

Centering lower-level interactions in multilevel models

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Longitudinal diary study on sexual behavior in Flanders

- ▶ Info on 66 heterosexual couples
 - ▶ Here, we only focus on men's data
- ▶ Daily measures during 3 weeks:
 - ▶ Daily morning reports on sexual and intimate behavior: amount of intimate acts (kissing, cuddling and caressing) measured on a 7-point scale
 - ▶ Daily evening reports on positive relationship feelings: average score of 9 items (happy, satisfied, understood, . . .) measured on a 7-point scale



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Question: What is the contribution of intimacy to next-day positive relationship feelings?

Standard multilevel modeling

In our example:

- ▶ X_{ij} : daily measurement of intimacy of individual j at time i
- ▶ Y_{ij} : next day's positive relational feelings

Standard analysis by a multilevel model with random intercept b_j :

$$E(Y_{ij}|X_{ij}, b_j) = \gamma_0 + \gamma X_{ij} + b_j \quad (1)$$



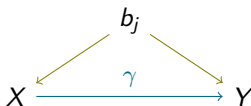
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Unfortunately, there may be **upper level endogeneity!**

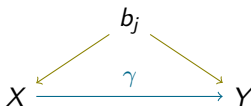
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(1) assumes that b_j and X_{ij} are independent
 \Rightarrow **biased estimator for γ** under upper-level endogeneity!

Centring of lower level effects

A solution to the upper-level endogeneity problem is to separate within-from between-effects:

$$E(Y_{ij} | X_{ij}, u_j) = \gamma_0 + \gamma_W X_{ij}^c + \gamma_B \bar{X}_j + b_j \quad (2)$$

$$\text{with } \bar{X}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} X_{ij} \text{ and } X_{ij}^c = X_{ij} - \bar{X}_j$$

- ▶ γ_W captures the **within-subject** effect
- ▶ γ_B captures the **between-subject** effect

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- ▶ As $X_{ij} - \bar{X}_j$ removes all upper level effects, it no longer depends on b_j in case of upper level endogeneity.
- ▶ The OLS-estimator for γ_W will converge to (in balanced designs):

$$\hat{\gamma}_W = \frac{\text{cov}(Y_{ij}, X_{ij} - \bar{X}_j)}{\text{var}(X_{ij} - \bar{X}_j)} \rightarrow \gamma \quad (3)$$

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⇒ No bias!

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(note that \bar{X}_j can also be excluded from (2), as $\bar{X}_j \perp X_{ij}^c$)



Revisited: Longitudinal study on sexual behavior in Flanders

New question: Does the effect of intimacy on next-day positive relationship feelings differ according to whether or not the participant has masturbated the previous day?

- ▶ X_{ij} : daily measurement of intimacy of individual j at time i
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Adjusted multilevel model:

$$E(Y_{ij} | X_{ij}, Z_{ij}, u_j) = \gamma_0 + \gamma_1 X_{ij} + \gamma_2 Z_{ij} + \gamma_3 X_{ij} Z_{ij} + b_j \quad (4)$$

⇒ Lower level interaction term!



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- ▶ b_j may again be correlated with X , Y and/or Z in case of upper level endogeneity
- ▶ The 'naive' model may then again yield biased estimators for the γ 's



Centring of lower-level interactions

There are two possible centring approaches:

- ▶ Product first, center next (P1C2):

$$X_{ij} * Z_{ij}$$

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 \overline{XZ}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} X_{ij} Z_{ij} & & (XZ)_{ij}^c = X_{ij} Z_{ij} - \overline{XZ}_j
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 & & & \swarrow & \searrow \\
 & & & X_{ij}^c * \overline{Z}_j & \overline{X}_j * Z_{ij}^c
 \end{array}$$

Question: which approach should we take? Do they differ in any way?



The P1C2-approach

► P1C2 model:

$$E(Y_{ij} | X_{ij}, Z_{ij}, u_j) = \gamma_0 + \gamma_1 X_{ij}^c + \gamma_2 Z_{ij}^c + \gamma_3 (XZ)_{ij}^c + b_j \quad (5)$$

with $(XZ)_{ij}^c = X_{ij}Z_{ij} - \overline{XZ}_j$ (and $\overline{XZ}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} X_{ij}Z_{ij}$)

⇒ $\hat{\gamma}_1$, $\hat{\gamma}_2$ and $\hat{\gamma}_3$ are unbiased estimators for γ_1 , γ_2 and γ_3

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- P1C2+ model:

$$E(Y_{ij} | X_{ij}, Z_{ij}, u_j) = \gamma_0 + \gamma_1 X_{ij}^c + \gamma_2 Z_{ij}^c + \gamma_3 (XZ)_{ij}^c + \gamma_4 \overline{X}_j + \gamma_5 \overline{Z}_j + \gamma_6 \overline{XZ}_j + b_j$$

⇒ in balanced designs, the estimated within-effects $\hat{\gamma}_1$, $\hat{\gamma}_2$ and $\hat{\gamma}_3$ are identical in both models



The C1P2-approach

- ▶ C1P2 model:

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- C1P2++ model:

$$E(Y_{ij} | X_{ij}, Z_{ij}, u_j) = \gamma_0 + \gamma_1 X_{ij}^c + \gamma_2 Z_{ij}^c + \gamma_3 X_{ij}^c Z_{ij}^c + \gamma_4 \bar{X}_j + \gamma_5 \bar{Z}_j \\ + \gamma_6 \bar{X}_j \bar{Z}_j + \gamma_7 \bar{X}_j Z_{ij}^c + \gamma_8 \bar{Z}_j X_{ij}^c + b_j$$

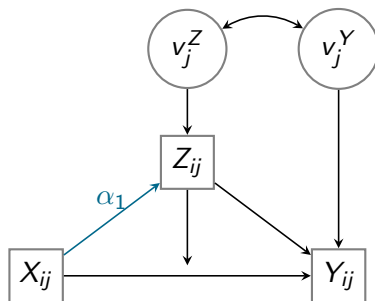
Example - Results

	<i>Intimacy = X</i>		<i>Masturbation = Z</i>		<i>Interaction = XZ</i>	
	Estimate	p-value	Estimate	p-value	Estimate	p-value
P1C2	0.079 (0.015)	< .001	-0.151 (0.079)	.057	-.0075 (0.039)	.054
P1C2+	0.079 (0.015)	< .001	-0.150 (0.079)	.059	-.0075 (0.039)	.054
C1P2	0.080 (0.015)	< .001	-0.163 (0.080)	.042	-.0102 (0.050)	.042
C1P2+	0.080 (0.015)	< .001	-0.160 (0.080)	.045	-.0098 (0.045)	.049
C1P2++	0.080 (0.015)	< .001	-0.167 (0.080)	.037	-.0096 (0.050)	.056

- ▶ Different approaches lead to **different estimates**
- ▶ Different approaches lead to **different conclusions** (at the 5% significance level)!



Simulation Study - Settings



Simulation	α_1	Distribution of X
Sim 1	0.0	$N(0, 1)$
Sim 2	0.0	$B(1, 0.5) - 0.5$
Sim 3	-0.2	$B(1, 0.5) - 0.5$
Sim 4	-1.5	$B(1, 0.5) - 0.5$



Simulation study - Results

- ▶ X and Z are grand-mean centred to facilitate interpretation
- ▶ Focus on within-effects only



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- ▶ Bias for interaction effect in $C1P2=C1P2+$ when Z is a mediator:

$$E(\hat{\gamma}_3) = \gamma_3 \frac{\text{cov}[X_{ij}Z_{ij}, X_{ij}^c Z_{ij}^c]}{\text{var}[X_{ij}^c Z_{ij}^c]} \quad (7)$$

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- ▶ No bias for $P1C2 = P1C2+$ and $C1P2++$



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- ▶ No bias for $P1C2 = P1C2+$ and $C1P2++$
- ▶ Precision of interaction effect estimator is about 30% smaller for $P1C2$ compared to $C1P2++$



Conclusions

1. P1C2 yields more precise estimators of the interaction effect compared to the C1P2-approaches
2. In contrast to C1P2, P1C2 is not affected by misspecification or omission of upper level effects (i.e. upper level endogeneity)



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1. P1C2 yields more precise estimators of the interaction effect compared to the C1P2-approaches
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Thank you!

Possible complication when the predictors are NOT centred

- ▶ Again consider C1P2 or C1P2+:

$$E(Y_{ij} | X_{ij}, Z_{ij}, u_j) = \gamma_0 + \gamma_1 X_{ij}^c + \gamma_2 Z_{ij}^c + \gamma_3 X_{ij}^c Z_{ij}^c + b_j$$

$$E(Y_{ij} | X_{ij}, Z_{ij}, u_j) = \gamma_0 + \gamma_1 X_{ij}^c + \gamma_2 Z_{ij}^c + \gamma_3 X_{ij}^c Z_{ij}^c \\ + \gamma_4 \bar{X}_j + \gamma_5 \bar{Z}_j + \gamma_6 \bar{X}_j \bar{Z}_j + b_j$$

- ▶ Bias in main effects for these approaches:

$$E[\hat{\gamma}_1] = \beta_1 + \beta_3 E(\bar{Z}_j)$$

$$E[\hat{\gamma}_2] = \beta_2 + \beta_3 E(\bar{X}_j)$$

Results - Simulation Study

Estimator		$\hat{\gamma}_1$				$\hat{\gamma}_2$				$\hat{\gamma}_3$			
		Estimate (sd_E)	se	Coverage	Power	Estimate (sd_E)	se	Coverage	Power	Estimate (sd_E)	se	Coverage	Power
Sim 1	P1C2	0.101 (0.031)	0.028	0.92	0.93	0.150 (0.027)	0.028	0.96	1.00	-0.101 (0.019)	0.020	0.95	1.00
	C1P2	0.101 (0.032)	0.028	0.92	0.92	0.150 (0.027)	0.028	0.96	1.00	-0.101 (0.029)	0.029	0.95	0.94
	C1P2++	0.101 (0.031)	0.028	0.93	0.93	0.150 (0.027)	0.028	0.96	1.00	-0.101 (0.029)	0.029	0.95	0.94
Sim 2	P1C2	0.103 (0.056)	0.055	0.95	0.46	0.150 (0.027)	0.028	0.96	1.00	-0.101 (0.040)	0.039	0.95	0.72
	C1P2	0.103 (0.056)	0.055	0.95	0.47	0.100 (0.027)	0.028	0.96	1.00	-0.099 (0.058)	0.058	0.95	0.41
	C1P2++	0.103 (0.056)	0.056	0.95	0.46	0.150 (0.027)	0.028	0.96	1.00	-0.099 (0.058)	0.058	0.95	0.41
Sim 3	P1C2	0.103 (0.056)	0.055	0.95	0.46	0.150 (0.027)	0.028	0.96	1.00	-0.101 (0.040)	0.039	0.95	0.72
	C1P2	0.103 (0.056)	0.055	0.95	0.46	0.150 (0.027)	0.028	0.96	1.00	-0.099 (0.058)	0.058	0.95	0.41
	C1P2++	0.103 (0.056)	0.055	0.95	0.46	0.150 (0.027)	0.028	0.96	1.00	-0.099 (0.058)	0.058	0.95	0.41
Sim 4	P1C2	0.103 (0.068)	0.069	0.95	0.33	0.150 (0.027)	0.028	0.96	1.00	-0.101 (0.040)	0.039	0.95	0.72
	C1P2	0.103 (0.068)	0.069	0.95	0.33	0.150 (0.027)	0.028	0.96	1.00	-0.090 (0.055)	0.055	0.95	0.37
	C1P2++	0.103 (0.068)	0.069	0.96	0.33	0.150 (0.027)	0.028	0.96	1.00	-0.099 (0.055)	0.056	0.96	0.41