The Principal Independent Components of Images

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Abstract

This paper proposes a new approach for the encoding of images by only a few important components. Classically, this is done by the Principal Component Analysis (PCA). Recently, the Independent Component Analysis (ICA) has found strong interest in the neural network community. Applied to images, we aim for the most important source patterns with the highest occurrence probability or highest information called *principal independent components* (PIC).

For the example of a synthetic image composed by characters this idea selects the salient ones. For natural images it does not lead to an acceptable reproduction error since no a-priori probabilities can be computed. Combining the traditional principal component criteria of PCA with the independence property of ICA we obtain a better encoding. It turns out that this definition of PIC implements the classical demand of Shannon's rate distortion theory.

1 Introduction

Classically, the encoding of images by only a few important components is done by Principal Component Analysis (PCA). One common solution is to cut the image into smaller patches or "subimages" which are transformed linearly by projecting them on the eigenvectors of their associated covariance matrix. It is well known that the transformed components with the highest variance (the *principal components*) yield an optimal reconstruction of the original subimages in the mean square error sense. However, for the criterion of minimal redundancy encoding, the PCA is suboptimal.

Recently, the Independent Component Analysis (ICA) has become subject to many research activities and several algorithms have been proposed by different authors, e.g. [1, 2, 3]. Here, the goal is to obtain linearly transformed components which are as independent as possible (the *independent components*). This corresponds to the minimisation of the mutual information between the transformed components and therefore reduces the overall encoding amount [1, 2].

Applied to image encoding, the ICA approach assumes that each observed signal vector $\mathbf{x} = (x_1,\ldots,x_n)^T$ (an image containing n pixels) is a linear mixture $\mathbf{x} = \mathbf{M}\mathbf{s}$ of n unknown independent source signals $\mathbf{s} = (s_1,\ldots,s_n)^T$. The unknown mixing matrix $\mathbf{M}$ must be non-singular; its columns can be viewed as "image primitives". To recover the sources signals, one has to determine a demixing matrix $\mathbf{B}$ with $\mathbf{s} = \mathbf{B}\mathbf{x}$.

There are several conditions involved in the demixing process [1]: in general, the recovered source signals (denoted by $\mathbf{y} = (y_1,\ldots,y_n)^T$ for clarity) are scaled and per-

muted versions of the original sources. Furthermore, at most one of the source signals **s** should have a Gaussian probability distribution or else the separation will become ambiguous. This is why the recovered sources **y** are conventionally assumed to be non-Gaussian random variables having unit variance.

As proposed in [1, 3] the determination of **B** reduces to the computation of an orthogonal matrix $\mathbf{W}_{ICA}$ if the observed signals **x** are prewhitened. This can be done by a simple PCA transform of the image vectors and scaling the obtained PCA components to unit variance. The corresponding prewhitening (or *sphering*) transform is denoted by the matrix $\mathbf{W}_{PCA}$.

Together with the convenient assumption that the recovered source signals are centered, i.e. $\langle \mathbf{y} \rangle \equiv \mathbf{0}$, we have the following ICA relation

$$\mathbf{y} = \mathbf{W}_{ICA}\mathbf{W}_{PCA}(\mathbf{x} - \langle \mathbf{x} \rangle) = \mathbf{B}\,(\mathbf{x} - \langle \mathbf{x} \rangle) = \mathbf{BM}\,(\mathbf{s} - \langle \mathbf{s} \rangle) = \mathbf{DP}\,(\mathbf{s} - \langle \mathbf{s} \rangle) \qquad (1)$$

where **D** is an unknown diagonal matrix and **P** an also unknown permutation matrix.

In this model the number of independent sources is assumed to be equal to the number of image pixels. Nevertheless, we expect that for a good representation covering most of the input data some of the sources are less important than others. Thus we aim for an ordering criterion which prefers the essential source signals called *principal independent components* (PIC).

2   An event-oriented image model

Due to the intuitive notion of "importance" we propose that principal independent components should have a high occurrence probability. Therefore, we consider images to be composed of the superposition of many small, independent image primitives, just like a single neuron of the retina sees the world by a limited focus, which appear with a certain probability. As a further restriction, we assume that only one of two possible states is assigned to each primitive: present in the superposition or not. This leads to the formulation of *image events* ω_i (denoting the presence of primitive i) and $\neg\omega_i$ (denoting its absence). The task consists now of determining the most important events, i.e. those with highest probability $P(\omega_i)$.

Applied to eq. (1), the image primitives are represented by the columns of the mixing matrix **M**, and the source signals s_i encode the associated image events by

$$s_i = \begin{cases} 1 & \text{for } \omega_i \quad (\text{primitive } i \text{ is present}) \\ 0 & \text{for } \neg\omega_i \quad (\text{primitive } i \text{ is not present}) \end{cases}$$

Thus, the average $\langle s_i \rangle \equiv \bar{s}_i$ of a source signal s_i and its variance σ_{is}^2 are given by

$$\bar{s}_i \equiv \langle s_i \rangle = P(s_i{=}1)\cdot 1 + P(s_i{=}0)\cdot 0 = P(s_i{=}1) = P(\omega_i) \qquad (2)$$

$$\sigma_{is}^2 = \langle s_i^2 \rangle - \bar{s}_i^{\,2} = P(s_i{=}1)\cdot 1^2 + P(s_i{=}0)\cdot 0^2 - \bar{s}_i^{\,2} = \bar{s}_i - \bar{s}_i^{\,2} = \bar{s}_i\,(1 - \bar{s}_i) \qquad (3)$$

Suppose that we have already computed the demixing matrix **B** in eq. (1). The recovered source signals y_i are scaled and permuted versions of the centered original sources s_i. Because the permutation **P** is unknown (and, in fact, of no interest) we assume $\mathbf{P} \equiv \mathbf{I}$ and concentrate on the non-zero scaling factors a_i satisfying

$$y_i = a_i (s_i - \bar{s}_i) \tag{4}$$

Since the recovered sources have zero mean and unit variance $\sigma_{iy}{}^2$ the following relation holds:

$$1 = \sigma_{iy}{}^2 = \langle y_i^2 \rangle = \langle (a_i (s_i - \bar{s}_i))^2 \rangle = a_i^2 (\langle s_i^2 \rangle - \bar{s}_i{}^2) = a_i^2 \sigma_{is}{}^2 = a_i^2 \bar{s}_i (1 - \bar{s}_i) \tag{5}$$

Now, if we ignore the centering terms in eq. (1), we can express the transformation of the source average $\langle \mathbf{s} \rangle$ to the observed average $\langle \mathbf{x} \rangle$ and to the recovered source average $\langle \mathbf{y} \rangle$ by

$$\langle \mathbf{x} \rangle = \mathbf{M} \langle \mathbf{s} \rangle \qquad \text{and} \qquad \langle \mathbf{y} \rangle = \mathbf{B} \langle \mathbf{x} \rangle = \mathbf{BM} \langle \mathbf{s} \rangle \tag{6}$$

Note that here $\langle \mathbf{y} \rangle$ is obviously non-zero unless for all i the probabilities $P(\omega_i)$ are zero. With eqs. (4), (6) we have

$$\langle y_i \rangle = a_i \bar{s}_i \tag{7}$$

Combining eqs. (5), (7) gives the desired relation for the occurrence probabilities

$$1 = (\langle y_i \rangle / \bar{s}_i)^2 \, \bar{s}_i (1 - \bar{s}_i) \qquad \text{or} \qquad P(\omega_i) = \bar{s}_i = \langle y_i \rangle^2 / (1 + \langle y_i \rangle^2) \tag{8}$$

By this we obtained a measure to order the observed ICA components according to their decreasing occurrence probabilities, i.e. $i \geq j \Leftrightarrow P(\omega_i) \geq P(\omega_j)$.

Furthermore, if $P(\omega_i) \leq 0.5$ holds for all i, the components y_i are ordered by their decreasing marginal entropy $H(y_i)$, because $H(y_i)$ is a convex function of the probability $P(\omega_i)$ and monotonically increasing up to its local maximum (located at $P(\omega_j) = 0.5$) [4].

3 Recovering the occurrence probabilities of events

To validate the theoretical results of the previous section, we computed a synthetic image according to the model in eq. (1). As image primitives we chose 16 pictures of 8×8 pixels visualising the letters 'A'…'P'. From these, 4096 different random linear mixtures were calculated and used as training samples. After prewhitening with the transform $\mathbf{W}_{PCA}$ we presented the samples to a hierarchical ICA network similar to the one proposed in [3] with *tanh* non-linearities. The image primitives along with the eigenimages and the recovered primitives are shown in Figure 1a-c.

For the whitened PCA components we observed near-Gaussian distributions (Figure 1d) while the distributions of the ICA components are slightly "blurred" versions of the original occurrence probabilities, see Figure 1e.

The initial and the estimated occurrence probabilities of the first four sources are listed in Table 1 (the error is due to the imperfectly learned demixing matrix $\mathbf{B}$). Also shown are their observed and their original marginal entropy (computed on 8 bit coefficients) compared to the marginal entropy of the first four whitened PCA components. Obviously, the single source information is reduced dramatically. Because of the "blurred" probability distributions, the marginal entropy of the recovered sources is still higher than the original entropy. However, by applying a rigorous quantization strategy we should be able to achieve further reduction [4].
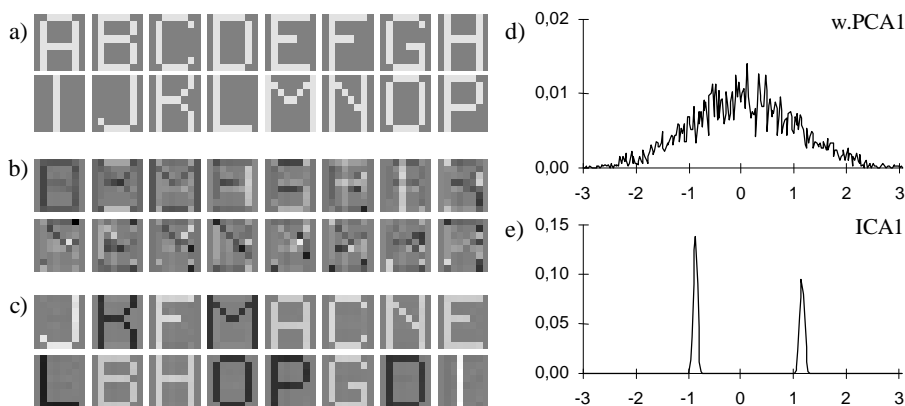
Figure 1: a) The image primitives, b) the eigenimages, and c) the recovered image primitives of the synthetic image. The probability distributions of the first whitened PCA component and of the first ICA component are shown in d) and e) respectively. To obtain the histograms the 4096 samples were quantified into 256 intervals on the horizontal axis.

source	probability initial	probability estim.	error	compo-nent	observed entropy	compo-nent	observed entropy	original entropy
'J'	0.444	0.463	-0.019	w.PCA1	7.398	ICA1 'J'	3.800	0.991
'K'	0.415	0.322	0.092	w.PCA2	7.408	ICA2 'K'	4.555	0.980
'F'	0.696	0.732	-0.036	w.PCA3	7.322	ICA3 'F'	4.745	0.886
'M'	0.624	0.618	0.006	w.PCA4	7.405	ICA4 'M'	4.164	0.955

Table 1: Four of the source letters, their associated initial and estimated occurrence probabilities. Also shown are the observed and original marginal entropy of the four recovered sources and the first four whitened PCA components (in *bits*).

4 Independent components of natural images

Since the initial goal of our examinations is the efficient encoding of images with only a few important components we searched for the PIC of natural images. In our simulations a picture called *Cactus* was divided into 4543 subimages (size: 8×8=64 pixels) which were randomly chosen as training samples [4]. After centering and prewhitening of the samples we determined the matrix **B**. The corresponding image primitives were very similar to those already known in the literature, see e.g. [5].

Here, the measured probability distributions of the sources were not bimodal. This excluded the event model of section 2 for calculating the occurrence probabilities and therefore prevented an order of the sources by most probable image events. Instead we calculated the marginal entropy of the recovered sources as a new ordering criterion which is closely related to the probability ordering (see section 2).

We found that especially all the ICA components had nearly the same information; there were no components which differed much from the others. Furthermore, the marginal entropy of the ICA components was just slightly smaller than the one of the whitened PCA components.
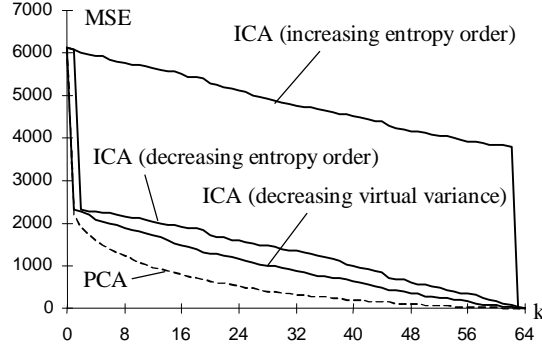
Figure 2: Decreasing the MSE by adding components.

Another measure for "importance" is the quality of the image restoration. Reconstructing the image by its first k components and comparing it with the original one gives the average error for neglecting the $n–k$ components. Therefore we compared the optimal MSE (mean square error) contribution of the PCA components (ordered by decreasing variance) to those of the ICA components (ordered by increasing and decreasing entropy). For the latter we defined a third ordering criterion called the *virtual variance*

$$\mathrm{var}^*(y_i) \equiv \mathrm{var}\left(\frac{\mathbf{b}_i}{\|\mathbf{b}_i\|}\left(\mathbf{x} - \langle\mathbf{x}\rangle\right)\right) = \frac{\mathrm{var}(y_i)}{\|\mathbf{b}_i\|^2} = \frac{1}{\|\mathbf{b}_i\|^2} \tag{9}$$

which considers the fact that the norm of a row $\mathbf{b}_i$ of the matrix $\mathbf{B}$ is in general not equal to unity. Consequently, an ICA component with higher virtual variance is assumed to be more important. Figure 2 shows the obtained error functions. In case of the ICA, ordering the components by their decreasing virtual variance gives the best results. However, our simulations showed that the subjective quality of image restoration by a few ICA components is not acceptable.

5   PIC and rate distortion theory

When the number of components in the transform approach for encoding images is reduced, the full space of image components (dimensions) is reduced to a subspace. The subspace of the ICA components is characterised by its information content whereas the subspace of the PCA components is characterised by its low MSE reconstruction error. Since the principal components of PCA cannot be replaced for obtaining a small MSE, their encoding information should be reduced by ICA. This idea can be performed in two ways:

1. Get the first k PCA components with an acceptable MSE. Then, by an ICA transform, we will get the same number of encoding coefficients but with less information, i.e. less encoding bits.

2. For the same amount of encoding information as the k PCA components take, we can also get p more ICA transformed PCA components. Since these $k+p$ base vectors of the ICA transform span the same space as the $k+p$ PCA components, the resulting image quality will be enhanced as if p more PCA components were added.

Thus the approach starting with the search for principal independent components leads to the error-bounded maximal information for each channel. This is classically known as the *rate distortion theory* [6] and has a broad range of applications in the telecommunication area.

The first one of the ideas above can be implemented if we order the k ICA components according to their decreasing virtual variance and encode only the first $k' < k$ components with low additional reconstruction error. This results in a further reduction of the number of encoding bits. To validate the latter idea we computed the ICA components of the first k PCA components for $k = 16,\ldots,21$. We found that for the same information rate about one additional ICA component can be encoded with an error reduction of 5%.

Finally, we examined the influence of quantization on the MSE and the overall information rate. For $k=16$ and $k=20$ the resolution of the PCA components and their associated ICA components was set to 5, 6, 7 and 8 bit. Lowering the resolution down to 6 bit did not increase the MSE significantly whereas the information rate decreased by about 43% (!). Again, the information gain of the ICA over the PCA was about one additional component.

As a remarkable result we observed that for $k=20$ and a resolution of 6 bit both the resulting MSE and the encoding amount of the ICA components were superior to the corresponding representation with $k=16$ components quantified to 8 bit (MSE $\approx$ 13%, information rate $\approx$ 54%). A systematic investigation of this behaviour is subject to future research.

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