1 **TITLE**

2 A multilevel society of herring-eating killer whales indicates adaptation to prey characteristics

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13 RUNNING TITLE

14 Multilevel society of herring-feeding killer whales

15

16 ABSTRACT

- 17 Non-social factors can influence animal social structure. In killer whales (Orcinus orca), fish-
- vs. mammal-eating ecological differences are regarded as key ecological drivers of their
- 19 multilevel society, including group size, but the potential importance of specific target prey

20 remains unclear. Here, we investigate the social structure of herring-eating killer whales in 21 Iceland and compare it to the described social structures of primarily salmon- and seal-eating populations in the Northeast Pacific, which form stable coherent basic units nested within a 22 hierarchical multilevel society. Using 29023 photographs collected over 6 years, we examined 23 24 the association patterns of 198 individuals combining clustering, social network structure and temporal patterns of association analysis. The Icelandic population had largely weak but non-25 26 random associations, which were not completely assorted by known ranging patterns. A fission-fusion dynamic of constant and temporary associations was observed but this was not 27 due to permanent units joining. The population-level society was significantly structured but 28 not in a clear hierarchical tier system. Social clusters were highly diverse in complexity and 29 there were indications of subsclusters. There was no indication of dispersal nor strong sex 30 differences in associations. These results indicate that the Icelandic herring-eating killer whale 31 32 population has a multilevel social structure without clear hierarchical tiers or nested coherent social units, different from other populations of killer whales. We suggest that local ecological 33 context, such as the characteristics of the specific target prey (e.g. predictability, biomass and 34 35 density) and subsequent foraging strategies may strongly influence killer whale social association patterns. 36

Key words: ecological context, hierarchical structure, multilevel societies, social structure,killer whale, orca

39 INTRODUCTION

40 The sociality of a group-living species is driven by a trade-off between its specific ecological, evolutionary and social contexts (Krause and Ruxton 2002). Non-social factors, particularly 41 predation risk, finding/catching food, defending resources and resource patchiness, can 42 43 strongly determine the social structure of simple social systems and provide the context for 44 the development of complex ones (Jarman 1974; Wrangham 1980; Whitehead 2008a). General socioecological frameworks have been developed for various taxa, characterizing 45 how such factors can affect sociality by using broad characteristics of a species/genera, such 46 47 as occurrence of group foraging, group size or mating system (Emlen and Oring 1977; 48 Wrangham 1980; Gowans et al. 2007). However, with the increase of within-species studies 49 (e.g. Barton et al. 1996, Sinha et al. 2005; Whitehead et al. 2012), it seems clear that it is important to emphasize intraspecific variation which likely reflects variability under different 50 ecological conditions. Investigating different populations of the same species across 51 ecological gradients is therefore valuable to evaluate the influence of ecological drivers. 52 53 Multilevel societies are among the social systems found on group-living species and have 54 been described as hierarchical structures of nested social levels (i.e. discrete social 55 stratification of associations among individuals into tiers) with at least one stable core unit 56 (Wittemyer et al. 2005; Grueter, Matsuda, et al. 2012; Grueter, Chapais, et al. 2012). Recently, de Silva and Wittemyer (2012) suggested that multilevel societies should be seen 57 along a continuum of nestedness and that some might present less clearly hierarchically 58 59 stratified social levels that transition more gradually. Commonly, multilevel societies exhibit fission-fusion dynamics, with frequent association, disassociation, and reassociation of 60

| 61 | groups of individuals (e.g. Connor et al. 1992). Although multilevel societies have been |
|----|---|
| 62 | studied more extensively in terrestrial mammals, particularly in primates (see Grueter, |
| 63 | Chapais, et al. 2012), such social systems are also observed in cetaceans and intraspecific |
| 64 | variation has been reported (Connor et al. 1998; Whitehead et al. 2012). For example, female |
| 65 | sperm whales (Physeter macrocephalus) form long-term stable social units which, in the |
| 66 | Pacific, temporarily group with other units with which they share part of the acoustic |
| 67 | repertoire, but rarely group in the North Atlantic, possibly due to differences in predation risk |
| 68 | (Whitehead et al. 2012). |
| 69 | One well-described tiered multilevel society among cetaceans is that of the 'resident' fish- |
| 70 | eating killer whale (Orcinus orca) population in the Northeast Pacific, hereafter termed |
| 71 | residents. The basic unit of this society is the matriline, consisting of an oldest surviving |
| 72 | female and her philopatric descendants, remaining associated with their mother for life (Bigg |
| 73 | et al. 1990; Baird and Whitehead 2000; Barrett-Lennard 2000). Within matrilineal units, |
| 74 | individuals associate strongly and at very similar levels, while matrilineal units can frequently |
| 75 | interact (Bigg et al. 1990; Baird and Whitehead 2000; Ford et al. 2000). Matrilines that share |
| 76 | at least part of their acoustic repertoire, probably due to common maternal ancestry, form the |
| 77 | next social level, the clan (Ford 1991). Different clans have no calls in common, and |
| 78 | matrilines from the same or different clans frequently travel together (Ford 1991). The next |
| 79 | and broadest social level (just under population) is the community, consisting of matrilines |
| 80 | that share a common area and associate periodically but not with those of another community |
| 81 | (Bigg et al. 1990). This multilevel society is based on distinct fission-fusion patterns of whole |
| 82 | coherent family-based units, where stable matrilineal units collectively associate more |

| 83 | frequently with other close kin units. The 'sub-pod' and 'pod' were traditionally considered |
|-----|---|
| 84 | intermediate social levels between the matriline and the clan, consisting of matrilines with |
| 85 | recent maternal ancestry that often (>95% and 50% of the time, respectively) travelled |
| 86 | together (Bigg et al. 1990; Ford 1991). However, recent studies have shown fluctuations in |
| 87 | the reoccurrence of associations between matrilines (Ford and Ellis 2002; Parsons et al. 2009), |
| 88 | as well as changes in the pods originally described (Ford et al. 2000), leading to suggestions |
| 89 | that the term 'pod' should only be used to designate aggregations of killer whales or as a |
| 90 | synonym for matriline (Ford and Ellis 2002). |
| 91 | Intraspecific variation in sociality among killer whales is believed to relate to prey-type. |
| 92 | Northeast Pacific resident killer whales mainly prey on salmon, especially Chinook |
| 93 | (Oncorhynchus tshawytscha) while mammal-eating killer whales (also referred to as |
| 94 | 'transients' or Bigg's killer whales) feed on marine mammals, especially harbour seals (Phoca |
| 95 | vitulina; Ford et al. 1998). Although sympatric, these two populations comprise two specialist |
| 96 | ecotypes that are socially and reproductively segregated (Bigg 1982; Barrett-Lennard 2000). |
| 97 | Both ecotypes exhibit coherent and stable matrilineal social units based on long-term kinship |
| 98 | associations but there are important distinctions between their social strategies. The resident |
| 99 | population forms larger matrilineal units than the mammal-eating population and while the |
| 100 | resident population is philopatric, there is some level of adult dispersal in the mammal-eating |
| 101 | population (Bigg et al. 1990; Baird and Whitehead 2000). For example, males may disperse to |
| 102 | briefly associate with other matrilines or live alone, randomly associating with other adult |
| 103 | males. Moreover, some females may disperse from the matriline and stay socially mobile, |
| 104 | associating strongly for short periods with different groups (Baird and Dill 1996; Baird and |

Whitehead 2000). This variation is considered to be due to the different foraging strategies of 105 the populations. Hunting marine mammal prey in large groups incurs greater costs by 106 increasing the probability of detection by the prey. Furthermore, the optimal energetic intake 107 for mammal-eating killer whales (preying upon medium-sized seals) declines for groups 108 109 larger than 3 individuals (Baird and Dill 1996). In contrast, resident killer whales spread out 110 and coordinate to locate salmon (Ford et al. 2000), potentially benefiting from larger group 111 sizes. With little or no predation risk, populations of this species apparently refine their social systems primarily in relation to foraging efficiency, particularly availability of resources and 112 113 competition for those resources.

114 In the North Atlantic, the only published study addressing sociality found greater similarities 115 between the Scottish mammal-eating population and Northeast Pacific mammal-eating population relative to residents, despite greater phylogenetic distance, suggesting that ecology 116 117 drives sociality more than phylogenetic inertia does (Beck et al. 2012). The study included a limited dataset from Icelandic herring-eating killer whales and their social structure was not 118 explored in detail. However, the study's hierarchical display of associations suggested that 119 120 social tiers were not clearly defined in this population and that associations at a variety of strengths existed. These features were not further addressed, nevertheless the study concluded 121 that the Icelandic fish-eating population is probably more similar to the Northeast Pacific 122 123 resident population than to mammal-eating populations.

Icelandic killer whales are believed to mainly prey upon Atlantic herring (*Clupea harengus*)
and follow the Icelandic summer-spawning (ISS) herring stock during its yearly migration

126 (Sigurjónsson et al. 1988) between overwintering, feeding and spawning grounds (Óskarsson

et al. 2009). Unlike the salmon prev of resident killer whales, herring form large and dense 127 128 schools as an antipredator strategy (Nøttestad and Axelsen 1999) and killer whales feeding on 129 herring schools use a coordinated group feeding strategy, encircling their prey to herd and capture it (Similä and Ugarte 1993). Feeding aggregations of killer whales are very common 130 131 in Iceland, making it difficult to discern isolated groups and confusing the determination of associations in the field (Sigurjónsson et al. 1988; Beck et al. 2012). In addition, herring can 132 133 undergo large variations in abundance and migration routes (Jakobsson and Stefánsson 1999; Óskarsson et al. 2009) making it a changeable food resource. In fact, recent research suggests 134 135 not all individuals specialize on ISS herring and follow it year-round. Other killer whales 136 observed only in one season or seasonally moving between Iceland and Scotland exhibited wider trophic niche width, suggesting diversity in foraging strategies (Samarra and Foote 137 2015; Samarra et al., in press; FIP Samarra et al. in prep). 138

139 In this study we investigate the social structure of herring-eating killer whales in Iceland, based upon patterns of association between photo-identified individuals in spawning and 140 141 overwintering grounds. We relate our results to the described societies of killer whales in the 142 Northeast Pacific. Specifically, we investigate: 1) the degree and diversity of associations 143 between pairs of individuals; 2) whether social structural units of individuals exist and are 144 hierarchically nested in the social structure; 3) how associations persist or change over time in 145 the population and depending on age-sex class and; 4) whether variations in movement and feeding strategy within the Iceland killer whale population influence sociality by promoting 146 147 social segregation. Given the differences in historical availability, migration patterns, and 148 anti-predator strategies of herring, salmon and seals, we hypothesize that broad ecology (fish-

vs. mammal-eating) alone cannot explain sociality and that local ecological conditions, such
as characteristics of prey schools and associated foraging strategy of the population, might
also strongly shape the social structure of killer whales.

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153 METHODS

154 Data collection

Photographs of killer whales were collected in July 2008-2010 and 2013-2015 in 155 Vestmannaeyjar (South Iceland), a spawning ground of ISS herring, and in February-March 156 157 2013-2014 and mid-February to mid-March 2015 in Grundarfjörður and Kolgrafafjörður (West Iceland), two fjords that were part of the ISS herring overwintering grounds. During 158 159 daylight hours, when killer whales were encountered, groups were approached and 160 photographs of all individuals surfacing together were taken using a variety of digital singlelens reflex cameras with telephoto lenses. On several occasions, more than one 161 photographer/camera was used. Sampling effort varied across years and seasons, due to 162 163 weather conditions, research effort priorities, and the number of research vessels used (Table 1). In the winters of 2014-2015 a whale-watching platform was also used. Due to the inherent 164 165 difficulty in approaching and photographing all individuals from whale watching platforms, only encounters when coverage of the groups present was considered complete (i.e. all 166 individuals in the group were identified) were included in the analysis. 167

168

169 **Photo-identification**

170 Killer whales were individually identified based on the size and shape of the dorsal fin, 171 patterns of the saddle patch and natural markings, such as nicks and scars, using left-side 172 pictures (Bigg 1982). For young animals without distinct natural markings, the eyepatch was also used for identification across seasons/years. The quality of photographs was judged based 173 174 upon focus, contrast, angle and overall quality assessment (adapted from Friday et al. 2000). Only high and medium quality photographs were used. To avoid false positives, matches were 175 176 confirmed if 3 distinct features of the individual were unambiguously identified. 177 To differentiate sex and stage of maturity four different categories were used: (1) Adult males 178 - adults that have reached sexual maturity and present distinguishable taller dorsal fin; (2) 179 Adult females – mature size individuals, with relatively smaller dorsal fin, seen during the 180 study period either consistently with a calf in echelon position, or without developing dorsal fin for at least 3 years, or that were matched to a preliminary catalogue from the Marine 181 Research Institute including photos taken between 1981 and 2007, without developing dorsal 182 fin; (3) Juveniles – identifiable individuals > 1 year old that have not reached mature size 183 (both sexes); (4) Other – whales of apparently larger size than juveniles but for which sex and 184 185 stage of maturity were impossible to determine. 186 Individuals that were only sighted in the summers of 2008-2010 were excluded from the

analysis to reduce bias resulting from including individuals that may have died during the first
years of the study and reduce the possibility of incomplete group coverage data from
fieldwork where photographic data collection was opportunistic. This procedure excluded 25
individuals from the study.

192 Association criterion

193 Due to the common observations of aggregations of individuals in Iceland (Sigurjónsson et al. 1988; Beck et al. 2012), spatiotemporally isolated groups in the field are unclear and it is 194 difficult to rigorously define a group. Despite this, the way in which animals are 195 196 photographed is related to their inherent social structure, as animals that prefer to associate 197 will undoubtedly be photographed together or in close proximity more often (Bigg et al. 1990). Using the capture time recorded in each photograph's metadata, we can discriminate 198 199 animals surfacing together in close proximity, since they are photographed within a very short 200 time frame. Individuals were considered associated for the day (sampling period) if 201 photographed by the same camera/photographer within 20 seconds. This value was quantitatively derived by maximum likelihood estimation of photographic bouts (Langton et 202 203 al. 1995; Luque and Guinet 2007; see Supplementary Material S1). The association criterion 204 matches our field observations that groupings of adjacent associated animals tended to surface 205 (and be available for photographing) close in time to each other and within 20 seconds, whereas non-contiguous animals were generally only available for photographing after a 206 207 longer time had passed. Shorter and longer temporal association criteria (5 seconds and 1 208 hour, respectively) were used to test the robustness of the observed association patterns to the temporal criterion used (Supplementary Material S2). These analyses suggested that the 209 210 association criterion value used in the study is likely meaningful to describe the animals' 211 social structure and appropriate to capture important associates without overloading the 212 analysis with random associations.

214 Analysis of associations

Only individuals seen on at least 5 different days were included in the analysis of

associations. This value is recommended by Whitehead (2008) as a minimum cut-off and it is

a more conservative restriction than several other studies (e.g. minimum cut-off of 4

sightings: Ottensmeyer and Whitehead 2003; Tosh et al. 2008; Beck et al. 2012; Esteban et al.

219 2016). We explored the consistency of the results under more restrictive thresholds (≥ 10 days

and \geq 20 days; see Supplementary Material S3). Due to similarity in the obtained results and

the fact that restricting the criterion for inclusion to a minimum of 5 sampling periods

significantly increased the number of individuals included in the analysis, this was considered

an appropriate threshold to describe the population dynamics of this social system.

All analyses described below were conducted using SOCPROG 2.6 (Whitehead 2009) in

225 MatLab 8.5 (MathWorks, Natick, MA, U.S.A.), except where noted. To quantify associations 226 between pairs of individuals we calculated the half-weight index (HWI), which estimates the proportion of time individuals spend together: HWI = 2AB/(A + B), where AB is the number 227 of times individuals A and B were identified associating with each other, and A and B are the 228 229 total number of times each individual was identified (Cairns and Schwager 1987; Whitehead 230 2008a). This symmetric association index was chosen since it minimizes sampling bias when 231 some individuals present were missed. This index was calculated per season (summer and 232 winter) and overall.

We used a permutation test, permuting the associations within samples (days), to test whether associations in the population were different from random, with the null hypothesis that between sampling periods there are no preferred/avoided associations (Bejder et al. 1998;

Whitehead 2008a). This test reveals whether or not an observed social structure is only due to properties of the dataset used (e.g. the size of aggregations of individuals and the number of encounters or sampling periods) when the associations are not different from random. The association matrix was permuted 10000 times, when the P value stabilized, with 1000 trials (inversion of part of the matrix of associations) per permutation. The random data obtained by this process were also used in the temporal analysis of associations.

To measure how diverse the associations were, we calculated the social differentiation (*S*) of the population. Social differentiation is the estimated coefficient of variation of association indices of the population. If *S* is close to 0, the associations are very homogenous, and if *S* > 1.0 the relationships are very diverse across dyads of animals (Whitehead 2008a; Whitehead 2009). The social differentiation was calculated using the likelihood method described by Whitehead (2008b), with non-parametric bootstrap for calculating its standard error (SE) and sampling periods chosen randomly for each of 1000 bootstrap samples.

249

250 Hierarchical stratification

Hierarchical clustering analysis using a dendogram display (tree diagram where individuals
are represented by nodes and the branching pattern represents the degree of associations) have
been used to visualize and interpret the social structure of killer whale populations (e.g. Bigg
et al. 1990; Baird and Whitehead 2000; Beck et al. 2012). This agglomerative technique
imposes a model where the social structure of the population is hierarchically structured: basic
social units (permanent or semipermanent social entities at high association values) are nested

within larger social units (permanent or semipermanent social entities – 'tiers' (Wittemyer et
al. 2005) - at low association values) in a stratified fashion (Whitehead 2008a; Whitehead
2009).

260 To investigate whether or not the Icelandic population exhibits clear hierarchical stratification 261 we combined the quantification of the distribution of HWI along with a hierarchical display of 262 associations and a visual exploration of the stratification of the population. We displayed associations as a dendogram, using the average-linkage clustering method. To identify the 263 264 association index at which significant divisions within the population occurred we used 265 modularity, defined by Newman (2004), controlling for differences in gregariousness 266 ('Modularity-G'; Lusseau 2007; Whitehead 2008a). To identify the degree of possible 267 stratification among individuals we explored the fragmentation of the population's social 268 network across lower values of HWI. We displayed the associations between individuals as a 269 social network, where nodes represent individuals and edges (links) between nodes represent 270 an existing association. Then, we sequentially removed edges in the population with 271 increasing HWI values to visualize the fragmentation of the network at each level, and 272 removed isolated (unconnected) nodes from the display for clarity. This was performed in R 273 3.2.3 (R Core Team 2015) using the package igraph (Csardi and Nepusz 2006).

274

275 Non-hierarchical structure and movement pattern assortative mixing

We used Newman's (2006) eigenvector-based clustering method to detect social clusterswithin the population. This clustering technique sequentially divides the population into

successive clusters and does not assume a hierarchical association between individuals. 278 Maximum modularity (Q) values higher than 0.3 describe a good division of the population 279 280 into clusters (Newman 2004). We described the composition, mean and maximum HWI, 281 movement pattern of adults and social differentiation (with and without juveniles) of each 282 cluster obtained by the division. The Pearson correlation test, calculated in MatLab 8.5 283 (MathWorks, Natick, MA, U.S.A.), was used to test for correlation between social 284 differentiation and cluster size, with and without juveniles. Associations for each cluster were displayed as sociograms (circular network diagrams), created in in R 3.2.3 (R Core Team 285 286 2015) using the package igraph (Csardi and Nepusz 2006). To study the possible substructure within clusters we applied Newman's (2006) eigenvector-based clustering method to each 287 cluster individually. 288

We investigated whether the different movement pattern of individuals determined 289 290 association patterns by examining the assortative mixing in the population, i.e. the tendency 291 for individuals with the same movement pattern to preferentially associate, using Newman's (2002) assortativity coefficient (r). Three different broad movement patterns were considered 292 293 based on individual sighting history (as in FIP Samarra et al. in prep): (1) only identified in 294 the winter season, (2) only identified in the summer season, (3) identified on both winter and summer seasons. This coefficient ranges from zero to one and high values of r indicate higher 295 296 assortativity of the population, i.e. individuals associate only with others of the same 'type'. The assortativity coefficient was calculated in R 3.2.3 (R Core Team 2015) using the package 297 298 assortnet (Farine 2014), for the whole population with and without juveniles, to account for

the possibility of juveniles only being identified later in the study period. The SE wascalculated using the jackknife method described by Newman (2003).

301

302 Temporal patterns of associations

303 To investigate how associations changed over time we calculated the standardized lagged 304 association rate (SLAR). All individuals, regardless of sighting frequency, were used in this 305 analysis to avoid positive bias (Whitehead 2008a). The SLAR is the estimate of the 306 probability that if two individuals are associated after a specified lag, the second individual is 307 a randomly chosen associate of the first (Whitehead 1995; Whitehead 2008a). Standard errors 308 were calculated using a temporal jackknife procedure with 15 day periods of data being 309 omitted in turn (Whitehead 1995; Whitehead 2007). To categorize how the relationships 310 between individuals changed over time, four different theoretical exponential models were 311 fitted (by maximum likelihood and binomial loss) to the full data set. The models are based in the presence/absence of constant and temporary associations (Whitehead 2008a). The model 312 313 that best fitted the data is indicated by the lowest quasilikelihood Akaike information criterion 314 (QAIC, Whitehead 2007). The difference between the QAIC of the best model and other models (Δ QAIC) indicates the degree of support for the less favored models: differences 0-2 315 316 indicating substantial support, 4-7 indicating less support and > 10 indicating essentially no 317 support for the alternative models (Burnham and Anderson 2002). Starting values of the 318 parameters for all models were obtained from the estimated parameters of the best-fitted 319 model in a preliminary fitting (with initial values of all parameters set to 0.5). The jackknife 320 method gives standard errors for the parameters of the model and for measures of social

| 321 | structure estimated from them: typical group size (as in Jarman (1974), number of associated |
|-----|--|
| 322 | individuals in groups, including the individual itself) and typical unit size (number of |
| 323 | individuals in permanent units), considering the case where permanent units temporarily |
| 324 | group (Whitehead 2008a). |
| 325 | |
| 326 | Sex differences in association patterns |
| 327 | Differences in patterns of association by sex were investigated as in Baird and Whitehead |
| 328 | (2000). The mean and maximum HWI within and between sexes were calculated for adults of |
| 329 | known sex (Females and Males) seen on 5 or more days. The mean HWI between A-B is an |
| 330 | estimate of the probability of a random individual of category A associating with any |
| 331 | individual of category B at any sampling period, so it is insensitive to different numbers of |
| 332 | individuals in different categories (Baird and Whitehead 2000). The maximum HWI of A-B is |
| 333 | the average maximum of association indices between each individual from category A and |
| 334 | any individual from category B. We tested the null hypothesis that associations between and |
| 335 | within sexes are similar using a Mantel test where associations between categories were |
| 336 | permuted 5000 times (Schnell et al. 1985). Variation in temporal patterns of associations were |
| 337 | analyzed using the SLAR for associations between all adults of known sex (Females and |
| 338 | Males) to avoid positive bias of the SLAR. |

340 Adult female-specific analysis

Mixing within aggregations of resident killer whales have been noted to differ between males 341 342 and females: adult males can temporarily travel away from their matrilines, possibly for 343 mating purposes (Bigg et al. 1990; Barrett-Lennard 2000), but adult females generally stay in their matrilineal units, which are spatially dispersed, particularly during foraging (Ford 1989). 344 345 We therefore separately examined the associations only between the most frequently 346 encountered adult females in the study, as they may show higher levels of association within 347 groups and a more clearly defined tier structure than observed in the overall population. This analysis was performed using 2 restricted datasets: 1) adult females encountered on more than 348 349 10 days and at least in 3 different years; 2) adult females encountered on at least 20 days and at least in 3 different years. Permutation tests and dendograms were performed as described 350 above, for both datasets. Associations between females were displayed in a sociogram created 351 352 in R 3.2.3 (R Core Team 2015) using the package igraph (Csardi and Nepusz 2006).

353

354 **RESULTS**

314 individuals (88 adult males, 94 adult females, 59 juveniles and 73 others) were identified
in a total of 29023 photographs taken on 110 different days. The mean ± standard deviation

(SD) number of individuals identified per day was 25.1 ± 20.6 individuals (range = 1-121)

total identifications per day).

359

360 Analysis of associations

198 individuals (56 adult males, 69 adult females, 41 juveniles and 32 others) were identified on at least 5 days (mean of 12.6 ± 7.1 days, range of 5-38 days) and used in the analysis of associations. 51 were only sighted in the winter season (including 8 juveniles), 32 only in the summer season (including 11 juveniles), and 115 on both seasons (including 22 juveniles). Most individuals were seen in several years (mean \pm SD of 3 ± 1.5 years, range of 1-6 different years).

The mean HWI of the population was low (mean \pm SD = 0.02 \pm 0.01, non-zero HWI mean \pm 367 368 $SD = 0.18 \pm 0.19$). Regardless of the season, the distribution of non-zero HWI values 369 observed showed a high proportion of low level associations and relatively fewer strong ties 370 at high HWI values (Figure 1). More than half of the pairs of associations were lower than 0.1 371 (51.4%, 1161 dyads). Only 9.9% (224 dyads) of the associations had $HWI \ge 0.5$ (individuals 372 associated more than half of the time). This was the value used by Baird and Whitehead 373 (2000) to define matrilines in the Pacific mammal-eating population and by Bigg et al. (1990) to define pods of matrilines that frequently associated. Only 0.9% of the associations (21 374 dyads) were higher than 0.8, the value used by Beck et al. (2012) to define primary social 375 376 tiers, equivalent to matrilines.

The SD and coefficient of variation (CV) of association indices were significantly higher in the real dataset than in the permuted data (real SD = 0.09, random SD = 0.05, P=0.0001; real CV = 4.12; random CV = 2.59; P = 0.0001). Hence, we could reject the null hypothesis that individuals associated randomly. The social differentiation of the population was close to 1 (*S* \pm SE = 0.98 \pm 0.03), revealing a highly diverse range of associations within the population.

383 Hierarchical stratification

| 384 | Applying the | ne hierarchical | dendogram | display | (cophe |
|-----|--------------|-----------------|-----------|---------|--------|
|-----|--------------|-----------------|-----------|---------|--------|

385 netic correlation coefficient [CCC] > 0.8; Figure 2), social clusters diverged at an extremely 386 low association index value (HWI of 0.02, maximum modularity of 0.68). The knot diagram presented an apparent constant rate of cumulative bifurcations, which only slightly increased 387 388 at very low association indices. This pattern was still visible using a very restrictive 389 association criterion (Figure S3 in Supplementary Material S2). The network of associations 390 was more interconnected at low HWI thresholds (Figure 3). However, without a larger 391 number of strong bonds the network started to fragment very quickly when links were 392 sequentially removed at low HWI thresholds. The network contained few stronger ties, as is 393 visible when HWI = 0.5, with very small sets differentiated and individuals detached from the network. Associations in the Icelandic killer whale population did not appear to be clearly 394 stratified into hierarchical tiers. Considering the wide range of association levels present, this 395 396 does not mean that individuals only associate with a small set of companions.

397

398 Examination of structure and movement pattern assortative mixing

Using Newman's (2006) clustering technique, the population could be significantly divided in 18 distinct clusters (Table 2; Q = 0.66). The social clusters obtained in the analysis were of mixed sex-age classes. The cluster sizes varied between 3 and 33 individuals, with a mean ± SD of 11 ± 7.8 individuals per cluster. As expected, mean HWI within clusters was higher

| 403 | than between clusters (within clusters mean HWI \pm SD = 0.27 \pm 0.17 and maximum HWI \pm |
|-----|---|
| 404 | SD = 0.65 \pm 0.17; between clusters mean HWI \pm SD = 0.01 \pm 0.01 and maximum HWI \pm SD |
| 405 | = 0.01 \pm 0.06). The assortativity coefficient of the network indicated some level of separation |
| 406 | of associations according to movement pattern (including juveniles $r \pm SE = 0.44 \pm 0.01$; not |
| 407 | including juveniles $r \pm SE = 0.49 \pm 0.01$) but much lower than would be expected if |
| 408 | individuals favored associations with others of equal movement pattern and/or avoided |
| 409 | associations with individuals with a different movement pattern. In fact, not all clusters were |
| 410 | discriminated by movement pattern: 5 clusters were composed of a mix of individuals sighted |
| 411 | in both seasons and individuals sighted in a single season. |
| 412 | Clusters were highly variable in their complexity (Table 2). There was a wide range of values |
| 413 | of social differentiation by cluster (with juveniles mean \pm SE = 0.52 \pm 0.1, min-max: 0-1.15; |
| 414 | without juveniles mean \pm SE = 0.49 \pm 0.1, min-max: 0-1.17). The Pearson's correlation test |
| 415 | showed that social differentiation was significantly correlated with unit size (with juveniles r |
| 416 | = 0.68, P = 0.002; without juveniles $r = 0.62$, P = 0.006). Within larger clusters not all |
| 417 | associations were strong (representing high social preference) or weak, and members |
| 418 | associated at many different degrees. In general, only a few individuals within each cluster |
| 419 | maintained strong associations (> 0.5 or 0.8) with other members and only 5 clusters had a |
| 420 | mean HWI > 0.5 . From the measures of social structure, inspection of photographs and direct |
| 421 | observations we concluded that we were not able to identify all companions of the members |
| 422 | of cluster F. This cluster was most likely incomplete and therefore was not included in further |
| 423 | descriptions. |

Our analysis distinguished three types of social clusters in the population: stable (C, G, K, N, 424 O and P; Figure 4), intermediate complexity (D, H, M, Q and R; Figure 5) and complex (A, B, 425 E, I, J and L; Figure 6) clusters. Stable clusters had high mean HWI values, very low social 426 differentiation and members with equal movement pattern. Only in cluster G two juveniles 427 428 were subclustered with a very low modularity value, likely because they were born during the 429 study period and only identified later in the study. Therefore, these clusters had no apparent 430 substructuring and associations between members were generally more homogeneous but not equal. 431 432 Intermediate complexity clusters had intermediate values of mean HWI and social 433 differentiation, showing potential but unclear subclustering (Q values generally < 0.3), since

434 individuals across potential subclusters also associated very frequently. In general, cluster

435 members had equal movement patterns, except for one cluster.

Complex clusters had very high values of social differentiation and very low mean HWI, but
high maximum HWI. In general, cluster members had different movement patterns, except for
two clusters. Complex clusters showed potential substructuring, although this was not clear
for all clusters (Q values of about 0.3 for cluster B and J). Associations between members of
complex clusters were diverse and only some members maintained strong associations, with
most associations being lower and at varying levels.

442

443 Temporal patterns of associations

444 The standardized lagged association rate SLAR ($g'(\tau)$) remained higher than would be 445 expected from random associations over the investigated time periods (τ ; Figure 7), indicating 446 that non-random associations persisted over time.

447 The two more complex models presented a reasonable fit to the data (see Supplementary

448 Material S4). The model SLAR3, labelled as 'constant companions plus casual acquaintances'

in Whitehead (2008a), had the lowest QAIC value, fitting the data best. Adding a second level

450 of dissociation (SLAR4), gave a similar curve and a very small difference of QAIC to SLAR3

451 indicating some support for this model. However, contrary to SLAR3, there was no

452 convergence and stable fit of SLAR4 when varying the parameters start values, which raised

doubt on the suitability of this model for the data. For this reason the simpler model SLAR3,

454 which has lowest QAIC and consistent parameters, was chosen to describe the temporal

455 patterning of associations. This model indicated that the population was driven by a

456 combination of longer-term relationships that last for many years, and temporary associations:

457 $g'(t) = 0.06 + 0.02e^{-0.0486t}$. Temporary associations decayed exponentially, with the model

458 suggesting important dissociations over scales of about 21 days (0.0486/days, SE = 0.09). The

459 proportion of long-term associations was 77%, with only 23% of temporary relationships.

460 This model's fit estimated a typical group size of 14.8 individuals (SE = 2.5) and a typical

461 unit size of 11.7 individuals (SE = 3.4).

462

463 Sex differences in association patterns

125 adults of known sex seen on 5 or more days were used in this analysis. Association levels 464 within and between adult sex classes were similar, with low mean association indices and 465 high maximum association indices (Table 3). The Mantel test did not reveal clear significant 466 467 differences in association between, relative to within adult sex classes (permutation test, P = 0.05). If the analysis was restricted to 75 adults of know sex seen on more than 10 days there 468 469 was no significant difference in association (permutation test, P = 0.13). The temporal 470 analysis suggests that Female-Male, Male-Male and Female-Female associations were somewhat stable across time and remained higher than random (Figure 8). For all types of 471 472 associations, the SLAR was higher than the SLAR between all individuals (higher probability of association). In general, all SLAR were relatively stable over time and no sex difference 473 was noticeable. 474

475

476 Adult female-specific analysis

| 477 | 32 adult females were sighted on more than 10 days over at least 3 years and only 12 of those |
|-----|--|
| 478 | were sighted on at least 20 days over at least 3 different years (Table 4). On both restriction |
| 479 | conditions, associations were non-random (32 females: real SD = 0.11 , random SD = 0.08 , P |
| 480 | = 0.0001 and real CV $= 2.33$, random CV $= 1.72$, P $= 0.0001$; 12 females: real SD $= 0.17$, |
| 481 | random SD = 0.14, P = 0.0001 and real CV = 1.51, random CV = 1.31, P < 0.0001). The |
| 482 | classical hierarchical clustering technique displayed dendograms with a varying level of |
| 483 | associations between females, with significant clusters discriminated at low HWI values (see |
| 484 | Table 4 and Supplementary Material S5). Although the cluster discrimination occurs at a |
| 485 | higher HWI value for the set of females with the more restrictive observational threshold, it is |

still a low value and mostly weak associations are present within the discriminated clusters.

487 The sociogram showed that, regardless the observational threshold, associations between

488 females are mainly weak even between most females from the same cluster (Figure 9). Also,

there are several weak associations between females from many different clusters.

490

491 **DISCUSSION**

Our results showed that associations within the Icelandic population of herring-eating killer whales were non-random but the number of strong associations was small. Although the dendogram display of associations presented a high cophenetic correlation coefficient, social clusters were differentiated at extremely low levels of association. With this technique, individuals were clustered together also by least preferred associations, i.e. weaker associations at very low HWI values, since not all individuals associated strongly within social units.

In a hierarchically structured society, transitions between structural tiers are clear because 499 500 individuals within a social cluster (nested in a tier) associate more strongly than individuals 501 within clusters at the level above. Societies without hierarchical nesting can still display a 502 dendogram with a cophenetic correlation coefficient > 0.8, indicating an acceptable match to the matrix of association indices (Bridge 1993), while being an inappropriate way of 503 realistically displaying associations (Whitehead 2008a; Whitehead 2009). When individuals 504 505 associate weakly overall the degree of potential hierarchical stratification is limited since an 506 individual cannot represent its social unit because associations within a social unit are not

equally strong. Our study showed this to be the case in this population. Thus, a non-stratified
way of studying the society was considered more appropriate than techniques that assume a
hierarchically-organized social structure.

The population could be significantly divided into social clusters, which were highly diverse 510 511 in complexity (even when using a more restrictive observation threshold; Supplementary 512 Material S3 – Figure S10). A small portion of the clusters presented more coherent associations between members, which might represent cohesive basic structures. The majority 513 514 of the clusters presented diverse association strengths and potential further subclustering. In 515 some social clusters, many individuals did not strongly associate with all other members. This 516 population presented both constant and temporary associations, not completely assorted by 517 movement pattern and with no clear differences between sexes. Together these results suggest that the Icelandic herring-eating killer whale population has a multilevel society with no clear 518 519 nested hierarchical structure of coherent social units, different from other populations of killer whales studied to date. 520

521 The evidence for non-random associations indicates that our results were not merely a

522 consequence of the quality or constraints of the dataset. It is possible that some of the HWI

values were negatively biased due to incomplete photographic coverage of

524 groupings/aggregations (Ottensmeyer and Whitehead 2003). However, this type of bias would

only increase the probability of not rejecting the null hypothesis of associations being random.

526 The analysis using the most encountered adult females aimed at reducing the potential

527 influence of recording sporadic associations, due to the observation that adult female resident

528 killer whales have lower levels of mixing with other groups than other age-sex classes. Thus,

529 a matrilineal structure may have been more clearly detectable among adult females than in the 530 overall population. However, our population-level results were instead strongly supported by 531 the adult female-specific analysis, with adult females also presenting an unclear hierarchical 532 structure but a complex sociality with rare strong associations and many weak associations 533 between females from the same cluster, and several associations between females from 534 different clusters. There are indications that the weakness of associations is due to a high 535 variability across years (associations on one year might not occur in a different year) but the small yearly number of sightings limits our ability to reach a definitive conclusion on the 536 537 stability of associations and yearly preferences between these individuals.

538

539 A complex multilevel society

540 In the Icelandic herring-eating killer whale population individuals clearly associated at 541 different levels, in some cases forming subcluster units. This society appears to tend towards 542 an incompletely nested multilevel society (as in Figure 6 in de Silva and Wittemyer 2012). 543 The levels of social stratification are not hierarchically distinct because transitions between 544 levels are gradual and may vary among individuals or sets of individuals, i.e. not all 545 individuals associate at similarly higher levels within social units and at distinctly lower levels 546 between social units. The variability in cluster complexity indicates diverse association 547 patterns among individuals and suggests different association strategies within the population. 548 Killer whale movement patterns did not assort their associations. In fact, individuals from 549 different subclusters and clusters with markedly different movement patterns were commonly

550 seen in tight groupings within less than one body length, a measure commonly used in other 551 killer whale social structure studies to define a group (e.g. Ivkovich et al. 2010, Esteban et al. 552 2016; Figure 10). Furthermore, complex cluster A (Figure 6) was formed by 3 highly 553 distinctive subclusters: subcluster A1, composed of individuals seen in Iceland year-round 554 following the movements of the ISS herring stock; subcluster A2 composed of individuals 555 that are only seen in Iceland in the winter; subcluster A3, composed of 5 individuals matched 556 to the Scottish population (only 2 Others and 1 Juvenile from this cluster were not matched) and sighted in Scotland in the summer (Samarra and Foote 2015). Combining social structure 557 558 analysis with genetics could help to clarify the underlying aspects of social contact reported 559 here between whales with different movement patterns and potentially different feeding ecologies in Iceland. It is worth noting that the individuals matched to the Scottish population 560 561 were not always sighted together in Scotland (Samarra and Foote 2015) nor in Iceland. It is 562 possible that individuals were missed in Scotland due to the opportunistic nature of data 563 collection. However, in our study we could confirm that these individuals were not always associating at close proximity. 564

The Icelandic multilevel society seems to be driven by a mix of both constant and temporary associations of mean duration of about 21 days. This temporal pattern of fission-fusion dynamics can occur in several types of social systems: 1) one in which constant permanent social units temporarily associate; 2) one in which individuals temporarily maintain casual but preferred associations and; 3) one in which permanent units exist but some individuals are 'floaters' who move between units (Whitehead 2008a). When full units of individuals collectively join, the typical group size should be twice the typical unit size, as in Pacific

sperm whales representing 2 temporal stable units joining (Whitehead et al. 1991) or larger, as 572 573 in Nova Scotia long-finned pilot whales where a group is comprised of several units 574 (Ottensmeyer and Whitehead 2003). Our study suggests that the temporal pattern did not result from permanent social units temporarily associating since the estimated typical group 575 576 size was less than double of the typical unit size. Also, there was no indication of 'floaters' 577 moving between units and no evidence of adult dispersal in the Icelandic population. Instead, 578 temporary associations are probably formed between preferred but casual associates or potentially by small sets of associates who temporarily associate with full permanent units, as 579 580 small sets of associated 'floaters'. Cluster members with weaker ties might represent these 581 casual but preferred temporary associates. It is unknown if this behavioral flexibility is only maintained when killer whales aggregate in herring grounds or if it is seasonally shaped, so 582 583 further studies will be necessary to understand this type of affiliation.

584

585 How can local ecological context shape killer whale social structure?

586 Methodological differences among studies (e.g. disparity in sampling procedures, definition of association, association index used) prevent a quantitative comparison of social structure 587 588 between the Icelandic and other killer whale populations. Nevertheless, overall social 589 structure comparisons can still be made. If sociality was determined by fish- vs. mammal-590 eating ecological differences alone (Beck et al. 2012), we would expect that the Icelandic 591 population would have a similar social structure to fish-eating resident killer whales. Indeed, 592 mammal-eating killer whales show dispersal of either sex from maternal groups and relatively 593 rare and unstable associations between adult males (Baird and Whitehead 2000) which we did

not observe in our study and is also not present in residents (Bigg et al. 1990). These specific
characteristics of the mammal-eating population are linked to optimal foraging group size
adjustment when feeding on seals (Baird and Dill 1996). However, the clear stable matrilineal
units (cohesive long-term groups) with members associating strongly and permanently (Bigg
et al. 1990; Baird and Whitehead 2000) common to both mammal-eating and residents, was
not found in the Icelandic herring-eating population.

Coherent basic social units have been described for other killer whale populations regardless 600 601 of targeted prey (in Alaska: Matkin et al. 1999; Marion Island: Tosh et al. 2008; Northwest 602 Pacific: Ivkovich et al. 2010; and Gibraltar: Esteban et al. 2016) and it has been considered a 603 firm characteristic of the species despite ecological differences. In the Icelandic herring-eating 604 population the possible existence of matrilineal units is not clear, but cannot be rejected. For 605 example, the potential subclustering of cluster D (Figure 5) is matched to direct observations 606 of constant close proximity associates, which could be more similar to basic matrilineal units. Yet, these subclusters were still strongly associated and were seen frequently switching 607 preference for close companions across days and years, as well as with individuals from other 608 609 clusters. Therefore, if matrilineal units are present in this population it is possible that these 610 are not entirely comparable to the ones present in other killer whale societies. An increase in the timespan of association data and genetic analysis, relating kinship and gene flow with the 611 612 underlying patterns of associations, will be crucial to inform on the presence and 613 characteristics of family bond-units in this population.

Further differences from the resident killer whale society were the lack of clear social tiersand hierarchical nesting in the Icelandic herring-eating society, which included fission-fusion

| 616 | dynamics at an individual (or sets of a few individuals) rather than at a group level (periodic |
|-----|--|
| 617 | merging of permanent social units). A parallel variation in multilevel structuring has been |
| 618 | quantified in elephant societies (de Silva and Wittemyer 2012). African elephants (Loxodonta |
| 619 | africana) maintain a clear multitiered society of coherent basic units that associate |
| 620 | hierarchically. In contrast, Asian elephants (Elephas maximus) have a complex multilevel |
| 621 | society without hierarchical structuring and nested units. Asian elephants do not maintain |
| 622 | clear core groups and associations can be either ephemeral or long-term. de Silva and |
| 623 | Wittemyer (2012) could not determine whether these differences were due to phylogenetic or |
| 624 | ecological factors, but there were significant environmental differences between the two |
| 625 | societies, such as differences in primary productivity and predation pressure. |
| 626 | Our study points to a different view of killer whale social structure, with a more dynamic and |
| 627 | fluid sociality than generally inferred from broad ecology. As argued by Beck et al. (2012), |
| 628 | ecology probably influences killer whale sociality rather than simply phylogenetic separation |
| 629 | of populations. However, considering only fish- vs. mammal-eating strategies as the |
| 630 | ecological condition influencing sociality ignores important particularities of local ecological |
| 631 | context. Herring-eating killer whales in Iceland target a prey with particular characteristics |
| 632 | different from salmon and seals, such as antipredator behaviors, unpredictability and patchy |
| 633 | distribution of high biomass. This shapes the feeding behavior of the population and probably |
| 634 | its social structure. |
| 635 | Herring is a schooling fish with a diverse repertoire of antipredator maneuvers (Nøttestad and |

635 Herring is a schooling fish with a diverse repertoire of antipredator maneuvers (Nøttestad and
636 Axelsen 1999). Feeding upon this prey requires a highly coordinated group feeding technique
637 to herd and catch herring (Similä and Ugarte 1993), unlike feeding techniques described for

other fish-eating killer whale populations. To efficiently hunt larger concentrations or school 638 639 sizes using a coordinated foraging technique, killer whales might benefit from larger group 640 sizes to encircle the herring school (Vabø and Nøttestad 1997; Nøttestad et al. 2002). Active adjustment of killer whale numbers hunting herring schools has been observed in Norway 641 642 (Nøttestad et al. 2002): on 4 observations of feeding groups (range of 22-46 individuals, mean 643 \pm SD = 33.5 \pm 10.6 individuals), the 2 largest groups (38 and 46 individuals in total) occurred 644 when the herring layer was larger (depth range of 150/160 meters to 350 meters) and were composed by different smaller groups of killer whales that gathered before starting to herd 645 646 herring, arriving from different directions. In these conditions, it might be important to maintain a fission-fusion society where associations are flexible and individuals can actively 647 adjust to these constantly changing requirements. 648

Herring can also undergo substantial changes in density and spatial distribution, particularly 649 650 in overwintering grounds (Óskarsson et al. 2009). The unpredictability of the prey may additionally promote the maintenance of a more fluid and flexible sociality. A socioecological 651 model proposed for dolphins suggests that when resources are unpredictable, dolphins will 652 653 present wide range movements, reduced competition by cooperative foraging and larger 654 groups to more effectively find and exploit large prey schools (Gowans et al. 2007). Dusky dolphins (Lagenorhynchus obscurus) in Argentina feed on schooling fish and present similar 655 656 basic herding techniques to herring-eating killer whales (Würsig and Würsig 1980). Their target prey is also unpredictably distributed. The population presents a strong fission-fusion 657 658 society with constantly fluctuating subgroup memberships (although some associations might 659 be constant) that split for feeding and social purposes (Würsig and Würsig 1980; Würsig and

Bastida 1986). This social structure is very different from dusky dolphins of New Zealand
(Markowitz 2004), whose target preys are more predictable.

Finally, feeding aggregations in Iceland are very common during summer and winter, in 662 grounds where herring are temporarily highly concentrated. The patchiness of a resource will 663 664 influence whether animals do aggregate and, although these aggregations for feeding are not 665 social structures (spatiotemporal clusters of individuals forced by non-social factors) they might act as catalysts for sociality (Whitehead 2008a). Recurring aggregations due to prev 666 667 behavior may offer a special local ecological context for the establishment of associations, 668 creating opportunities for social interactions with other individuals and somehow shaping the 669 social structure of this population. The dynamic nature of the society described here may have 670 been uncovered because our data collection took place mostly during periods when large aggregations of whales can occur, due to this particular ecological context. Future work 671 672 focusing on social associations of herring-eating killer whales during periods when herring are more dispersed may reveal stronger social bonds, and clear long-term stable matrilineal 673 674 groups, if group sizes are substantially lower than observed during the herring spawning and 675 overwintering periods.

Other ecological differences such as habitat characteristics or historical capture might have also shaped the social structure of this population but we lack sufficient information to determine their influence at present. Furthermore, the Icelandic population is comprised of individuals with different seasonal movement patterns that associate at least seasonally. This alone can influence the social structure of the population, since different movement patterns within the same population suggest exposure to different environmental conditions. This

might also lead to variation in social factors within the population, e.g. mating competition or
avoidance, which can influence the structuring of basic and high-order groups in mammals
(Silk 2007). More information on the genetic relatedness of whales with different movement
patterns is needed to understand how it may affect the resulting society.

686 We have shown that the Icelandic herring-eating killer whale population has a complex 687 multilevel social structure with no clear hierarchical nesting and no strong social segregation by movement pattern. This social system appears to be different from other populations of 688 killer whales worldwide, but continued photo-identification data will be crucial to investigate 689 690 these questions over longer time scales and under different seasonal, spatial and prey 691 behavioral contexts. The differences observed suggest that fish vs. marine mammal prey-type 692 alone does not define killer whale social structure and local ecological context, such as prey 693 characteristics and foraging strategy, are probably strong drivers of sociality. The factors 694 constraining hierarchical stratification of societies are little understood and to our knowledge are not addressed in socioecological frameworks (e.g. Emlen and Oring 1977; Wrangham 695 696 1980; Gowans et al. 2007). Comparative studies of populations targeting similar prey will be 697 extremely important to quantitatively assess the degree of variation in multilevel social 698 structuring with local ecological context.

699

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719

720 Data accessibility statement

Analyses reported in this article can be reproduced using the data provided by Tavares et al.(2016).

724 **REFERENCES**

- 725 Baird RW, Dill LM. 1996. Ecological and social determinants of group size in transient killer
- whales. Behav. Ecol. 7:408–416.
- 727 Baird RW, Whitehead H. 2000. Social organization of mammal-eating killer whales: group
- stability and dispersal patterns. Can. J. Zool. 78:2096–2105.
- 729 Barrett-Lennard LG. 2000. Population structure and mating patterns of killer whales (Orcinus
- *orca*) as revealed by DNA analysis [PhD thesis]. University of British Columbia.
- 731 Barton RA, Byrne RW, Whiten A. 1996. Ecology, feeding competition and social structure in
- baboons. Behav. Ecol. Sociobiol. 38:321–329.
- 733 Beck S, Kuningas S, Esteban R, Foote AD. 2012. The influence of ecology on sociality in the
- killer whale (*Orcinus orca*). Behav. Ecol. 23:246–253.
- Bejder L, Fletcher D, Bräger S. 1998. A method for testing association patterns of social
- animals. Anim. Behav. 56:719–725.
- 737 Bigg MA. 1982. An assessment of killer whale (Orcinus orca) stocks off Vancouver Island,
- 738 British Columbia. Rep. Int. Whal. Comm. 32:655–666.
- 739 Bigg MA, Olesiuk P, Ellis GM, Ford JKB, Balcomb KC. 1990. Social organization and
- 740 genealogy of resident killer whales (Orcinus orca) in the coastal waters of British Columbia
- and Washington State. Reports Int. Whal. Commission Spec. Issue 12:383–405.
- 742 Bridge PD. 1993. Classification. In: Fry JC, editor. Biological data analysis. Oxford, UK:
- 743 Oxford University Press. p. 219–242.

- 744 Burnham KP, Anderson DR. 2002. Model selection and multimodel inference: a practical
- r45 information-theoretic approach. New York, USA: Springe-Verlag.
- Cairns SJ, Schwager SJ. 1987. A comparison of association indices. Anim. Behav. 35:1454–
 1469.
- 748 Connor RC, Mann J, Tyack PL, Whitehead H. 1998. Social evolution in toothed whales.
- 749 Trends Ecol. Evol. 13:228–232.
- 750 Connor RC, Smolker RA, Richards AF. 1992. Two levels of alliance formation among male
- bottlenose dolphins (*Tursiops sp.*). Proc. Natl. Acad. Sci. 89:987–990.
- 752 Csardi G, Nepusz T. 2006. The igraph software package for complex network research.
- 753 InterJournal, Complex Syst. 1695:1–9.
- Emlen ST, Oring LW. 1977. Ecology, sexual selection, and the evolution of mating systems.
- 755 Science. 197:215–223.
- Esteban R, Verborgh P, Gauffier P, Giménez J, Foote AD, de Stephanis R. 2016. Maternal
- 757 kinship and fisheries interaction influence killer whale social structure. Behav. Ecol.
- 758 Sociobiol.:111–122.
- 759 Farine DR. 2014. Measuring phenotypic assortment in animal social networks: weighted
- associations are more robust than binary edges. Anim. Behav. 89:141–153.
- Ford JKB. 1989. Acoustic behaviour of resident killer whales (Orcinus orca) off Vancouver
- 762 Island, British Columbia. Can. J. Zool. 67:727–745.
- Ford JKB. 1991. Vocal traditions among resident killer whales (Orcinus orca) in coastal

- vaters of British Columbia. Can. J. Zool. 69:1454–1483.
- Ford JKB, Ellis GM, Balcomb KC. 2000. Killer whales: the natural history and genealogy of
- 766 Orcinus orca in British Columbia and Washington. Vancouver, Canada: UBC Press.
- Ford JKB, Ellis GM, Barrett-Lennard LG, Morton AB, Palm RS, Balcomb KC. 1998. Dietary
- specialization in two sympatric populations of killer whales (Orcinus orca) in coastal British
- 769 Columbia and adjacent waters. Can. J. Zool. 76:1456–1471.
- 770 Friday N, Smith TD, Stevick PT, Allen J. 2000. Measurement of photographic quality and
- individual distinctiveness for the photographic identification of humpback whales, *Megaptera*
- *novaeangliae*. Mar. Mammal Sci. 16:355–374.
- Fruchterman TMJ, Reingold EM. 1991. Graph drawing by force-directed placement. Softw.
- 774 Pract. Exp. 21:1129–1164.
- Gowans S, Würsig B, Karczmarski L. 2007. The social structure and strategies of delphinids:
- predictions based on an ecological framework. Adv. Mar. Biol. 53:195–294.
- 777 Grueter CC, Chapais B, Zinner D. 2012. Evolution of multilevel social systems in nonhuman
- primates and humans. Int. J. Primatol. 33:1002–1037.
- Grueter CC, Matsuda I, Zhang P, Zinner D. 2012. Multilevel societies in primates and other
- mammals: introduction to the special issue. Int. J. Primatol. 33:993–1001.
- 781 Ivkovich TV, Filatova OA, Burdin AM, Sato H, Hoyt E. 2010. The social organization of
- resident-type killer whales (Orcinus orca) in Avacha Gulf, Northwest Pacific, as revealed
- through association patterns and acoustic similarity. Mamm. Biol. Zeitschrift fur Saugetierkd.

- 784 75:198–210.
- Jarman PJ. 1974. The social organisation of antelope in relation to their ecology. Behaviour
 48:215–267.
- 787 Krause J, Ruxton GD. 2002. Living in Groups. New York, USA: Oxford University Press.
- Langton SD, Collett D, Sibly RM. 1995. Splitting behaviour into bouts; a maximum
- 789 likelihood approach. Behaviour 132:781–799.
- 790 Luque SP, Guinet C. 2007. A maximum likelihood approach for identifying dive bouts
- improves accuracy, precision and objectivity. Behaviour 144:1315–1332.
- Lusseau D. 2007. Why are male social relationships complex in the Doubtful Sound
- bottlenose dolphin population? PLoS One 2:e348.
- Markowitz TM. 2004. Social organization of the New Zealand dusky dolphin. Texas A&MUniversity.
- 796 Matkin CO, Ellis GM, Olesiuk P, Saulitis EL. 1999. Association patterns and inferred
- 797 genealogies of resident killer whales, *Orcinus orca*, in Prince William Sound, Alaska. Fish.
- 798 Bull. 97:900–919.
- Newman MEJ. 2002. Assortative mixing in networks. Phys. Rev. Lett. 89:208701.
- Newman MEJ. 2003. Mixing patterns in networks. Phys. Rev. E 67:026126.
- Newman MEJ. 2004. Analysis of weighted networks. Phys. Rev. E 70:056131.
- Newman MEJ. 2006. Modularity and community structure in networks. Proc. Natl. Acad. Sci.

- 803 U. S. A. 103:8577–8582.
- 804 Nøttestad L, Axelsen BE. 1999. Herring schooling manoeuvres in response to killer whale
- attacks. Can. J. Zool. 77:1540–1546.
- 806 Nøttestad L, Fernö A, Axelsen BE. 2002. Digging in the deep: killer whales' advanced
- hunting tactic. Polar Biol. 25:939–941.
- 608 Óskarsson GJ, Gudmundsdottir A, Sigurdsson T. 2009. Variation in spatial distribution and
- 809 migration of Icelandic summer-spawning herring. ICES J. Mar. Sci. J. du Cons. 66:1762–
- 810 1767.
- 811 Ottensmeyer CA, Whitehead H. 2003. Behavioural evidence for social units in long-finned
- 812 pilot whales. Can. J. Zool. 1338:1327–1338.
- 813 R Core Team. 2015. R: A language and environment for statistical computing.
- 814 Samarra FIP, Foote AD. 2015. Seasonal movements of killer whales between Iceland and
- 815 Scotland. Aquat. Biol. 24:75–79.
- 816 Samarra FIP, Vighi M, Aguilar A and Víkingsson GA. In press. Intra-population variation in
- 817 isotopic niche in herring-eating killer whales. Mar. Ecol. Prog. Ser.
- 818 Schnell GD, Watt DJ, Douglas ME. 1985. Statistical comparison of proximity matrices:
- applications in animal behaviour. Anim. Behav. 33:239–253.
- 820 Sigurjónsson J, Lyrholm T, Leatherwood S, Jónsson E, Víkingsson GA. 1988.
- Photoidentification of killer whales off Iceland 1981 through 1986. Rit Fiskid. 11:99–114.
- 822 Silk JB. 2007. The adaptive value of sociality in mammalian groups. Philos. Trans. R. Soc.

- 823 London B Biol. Sci. 362:539–559.
- de Silva S, Wittemyer G. 2012. A comparison of social organization in Asian elephants and
- African savannah elephants. Int. J. Primatol. 33:1125–1141.
- 826 Similä T, Ugarte F. 1993. Surface and underwater observations of cooperatively feeding killer
- whales in northern Norway. Can. J. Zool. 71:1494–1499.
- 828 Sinha A, Mukhopadhyay K, Datta-Roy A, Ram S. 2005. Ecology proposes, behaviour
- disposes: Ecological variability in social organization and male behavioural strategies among
- wild bonnet macaques. Curr. Sci. 89:1166–1179.
- Tavares SB, Samarra FIP, Miller PJO. 2016. Data from: A multilevel society of herring-eating
- killer whales indicates adaptation to prey characteristics. Behavioral Ecology.
- 833 http://dx.doi.org/10.5061/dryad.j619s.
- Tosh CA, De Bruyn PJN, Bester MN. 2008. Preliminary analysis of the social structure of
- killer whales, *Orcinus orca*, at subantarctic Marion Island. Mar. Mammal Sci. 24:929–940.
- 836 Vabø R, Nøttestad L. 1997. An individual based model of fish school reactions: predicting
- antipredator behaviour as observed in nature. Fish. Oceanogr. 6:155–171.
- 838 Whitehead H. 1995. Investigating structure and temporal scale in social organizations using
- 839 identified individuals. Behav. Ecol. 6:199–208.
- 840 Whitehead H. 2007. Selection of models of lagged identification rates and lagged association
- rates using AIC and QAIC. Commun. Stat. Simul. Comput. 36:1233–1246.
- 842 Whitehead H. 2008a. Analyzing animal societies: quantitative methods for vertebrate social

- analysis. Chicago, USA: University Of Chicago Press.
- 844 Whitehead H. 2008b. Precision and power in the analysis of social structure using
- associations. Anim. Behav. 75:1093–1099.
- 846 Whitehead H. 2009. SOCPROG programs: analysing animal social structures. Behav. Ecol.
- 847 Sociobiol. 63:765–778.
- 848 Whitehead H, Antunes R, Gero S, Wong SNP, Engelhaupt D, Rendell L. 2012. Multilevel
- societies of female sperm whales (*Physeter macrocephalus*) in the Atlantic and Pacific: why
- are they so different? Int. J. Primatol. 33:1142–1164.
- 851 Whitehead H, Waters S, Lyrholm T. 1991. Social organization of female sperm whales and
- their offspring: constant companions and casual acquaintances. Behav. Ecol. Sociobiol.
- **853** 29:385–389.
- 854 Wittemyer G, Douglas-Hamilton I, Getz WM. 2005. The socioecology of elephants: analysis
- of the processes creating multitiered social structures. Anim. Behav. 69:1357–1371.
- Wrangham RW. 1980. An ecological model of female-bonded primate groups. Behaviour
 75:262–300.
- 858 Würsig B, Bastida R. 1986. Long-range movement and individual associations of two dusky
- dolphins (*Lagenorhynchus obscurus*) off Argentina. J. Mammal. 67:773–774.
- 860 Würsig B, Würsig M. 1980. Behavior and ecology of the dusky dolphin, *Lagenorhynchus*
- *obscurus*, in the South Atlantic. Fish. Bull 77:871–890.





Figure 1. Distribution of non-zero half-weight index (HWI) values in the population using the



Figure 2. Average-linkage cluster analysis. (a) Dendogram of 198 individuals encountered on

at least 5 days (cophenetic correlation coefficient [CCC] = 0.94). (b) Knot diagram of

866 full dataset and by season.

867

- cumulative number of bifurcations across HWI levels. (c) A maximum modularity-G, within
 hierarchical clustering, of 0.68 suggests a division into distinct clusters at an HWI of 0.02
- 873 (dashed line).



875

Figure 3. Network fragmentation with increasing HWI threshold. Isolated individuals are
removed from the network (n indicates the number of individuals present). Note that at
HWI > 0.1 the network starts fragmenting quickly and more individuals become isolated from
the network. Plotted using Fruchterman-Reingold force-directed layout (Fruchterman and
Reingold 1991).



Figure 4. Sociogram of stable clusters. The thickness of the edges is related to the HWI value
of association. Nodes represent individuals and are shaped/colored based on age-sex class
(black circle: Female; grey circle: Male; black square: Other; grey square: Juvenile). There
was no apparent subcluster division. Two juveniles were subclustered in cluster G but with a
very low modularity value (Q), likely because the juveniles were only identified on the later
years of the study, contrary to other members.



890

Figure 5. Sociogram of intermediate complexity clusters. The thickness of the edges is related
to the HWI value of association. Nodes represent individuals and are shaped/coloured based

- 893 on age-sex class (black circle: Female; grey circle: Male; black square: Other; grey square:
- Juvenile). Q indicates the modularity of potential subcluster division.



Figure 6. Sociogram of complex clusters. The thickness of the edges is related to the HWI
value of association. Nodes represent individuals and are shaped/colored based on age-sex
class (black circle: Female; grey circle: Male; black square: Other; grey square: Juvenile). Q
indicates the modularity of potential subcluster division. A1, A2 and A3 indicate the 3
subclusters of Cluster A.



Figure 7. Standardized random and lagged association rates (SLAR, curve smoothed with
30000 moving average). Vertical bars represent temporal jackknife standard errors. The two

models of the exponential family with the lowest QAIC values, SLAR3 and SLAR4, are

shown (see Supplementary Material S4 for formulas and QAIC values).

908



909

910 Figure 8. Standardized lagged association rates (SLAR) for different associations between

911 adults. A different moving average was chosen accordingly to smooth lines. Jackknife

912 grouping factor of 15, shown as vertical bars. SLAR between Females and Males (F-M) and

913 between Males and Males (M-M) are high and relatively stable. Although lower, SLAR

between Females and Females (F-F) are also high and much higher than the SLAR between

all individuals (All) or if individuals had a random chance of associating (Random).



Figure 9. Sociograms of associations for the 32 most frequently encountered adult females (on
more than 10 days over at least 3 years) from 12 different clusters. Nodes represent each
female and are colored black if the individual was also seen on at least 20 days over at least 3
different years. Members of the same cluster were included within the same grey shading.
Note the lack of strong associations and that there are many weak associations between
females from different clusters. This is observed regardless of the minimum number of
sightings, with predominantly weak associations across black nodes.



927 Figure 10. Examples of close associations between individuals from different subclusters and

- 928 clusters: (a) IF-4 (Female, subcluster A3, Scotland ID 21) in close association with IS121
- 929 (Other, subcluster A1); (b) 997 (Female, subcluster A3, Scotland ID 19) in close association
- 930 with IS041 (Female, cluster L); (c) IS172 (Other, subcluster A3) associating with IS049
- 931 (Female, cluster D); (d) IS229 (Other, subcluster A3) associating with IS030 (Female, cluster

932 D).

933 TABLES

934 Table 1

935 Summary of the photo-identification sampling effort included in this study.

| | | Sampling pe | riods used | | | | | |
|--------|--------|-------------|------------|---|-------------|-------------|--|--|
| (days) | | | vs) | | Nur | Number of | | |
| | - | Research | WW | Start-end of | Dhotographs | Identified | | |
| Year | Season | vessels | platform | sampling periods | rnotographs | individuals | | |
| 2008 | Summer | 6 | - | $8^{th} - 20^{th}$ July | 382 | 29 | | |
| 2009 | Summer | 16 | - | $7^{th} - 29^{th}$ July | 2552 | 65 | | |
| 2010 | Summer | 6 | - | $4^{th}-10^{th} \ July$ | 748 | 70 | | |
| 2013 | Winter | 23 | - | 10th February – 24th March | 5649 | 211 | | |
| 2013 | Summer | 4 | - | 17 th -29 th July | 1980 | 51 | | |
| 2014 | Winter | 19 | 1 | 13th February – 31st March | 5510 | 115 | | |
| 2014 | Summer | 15 | - | $6^{th}-27^{th} \ July$ | 5265 | 149 | | |
| 2015 | Winter | - | 1 | 1 st March | 118 | 3 | | |
| 2015 | Summer | 19 | - | $7^{th}-29^{th} \ July$ | 6819 | 131 | | |

936

5 Days of sampling are discriminated by type of platform: research vessels and whale-watching (WW) boat.

937 Table 2

| | | Movement | | | | | | excluding |
|---------|--------|----------------------|------|------------------------------|---------------|--------------|--------------------|-------------|
| Cluster | n | pattern ¹ | Days | Identifications ² | Mean HWI (SD) | Max HWI (SD) | S (SE) | juveniles |
| А | 24 (5) | WB | 34 | 2279 | 0.178 (0.04) | 0.67 (0.15) | 0.88 (0.1) | 0.85 (0.1) |
| В | 33 (3) | W | 31 | 4112 | 0.12 (0.06) | 0.62 (0.15) | 1.01 (0.04) | 1.01 (0.04) |
| С | 9 (3) | В | 15 | 798 | 0.54 (0.09) | 0.73 (0.11) | 0.08 (0.12) | 0.05 (0.13) |
| D | 11 (3) | В | 60 | 3883 | 0.27 (0.06) | 0.65 (0.13) | 0.66 (0.06) | 0.57 (0.08) |
| Е | 10 (4) | WB | 24 | 549 | 0.2 (0.04) | 0.55 (0.14) | 0.96 (0.05) | 0.86 (0.08) |
| F | 3 | В | 10 | 91 | 0.12 (0.05) | 0.18 (0) | 0 (-) ³ | - |
| G | 8 (4) | В | 27 | 1918 | 0.49 (0.07) | 0.67 (0.08) | 0.17 (0.11) | 0 (0.07) |
| Н | 13 | В | 31 | 1817 | 0.34 (0.04) | 0.74 (0.07) | 0.36 (0.12) | - |
| I | 17 (4) | SB | 65 | 2754 | 0.13 (0.05) | 0.56 (0.19) | 1.15 (0.03) | 1.17 (0.04) |
| J | 11 (3) | SB | 33 | 1344 | 0.31 (0.16) | 0.81 (0.25) | 0.95 (0.06) | 1.08 (0.05) |
| K | 4 (1) | S | 19 | 675 | 0.6 (0.08) | 0.69 (0.08) | 0 (0.11) | 0 (0.03) |
| L | 18 (5) | В | 54 | 2825 | 0.16 (0.02) | 0.59 (0.15) | 1 (0.06) | 0.91 (0.1) |
| М | 4 | В | 8 | 142 | 0.27 (0.09) | 0.5 (0.12) | 0.47 (0.32) | - |
| N | 7 (2) | SB | 9 | 464 | 0.54 (0.08) | 0.73 (0.1) | 0 (0.19) | 0 (0.12) |
| 0 | 5 | В | 8 | 456 | 0.6 (0.1) | 0.74 (0.12) | 0 (0.2) | - |
| Р | 5 (2) | В | 22 | 680 | 0.53 (0.09) | 0.68 (0.11) | 0.17 (0.09) | 0 (0.09) |
| Q | 7 (1) | В | 31 | 1137 | 0.37 (0.1) | 0.75 (0.13) | 0.67 (0.14) | 0.68 (0.14) |
| R | 9 (1) | В | 18 | 511 | 0.28 (0.13) | 0.64 (0.18) | 0.79 (0.09) | 0.86 (0.09) |

938 Summary of different clusters identified using Newman's (2006) clustering technique.

939 n, number of members with number of juveniles in brackets; HWI, half-weight index of association; SD, standard deviation;

940 *S*, social differentiation; SE, standard error.

941 ¹Movement pattern of non-juvenile members: W – only seen in the winter, S – only seen in the summer, B – seen in both

seasons, WB – seen only in the winter or in both seasons, SB – seen only in the summer or in both seasons.

943 ² Total number of photographic record of identified individuals of each cluster.

³ There was insufficient association data to calculate SE of *S* for Cluster F.

S (SE)

- 945 Table 3
- 946 Distribution of HWI for adult individuals seen at least 5 times, between and within sex
- 947 classes.

| | | × × |
|-----------------|-------------|-------------|
| Females-All | 0.02 (0.01) | 0.59 (0.18) |
| Males-All | 0.02 (0.01) | 0.62 (0.21) |
| Females-Females | 0.02 (0.01) | 0.44 (0.21) |
| Females-Males | 0.02 (0.01) | 0.48 (0.24) |
| Males-Females | 0.02 (0.01) | 0.56 (0.23) |
| Males-Males | 0.02 (0.01) | 0.48 (0.26) |
| Within classes | 0.02 (0.01) | 0.46 (0.24) |
| Between classes | 0.02 (0.01) | 0.52 (0.23) |
| All-All | 0.02 (0.01) | 0.6 (0.19) |

Adult sex classes Mean HWI (SD) Max HWI (SD)

948 HWI, half-weight index of association; SD, standard deviation.

- 949 Table 4
- 950 Summary of the results of the adult female-specific analysis under two different observational
- 951 thresholds.

| >10 days,≥3 years | \geq 20 days, \geq 3 years | |
|--|---|--|
| 32 | 12 | |
| 19.1 ± 6.6 days (range of 11-34 days) over | 26.3 ± 4.5 days (range of 20-34 days) over | |
| 4.6 ± 1 years (range of 3-6 years) | 5.3 ± 0.5 years (from 5-6 years) | |
| Yes | Yes | |
| Figure S10 - Supplementary Material S5 | Figure S11 - Supplementary Material S5 | |
| 0.04 | 0.21 | |
| 0.48 | 0.39 | |
| | > 10 days, ≥ 3 years 32 19.1 ± 6.6 days (range of 11-34 days) over 4.6 ± 1 years (range of 3-6 years) Yes Figure S10 - Supplementary Material S5 0.04 0.48 | |

952 n, number of adult females in the analysis; HWI, half-weight index of association; SD, standard deviation.