

# Probabilistic version of Trotter's exponential product formula in Banach algebras

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## 1. Introduction and results

It is an elementary fact that the exponential function may be defined by the equivalent formulae

$$\exp(x) = \lim_{n \rightarrow \infty} \left( 1 + \frac{1}{n} x \right)^n = \sum_{n=0}^{\infty} \frac{1}{n!} x^n$$

not only when  $x$  is a real or complex number but also when it is a matrix with real or complex entries or a bounded operator acting on a Hilbert space or a Banach space or, even when it is an element of an abstract Banach algebra  $\mathfrak{B}$  with identity 1 (for a definition of a Banach algebra see for instance [1]). If  $\mathfrak{B}$  is not commutative then in general  $\exp(x)\exp(y) \neq \exp(x+y)$ . There is, however, a formula which replaces the addition law of the exponential function, namely

$$(1) \quad \lim_{n \rightarrow \infty} \left( \exp\left(\frac{x}{n}\right) \exp\left(\frac{y}{n}\right) \right)^n = \exp(x+y)$$

and this holds regardless whether  $x$  and  $y$  commute or not. Formula (1) is capable of further generalization; see TROTTER [2]. Specifically,  $x$  and  $y$  may be unbounded operators of a certain type, namely generators of continuous one-parameter operator semi-groups. In the present paper we are not concerned with Trotter's generalization, but we shall still refer to (1) as the Trotter product formula. The symbols  $x, y, \dots, a, b, \dots$  shall generally denote elements of the Banach algebra  $\mathfrak{B}$ . The norm of  $x \in \mathfrak{B}$  is written  $\|x\|$ .

Let  $\mathbf{X} = (x_1, x_2, \dots, x_m)$  be any finite sequence of elements of  $\mathfrak{B}$ . With any such sequence we associate the product

$$T(\mathbf{X}) = \exp\left(\frac{x_1}{m}\right) \exp\left(\frac{x_2}{m}\right) \dots \exp\left(\frac{x_m}{m}\right)$$

which will be called its Trotter product. Note that it depends essentially on the

order of the factors, i.e. on  $\mathbf{X}$  as a sequence, not merely as a set. We also write, for the mean of the elements of  $\mathbf{X}$ ,  $M(\mathbf{X}) = \frac{1}{m} (x_1 + x_2 + \dots + x_m)$ . Using this notation we can express the Trotter product formula as follows: If  $\mathbf{X}_k$  (for  $k=1, 2, \dots$ ) is the sequence of length  $2k$  whose elements are alternatingly  $x$  and  $y$ , then

$$(2) \quad \lim_{k \rightarrow \infty} T(\mathbf{X}_k) = \exp(\lim_{k \rightarrow \infty} M(\mathbf{X}_k)).$$

We now raise the following question. Under what natural conditions on a sequence  $\mathbf{X}_1, \mathbf{X}_2, \dots$  (subject to the requirement that their lengths tend to infinity) will (2) hold?

We shall prove two theorems relevant to this question. The first theorem gives a rather general sufficient condition for an infinite sequence  $\mathbf{X}_1, \mathbf{X}_2, \dots$  to satisfy (2). The original Trotter formula is an obvious consequence of this condition. But our theorem also shows that to a certain extent the order of the factors in the Trotter product may be made subject to considerable rearrangement without destroying the validity of (2).

For any  $\mathbf{X} = (x_1, x_2, \dots, x_m)$  let us write  $\varrho = \varrho(\mathbf{X}) = \text{Max}_{1 \leq j \leq m} \|x_j\|$ . Let  $\pi$  denote a partition of the sequence  $(1, 2, \dots, m)$  into successive subsequences  $(1, 2, \dots, m_1)$ ,  $(m_1 + 1, m_1 + 2, \dots, m_1 + m_2)$ ,  $\dots$ ,  $(m - m_s + 1, m - m_s + 2, \dots, m)$ , and let  $\mathbf{Y}_1, \mathbf{Y}_2, \dots, \mathbf{Y}_s$  be the corresponding subsequences of  $\mathbf{X}$ . For any element  $g \in \mathfrak{B}$  we introduce a quantity  $\delta = \delta(\mathbf{X}, \pi, g)$  which measures the closeness to which  $g$  uniformly approximates the "partial means" of  $\mathbf{X}$  induced by the partition  $\pi$

$$(3) \quad \delta(\mathbf{X}, \pi, g) = \text{Max}_{1 \leq j \leq s} \|M(\mathbf{Y}_j) - g\|.$$

We also define a quantity

$$(4) \quad \eta = \eta(\pi) = \sum_{j=1}^s \left( \frac{m_j}{m} \right)^2.$$

If  $\eta_0 = \max_{1 \leq j \leq s} \left( \frac{m_j}{m} \right)$ , we have clearly

$$(5) \quad \eta_0^2 \leq \eta \leq \eta_0,$$

so that  $\eta$  is a kind of a measure for the relative fineness of  $\pi$ .

**Theorem 1.** *Let  $\mathbf{X}_1, \mathbf{X}_2, \dots$  be an infinite sequence of finite sequences of elements of  $\mathfrak{B}$ . Suppose that  $\varrho(\mathbf{X}_k)$  is bounded. Suppose that  $g \in \mathfrak{B}$  exists and a sequence of partitions  $\pi_k$  of  $\mathbf{X}_k$  into successive subsequences exists such that  $\eta(\pi_k) \rightarrow 0$  and  $\delta(\mathbf{X}_k, \pi_k, g) \rightarrow 0$  as  $k \rightarrow \infty$ . Then  $M(\mathbf{X}_k) \rightarrow g$  and  $T(\mathbf{X}_k) \rightarrow \exp(g)$  as  $k \rightarrow \infty$ .*

Note that the quantity  $\delta(\mathbf{X}, \pi, g)$  is independent of the order of the elements  $x_i$  within each of the subsequences  $\mathbf{Y}_j$  ( $j=1, 2, \dots, s$ ). Thus our theorem formulates

mathematically the intuitive notion that the order of the factors in the Trotter product  $T(\mathbf{X})$  is irrelevant "locally" and only essential "in the large".

Our second Theorem deals with the following problem. Suppose we consider infinite sequences  $\mathbf{Z}=(x_1, x_2, \dots)$  whose elements are all taken from a fixed finite subset  $\{a_1, a_2, \dots, a_\sigma\}$  of  $\mathfrak{B}$ . With any such infinite sequence we may consider the sequence of its finite sections  $\mathbf{X}_1=(x_1), \mathbf{X}_2=(x_1, x_2), \dots$ . Is it possible to characterize the extent of the set of those infinite sequences  $\mathbf{Z}$  for which the generalized Trotter product formula (2) holds?

This question leads to measure-theoretic considerations, and can be naturally formulated in probabilistic terminology.

**Theorem 2.** *Let  $a_1, a_2, \dots, a_\sigma \in \mathfrak{B}$ , and let  $p_1, p_2, \dots, p_\sigma \geq 0, \sum p_\nu = 1$ . Suppose an infinite (random) sequence  $\mathbf{Z}=(x_1, x_2, \dots)$  is formed by choosing, independently for each  $j \geq 1, x_j = a_\sigma$  with probability  $p_\sigma$ . Let  $g = p_1 a_1 + p_2 a_2 + \dots + p_\sigma a_\sigma$ . Then the probability is unity that  $M(\mathbf{X}_k) \rightarrow g$  and  $T(\mathbf{X}_k) \rightarrow \exp(g)$  as  $k \rightarrow \infty$ .*

Thus in the sense of the given probability measure defined on the set of sequences  $\mathbf{Z}$ , for almost every sequence the generalized Trotter product formula holds.

### 2. An auxiliary inequality

Both theorems derive from an elementary estimate formulated as follows:

**Lemma.** *Let  $\mathbf{X}$  be a finite sequence of elements of  $\mathfrak{B}$ ,  $\pi$  any partition of it into successive subsequences, and  $g \in \mathfrak{B}$ . Let  $q = q(\mathbf{X}), \eta = \eta(\pi)$  and  $\delta = \delta(\mathbf{X}, \pi, g)$  be defined as above. Then*

$$\|T(\mathbf{X}) - \exp(g)\| \leq e^{\|g\|} \left( e^{\delta + \frac{1}{2} \eta e^2 e^2} - 2e^{-\frac{1}{2} \eta \|g\|^2} + 1 \right).$$

In the proof of the Lemma we shall make use of the following facts:

(A) If  $P(x_1, x_2, \dots, x_q)$  is a polynomial or a power series with non-negative real coefficients then

$$\|P(x_1, x_2, \dots, x_q)\| \leq P(\|x_1\|, \|x_2\|, \dots, \|x_q\|)$$

for all  $x_j \in \mathfrak{B}$  such that the right hand side is finite.

(B) For any  $t \geq 0, e^t - 1 - t \leq \frac{1}{2} t^2 e^t$

(C) For any  $t \geq 0, e^{t-2t^2} \leq 1 + t \leq e^t$ .

Let  $\mathbf{Y}_j (j=1, 2, \dots, s)$  be the subsequences of  $\mathbf{X}$  produced by the partition  $\pi$ , and let  $m_j$  be their respective length,  $m = \sum_j m_j$ . Let  $T(\mathbf{X}) = y_1 y_2 \dots y_s$  where  $y_j$  is the product of those factors in the product, taken in their proper order, which

involve the elements  $x_k \in Y_j$ . Write  $y_j = 1 + \frac{m_j}{m}g + r_j$ , thus defining  $r_j$ . For notational convenience we now consider  $j=1$ . We have

$$r_1 = \left[ y_1 - 1 - \frac{m_1}{m} M(Y_1) \right] + \frac{m_1}{m} [M(Y_1) - g].$$

The norm of the second term is bounded by  $\frac{m_1}{m} \delta$ . To obtain a bound on the norm of the first term we note that if

$$y_1 - 1 - \frac{m_1}{m} M(Y_1) = \exp\left(\frac{x_1}{m}\right) \exp\left(\frac{x_2}{m}\right) \dots \exp\left(\frac{x_{m_1}}{m}\right) - 1 - \frac{1}{m} (x_1 + x_2 + \dots + x_{m_1})$$

is regarded as a power series in the  $x_j$  ( $j=1, 2, \dots, m_1$ ) it has non-negative coefficients (the negative terms cancel!). Thus by principle (A) above we may replace  $x_j$  by  $\|x_j\|$ , and then taking (B) and the definition of  $\varrho$  into account we get

$$\left\| y_1 - 1 - \frac{m_1}{m} M(Y_1) \right\| \leq \frac{1}{2} \left( \frac{m_1}{m} \right)^2 \varrho^2 e^\varrho.$$

Thus we have for  $j=1, 2, \dots, s$

$$(6) \quad \|r_j\| \leq \frac{1}{2} \left( \frac{m_j}{m} \right)^2 \varrho^2 e^\varrho + \frac{m_j}{m} \delta.$$

Next, let  $z_j = 1 + \frac{m_j}{m}g$ , and consider the difference  $T(X) - z_1 z_2 \dots z_s = y_1 y_2 \dots y_s - z_1 z_2 \dots z_s$ . As a polynomial in  $g$  and the  $r_j$ , it has again non-negative coefficients, so we apply principle (A). The norms of  $r_j$  are majorized by (6), and so by the inequality (C)

$$1 + \frac{m_j}{m} \|g\| + \|r_j\| \leq e^{\frac{m_j}{m} \|g\| + \frac{1}{2} \left( \frac{m_j}{m} \right)^2 \varrho^2 e^\varrho + \frac{m_j}{m} \delta},$$

and

$$-\left( 1 + \frac{m_j}{m} \|g\| \right) \leq -e^{\frac{m_j}{m} \|g\| - \frac{1}{2} \left( \frac{m_j}{m} \right)^2 \varrho^2 e^\varrho},$$

where in the last step we used  $\frac{m_j}{m} \|g\| \leq \sqrt{\eta} \|g\| \leq 1$ . Thus we see that

$$\|y_1 y_2 \dots y_s - z_1 z_2 \dots z_s\| \leq e^{\|g\|} \left( e^{\delta + \frac{1}{2} \eta \varrho^2 e^\varrho} - e^{-\frac{1}{2} \eta \|g\|^2} \right).$$

Arguing analogously, we have also

$$\|\exp(g) - z_1 z_2 \dots z_s\| \leq e^{\|g\|} \left( 1 - e^{-\frac{1}{2} \eta \|g\|^2} \right).$$

The last two inequalities together imply the conclusion of the lemma.

We note that Theorem 1 is an immediate consequence of the lemma, since the estimate for  $\|T(X_k) - \exp(g)\|$  supplied by the lemma tends to zero as  $k \rightarrow \infty$  if the hypotheses of the theorem are fulfilled.

### 3. Proof of Theorem 2

The idea of the proof of Theorem 2 is to find an appropriate sequence of partitions  $\pi_k$  ( $k=1, 2, 3, \dots$ ) such that if we let  $\delta_k = \delta(X_k, \pi_k, g)$  then

$$(7) \quad \mathbf{P}\{\lim_{k \rightarrow \infty} \delta_k = 0\} = 1,$$

and at the same time such that

$$(8) \quad \lim_{k \rightarrow \infty} \eta(\pi_k) = 0.$$

Indeed, if (7) and (8) are fulfilled then the conclusion of Theorem 2 follows from Theorem 1.

Let  $C_1 < C_2$  and  $\beta < 1$  be three positive constants. We define the partition  $\pi_k$  of  $(1, 2, 3, \dots, k)$  into successive subsequences of lengths  $m_j = m_j(k)$  ( $j=1, 2, \dots, s=s(k)$ ) in such a manner that for all  $j$  and  $k$

$$(9) \quad C_1 k^\beta < m_j < C_2 k^\beta.$$

Since  $\sum_j m_j = k$ , it follows that

$$(10) \quad s = s(k) = O(k^{1-\beta}),$$

and therefore

$$(11) \quad \eta_k = \eta(\pi_k) = \sum_{j=1}^{s(k)} \left(\frac{m_j}{k}\right)^2 = O(k^{\beta-1}),$$

so that (8) holds. Note that in our probabilistic set-up the partitions  $\pi_k$  are not random variables (i.e.  $\pi_k$  is constant over the whole probability space).

Next, we remark that, by virtue of the Borel—Cantelli lemma, in order to prove (7) it is sufficient to prove that

$$(12) \quad \sum_{k=1}^{\infty} \mathbf{P}\{\delta_k \geq \varepsilon\} < \infty$$

for any positive  $\varepsilon$ .

Let  $Y_1, Y_2, \dots, Y_s$  be the successive subsequences of  $X_k$  produced by the partition  $\pi_k$ . Let  $N_{jv}$  ( $j=1, 2, \dots, s$ ;  $v=1, 2, \dots, \sigma$ ) be the number of occurrences of  $a_v$  among the elements of the subsequence  $Y_j$ . The  $N_{jv}$  are random variables subject to the multinomial distribution determined by the probabilities  $p_v$ , and for different  $j$  they are independent. We have

$$\delta_k = \text{Max}_{1 \leq j \leq s} \left\| \sum_{v=1}^{\sigma} \left(\frac{N_{jv}}{m_j} - p_v\right) a_v \right\| \leq AM_k,$$

where

$$A = \text{Max}_v \|a_v\| \quad \text{and} \quad M_k = \text{Max}_{1 \leq j \leq s} \sum_{v=1}^{\sigma} \left| \frac{N_{jv}}{m_j} - p_v \right|.$$

Thus we need to show

$$(13) \quad \sum_{k=1}^{\infty} \mathbf{P}\{M_k \geq \varepsilon\} < \infty.$$

Since the following inclusion (implication) of events holds

$$\{M_k \geq \varepsilon\} = \bigcup_{j=1}^{s(k)} \left\{ \sum_{v=1}^{\sigma} \left| \frac{N_{jv}}{m_j} - p_v \right| \geq \varepsilon \right\} \subseteq \bigcup_{j=1}^{s(k)} \bigcup_{v=1}^{\sigma} \left\{ \left| \frac{N_{jv}}{m_j} - p_v \right| \geq \frac{\varepsilon}{\sigma} \right\},$$

we obtain for the probabilities of the complementary events

$$(14) \quad \mathbf{P}\{M_k < \varepsilon\} \geq \mathbf{P} \left\{ \bigcap_{j=1}^{s(k)} \bigcap_{v=1}^{\sigma} \left\{ \left| \frac{N_{jv}}{m_j} - p_v \right| < \frac{\varepsilon}{\sigma} \right\} \right\} = \\ = \prod_{j=1}^{s(k)} \mathbf{P} \left\{ \bigcap_{v=1}^{\sigma} \left\{ \left| \frac{N_{jv}}{m_j} - p_v \right| < \frac{\varepsilon}{\sigma} \right\} \right\} \geq \prod_{j=1}^{s(k)} \left[ 1 - \sum_{v=1}^{\sigma} \mathbf{P} \left\{ \left| \frac{N_{jv}}{m_j} - p_v \right| \geq \frac{\varepsilon}{\sigma} \right\} \right].$$

The equality in (14) is due to the fact that we are dealing with the intersection (conjunction) of independent events.

Suppose  $N$  is the number of "successes" in a sequence of  $m$  Bernoulli trials with probability  $p$  for success. We have then the following fact [3]: given any  $\alpha > 1$ , for all sufficiently large  $m$  (depending only on  $\alpha$  and  $p$ )

$$\mathbf{P} \left\{ \frac{|N - mp|}{[mp(1-p)]^{1/2}} \geq (2\alpha \log m)^{1/2} \right\} < \frac{1}{m^{\alpha}}.$$

It follows that for any  $\varepsilon > 0$

$$(15) \quad \mathbf{P} \left\{ \left| \frac{N}{m} - p \right| \geq \varepsilon \right\} < \frac{1}{m^{\alpha}}$$

for all sufficiently large  $m$  (depending only on  $\varepsilon$ ,  $\alpha$  and  $p$ ). If the inequality (15) is used for the probabilities on the right hand side of (14) we obtain

$$\mathbf{P}\{M_k < \varepsilon\} \geq \prod_{j=1}^{s(k)} \left[ 1 - \frac{\sigma}{m_j^{\alpha}} \right]$$

which is valid for all sufficiently large  $k$ , since (9) implies that then all the  $m_j$  are large enough. But (9) and (10) show that for suitable constants  $C$  and  $C'$

$$(16) \quad \mathbf{P}\{M_k < \varepsilon\} \geq (1 - Ck^{-\alpha\beta})^{C'k^{1-\beta}} = 1 - O(k^{-\gamma})$$

where  $\gamma = \alpha\beta + \beta - 1$ . Since  $\alpha > 1$  was arbitrary we may suppose  $\gamma > 1$ , so that  $\mathbf{P}\{M_k \geq \varepsilon\} = O(k^{-\gamma})$  and (13) follows. This concludes the proof of Theorem 2.

**References**

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- [3] W. FELLER, *Introduction to Probability Theory and its Applications*. I (New York, 1957). See Section VIII. 4, especially equ. (4. 5).

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