\}$

 $\{1, 2, 5, 10\}$. Hence, there are a total of 768 settings, and we believe that this covers a wide range of realistic situations. In addition, it is noteworthy that all simulations are obtained for TCP-like parameters, and not for the limit small values. Total simulation time is set equal to 500 the largest RTT. Each average value is obtained over four simulation runs, excluding the initial 20% of the trace to eliminate initial transient. All confidence intervals are computed as 95\% of confidence. For all simulation results, the claim of [6] that the results substantially deviate for case is confirmed, therefore, these results are plotted differently. Hereinafter, we refer to the parking-lot with all access delays equal to zero as a HETRTT; formally $\delta_i^{in} = \delta_i^{out} = 0$, for all $i = 0, 1, \dots, I$. In the sequel, scatter plots of the ${\cal F}_A^h$ analytical throughput against the corresponding simulation result for class 0 flows are shown.

In Fig. 4 and 5 the HOMRTT network is considered with stolen lag and no stolen lag assumption, respectively. The simulations confirm what follows from the model (Section II-D) that rates are distributed according to F_A^h regardless wheather the feedback is based on the last rate updating or on the value that is one lag behind.

In Fig. 6 and 7, for the HETRTT scenario, the F_A and F_A^h analytical results are plotted against respective simulation counterparts. It is evident that the F_A^h matches the simulation results by taking into account diversity of the RTTs.

Fig. 8 demonstrates how the additive–increase proportional to the RTT [1], for the HETRTT network, corrects the bias against the connections with long RTTs, thus the simulation results conform to both F_A and F_A^h fairness.

In Fig. 9 we consider the parking-lot with 2 links, such that all access delays are zero except the egress link of class 1 flows and ingress link of class 2 flows, which are set equal and varied such that the network setting gradually shifts from the HETRTT case (the leftmost point, Fig. 9) to the case where τ_1 and τ_2 are double of τ_0 (the rightmost point, Fig. 9). The results substantiate the validity of F_4^h

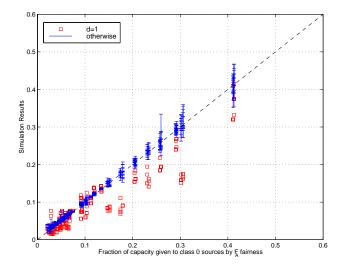


Fig. 4. HOMRTT with the stolen lag, $\tau_i = 0.2$ s, $r_i = 0.5$ for all flows.

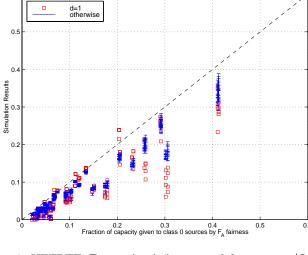


Fig. 6. HETRTT, F_A vs. simulation, $\tau_0 = 0.2$ s, $\tau_i = \tau_0/I$, for i = 1, 2, ..., I, and $r_i = 5$, $\eta_i = 0.5$ for all flows.

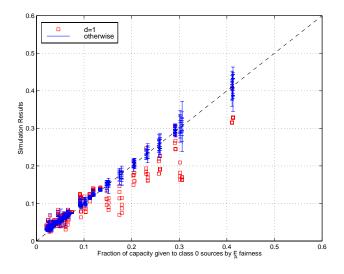


Fig. 5. HOMRTT with no stolen lag, $\tau_i \implies .2$ s, $r_i \implies \eta_i = 0.5$ for all flows.

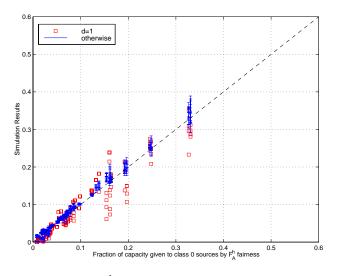


Fig. 7. HETRTT, F_A^h vs. simulation, $\tau_0=0.2$ s, $\tau_i=\tau_0/I$, for $i=1,2,\ldots,I$, and $r_i=5,\,\eta_i=0.5$ for all flows.

and illustrate how the system operates according to different fairness criteria depending on RTT.

IV. DISCUSSION AND RELATED WORK

To simplify the presentation let us consider the HOM-RTT network with synchronous rate updating and common RTT, τ , where $\{x_k\}_{k\geq 0}$ represents a sequence of rate vectors. In Section II-A we argued that n+1-th rate updating x_{n+1} is based on the feedback depending on x_{n-1} , and not on x_n , as is commonly assumed in the existing work, e.g. [4]–[6]. Recall that we refer to this as the stolen lag. From the weak convergence it follows that the stolen lag is not relevant in respect to the limit mean rate distribution. However, the rate dynamics with no stolen lag has a higher variance.

Intuitively, this can be explained as follows. Let us assume that the network operates in the steady state and the packet–loss ratio is small and constant. Then, negative feedback $I_{i,n}$ is mainly driven by either $x_{i,n}$ or $x_{i,n-1}$, for the stolen lag and no stolen lag, respectively. During the period of rate decreasing, at n+1-th rate updating, rate $x_{i,n}$ is further decreased with probability proportional to either $x_{i,n}$, or $x_{i,n-1}$, for the stolen lag and no stolen lag, respectively. Since $x_{i,n-1} \geq x_{i,n}$, it follows that with a higher probability the rate is further decreased with no stolen lag, compared to the stolen lag case. Analogous reasoning applies also for a period of the rate increasing. This implies a higher variance of the rate for the case with no stolen lag. We substantiated the above argument by Markov chain modeling of a discrete process x, for which we computed

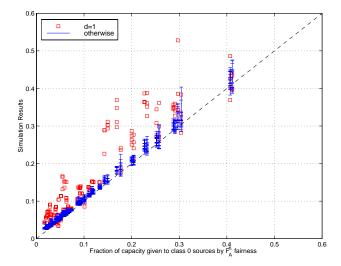


Fig. 8. HETRTT, additive-increment proportional to RTT, $r_i = K\tau_i$, K = 25, $\eta_i = .5$, $\tau_0 = .2$ s, $\tau_i = \tau_0/I$, $i = 1, 2, \ldots, I$.

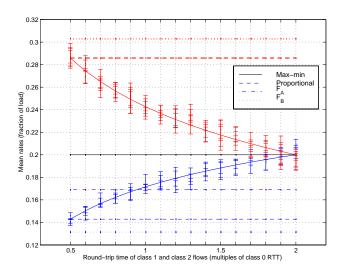


Fig. 9. Impact of the RTT on rate distribution, c =250 , I = 2, v = 3, w = 2, τ_0 = 0.2, τ_i = 0.1 – 0.4, δ_1 = δ_2 = δ = 5 ε = 0.05.

stationary distribution, for both cases. However, due to the space limitations these results are not shown here. Instead, in Fig. 10 and 11, the HOMRTT simulation results are shown for both the stolen lag and no stolen lag case, respectively. Traces of the utilization of link 1, in Fig. 10c and 11c, particularly demonstrate observed phenomena.

Furthermore, it turns out that for the parking-lot topology, analyzed in Section III, the throughput of flows traversing multiple bottlenecks (34) fully conforms to the results obtained in [1]. Let $W(\cdot)$ be a real-valued strictly positive function, then for TCP it generally holds r_i

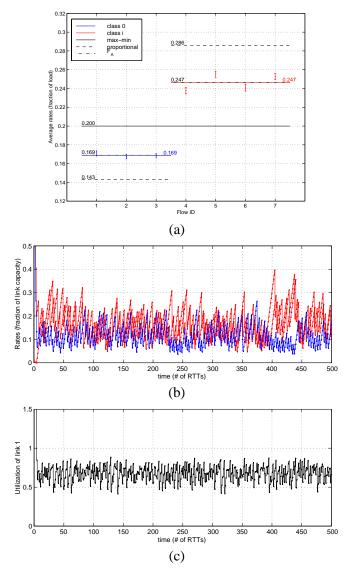


Fig. 10. HOMRTT with the stolen lag, c=250, $\tau=0.2$, $r_i=5$, $\eta_i=0.5$, I=2, v=3, w=2, d=0.5, p=5, (a) rate distribution, (b) rate traces, and (c) utilization of link 1.

 $W(\tau_i)/\tau_i$, thus (34) becomes

$$\frac{x_0}{c} = \frac{1}{v + w\sqrt{I}\frac{\tau_0}{\tau_I}\sqrt{\frac{W(\tau_I)}{W(\tau_0)}}},$$

which is exactly the same form, for both a single bottleneck ($I \dashv$) with different RTTs, and multiple bottlenecks, as obtained by simplified modeling in [1].

V. APPLICATION TO THE INTERNET

A. Bias Against Connections with Long RTTs

Our fundamental result allows us to better understand the bias of TCP against connections with long round trip times.

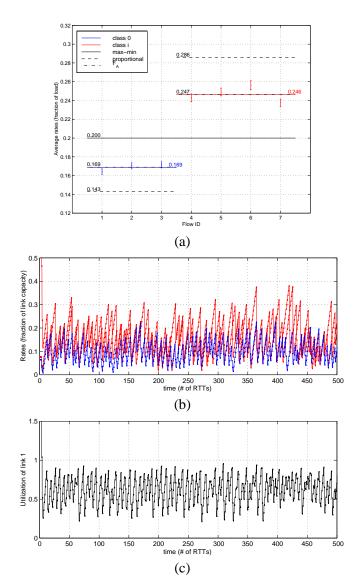


Fig. 11. HOMRTT with no stolen lag, c=2500, $\tau=0.2$, $r_i=5$, $\eta_i=0.5$, I=2, v=3, w=2, d=0.5, p=5, (a) rate distribution, (b) rate traces, and (c) utilization of link 1.

First, for any congestion control mechanism, if the distribution of rates tends to maximize a concave utility function, then flows with many hops are likely to receive a small rate [5]. Since F_A^h is concave with x_i , this is true with our system, whatever the rate adaptation parameters r_i and η_i are. This is probably a desired bias, since flows with many hops use more network resources. In practice, many hops often mean larger RTT, but not always.

Second, both the specific values of the rate adaptation parameters, r_i and η_i , and the update frequency $1/\tau_i$ also play a role (τ_i is the RTT for source i). With TCP–Reno, with no delayed ACKs, we have:

• $r_i = \text{MSS}_i/\tau_i$ – in the congestion avoidance phase, with no delayed ACKs, the window is effectively increased by one packet per RTT (resp. 1/2 packet for delayed ACKs).

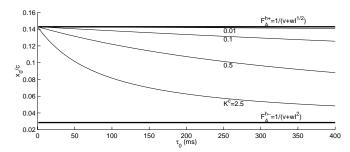


Fig. 12. Residual bias against long RTT connections, HETRTT, $\tau_I = \tau_0/I$, I = 4, v = 3, w = 2, $K^c = K/c$.

In the formula, MSS_i is the packet size for source i.

• $\eta_i = 0.5$ - when a loss is detected, the target window size is divided by a factor 2.

This results in an obvious bias against long RTTs: the increase element r_i is smaller, and less frequent for long RTTs.

A fix to this undesired bias has been proposed in [1]; it consists in setting the increase element to $r_i = K\tau_i$, where K is a common constant. This is actually done by multiplying the window increase parameter by τ_i^2 , considering that the rate is equal to the window size divided by τ_i . In the limiting case where rates x_i are large, we can apply the first order development (26) and obtain that the rates tend to maximize $-\sum_{i\in\mathcal{S}}\frac{1}{x_i}$, thus there is no bias against long RTTs in this case (but the bias against many hops does exist). However, the non-linear nature of F_A^h indicates that this is not always true. If some rates are small, for example in a very loaded network, then their utility function equals to $\frac{1}{\tau_i}(-\log(K\tau_i) + \log x_i)$, or if this holds for all flows, then the distribution of rates tends towards weighted proportional fairness (27), with a weight equal to the inverse of the RTT. Thus there might still remain an undesirable negative bias against long RTTs for precisely those sources that are not able to obtain a large rate. We can verify this bias in the results illustrated in Fig. 12. However, note that in practice, the case where sources have small rates might not correspond to the regime of rare negative feedback, which is assumed in our paper. A verification by simulation is thus required, and remains to be done.

B. TCP throughput-loss formula

In the steady–state, supposing that $g_l(\cdot) << 1$, $q_i \triangleq \sum_{l \in \mathcal{L}} A_{l,i} g_l(f_l)$ corresponds to the packet loss–ratio. In the limit case $\dot{x^*} = 0$, as $t \to \infty$, from the limit mean ODE (20) one can obtain

$$x_i^* = \sqrt{\frac{r_i}{\eta_i \tau_i q_i^*}}, \text{ as } q_i^* \to 0.$$
 (35)

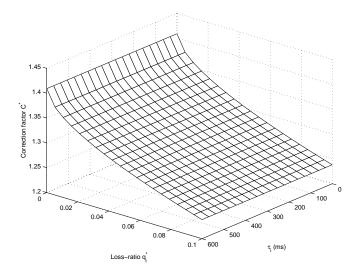


Fig. 13. Correction of the TCP throughput-loss formula (not delayed ACKs).

For a TCP connection, (35) is referred to as a TCP throughput–loss formula

$$x_i^* = \frac{C}{\tau_i \sqrt{q_i^*}} \, \text{pdx /s}, \tag{36}$$

where, from (35), $C=\sqrt{2}$ (resp. C=1), for not delayed ACKs $r_i = 1/\tau_i$ (resp. for delayed ACKs $r_i = 1/\tau_i$), and in both cases $\eta_i = 1/\tau_i$. The reported measurements [2] indicate that C varies substantially for different TCP implementations.

TABLE II LSE CORRECTION OF THE TCP THROUGHPUT-LOSS FORMULA (v er $q_i^* \in [10^{-4}, 10^{-1}]$ for $\tau_i \in [10$ ms600m).

ACK Type	$LSE(\hat{C}), \alpha = 0.5$	$LSE(\hat{C}, \alpha)$
not delayed ACK	1.2997	1.1863, 0.5274
delayed ACK	0.9377	0.8731, 0.5215

In the derivation of (35) we used a conditional probability of negative feedback equal to $x_i\tau_i\sum_{l:A_{l,i}>0}g_l(f_l)$. However, in Section II-B it is shown that, in general, conditional probability of negative feedback has the form $1-\prod_{l:A_{l,i}>0}\left[1-g_l(f_l)\right]^{x_i\tau_i}$, from which the former expression follows for $g_l(\cdot)\to 0$. In reference to [2], we consider a single bottleneck so that in the steady state $g_l\equiv q_i^*$ is the packet loss–ratio. Hence, using the exponential form of the probability of negative feedback in the limit mean ODE, $\dot{x}^*=0$, as $t\to\infty$, yields a throughput–loss formula of the form

$$r_i - (r_i + \eta_i x_i^*)[1 - (1 - q_i^*)^{x_i^* \tau_i}] = 0.$$
 (37)

Let y_i^* be a solution of (37) and x_i^* be given by (35). Then we compute a correction factor $C^* = y_i^* C/x_i^*$, which is plotted in Fig. 13. It is evident that C^* depends significantly on the loss–ratio q_i^* . We computed Least Square Error (LSE) fitting of the function $x_i = \hat{C}/(\tau_i q_i^\alpha)$ to the solution of (37), in respect to parameters $(\hat{C}, \alpha = 0.5)$, and (\hat{C}, α) , over values of q_i^* and τ_i as indicated in Table II. Finally, it turns out that the fitted values (Table II) are in agreement with the parameters fitted to the measured data of a set of TCP implementations [2].

VI. CONCLUSION

We have shown that under certain reasonable assumptions, congestion avoidance based on additive-increase and multiplicative-decrease shares throughput according to a generalization of F_A fairness, which explicitly accounts for communication delays and differences in round trip times. It is important to note that this influence does not come into play because of the delayed feedback to the end-systems, but it is a fundamental consequence of the diversity of the rate updating intervals. It is demonstrated that different fairness objectives can be achieved depending on the values of additive increments and RTTs. It is shown that throughput results [1] are an asymptotic case of F_A^h , for small additive–increase/multiplicative–decrease ratio relatively to connection throughput. We also gave some insight into the bias against flows with long round trip times. Further research should concentrate on analyzing the impact of RTT variations due to queueing.

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APPENDIX

I. THE LIMIT MEAN ODE

Theorem 1 ([8], Ch. 12.3, p. 364–365) Assume (12)-(16). Then

$$\{x_i^{\gamma}(\cdot), \tau_i^{\gamma}(\cdot), \hat{x}^{\gamma}(\cdot), N_i^{\gamma}(\cdot), i \in \mathcal{S} \}$$

is tight in $D^{4|S|}[0,\infty)$. Let γ index a weakly convergent subsequence whose weak sense limit we denote by

$$(x_i(\cdot), \tau_i(\cdot), \hat{x}(\cdot), N_i(\cdot), i \in \mathcal{S}).$$

Then the limits are Lipschitz continuous with probability one and

$$x_i(t) = \hat{x}_i(\tau_i(t)), \ \hat{x}_i(t) = x_i(N_i(t))$$
 (38)

$$N_i(\tau_i(t)) = t.$$
 (39)

Also,

$$\tau_i(t) = \int_0^t \bar{u}_i(\hat{x}(\tau_i(s)))ds, \tag{40}$$

$$\dot{x}_i(t) = \bar{h}_i(\hat{x}(\tau_i(t))) + z_i(t), \tag{41}$$

$$\dot{\hat{x}}_i = \frac{\bar{h}_i(\hat{x})}{\bar{u}_i(\hat{x})} + \hat{z}_i, i \in \mathcal{S}, \tag{42}$$

where the z_i and \hat{z}_i serve the purpose of keeping the paths in the interval $[a_i, b_i]$.

If L_C is asymptotically stable in the sense of Lyapunov, then $\forall \mu > 0 \; \exists \; T_{\mu} > 0 \; \text{s.t.}$ for $t \geq T_{\mu}, \; ||\hat{x}(t), L_C|| < \mu$. Also, $\forall T > T_{\mu}$,

$$\limsup_{\gamma} P \left\{ \sup_{T_{\mu} \le t \le T} || \hat{x}^{\gamma}(t), L_C| \ge \mu \right\} = 0.$$
 (43)

For large T and $T_1 \not\supset T$, $\hat{x}^{\gamma}(\cdot)$, $t \in [T,T_1]$, spends nearly all of its time in a small neighborhood of L_C . Now, drop the constraint set C and suppose that $\{\hat{x}_n^{\gamma}; \gamma, n\}$ is bounded with probability one. Then the above conclusions continue to hold with $z_i(t) = \hat{z}_i(t) = 0$ and L_C replaced by some limit set of (42).