

**THE DAVIS-GUT LAW FOR INDEPENDENT AND  
IDENTICALLY DISTRIBUTED BANACH SPACE VALUED  
RANDOM ELEMENTS**

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ABSTRACT. An analog of the Davis-Gut law for a sequence of independent and identically distributed Banach space valued random elements is obtained, which extends the result of Li and Rosalsky (A supplement to the Davis-Gut law. J. Math. Anal. Appl. 330 (2007), 1488–1493).

1. INTRODUCTION

Let  $\{X, X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables. The following theorem, which is related to the classical Hartman-Wintner law of the iterated logarithm (see, Hartman and Wintner, [6]), is well known. As usual we let  $\log t = \log_e \max\{e, t\}$  for  $t \geq 0$ .

THEOREM 1.1. *The following three statements are equivalent:*

$$(1.1) \quad EX = 0 \text{ and } EX^2 = 1,$$

$$(1.2) \quad \sum_{n=1}^{\infty} \frac{1}{n} P \left\{ \left| \sum_{k=1}^n X_k \right| > (1 + \varepsilon) \sqrt{2n \log \log n} \right\} \begin{cases} < \infty, & \text{if } \varepsilon > 0 \\ = \infty, & \text{if } \varepsilon < 0, \end{cases}$$

$$(1.3) \quad \sum_{n=1}^{\infty} \frac{\log \log n}{n} P \left\{ \left| \sum_{k=1}^n X_k \right| > (1 + \varepsilon) \sqrt{2n \log \log n} \right\} \begin{cases} < \infty, & \text{if } \varepsilon > 0 \\ = \infty, & \text{if } \varepsilon < 0. \end{cases}$$

This result is referred to as the Davis-Gut law. The implication “(1.1) $\Rightarrow$ (1.2)” was formulated by Davis ([3]) with an invalid proof which was corrected by Li et al. ([11]). The implication “(1.2) $\Rightarrow$ (1.1)” was obtained

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by Gut ([5]). The equivalence between (1.1) and (1.3) was established by Li ([9]). Necessary and sufficient conditions for (1.3) in a Banach space setting were obtained by Li ([9]). For moving average processes, the implications “(1.1) $\Rightarrow$ (1.2)” and “(1.1) $\Rightarrow$ (1.3)” were obtained by Chen and Wang ([1]).

Li and Rosalsky ([10]) provided the following supplement to the Davis-Gut law. When  $h(t) \equiv 1$ , it yields the equivalence between (1.1) and (1.2).

**THEOREM 1.2.** *Let  $h(\cdot)$  be a positive nondecreasing function on  $(0, \infty)$  such that  $\int_1^\infty (th(t))^{-1} dt = \infty$ . Write  $\Psi(t) = \int_1^t (sh(s))^{-1} ds, t \geq 1$ . Then (1.1) and*

$$(1.4) \quad \sum_{n=1}^{\infty} \frac{1}{nh(n)} P \left\{ \left| \sum_{k=1}^n X_k \right| > (1 + \varepsilon) \sqrt{2n \log \Psi(n)} \right\} \begin{cases} < \infty, & \text{if } \varepsilon > 0 \\ = \infty, & \text{if } \varepsilon < 0 \end{cases}$$

are equivalent.

Recently, Liu et al. ([12]) extended Theorem 1.2 to moving average processes which then extends the work of Chen and Wang ([1]) by establishing the implication “(1.2)  $\Rightarrow$  (1.1)” for moving average processes.

In this paper, we will extend Theorem 1.2 for a sequence of independent and identically distributed Banach space valued random elements.

## 2. PRELIMINARIES AND LEMMAS

Let  $B$  be a real separable Banach space with norm  $\|\cdot\|$  and let  $B^*$  denote the topological dual space of  $B$ . We let  $B_1^*$  denote the unit ball of  $B^*$ . Let  $(\Omega, \mathcal{F}, P)$  be a probability space. A *random element*  $X$  taking values in  $B$  is defined as an  $\mathcal{F}$ -measurable function from  $(\Omega, \mathcal{F})$  into  $B$  equipped with the Borel sigma-algebra; we call it a  $B$ -valued random element for short. The *expected value* or *mean* of a  $B$ -valued random element  $X$  is defined to be the Bochner integral and is denoted by  $EX$ .

**LEMMA 2.1.** *Let  $\{k_n, n \geq 1\}$  be a sequence of positive integers and  $\{X_{nk}, 1 \leq k \leq k_n, n \geq 1\}$  an array of rowwise independent  $B$ -valued random elements. Suppose that there exists  $\delta > 0$  such that  $\|X_{nk}\| \leq \delta$  a.s. for all  $1 \leq k \leq k_n, n \geq 1$ . If  $\sum_{k=1}^{k_n} X_{nk} \rightarrow 0$  in probability, then  $E\|\sum_{k=1}^{k_n} X_{nk}\| \rightarrow 0$  as  $n \rightarrow \infty$ .*

**PROOF.** Let  $\{X'_{nk}, 1 \leq k \leq k_n, n \geq 1\}$  be an independent copy of  $\{X_{nk}, 1 \leq k \leq k_n, n \geq 1\}$ . Then by Lemma 2.2 in Chen and Wang ([2]), it suffices to show that

$$(2.1) \quad E \left\| \sum_{k=1}^{k_n} (X_{nk} - X'_{nk}) \right\| \rightarrow 0 \text{ as } n \rightarrow \infty.$$

It is easy to show that

$$\sum_{k=1}^{k_n} (X_{nk} - X'_{nk}) \rightarrow 0 \text{ in probability}$$

and  $\|X_{nk} - X'_{nk}\| \leq 2\delta$ . Therefore by Lemma 2.1 in Hu et al. ([7]), (2.1) holds and the proof is completed.  $\square$

LEMMA 2.2. *Let  $0 < b_n \uparrow \infty$ , and  $\{X, X_n, n \geq 1\}$  a sequence of independent and identically distributed  $B$ -valued random elements. If  $b_n^{-1} \sum_{k=1}^n X_k \rightarrow 0$  in probability, then  $E\|b_n^{-1} \sum_{k=1}^n X_k I(\|X_k\| \leq b_n)\| \rightarrow 0$  as  $n \rightarrow \infty$ .*

PROOF. Let  $\{X, X'_n, n \geq 1\}$  be an independent copy of  $\{X, X_n, n \geq 1\}$ . Then

$$(2.2) \quad b_n^{-1} \sum_{k=1}^n (X_k - X'_k) \rightarrow 0 \text{ in probability.}$$

By Lévy’s inequality (see display (2.7) in Ledoux and Talagrand [8, p. 47]), for every  $t > 0$ ,

$$P\left\{\max_{1 \leq k \leq n} \|X_k - X'_k\| > t\right\} \leq 2P\left\{\left\|\sum_{k=1}^n (X_k - X'_k)\right\| > t\right\},$$

which by (2.2) ensures that

$$(2.3) \quad P\left\{\max_{1 \leq k \leq n} \|X_k - X'_k\| > b_n/2\right\} \rightarrow 0 \text{ as } n \rightarrow \infty.$$

By Lemma 2.6 of Ledoux and Talagrand [8, p. 51],

$$(2.4) \quad \begin{aligned} nP\{\|X - X'\| > b_n/2\} &= \sum_{k=1}^n P\{\|X_k - X'_k\| > b_n/2\} \\ &\leq 2P\left\{\max_{1 \leq k \leq n} \|X_k - X'_k\| > b_n/2\right\} \end{aligned}$$

when  $n$  is sufficiently large. By display (6.1) in Ledoux and Talagrand [8, p. 150],

$$(2.5) \quad P\{\|X\| > b_n\} \leq 2P\{\|X - X'\| > b_n/2\}$$

when  $n$  is sufficiently large. Therefore by (2.3), (2.4), and (2.5),

$$(2.6) \quad nP\{\|X\| > b_n\} \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Note that for any  $\varepsilon > 0$

$$P\left\{\left\|\sum_{k=1}^n X_k I(\|X_k\| \leq b_n)\right\| > \varepsilon b_n\right\} \leq nP\{\|X\| > b_n\} + P\left\{\left\|\sum_{k=1}^n X_k\right\| > \varepsilon b_n\right\}.$$

Then by (2.6) and  $b_n^{-1} \sum_{k=1}^n X_k \rightarrow 0$  in probability, it follows that

$$b_n^{-1} \sum_{k=1}^n X_k I(\|X_k\| \leq b_n) \rightarrow 0$$

in probability. The conclusion then follows from Lemma 2.1. □

The following lemma is due to Einmahl and Li ([4]).

LEMMA 2.3. *Let  $Z_1, \dots, Z_n$  be independent  $B$ -valued random elements with mean zero such that for some  $s > 2$ ,  $E\|Z_k\|^s < \infty$ ,  $1 \leq k \leq n$ . Then we have for  $0 < \eta \leq 1$ ,  $\delta > 0$ , and  $t > 0$ ,*

$$P \left\{ \max_{1 \leq m \leq n} \left\| \sum_{k=1}^m Z_k \right\| \geq (1 + \eta) E \left\| \sum_{k=1}^n Z_k \right\| + t \right\} \leq \exp \left\{ -\frac{t^2}{(2 + \delta)\Lambda_n^2} \right\} + C \sum_{k=1}^n E\|Z_k\|^s / t^s,$$

where  $\Lambda_n^2 = \sup\{\sum_{k=1}^n E f^2(Z_k) : f \in B_1^*\}$  and  $C$  is a positive constant depending on  $\eta, \delta$  and  $s$ .

LEMMA 2.4. *Let  $h(t)$  and  $\Psi(t)$  be as in Theorem 1.2. Suppose that  $X$  is a  $B$ -valued random element with*

$$(2.7) \quad \sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > \sqrt{n \log \Psi(n)}\} < \infty.$$

Then for any  $s > 2$ ,

$$\sum_{n=1}^{\infty} \frac{1}{h(n)} \cdot \frac{1}{(n \log \Psi(n))^{s/2}} \cdot E\|X\|^s I(\|X\| \leq \sqrt{n \log \Psi(n)}) < \infty.$$

PROOF. Set  $b_0 = 0$  and  $b_n = \sqrt{n \log \Psi(n)}$ ,  $n \geq 1$ . Note that  $\Psi(n) \uparrow$  and therefore  $b_n/\sqrt{n} \uparrow$ . Then  $b_k/b_n \leq \sqrt{k/n}$  whenever  $1 \leq k \leq n$ . Hence,

$$\begin{aligned} & \sum_{n=1}^{\infty} \frac{1}{h(n)} \cdot \frac{1}{(n \log \Psi(n))^{s/2}} \cdot E\|X\|^s I(\|X\| \leq \sqrt{n \log \Psi(n)}) \\ &= \sum_{n=1}^{\infty} \frac{1}{h(n)b_n^s} \sum_{k=1}^n E\|X\|^s I(b_{k-1} < \|X\| \leq b_k) \\ &\leq \sum_{n=1}^{\infty} \frac{1}{h(n)b_n^s} \sum_{k=1}^n b_k^s P\{b_{k-1} < \|X\| \leq b_k\} \\ &= \sum_{k=1}^{\infty} b_k^s P\{b_{k-1} < \|X\| \leq b_k\} \sum_{n=k}^{\infty} \frac{1}{h(n)b_n^s} \end{aligned}$$

$$\begin{aligned}
 &\leq \sum_{k=1}^{\infty} k^{s/2} P\{b_{k-1} < \|X\| \leq b_k\} \sum_{n=k}^{\infty} \frac{1}{n^{s/2}h(n)} \\
 &\leq C \sum_{k=1}^{\infty} \frac{k}{h(k)} P\{b_{k-1} < \|X\| \leq b_k\} \\
 &\leq \frac{C}{h(1)} + C \sum_{k=1}^{\infty} \left[ \frac{k+1}{h(k+1)} - \frac{k}{h(k)} \right] P\{\|X\| > b_k\} \\
 &\leq \frac{C}{h(1)} + C \sum_{k=1}^{\infty} \frac{1}{h(k)} P\{\|X\| > b_k\} < \infty,
 \end{aligned}$$

where  $C = (s/2 - 1)^{-1}$ . The proof is completed. □

LEMMA 2.5. *Let  $h(n), \Psi(n)$  be as in Theorem 1.2. Then for any  $B$ -valued random element  $X$ , (2.7) is equivalent to*

$$(2.8) \quad \sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > M\sqrt{n \log \Psi(n)}\} < \infty$$

for some  $M > 0$ .

PROOF. It suffices to prove that (2.7) implies (2.8) for all  $0 < M < 1$ . Set  $b_n = \sqrt{n \log \Psi(n)}, n \geq 1$ . Note that  $\Psi(n) \uparrow$  and therefore  $b_n/\sqrt{n} \uparrow$ . Then  $b_n \leq 2^{-1/2}b_{2n}$  for  $n \geq 1$ . Hence,

$$\frac{1}{h(2n)} P\{\|X\| > 2^{-1/2}b_{2n}\} \leq \frac{1}{h(n)} P\{\|X\| > b_n\}$$

and

$$\begin{aligned}
 \frac{1}{h(2n+1)} P\{\|X\| > 2^{-1/2}b_{2n+1}\} &\leq \frac{1}{h(2n)} P\{\|X\| > 2^{-1/2}b_{2n}\} \\
 &\leq \frac{1}{h(n)} P\{\|X\| > b_n\},
 \end{aligned}$$

which ensures that

$$\begin{aligned}
 \sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > 2^{-1/2}b_n\} &= \frac{1}{h(1)} P\{\|X\| > 2^{-1/2}b_1\} \\
 &+ \sum_{n=1}^{\infty} \frac{1}{h(2n)} P\{\|X\| > 2^{-1/2}b_{2n}\} + \sum_{n=1}^{\infty} \frac{1}{h(2n+1)} P\{\|X\| > 2^{-1/2}b_{2n+1}\} \\
 &\leq \frac{1}{h(1)} P\{\|X\| > 2^{-1/2}b_1\} + 2 \sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > \sqrt{n \log \Psi(n)}\} < \infty.
 \end{aligned}$$

Then by mathematical induction, for any integer  $k \geq 1$ ,

$$\sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > 2^{-k/2} b_n\} < \infty.$$

The proof is completed. □

### 3. THE MAIN RESULT AND ITS PROOF

We now state and prove the main result.

**THEOREM 3.1.** *Let  $h(t)$  and  $\Psi(t)$  be as in Theorem 1.2. Let  $\{X, X_n, n \geq 1\}$  be a sequence of independent and identically distributed  $B$ -valued random elements. Suppose that*

$$(\sqrt{n \log \Psi(n)})^{-1} \sum_{k=1}^n X_k \rightarrow 0 \text{ in probability.}$$

(i) *Suppose that (2.7) holds and*

$$(3.1) \quad EX = 0, \quad Ef^2(X) < \infty \quad \forall f \in B^*.$$

*Then*

$$(3.2) \quad \sum_{n=1}^{\infty} \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n X_k \right\| > (1 + \varepsilon) \sqrt{2\sigma^2 n \log \Psi(n)} \right\} \begin{cases} < \infty, & \text{if } \varepsilon > 0 \\ = \infty, & \text{if } \varepsilon < 0, \end{cases}$$

*where  $\sigma^2 = \sup\{Ef^2(X) : f \in B_1^*\}$ .*

(ii) *Conversely, suppose that*

$$(3.3) \quad \sum_{n=1}^{\infty} \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n X_k \right\| > M \sqrt{n \log \Psi(n)} \right\} < \infty$$

*holds for some  $M > 0$ . Then (2.7) and (3.1) hold.*

**PROOF.** Set  $a_n = \sqrt{2\sigma^2 n \log \Psi(n)}$ ,  $b_n = \sqrt{n \log \Psi(n)}$ ,  $n \geq 1$  and

$$X_{nk} = X_k I(\|X_k\| \leq b_n), \quad Z_{nk} = X_{nk} - EX_{nk}, \quad 1 \leq k \leq n, \quad n \geq 1.$$

(i) Suppose that (2.7) and (3.1) hold. We first prove that

$$(3.4) \quad \sum_{n=1}^{\infty} \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n X_k \right\| > (1 + \varepsilon) a_n \right\} < \infty \quad \forall \varepsilon > 0.$$

Note that for any  $\varepsilon > 0$ ,

$$P \left\{ \left\| \sum_{k=1}^n X_k \right\| > (1 + \varepsilon) a_n \right\} \leq nP\{\|X\| > b_n\} + P \left\{ \left\| \sum_{k=1}^n X_{nk} \right\| > (1 + \varepsilon) a_n \right\}.$$

Hence, by (2.7), to prove (3.4), it suffices to prove that

$$(3.5) \quad \sum_{n=1}^{\infty} \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n X_{nk} \right\| > (1 + \varepsilon)a_n \right\} < \infty \quad \forall \varepsilon > 0.$$

By Lemma 2.2,

$$\frac{1}{b_n} \left\| \sum_{k=1}^n EX_{nk} \right\| \leq \frac{1}{b_n} E \left\| \sum_{k=1}^n X_{nk} \right\| \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

and

$$\frac{1}{b_n} E \left\| \sum_{k=1}^n Z_{nk} \right\| \leq \frac{2}{b_n} E \left\| \sum_{k=1}^n X_{nk} \right\| \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Then to prove (3.5), it suffices to prove that

$$(3.6) \quad \sum_{n=1}^{\infty} \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n Z_{nk} \right\| > 2E \left\| \sum_{k=1}^n Z_{nk} \right\| + (1 + \varepsilon)a_n \right\} < \infty \quad \forall \varepsilon > 0.$$

By Lemma 2.3, for some  $s > 2$  and any  $\delta > 0$

$$(3.7) \quad \begin{aligned} & P \left\{ \left\| \sum_{k=1}^n Z_{nk} \right\| > 2E \left\| \sum_{k=1}^n Z_{nk} \right\| + (1 + \varepsilon)a_n \right\} \\ & \leq \exp \left\{ -\frac{(1 + \varepsilon)^2 a_n^2}{(2 + \delta)\Lambda_n^2} \right\} + \frac{C}{b_n^s} \sum_{k=1}^n E \|Z_{nk}\|^s, \end{aligned}$$

where  $\Lambda_n^2 = \sup \{ \sum_{k=1}^n E f^2(Z_{nk}) : f \in B_1^* \}$ . Note that for all  $f \in B_1^*$ ,

$$\begin{aligned} E f^2(Z_{nk}) &= E f^2(X_{nk}) - (E f(X_{nk}))^2 \leq E f^2(X_{nk}) \\ &\leq E f^2(X), \quad 1 \leq k \leq n, \quad n \geq 1. \end{aligned}$$

Therefore  $\Lambda_n^2 \leq n\sigma^2, n \geq 1$ . Choose  $\delta > 0$  small enough so that  $t = 2(1 + \varepsilon)^2 / (2 + \delta) > 1$ . Then

$$(3.8) \quad \begin{aligned} \sum_{n=1}^{\infty} \frac{1}{nh(n)} \exp \left\{ -\frac{(1 + \varepsilon)^2 a_n^2}{(2 + \delta)\Lambda_n^2} \right\} &\leq \sum_{n=1}^{\infty} \frac{1}{nh(n)} \exp \left\{ -\frac{(1 + \varepsilon)^2 a_n^2}{(2 + \delta)\Lambda_n^2} \right\} \\ &\leq \sum_{n=1}^{\infty} \frac{1}{nh(n)} \exp \{ -t \log \Psi(n) \} \\ &\leq \sum_{n=1}^{\infty} \frac{1}{nh(n)} \cdot \frac{1}{(\Psi(n))^t} < \infty, \end{aligned}$$

since  $\int_1^\infty dx/[xh(x)\Psi^t(x)] < \infty$ . By the  $C_r$ -inequality, Hölder's inequality, and Lemma 2.4,

$$(3.9) \quad \sum_{n=1}^\infty \frac{1}{nh(n)} \cdot \frac{1}{b_n^s} \sum_{k=1}^n E\|Z_{nk}\|^s \leq \sum_{n=1}^\infty \frac{1}{h(n)} \cdot \frac{1}{(n \log \Psi(n))^{s/2}} \cdot E\|X\|^s I(\|X\| \leq \sqrt{n \log \Psi(n)}) < \infty.$$

By (3.7), (3.8), and (3.9), (3.6) holds and hence (3.4) holds as was argued above.

Now we prove that

$$(3.10) \quad \sum_{n=1}^\infty \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n X_k \right\| > (1 + \varepsilon)a_n \right\} = \infty \quad \forall \varepsilon < 0.$$

For any  $f \in B^*$ , by (3.1),  $Ef(X) = 0$  and  $Ef^2(X) < \infty$ . Then by the implication “(1.1)  $\Rightarrow$  (1.4)” in Theorem 1.2, for all  $\varepsilon < 0$

$$(3.11) \quad \sum_{n=1}^\infty \frac{1}{nh(n)} P \left\{ \left| \sum_{k=1}^n f(X_k) \right| > (1 + \varepsilon)\sqrt{2Ef^2(X)n \log \Psi(n)} \right\} = \infty.$$

Note that for any  $f \in B_1^*$ ,  $|\sum_{k=1}^n f(X_k)| \leq \|\sum_{k=1}^n X_k\|$  and so it follows from (3.11) that for all  $f \in B_1^*$ , for all  $\varepsilon < 0$

$$(3.12) \quad \sum_{n=1}^\infty \frac{1}{nh(n)} P \left\{ \left\| \sum_{k=1}^n X_k \right\| > (1 + \varepsilon)\sqrt{2Ef^2(X)n \log \Psi(n)} \right\} = \infty.$$

Hence (3.10) holds by (3.12) and  $\sigma^2 = \sup\{Ef^2(X) : f \in B_1^*\}$ . Combining (3.4) and (3.10) yields (3.2).

(ii) Assume that (3.3) holds for some  $M > 0$ . Then for any  $f \in B_1^*$ ,

$$\sum_{n=1}^\infty \frac{1}{nh(n)} P \left\{ \left| \sum_{k=1}^n f(X_k) \right| > Mb_n \right\} < \infty.$$

Then by the implication “(2.3)  $\Rightarrow$  (2.4)” of Li and Rosalsky ([10]), it follows that  $Ef(X) = 0$  and  $Ef^2(X) < \infty$ . Hence (3.1) holds.

Let  $\{X', X'_n, n \geq 1\}$  be an independent copy of  $\{X, X_n, n \geq 1\}$ . Then by the same argument as in the proof of Lemma 2.2,

$$\begin{aligned} nP\{\|X\| > 4Mb_n\} &\leq 8P \left\{ \left\| \sum_{k=1}^n (X_k - X'_k) \right\| > 2Mb_n \right\} \\ &\leq 16P \left\{ \left\| \sum_{k=1}^n X_k \right\| > Mb_n \right\}, \end{aligned}$$



which by (3.3) ensures that

$$\sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > 4Mb_n\} < \infty$$

and so (2.7) holds by Lemma 2.5. The proof is completed. □

REMARK 3.2. A sufficient condition for (2.7) is  $E\|X\|^2 < \infty$ . Indeed,

$$\begin{aligned} \sum_{n=1}^{\infty} \frac{1}{h(n)} P\{\|X\| > \sqrt{n \log \Psi(n)}\} &\leq \frac{1}{h(1)} \sum_{n=1}^{\infty} P\{\|X\| > \sqrt{n}\} \\ &\leq \frac{1}{h(1)} E\|X\|^2 < \infty. \end{aligned}$$

REMARK 3.3. Some examples of moment conditions which are equivalent to (2.7) for various choices of  $h(\cdot)$  will now be given.

CASE (i). Set  $h(t) = (\log \log t)^b$  where  $b \geq 0$ . Then  $\log \Psi(t) \sim \log \log t$  as  $t \rightarrow \infty$  and (2.7) is equivalent to  $E\|X\|^2 / (\log \log \|X\|)^{b+1} < \infty$ .

CASE (ii). Set  $h(t) = (\log t)^r$  where  $0 \leq r < 1$ . Then  $\log \Psi(t) \sim (r - 1) \log \log t$  as  $t \rightarrow \infty$  and (2.7) is equivalent to  $E\|X\|^2 / [(\log \|X\|)^r \log \log \|X\|] < \infty$ .

CASE (iii). Set  $h(t) = \log t$ . Then  $\log \Psi(t) \sim \log \log \log t$  as  $t \rightarrow \infty$  and (2.7) is equivalent to  $E\|X\|^2 / [(\log \|X\|) \log \log \log \|X\|] < \infty$ .

CASE (iv). In Case (i), take  $b = 0$ , or in Case (ii), take  $r = 0$ . Then (2.7) is equivalent to  $E\|X\|^2 / \log \log \|X\| < \infty$ .

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