1	Distinct clusters of stunted children in India: An observational study
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3	Green MA ¹ * PhD, Corsi DJ ² PhD, Mejía-Guevara I ³ PhD, Subramanian SV ⁴ PhD.
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5	¹ Department of Geography & Planning, University of Liverpool, Liverpool, UK.
6	² Ottawa Hospital Research Institute, Ottawa, Canada.
7	³ Centre for Population Health Sciences, Stanford University, USA
8	⁴ School of Public Health, Harvard University, Boston, USA.
9	
10	* Corresponding autor. Address: Department of Geography & Planning, Roxby Building,
11	University of Liverpool, Liverpool, L69 7ZT, UK. Email: <u>mark.green@liverpool.ac.uk</u> . Tel:
12	+44 151 794 2854.
13	
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21	

22 ABSTRACT

<u>Background</u>: Childhood stunting is often conceptualised as a singular concept (i.e. stunted or
not) and such an approach implies similarity in the experiences of children who are stunted.
Furthermore, risk factors for stunting are often treated in isolation and limited research has
examined how multiple risk factors interact together. Our aim was to examine whether there
are subgroups among stunted children, and if parental characteristics influence the likelihood
of these subgroups among children.

29 <u>Methods</u>: Children who were stunted were identified from the 2005-06 Indian National

30 Family Health Survey (n=12 417). Latent class analysis was used to explore the existence of

31 subgroups among stunted children by their social, demographic and health characteristics.

32 We examined whether parental characteristics predicted the likelihood of a child belonging to

ach latent class using a multinomial logit regression model.

<u>Results</u>: We found there to be five distinct groups of stunted children; 'poor, older and poor
health-related outcomes', 'poor, young and poorest health-related outcomes', 'poor with
mixed health-related outcomes', 'wealthy and good health-related outcomes', and 'typical
traits'. Both mother and father's educational attainment, body mass index and height were
important predictors of class membership.

39 <u>Conclusions</u>: Our findings demonstrate evidence that there is heterogeneity of the risk factors

40 and behaviours among children who are stunted. It suggests that stunting is not a singular

41 concept; rather there are multiple experiences represented by our 'types' of stunting.

42 Adopting a multi-dimensional approach to conceptualising stunting may be important for

43 improving the design and targeting of interventions for managing stunting.

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45 KEYWORDS

46 Stunting; India; Latent Class Analysis; Children; Undernutrition; Socioeconomic Factors.

47 INTRODUCTION

Stunting is the chronic retardation of child growth as a result of nutritional inadequacies and 48 defined by low height for age. The World Health Organisation (WHO) estimated that in 2016 49 50 there were 155 million children aged under 5 years (23.8%) globally who were stunted (WHO, 2017). While India has seen large improvements towards tackling stunting with the 51 estimated prevalence declining from 51% of all Indian children under 5 in 2005-06 (Black et 52 al., 2008) to 33% in 2015-16 (Corsi, Meija, & Subramanian, 2016a), India still has the largest 53 global burden of childhood undernutrition. While there has been progress in tackling stunting 54 55 both internationally and in India recently, aided by investment towards achieving Millennium Development Goal 1, India is unlikely to achieve the goal set by the UN (Ministry of Women 56 and Child Development, 2014; United Nations, 2015) or future goals such as the Sustainable 57 58 Development Goals (de Onis et al., 2013).

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Stunting reflects the accumulation of micronutrient deficiencies typically as a result of 60 61 undernutrition. A lack of specific nutrients can lead to stunted children having reduced immunological capacity to fight against (infectious) disease(s) (Black et al., 2008; Schaible & 62 Kaufmann, 2007). Stunting itself can lead to longer term implications for child and 63 subsequent adult health (UNICEF, 2013). It can also produce social implications as well. For 64 65 example, children who are stunted are more likely to have poorer cognitive development 66 which can restrict their educational achievement and future employment prospects (Crookston et al., 2011; Grantham-McGregor et al., 2007; Martorell et al., 2010). The impact 67 of stunting can be intergenerational with stunted mothers more likely to have premature 68 69 children who then typically suffer from retarded growth (Grantham-McGregor et al., 2007; Subramanian, Ackerson, & Davey Smith, 2010; UNICEF, 2013). The combination of the 70

high prevalence of stunting alongside these associated health and social implications, makes
stunting an important policy consideration in India.

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74 Current epidemiological and public health studies tend to focus on single risk factors of stunting rather than explore the complex interplay between multiple factors. While there is 75 growing evidence that childhood stunting is influenced by multi-factorial drivers (Corsi, 76 Meija, & Subramanian, 2016b; Danaei et al., 2016; Fenske, Burns, Hothorn, & Rehfuess, 77 2013), there have been no studies to our knowledge that have sought to explore the existence 78 79 of heterogeneity among stunted children. Previous reviews on the effectiveness of nutritional interventions to prevent stunting have only reported limited success in reducing the 80 prevalence of stunting (Bhutta et al., 2008, 2013; Dewey & Adu-Afarwuah, 2008). One 81 82 possible explanation for this is that policy interventions are often delivered to all children 83 who are stunted together, who are effectively treated as a single homogenous entity. This approach may not be an efficient distribution of resources if individual characteristics (and 84 85 how individuals may respond to an intervention) are not similar. Failure to incorporate multidimensional explanations for stunting characteristics will also miss out on understanding the 86 wider determinants of stunting and hence limit our ability to design effective interventions. 87 88

Our study presents an alternative approach to exploring stunting in children. Using latent class analysis to explore similarities in multivariable associations across observations, we identify five 'types' of stunted children in a representative survey of India (2005-06). We also analyse the role of parental characteristics on latent class membership. Previous research has demonstrated the importance of parental characteristics such as education (Bhutta et al., 2013), body mass index (BMI) (Subramanian et al., 2010), height (Subramanian, Ackerson,

95	Smith, & John, 2009) and whether the mother was married as a child (Raj et al., 2010) on the
96	risk of their children being stunted (also see Corsi et al., 2016b; Danaei et al., 2016).
97	
98	KEY MESSAGES
99	• The determinants of stunting are numerous, complex and interacting, however current
100	research fails to consider the multidimensional nature of stunting rather treating it as a
101	singular concept.
102	• Our paper demonstrates the existence of five distinct types of stunted individuals in
103	terms of health, social and parental characteristics.
104	• Applying our multidimensional approach will help improve our understanding of the
105	condition, as well as how to design effective interventions.
106	
107	METHODS
108	Data
109	Data were taken from the 2005-06 Indian National Family Health Survey (NFHS). The
110	NFHS is the largest and most recent representative survey currently available that includes
111	objective data on child anthropometry. We focused on children aged between 6 months and 5
112	years. We selected all singleton children whose mothers (the survey does not allow for proxy
113	reporting i.e. by other caregivers) had fully completed a 24 hour dietary assessment (n =
114	32,360). Observations were excluded if children had missing data or values of height and
115	weight were implausible (defined as ± 6 standard deviations WHO growth standards; WHO,
116	2006) resulting in a sample size of 28,984. Ethical approval was not required due to the study
117	being secondary data analysis.

119	We identified whether children were stunted based on WHO guidelines. Height and weight
120	measurements were converted into age- and sex-specific z-scores based on WHO child
121	growth standards (WHO, 2006). Stunting was defined as any z-score below -2 standard
122	deviations. All children who were stunted were included in the analysis resulting in a final
123	sample size of 12,417 (43%).
124	
125	The selection of risk factors related to stunting, to be included as variables in our analysis,
126	was based on the approach taken in a previous study (Corsi et al., 2016b). Each variable
127	selected below was adapted from the UNICEF conceptual framework on the determinants of
128	child undernutrition (Bhutta et al., 2008). We included variables to measure both social and
129	health characteristics of children to account for different aspects of stunting. Most of the
130	variables we used have standard definitions and have been described in more detail elsewhere
131	(Barros et al., 2012; Corsi et al., 2016b). Variables included were:
132	• Sex (male or female)
133	• Household wealth (split into quintiles)
134	• Life-stage (categorised as: 6 to 11 months, 12 to 23 months, 24 to 35 months, 36 to 47
135	months, and 48 to 59 months)
136	• Diet diversity based on a scoring system designed by Ruel and Menon (Ruel &
137	Menon, 2002) (categorised by quintile)
138	• Child was breastfed within one hour of birth or not
139	• Child had an infectious disease in previous two weeks or not
140	• Water source was through a piped connection to the dwelling or not
141	• Stools were safely disposed in the house or not

142	•	Sanitation facility was improved (i.e. the hygienic separation of human excreta from
143		contact with individuals using facilities such as a latrine flushing to a sewer or septic
144		tank) or not

- House air quality (defined as; use of non-solid fuels, solid fuels in a separate kitchen,
 and solid fuels in a non-separate kitchen)Iodized salt used in household or not
- Child was fully vaccinated or not

• Vitamin A supplements taken or not

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- Seven variables of parental characteristics were also included as predictors of latent class 150 membership. Mother's and father's education were included separately and divided into the 151 following categories; no schooling, primary education, secondary education and post-152 secondary education. An issue with the inclusion of education is that household wealth is 153 included in the latent class input variables and education may be endogenous to household 154 wealth. While they are capturing slightly different concepts (i.e. material resources versus 155 156 parental cognition), the issue should be considered alongside the interpretation of our results. The height of the mother was also included and divided into the following categories (cm); 157 greater than or equal to 160.0, 155.0-159.9, 150.0-154.9, 145.0-149.9, and less than 145.0. 158 159 For fathers' height the categories were: greater than or equal to 170.0, 165.0-169.9, 160.0-164.9, 155.0-159.9, and less than 155.0. The height categories are different to reflect the 160 different distribution of values between males and females (i.e. men were taller on average). 161 Parental BMI was calculated separately by dividing body mass (kg) by height-squared (m^2) 162 and further split into groups based on WHO defined cut offs; underweight (<18.5), normal 163 164 (18.5-24.9) and overweight (>=25). Finally, we included whether the mother was married before the age of 18 years. 165
- 166

167 Statistical Analysis

Latent class analysis (LCA) was used to explore the existence of homogenous groups within children who are stunted. LCA is a finite mixture model which seeks to identify a latent structure within data through a probabilistic model (Collins & Lanza, 2010). The aim is to identify a categorical latent variable which is not directly measured but captured through other observed variables. Groups are identified based on the multi-dimensional distribution of variables.

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As LCA is an exploratory method; a decision must be made on the number of latent classes
that best describes the underlying structure of the data. To identify the most appropriate
number of classes we ran several models for a range of solutions between 2 and 10. We did
not consider a larger number of latent classes since we wanted the chosen model to be
parsimonious. Model fit was assessed using the adjusted Bayesian Information Criterion
(BIC), consistent Akaike Information Criterion (AIC) and G-squared statistic (Collins &
Lanza, 2010).

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One strength of LCA is that covariates can be included in the model to predict how factors are associated with class membership (Collins & Lanza, 2010). Covariates were modelled using a multinomial logit model. Mother's and father's education level, body mass index category, and height were each included as covariates, as well as whether the mother was married under the age of 18. We report odds ratios for the model and the 95% confidence intervals for these estimates.

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All analyses were undertaken using SAS v9.3 and the PROC LCA procedure (Lanza et al.,

191 2007). Sample weights were included in the analysis allowing our observations to be

representative and to account for the survey design (although PROC LCA cannot account for
the stratified multi-stage cluster design for how the data were collected, limiting the
representativeness of the analyses).

195

196 **RESULTS**

197 Table 1 presents summary statistics of the characteristics of children who were stunted and for the whole sample to help contextualise the data. There were slightly more males 198 compared to females in our sample of stunted children, although this only differed slightly 199 200 from all children. There was a greater proportion of stunted children from poorer households compared to the whole sample. Stunted children were on average older than the average for 201 202 the entire sample. Stunted children had less diverse diets and lower prevalence of some 203 health-related measures (e.g. vaccinations, safe disposal of stools) in comparison to the whole sample. The characteristics of parents also followed these patterns. 204

205

Table 1: Sample characteristics of stunted children (sample weights were applied).

		Stunted		
		children	All children	
			(percentage)	
		(percentage)		
Panel A: La	tent class analysis input variables			
Malaa		51.0	516	
Males		51.9	51.0	
Females		48.1	48.5	
Household	1 (Lowest)	34.7	28.6	
wealth	2	24.9	21.6	
weath	2	21.9	21.0	
quintile	3	19.1	18.9	

	4	14.9	17.6
	5 (Highest)	6.9	13.4
	6-11 months	8.7	14.9
	12-23 months	30.2	29.2
Life stage	24-35 months	30.7	28.1
	36-47 months	14.9	11.8
	48-59 months	15.5	9.3
Diat	1 (Lowest)	35.6	31.4
diversity	2	25.5	25.3
diversity	3	21.6	22.3
quintile	4	10.8	11.8
quintile	5 (Highest)	6.6	9.3
Breast fed within one hour of birth		28.1	23.4
Infectious disease in past two weeks		12.9	28.5
Water drink	ing source was through a piped	11.6	17.8
connection		11.0	17.0
Safe dispos	al of stools	11.6	15.5
Improved sa	anitary facility	15.5	22.4
	Non-solid fuels	11.9	18.3
Household	Solid fuels in separate kitchen	52.8	51.1
air quality	Solid fuels in non-separate	35.3	30.6
	kitchen	55.5	50.0
Iodized salt	used	41.9	46.2
Fully vaccin	nated	37.3	40.9
Vitamin A s	supplement taken	18.5	20.1

Panel B: Covariates for explaining class

<u>membership</u>

	No schooling	59.3	50.0
Mothers	Primary	14.1	14.0
education	Secondary	22.2	26.6
	Post-secondary	4.5	9.4
Mothers	Underweight	44.8	41.1
body mass	Normal	51.5	53.1
index	Overweight	3.7	5.8
	<145	15.8	12.1
Mothers	145-149.99	31.6	27.0
height	150-154.99	32.4	33.5
(cm)	155-159.99	15.8	19.8
	>=160	4.5	7.6
Mother was	married before 18	66.8	60.7
	No schooling	34.6	29.3
Fathers	Primary	16.8	15.2
education	Secondary	36.4	37.6
	Post-secondary	12.2	17.9
Fathers	Underweight	36.2	31.4
body mass	Normal	58.0	59.5
index	Overweight	5.8	9.1
Fathers	<155	9.0	7.0
height	155-159.99	20.8	17.5
(cm)	160-164.99	31.5	29.9

165-169.99	24.1	26.5
>=170	14.6	19.1

208

Figure 1 presents the analysis of the number of latent classes that best captures the underlying 209 structure of the data. While an increasing number of groups typically resulted in an improved 210 model across each measure of model fit, the rate of these improvements decreased with 211 212 increasing number of solutions. Both the consistent AIC and adjusted BIC have a clear 'knee point' at 5 classes whereby increasing number of classes adds little additional information. 213 214 The G-squared also has a kink at 5, but continues decreasing. However, the measure is less useful with large samples (Collins & Lanza, 2010). We selected a 5 class solution for our 215 216 analysis.



Figure 1: Examining model fit statistics by the number of latent classes.

219

Table 2 presents the conditional probabilities of each latent class and the latent class
prevalences (Figure 2 presents the conditional probabilities using a radial plot to aid
interpretation). We named each latent class and described their characteristics below:

223

1. Poor, older and poor health-related outcomes: While the characteristics are mostly 224 similar to 'poor, young and poorest health-related outcomes', there are some key 225 226 differences. The children mainly differ based on life stage being older. The latent class also have a higher prevalence of children fully vaccinated, fewer children 227 228 suffering from infectious diseases and a bimodal distribution for the diet variable. 2. *Poor, young and poorest health-related outcomes*: The latent class contained the 229 largest probability of the lowest two quintiles of household wealth. They are also the 230 231 youngest class compared to the other classes. Characteristics were largely the worst in comparison to the other classes. Diet diversity was low and while 'poor, older and 232 poor health-related outcomes' had the largest probability for quintile 1, the class has 233 the largest combined probability for quintiles 1 and 2. The class also had some of the 234 poorest health-related characteristics especially for hygiene and sanitation, household 235 air quality and full vaccinations. It is the largest class. 236

3. *Poor with mixed health-related outcomes*: There are higher probabilities of children in
the lower quintiles of household wealth. The class displays relatively good healthrelated characteristics, with the highest probability for vitamin A supplements, full
vaccinations and breast feeding. However, it also has the highest probability for
infectious disease and low probabilities for the hygiene, sanitation and household air
quality variables. The majority were aged between one and three years old. It is the
second smallest class as well.

244	4.	Wealthy and good health-related outcomes: Most children in the class are in the top
245		two quintiles of household wealth. They displayed the best health-related outcomes
246		characteristics in comparison to the other clusters. However, they did not perform the
247		best for every health-related outcome variable e.g. fully vaccinated or vitamin A
248		supplements. The majority were aged between one and three years old. It was the
249		smallest class.
250	5.	Typical traits: The conditional probabilities largely fall in the middle in comparison to
251		the other latent classes. There are few features that make the class distinctive other
252		than this.
253		
254	There	was little variation in proportion of males and females in each latent class suggesting
255	that ou	r latent classes are largely independent of sex.
256		



Figure 2: A radial plot of the conditional probabilities for each latent class.

		Class 1	Class 2	Class 3	Class 4	Class 5
Prevalence	(Y)	0.190	0.349	0.103	0.107	0.251
Conditional	Probabilities (ρ)					
Males		0.485	0.527	0.545	0.576	0.501
Females		0.515	0.473	0.455	0.424	0.499
	1 (Lowest)	0.554	0.591	0.340	0.000	0.001
Household	2	0.321	0.309	0.331	0.001	0.167
wealth	3	0.106	0.092	0.229	0.040	0.442
quintile	4	0.019	0.009	0.099	0.420	0.347
	5 (Highest)	0.001	0.000	0.001	0.539	0.043
	6-11 months	0.000	0.139	0.065	0.081	0.091
	12-23 months	0.002	0.420	0.419	0.368	0.291
Life stage	24-35 months	0.237	0.296	0.385	0.355	0.322
	36-47 months	0.194	0.146	0.094	0.095	0.166
	48-59 months	0.567	0.000	0.036	0.101	0.131
	1 (Lowest)	0.533	0.372	0.261	0.186	0.310
Diet	2	0.000	0.396	0.270	0.223	0.259
diversity	3	0.255	0.160	0.237	0.264	0.235
score	4	0.145	0.057	0.129	0.155	0.120
quintile	5 (Highest)	0.067	0.015	0.103	0.171	0.076
Breast fed within one hour of birth		0.132	0.118	0.362	0.348	0.262
Infectious d	isease in past two weeks	0.143	0.333	0.360	0.287	0.275

Table 2: The characteristics of a 5 group latent class analysis of stunted children.

connection	0.026	0.019	0.057	0.583	0.237	
Safe disposa	0.056	0.028	0.070	0.551	0.117	
Improved sa	unitary facility	0.036	0.023	0.085	0.696	0.227
	Non-solid fuels	0.005	0.003	0.000	0.735	0.154
Household	Solid fuels in separate kitchen	0.544	0.497	0.713	0.247	0.604
air quality Solid fuels in non-separate						
	kitchen	0.451	0.500	0.286	0.018	0.242
Iodized salt used		0.312	0.300	0.478	0.785	0.486
Fully vaccinated		0.275	0.139	0.882	0.682	0.432
Vitamin A supplement taken		0.077	0.118	0.690	0.283	0.131

Water drinking source was through a piped

Table 3 presents the results exploring the association between parental characteristics and 262 latent class membership. The group 'typical traits' (class 5) were selected as the comparator 263 class for interpreting the estimates. Mothers who were married before the age of 18, lower 264 educational attainment and underweight parents were positively associated with membership 265 of classes 1 to 3 compared to 'typical traits'. Relationships were fairly consistent between 266 mothers and fathers, other than for 'poor with mixed health-related outcomes' (class 3) where 267 mother's education was not important. The direction of these associations to 'wealthy and 268 good health-related outcomes' (class 4) were opposite (although there was little association 269 for underweight fathers). 270

271

Height displayed less certainty in the direction of estimates for both mothers and fathers with

273 most confidence intervals crossing a value of 1. While taller mothers were positively

associated with class membership of 'wealthy and good health-related outcomes' (class 4),

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the result for father's height was contrary. Taller fathers were negatively associated with
membership of 'poor, young and poorest health-related outcomes' (class 2) compared to
'typical traits' (class 5).

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Table 3: A multinomial logit model of the association between characteristics of parents

and class membership. Note: We present odds ratios with 95% confidence intervals in

281 brackets.

						Class 5
		Class 1	Class 2	Class 3	Class 4	(Reference
	No	2.99 (2.19-	3.02 (2.27-	0.67 (0.37-	0.56 (0.33-	
	schooling	4.09)	4.03)	1.22)	0.95)	
M = (1	Primary			Reference		
Mothers	Secondary	0.35 (0.23-	0.32 (0.22-	0.72 (0.36-	3.12 (2.13-	
education		0.54)	0.48)	1.46)	4.56)	
	Post-	0.17 (0.03-	0.11 (0.03-	0.44 (0.08-	10.29 (6.33-	
	secondary	0.87)	0.47)	2.42)	16.71)	
Mothers	Underweight	1.32 (1.08-	1.59 (1.34-	1.75 (1.23-	0.73 (0.58-	
hada	Underweight	1.61)	1.90)	2.50)	0.93)	
body	Normal			Reference		
mass	0	0.53 (0.28-	0.38 (0.21-	0.46 (0.12-	3.40 (2.30-	
index	Overweight	1.01)	0.69)	0.169)	5.02)	
	-1.45	1.07 (0.78-	1.25 (0.94-	1.16 (0.70-	1.41 (0.98-	
	<143	1.48)	1.66)	1.91)	2.02)	

	145-149.99	1.05 (0.83-	1.02 (0.83-	1.33 (0.91-	0.83 (0.63-	
		1.34)	1.27)	1.94)	1.09)	
Mothers	150-154.99			Reference		
height	155-159.99	0.98 (0.73-	0.84 (0.65-	0.82 (0.53-	1.60 (1.19-	
(cm)		1.31)	1.09)	1.25)	2.33)	
	>=160	0.70 (0.45-	0.80 (0.54-	0.78 (0.38-	1.50 (0.96-	
		1.08)	1.19)	1.58)	1.33)	
Mother was married		1.56 (1.25-	1.72 (1.40-	1.16 (0.83-	0.52 (0.42-	
before 18		1.95)	2.11)	1.63)	0.65)	
Fathers education	No	3.02 (2.26-	3.14 (2.41-	1.31 (0.77-	0.51 (0.27-	
	schooling	4.03)	4.11)	2.22)	0.98)	
	Primary	Reference				
	Secondary	0.69 (0.52-	0.68 (0.53-	0.48 (0.33-	1.82 (1.23-	
		0.91)	0.86)	0.70)	2.71)	
	Post-	0.58 (0.35-	0.44 (0.27-	0.34 (0.17-	2.32 (1.46-	
	secondary	0.94)	0.70)	0.70)	3.68)	
Fathers body mass index	Underweight	1.84 (1.36-	1.62 (1.21-	1.47 (0.88-	0.96 (0.65-	
		2.49)	2.16)	2.47)	1.41)	
	Normal	Reference				
	Overweight	0.30 (0.11-	0.30 (0.12-	0.25 (0.04-	4.72 (2.89-	
		0.83)	2.16)	1.50)	7.70)	
Fathers height (cm)	<155	1.25 (0.73-	1.19 (0.74-	1.17 (0.56-	0.18 (0.06-	
		2.15)	1.92)	2.47)	0.54)	
	155-159.99	0.90 (0.63-	0.91 (0.67-	0.64 (0.36-	0.67 (0.44-	
		1.28)	1.24)	1.14)	1.03)	

160-164.99		Reference			
165 160 00	0.83 (0.60-	0.66 (0.48-	0.70 (0.43-	1.03 (0.74-	
105-109.99	1.14)	0.89)	1.14)	1.43)	
> -170	0.65 (0.42-	0.61 (0.41-	0.55 (0.26-	0.65 (0.44-	
>=170	1.14)	0.93)	1.19)	0.98)	

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284

285 **DISCUSSION**

Our study revealed five distinct 'types' of stunted children in India demonstrating clear 286 heterogeneity of stunting. An important contribution of our study is the examination of 287 288 multiple risk factors and determinants within a mutually adjusted casual framework. A major limitation of many high profile studies in this field is the continued examination of single risk 289 factors in isolation leading to biased effect estimates (Danaei et al., 2016). The high 290 prevalence of stunting in India (and globally), combined with the associated health and social 291 implications of stunting (Black et al., 2008; Schaible & Kaufmann, 2007; UNICEF, 2013), 292 has seen considerable interest in how to effectively design interventions to tackle stunting 293 (Bhutta et al., 2008, 2013). Treating childhood stunting as a singular concept may lead to a 294 false dichotomy of the determinants and experiences of stunting, limiting our understanding 295 296 of the issue and our ability to tackle it.

297

The multidimensional nature of childhood stunting identified in our study has important implications for future policy. The distinct characteristics of each latent class suggest differing strategies are required to tackle the issue. Ignoring the complex combination of characteristics that constitute each class may restrict the effectiveness of policies or lead to inefficient targeting of resources (Dewey & Adu-Afarwuah, 2008; Fenske et al., 2013). For example, improving vaccination uptake among stunted children to protect against infectious
diseases would appear important given that only 37.3% of stunted children were fully
vaccinated. However, the policy would be less appropriate for 'poor with mixed healthrelated outcomes' (class 3) who had a high prevalence of vaccination, where it may be better
to focus on other issues such as improving access to clean water. Future research should
explore the application of our latent classes within different interventions to assess how
useful they are for delivering policy more efficiently.

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311 The discovery of a 'poor but mixed health-related outcomes' cluster was a particularly important finding since it demonstrates the possibility to achieve good health-related 312 outcomes despite the imposing forces of a lack of wealth. Children in the class had the 313 314 highest prevalence of children fully vaccinated and vitamin A supplements taken, but low prevalence of improved drinking water source or safe sanitation. The finding may suggest 315 that it is useful when implementing interventions aimed at improving health-related outcomes 316 in poor and stunted children to improve education and awareness about access to health 317 services as these represent potential successes. Wealth and neighbourhood conditions are less 318 modifiable, and therefore harder to address (Bhutta et al., 2008; Poel, Hosseinpoor, 319 Speybroeck, Ourti, & Vega, 2008). However, the finding may simply relate to the successful 320 delivery of programmes targeting low socioeconomic status areas or regions (e.g. to increase 321 322 vaccinations).

323

The associations found with the covariates (as well as the relative importance of household wealth in determining the latent classes) are indicative of the importance of the social gradient in understanding latent classes. Whilst most stunting related interventions are aimed at improving nutrition (Bhutta et al., 2013), it is clear that for any intervention to be

328 successful they must be combined with the wider social context (Corsi et al., 2016b; Dewey & Adu-Afarwuah, 2008). Tackling the social gradient will need to be more targeted than 329 simply encouraging economic growth, since economic growth alone is not associated with 330 331 improved child nutrition (Subramanyam, Kawachi, Berkman, & Subramanian, 2011). While wealthy children have been previously shown to be less likely to be stunted (Corsi et al., 332 2016b; Fenske et al., 2013; Poel et al., 2008), our findings indicate that they make up their 333 334 own latent class with more favourable health-related outcomes compared to the other groups. They were, however, a small class but nonetheless it underscores that certain "better off" 335 336 children may be at risk for stunting despite relatively higher socioeconomic status.

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Parental BMI was associated with class membership, although more consistently for mothers 338 339 compared to fathers. This association may partly reflect the social gradient as well, since 340 poorer individuals have been previously demonstrated to be more likely to be underweight (Bhutta et al., 2008; Subramanian, Corsi, Subramanyam, & Smith, 2013). However, there 341 342 may also be an independent association. It is plausible that underweight parents pass on similar habits to the children. Previous research has shown that stunting displays an 343 intergenerational aspect (Grantham-McGregor et al., 2007; Subramanian et al., 2009; 344 UNICEF, 2013). A similar interpretation may be derived for height as well although the 345 346 results were inconsistent. Our finding that fathers characteristics explain the type of stunting 347 supports calls to move away from a 'maternal' to a 'household' understanding of the determinants of stunting (Corsi et al., 2016a). 348

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There are several limitations to our approach. The data were cross-sectional and therefore our ability to draw inferences about our groupings is limited. It will be important to replicate the approach using longitudinal observations to explore how the groups develop and change over

353 time. We only included seven predictors of class membership and it will be important to examine the association of additional factors that have been shown to be related to stunting 354 such as the socio-economic context of local regions (Subramanyam et al., 2011). LCA can 355 356 only make observations about how factors are co-associated within stunted children, and it does not show the strength of an association to the risk of stunting. The data we used were 357 collected in 2005-06 and are therefore outdated. They are, however, the most recent data 358 available at the time of analysis. Irrespective of the date, it does not change the notion that the 359 strength of the paper is its conceptual approach of exploring a typology of stunting. The 360 361 NFHS is currently processing the data collected in the fourth survey wave (2015-16) and it will be important to update our study using these newer data when they are released to 362 examine how the situation has changed. The risk factor data are self-reported and therefore 363 364 potentially subject to bias. There were no alternative large representative data sets which 365 included objective measures limiting the quality of our observations. The data used for the characteristics of fathers correspond to the mother's partner at the time of the survey. While 366 367 the majority of partners are the biological father of the child, it may introduce a small amount of bias in estimates. However, this should not distract from the importance of how adults 368 within the household may influence experiences. 369

370

In conclusion, the causes of childhood stunting are complex and multidimensional. Our paper
contributes to a literature that has largely examined stunting as a singular concept,
demonstrating heterogeneity among stunted children. We hope that the approach outlined in
this paper will help policy makers in designing effective interventions as opposed to more

375 simplistic approaches that do not differentiate in terms of the individuals they target.

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