



University of Pennsylvania  
ScholarlyCommons

---

Statistics Papers

Wharton Faculty Research

---

1964

## A Limit Law Concerning Moving Averages

Larry A. Shepp  
*University of Pennsylvania*

Follow this and additional works at: [http://repository.upenn.edu/statistics\\_papers](http://repository.upenn.edu/statistics_papers)



Part of the [Applied Statistics Commons](#)

---

### Recommended Citation

Shepp, L. A. (1964). A Limit Law Concerning Moving Averages. *The Annals of Mathematical Statistics*, 35 (1), 424-428.  
<http://dx.doi.org/10.1214/aoms/1177703767>

---

This paper is posted at ScholarlyCommons. [http://repository.upenn.edu/statistics\\_papers/342](http://repository.upenn.edu/statistics_papers/342)  
For more information, please contact [repository@pobox.upenn.edu](mailto:repository@pobox.upenn.edu).

---

# A Limit Law Concerning Moving Averages

**Disciplines**

Applied Statistics

# A LIMIT LAW CONCERNING MOVING AVERAGES<sup>1</sup>

BY L. A. SHEPP

*University of California, Berkeley and Bell Telephone Laboratories*

**0. Introduction.** This note deals with the limiting values of the averages,  $T_n = (S_{n+f(n)} - S_n)/f(n)$  where  $S_n = X_1 + \dots + X_n$ ,  $n = 1, 2, \dots$ , is a sum of mutually independent and identically distributed random variables. The function  $f$  takes positive integer values and nondecreasingly tends to infinity. We shall determine the almost everywhere constant  $\limsup T_n$  in terms of the moment generating function of  $X_1$  and the radius of convergence of  $\sum x^{f(n)}$ , denoted  $r$ .

If we define  $N_n$  as the number of consecutive successes beginning at trial  $n+1$  in a sequence of Bernoulli trials with success probability  $p$ , we let  $U_n = N_n/f(n)$ . It is shown that, almost surely,  $\limsup U_n = \log_p r$ . This result will be compared with  $\limsup T_n$  for the case of Bernoulli trials. The author is indebted to Professor D. J. Newman for drawing his attention to this problem.

**1. Statement of results.** The events  $\{T_n > c \text{ infinitely often (i.o.)}\}$  are tail events,  $-\infty < c < \infty$ , and so  $T = \limsup T_n$  is constant with probability one ([3], p. 229).

We form the series  $\sum x^{f(n)} = \sum A_k x^k$  and determine its radius of convergence  $r$ ,  $0 \leq r \leq 1$ . The constant  $T = T(f)$  will depend only on  $r$ , and when this is to be emphasized, will be denoted  $T_r$ .

Let  $M(t) = Ee^{X_1 t} \leq +\infty$  be defined for all  $t$ , and set [1], for any  $a$ ,  $m(a) = \min M(t)e^{-at} \leq M(0) = 1$ . If  $M(t) < \infty$  for some  $t > 0$ , then  $EX_1^+ < \infty$ , and so  $EX_1$  is well defined,  $-\infty \leq EX_1 < \infty$ . Let  $\text{ess sup } X_1 = A \leq +\infty$ . We assume  $X_1$  is nonconstant a.s., so  $EX_1 < A$ .

**THEOREM 1.** *If  $M(t) < \infty$  for some  $t > 0$ , then*

$$(1.1) \quad T_r = A \quad \text{for } 0 \leq r \leq P\{X_1 = A\}.$$

$$(1.2) \quad T_r = a_r \quad \text{for } P\{X_1 = A\} < r < 1$$

where  $a = a_r$  is the unique solution (inverse) of  $m(a) = r$ ,  $EX_1 < a < A$ .

$$(1.3) \quad T_r = EX_1 \quad \text{for } r = 1.$$

Moreover,  $T_r$  decreases strictly and continuously for  $P\{X_1 = A\} \leq r \leq 1$ .

The consideration of the sequence  $-X_i$  leads to a similar theorem for  $\liminf T_n$ . If  $f$  is strictly increasing then  $r = 1$ . The existence of  $\lim T_n$  is assured if and only if  $r = 1$ , whence  $\lim T_n = EX_1$  a.s.,  $-\infty < EX_1 < \infty$ . Proofs of these assertions are in Section 3. Our methods break down in case  $M(t) = \infty$  for all  $t > 0$ , and this case is not discussed.

Received 27 September 1961; revised 11 September 1963.

<sup>1</sup> Prepared with partial support of National Science Foundation Grant G-14648.

An immediate corollary of this theorem is that if  $f_1$  and  $f_2$  are two functions, as above, and  $f_1(n) \leq f_2(n)$ ,  $n = 1, 2, \dots$ , then  $T(f_2) \leq T(f_1)$ . A class of  $f$  in central position are  $f_c(n) = 1 + [c \log n]$ ,  $c > 0$ ,  $n = 1, 2, \dots$ , whence  $r = \exp(-1/c)$ .

**2. Chernoff's theorem.** It is easily seen that  $\log P\{S_n \geq na\}$  is subadditive in  $n$  for fixed  $a$ , and hence  $\lim - (1/n) \log P\{S_n \geq na\}$  exists. Chernoff [1], has found this limit in terms of  $m(a)$ . We suppress the subscript of  $X_1$  in this discussion.

**THEOREM 2.** (Chernoff) *We suppose  $M(t) < \infty$  for some  $t > 0$ . We have  $-\infty \leq EX < \infty$ , and  $m(a) = \min_{t \geq 0} Ee^{(X-a)t} < 1$ , if  $a > EX$ . Moreover, for  $n = 1, 2, \dots$*

$$(2.1) \quad P\{S_n \geq na\} \leq (m(a))^n,$$

and for  $0 < \epsilon < m(a)$ , and  $n > n_0(\epsilon)$ ,

$$(2.2) \quad P\{S_n \geq na\} > (m(a) - \epsilon)^n.$$

This is a slight variant of [1], and we should complete the proof in [1] by noting that for  $0 < t < \delta$ , some  $\delta > 0$ ,  $r(t) = M(t)e^{-at}$  is convex, and so  $r'(t) = E(X - a)e^{(X-a)t}$  decreases to  $E(X - a) < 0$  as  $t$  decreases to zero (*dominated convergence*). It follows that  $r'(t) < 0$ ,  $0 < t < \delta_1$ , some  $\delta_1 > 0$ , and so  $r(t) = r(0^+) + \int_0^t r'(u) du < r(0^+) = 1$ ,  $0 < t < \delta_1$ , whence  $m(a) < 1$ . To prove the validity of the statement  $m(a) = \min_{t \geq 0} Ee^{(X-a)t}$  note that this is trivial if  $M(t) = \infty$  for all  $t < 0$ , and if  $M(t) < \infty$  for some  $t < 0$  then  $r'(0) = EX - a < 0$  and so the minimum occurs to the right of zero, by convexity.

If  $A = \text{ess sup } X = \infty$ , set  $m(\infty) = 0$ ; if  $EX = -\infty$ , set  $m(-\infty) = 1$ . The function  $m(a)$  is then defined for extended real numbers in the interval  $EX \leq a \leq A$ .

(2.3) **LEMMA.** *If  $M(t) < \infty$  for some  $t > 0$ , then for  $EX \leq a \leq A$ ,  $m(a)$  is continuous and strictly monotonic. Moreover,  $m(EX) = 1$ .*

**PROOF.** For  $h > 0$  and  $EX < a < a + h < A$ , we have by the first part of Theorem 2,

$$(2.4) \quad m(a + h) = \min_{t \geq 0} Ee^{(X-a)t} e^{-ht} < m(a).$$

Continuity of  $m(a)$  at finite  $a$ ,  $EX \leq a < A$  is easily shown using the fact that  $m(a) = Ee^{(X-a)t_1}$  for  $0 \leq t_1 < \infty$ . If  $A < \infty$ ,  $\epsilon > 0$ , we set  $t = \epsilon^{-\frac{1}{2}}$ , obtaining

$$(2.5) \quad m(A - 0) = \lim_{\epsilon \downarrow 0} \min_{t \geq 0} Ee^{(X-A+\epsilon)t} \leq \lim_{\epsilon \downarrow 0} E \exp [(X - A)\epsilon^{-\frac{1}{2}} + \epsilon^{\frac{1}{2}}].$$

It follows that  $m(A - 0) \leq P\{X = A\} = m(A)$ . However, for  $\epsilon > 0$ ,

$$(2.6) \quad m(A - \epsilon) \geq \min_{t \geq 0} Ee^{(X-A)t} \min_{t \geq 0} e^{\epsilon t} = m(A).$$

Thus  $m$  is continuous at  $A$ , if  $A < \infty$ . If  $A = \infty$ , then for some fixed  $t > 0$ ,

$$(2.7) \quad 0 \leq \lim_{a \rightarrow \infty} m(a) \leq \lim_{a \rightarrow \infty} Ee^{(X-a)t} = 0 = m(\infty).$$

If  $EX = -\infty$ , we have  $1 \geq \lim_{a \rightarrow -\infty} m(a) \geq \lim_{a \rightarrow -\infty} P\{X \geq a\} = 1 = m(-\infty)$ . This proves the continuity assertions. If  $EX > -\infty$ , then  $M'(0^+) = EX$ , and  $m(EX) = 1$  follows. This proves (2.3).

**3. Proof of Theorem 1.** We shall need the following simple lemma for power series. Here  $f$  is a function of the type described in the introduction.

(3.1) LEMMA. *If  $0 < x < y < 1, b > 0$  and  $\sum x^{f(n)} < \infty$ , then  $\sum y^{f(n_k)} = \infty$  where  $[bf(n_1)] \geq 1$  and for  $k = 1, 2, \dots$ ,*

$$(3.2) \quad n_{k+1} = n_k + [bf(n_k)].$$

PROOF. Choose  $K$  so that for  $n \geq n_K$ ,

$$(3.3) \quad (x/y)^{f(n)} f(n) < 1.$$

Since  $f$  is nondecreasing, we have for  $k = 1, 2, \dots$ ,

$$(3.4) \quad \sum_{j=0}^{n_{k+1}-n_k-1} y^{f(n_k+j)} / f(n_k + j) \leq y^{f(n_k)} (n_{k+1} - n_k) / f(n_k) \leq b y^{f(n_k)}.$$

Using (3.3) and the fact that  $n_k \rightarrow \infty$  (since  $[bf(n_k)] \geq 1$  in (3.2)), we get

$$(3.5) \quad \sum_{k=K}^{\infty} y^{f(n_k)} \geq b^{-1} \sum_{k=K}^{\infty} \sum_{j=n_k}^{n_{k+1}-1} y^{f(j)} / f(j) \geq b^{-1} \sum_{n \geq n_K} x^{f(n)} = \infty,$$

proving (3.1).

To prove Theorem 1, let  $A_n(a)$  be the event  $\{T_n \geq a\}$ . Since we are dealing with identically distributed random variables, we have

$$(3.6) \quad P\{A_n(a)\} = P\{S_{f(n)} \geq af(n)\}.$$

In proving (1.1), we assume first  $A < \infty$ . If  $r \leq m(A)$ , then for  $h > 0$ , sufficiently small,  $m(A-h) > r$  and we choose  $\epsilon > 0$  so that  $\alpha = m(A-h) - 2\epsilon > r$ . Since  $\sum \alpha^{f(n)} = \infty$ , we may apply (3.1) with  $b = 1, x = \alpha, y = \alpha + \epsilon$ . For the sequence  $n_k$  of (3.1), we have  $\sum y^{f(n_k)} = \infty$ . By (2.2) and (3.6),  $P\{A_n(A-h)\} \geq y^{f(n)}$ , for  $n$  sufficiently large. It follows that  $\sum P\{A_{n_k}(A-h)\} = \infty$ . Since  $A_{n_k}$  depends only on that segment of  $\{X_n\}$  for which  $n_k < n \leq n_k + f(n_k) = n_{k+1}$ , and these segments are disjoint, the events  $A_{n_k}$  are independent,  $k = 1, 2, \dots$ . By Borel's lemma ([3], p. 228),  $P\{A_{n_k}(A-h)\text{i.o.}\} = 1$  and so  $T \geq A-h$ . Since  $T_n \leq A$ , a.s. for each  $n$ , we obtain  $T = A$ . If now  $A = \infty$  and  $r \leq m(A)$ , then  $r = 0$ . Since  $m(a) > 0$  for finite  $a > EX_1$  we have for  $2\epsilon < m(a)$ ,  $\sum P\{A_n(a)\} \geq \sum (m(a) - \epsilon)^{f(n)} = \infty$  and arguing as above,  $P\{A_{n_k}(a)\text{ i.o.}\} = 1$ . Hence  $T \geq a$  for every  $a$ , and  $T = \infty = A$ . This proves (1.1).

If  $m(A) < r < 1$ , then there is a unique  $a, EX_1 < a < A$ , for which  $m(a) = r$  by (2.3). We shall show  $T_r = a$ . Let  $h > 0$  be sufficiently small in the following. Using (2.1) we have, since  $m(a+h) < m(a) = r$ ,

$$(3.7) \quad \sum P\{A_n(a+h)\} \leq \sum (m(a+h))^{f(n)} < \infty.$$

By Cantelli's lemma ([3], p. 228) we have  $T_r \leq a + h$  so  $T_r \leq a$ . On the other hand, by (2.2),  $P\{A_n(a - h)\} \geq (m(a - h) - \epsilon)^{f(n)}$  for  $0 < \epsilon < m(a - h)$ . If  $\epsilon$  is so small that  $m(a - h) - 2\epsilon > m(a) = r$ , then  $\sum (m(a - h) - 2\epsilon)^{f(n)} = \infty$ . By (3.1), again with  $b = 1$ , we obtain a sequence  $n_k$  for which  $\sum P\{A_{n_k}(a - h)\} \geq \sum (m(a - h) - \epsilon)^{f(n_k)} = \infty$ . As before we conclude that  $T_r \geq a - h$  and so  $T_r = a$  and (1.2) is proved. To prove (1.3) assume that  $y > EX_1$ . By Theorem 2,  $m(y) < 1$  and so

$$(3.8) \quad \sum P\{A_n(y)\} \leq m(y)^{f(n)} < \infty,$$

since  $r = 1$ . It follows that  $T_1 \leq EX_1$ . The weak law of large numbers, ([2], p. 228) shows that for  $E|X_1| < \infty$  and  $\epsilon > 0$ ,  $P\{S_n > (EX_1 - \epsilon)n\} \rightarrow 1$ . Thus  $P\{A_n(EX_1 - \epsilon)\} \rightarrow 1$  and so for  $0 \leq r \leq 1$ ,  $T_r \geq EX_1$ . We have  $T_1 = EX_1$  in any case and (1.3) is proved. In this last case we have  $\lim T_n$  existing, since

$$(3.9) \quad \liminf T_n = -\limsup -T_n = -E(-X_1) = EX_1 = \limsup T_n$$

almost surely. In (3.9) we have tacitly assumed that  $M(t) < \infty$  also for some  $t < 0$  and so  $E|X_1| < \infty$  here. Thus Theorem 1 and the assertions following are proved.

**4. An example.** As an illustration of Theorem 1 we consider the Bernoulli case,  $P\{X_1 = 1\} = p = 1 - P\{X_1 = 0\}$  for  $0 < p < 1$ . We have for  $p \leq a \leq 1$ ,

$$(4.1) \quad m(a) = (p/a)^a((1-p)/(1-a))^{1-a}$$

and thus, for  $p \leq r \leq 1$ ,  $T_r = a$  where  $m(a) = r$ , and for  $0 \leq r \leq p$ ,  $T_r = 1$ .

**5. A related problem on success runs.** If we set  $N_n = k$  if  $X_{n+1} = \dots = X_{n+k-1} = 1$  and  $X_{n+k} = 0$ ;  $k = 1, 2, \dots$ , in the example, then  $N_n$  counts the waiting time until the first failure following the  $n$ th Bernoulli trial (success runs [2] p. 197, Example 5, D. J. Newman). If  $f$  is a function of the type considered above, we set  $U_n = N_n/f(n)$ ,  $n = 1, 2, \dots$  and  $U = \limsup U_n$ . It will be shown that  $U$  is a constant a.s., depending on  $f$  only through  $r$ , the radius of convergence of  $\sum x^{f(n)}$ .

(5.1) **COROLLARY.** For  $0 \leq r \leq 1$ ,  $U = \limsup U_n = \log_p r$ .

**PROOF.** Let  $B_n(b)$  denote the event  $\{U_n > b\}$ ,  $0 < b$  and  $n = 1, 2, \dots$ . We have  $P\{B_n(b)\} = p^{[bf(n)]+1}$ , and if  $p^b < r$ ,

$$(5.2) \quad \sum P\{B_n(b)\} \leq \sum p^{bf(n)} < \infty.$$

It follows that  $P\{B_n(b) \text{ i.o.}\} = 0$  and that  $U \leq \log_p r$ . If  $p^b > r$ , then finding an  $\epsilon > 0$  for which  $p^{b+\epsilon} > r$ ,  $\sum p^{(b+\epsilon)f(n)} = \infty$ . Using (3.1) with  $x = p^{b+\epsilon}$ ,  $y = p^b$  we obtain a sequence  $n_k$  for which

$$(5.3) \quad \sum p^{bf(n_k)} = \infty$$

$$(5.4) \quad n_{k+1} = n_k + [bf(n_k)].$$

The event  $B_{n_k}(b)$  refers to the section of  $\{X_n\}$  for  $n_k < n \leq n_k + [bf(n_k)] = n_{k+1}$

and thus  $B_{n_k}(b)$  are independent,  $k = 1, 2, \dots$ . Moreover,  $\sum P\{B_{n_k}(b)\} \geq p \sum p^{bf(n_k)} = \infty$  by (5.3). It follows that  $U \geq b$  and so  $U \geq \log_p r$ . This proves (5.1).

Comparing  $T$  and  $U$  in the Bernoulli case, it is easily seen that whenever  $U < 1$ , we have  $U_n \leq T_n$  eventually and so  $U \leq T \leq 1$ . If  $T = a$  and  $U = b$  we have the relationship, from (4.1) and (5.1)

$$(5.5) \quad (p/a)^a((1-p)/(1-a))^{1-a} = p^b = r$$

for  $p \leq a \leq 1$  and  $0 \leq b \leq 1$ .

#### REFERENCES

- [1] CHERNOFF, H. (1952). A measure of asymptotic efficiency for tests of a hypothesis based on the sum of observations. *Ann. Math. Statist.* **23** 493–507.
- [2] FELLER, W. (1957). *An Introduction to Probability Theory and Its Applications*, 1 (2nd ed.). Wiley, New York.
- [3] LOÈVE, M. (1955). *Probability Theory*. Van Nostrand, New York.