

Lee, W; Bell, ML; Gasparrini, A; Armstrong, BG; Sera, F; Hwang, S; Lavigne, E; Zanobetti, A; Coelho, MSZS; Saldiva, PHN; Osorio, S; Tobias, A; Zeka, A; Goodman, PG; Forsberg, B; Rocklv, J; Hashizume, M; Honda, Y; Guo, YL; Seposo, X; Van Dung, D; Dang, TN; Tong, S; Guo, Y; Kim, H (2017) Mortality burden of diurnal temperature range and its temporal changes: A multi-country study. Environment international. ISSN 0160-4120 DOI: https://doi.org/10.1016/j.envint.2017.10.018

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Title: Mortality Burden of Diurnal Temperature Range and Its Temporal Changes: A Multi-Country Study

3 Author names and Affiliations:

4 Whanhee Lee1, Michelle L. Bell2, Antonio Gasparrini3, Ben G. Armstrong3, Francesco Sera3,

5 Sunghee Hwang1, Eric Lavigne4, Antonella Zanobetti5, Micheline de Sousa Zanotti Stagliorio

6 Coelho6, Paulo Hilario Nascimento Saldiva6, Samuel Osorio7, Aurelio Tobias8, Ariana Zeka9,

7 Patrick G. Goodman10, Bertil Forsberg11, Joacim Rocklov11, Masahiro Hashizume12,

8 Yasushi Honda13, Yue-Liang Leon Guo14, Xerxes Seposo13, Do Van Dung15, Tran Ngoc

9 Dang13, Shilu Tong16, Yuming Guo17, and Ho Kim1,*

¹ Graduate School of Public Health, Seoul National University, Seoul, Republic of Korea.

² School of Forestry and Environmental Studies, Yale University, New Haven, Connecticut,
USA.

³ Department of Social and Environmental Health Research, London School of Hygiene and
 Tropical Medicine, London, United Kingdom.

⁴ School of Epidemiology, Public Health and Preventive Medicine, University of Ottawa,
Ottawa, Canada.

⁵ Department of Environmental Health, Harvard T.H. Chan School of Public Health, Boston,
Massachusetts, USA.

⁶ Laboratory of Experimental Air Pollution, Department of Pathology, School of Medicine,
University of Sao Paulo, Sao Paulo, Brazil.

²¹ ⁷ University of Los Andes, Colombia.

⁸ Institute of Environmental Assessment and Water Research, Spanish Council for Scientific

| 23 | Research, | Barcel | lona, | Spain. |
|----|-----------|--------|-------|--------|
|----|-----------|--------|-------|--------|

| 24 | ⁹ Sciences Institute, Dublin Institute of Technology, Dublin, Ireland. |
|----|------------------------------------------------------------------------------------------------------|
| 25 | ¹⁰ Institute for the Environment, Brunel University London, London, UK. |
| 26 | ¹¹ Department of Public Health and Clinical Medicine, Ume University, Ume, Sweden. |
| 27 | ¹² Department of Pediatric Infectious Diseases, Institute of Tropical Medicine, Nagasaki |
| 28 | University, Nagasaki, Japan. |
| 29 | ¹³ Faculty of Health and Sport Sciences, University of Tsukuba, Tsukuba, Japan. |
| 30 | ¹⁴ Department of Environmental and Occupational Medicine, National Taiwan University, |
| 31 | Taipei, Taiwan. |
| 32 | ¹⁵ Department of Medical Statistics, University of Medicine and Pharmacy at Ho Chi Minh |
| 33 | City, Ho Chi Minh City, Vietnam. |
| 34 | ¹⁶ School of Population Health, Institute of Health and Biomedical Innovation, Queensland |
| 35 | University of Technology, Brisbane, Australia. |
| 36 | ¹⁷ School of Population Health, University of Queensland, Brisbane, Australia. |
| 37 | |
| 38 | Corresponding author: Ho Kim, Graduate School of Public Health, Seoul National University, |
| 39 | 1 Gwanak-ro, Gwanak-gu, Seoul 151-742, Republic of Korea. Tel: (82) 2 880-2702. E-mail: |
| 40 | hokim@snu.ac.kr |
| 41 | Manuscript Types: Research Papers |

42 ABSTRACT

Although diurnal temperature range (DTR) is a key index of climate change, far few studies 43 have reported the health burden of DTR and its temporal changes at a multi-country scale. 44 Therefore, we assessed the attributable risk fraction of DTR on mortality and its temporal 45 variations in a multi-country data set. We collected time-series data covering mortality and 46 weather variables from 308 cities in 10 countries from 1972 to 2013. The temporal change in 47 DTR-related mortality was estimated for each city with a time-varying distributed lag model. 48 Estimates of each city were pooled using a multivariate meta-analysis. The results showed that 49 the attributable fraction of total mortality to DTR was 2.5% (95% eCI: 2.3-2.7%) over the entire 50 study period. In overall countries, the attributable fraction has increased from 2.4% (2.1-2.7%) 51 to 2.7% (2.4-2.9%) between the first and last study years. This study found that DTR has 52 significantly attributed to mortality in overall countries, and this attributable fraction has 53 54 significantly increased overtime in the USA, UK, Spain, and South Korea. Therefore, because the health burden of DTR is likely to increase in future, countermeasures are needed against 55 the increase. 56

- 57 Keywords: Diurnal temperature range, Attributable mortality risk fraction, Time-varying
 58 Effect, Climate Change.
- 59 Abbreviations: Attributable risk fraction (ARF), Distributed Lag Non-linear Model (DLNM).
- 60 Funding: This study was funded by the Global Research Lab (#K21004000001-10A0500-
- 61 00710) through the National Research Foundation of Korea.

62 **1. Introduction**

Diurnal temperature range (DTR, i.e., the intra-day temperature change) is a well-known risk 63 factor of weather-related human health. Numerous studies have described a positive association 64 between DTR and mortality (Cao et al. 2009; Lim et al. 2015; Tam et al. 2009; Vutcovici et al. 65 2014; Yang et al. 2013a), and have reported that people who are elderly, less educated, female 66 or have cardiovascular or respiratory disease are more susceptible to DTR than others (Kan et 67 al. 2007b; Lim et al. 2012a; Yang et al. 2013b). In addition, because the DTR has been reported 68 69 as an important meteorological indicator closely related with global climate change (Braganza 70 et al. 2004; Kan et al. 2007b; Yang et al. 2013b), an in-depth investigation of the DTR-mortality relationship becomes important as it helps to assess the future health impact of climate change 71 72 more comprehensively.

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Biological mechanisms through which a sudden change in absolute temperature might affect 74 mortality have been described in previous medical and epidemiological studies (Garrett et al. 75 76 2009; Garrett et al. 2011; Greenberg et al. 1983; Keatinge et al. 1984a; Martinez-Nicolas et al. 77 2015; Qiu et al. 2013). Sudden changes in within-day temperature may cause physiological health problems (Garrett et al. 2009; Garrett et al. 2011); unstable weather or temperature 78 changes can lead to the onset of cardiovascular events brought on by increased workload. This 79 can affect the respiratory system by triggering inflammatory nasal responses (Ballester et al. 80 1997; Carder et al. 2005; Graudenz et al. 2006; Hashimoto et al. 2004; Imai et al. 1999; Luurila 81 1980). These mechanisms have been suggested as potential causes of increasing human 82 mortality (Buguet 2007; Guo et al. 2016). 83

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Based on this biological evidence, previous studies have tried to estimate the risk of DTR on mortality (Lim et al. 2015; Tam et al. 2009; Vutcovici et al. 2014). However, most previous studies assessed the risk of DTR using only terms of relative risk (RR), not attributable risk fraction which can quantify the mortality burden. Furthermore, because a majority of the previous studies were conducted in single cities or single countries and used statistically different methods (Kan et al. 2007b; Lim et al. 2012a; Yang et al. 2013b), results of these studies might have limited applicability to a multi-country scale.

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93 Most previous studies estimated the risk of DTR on mortality using historical data (Kan et al. 94 2007b; Lim et al. 2012a) and the estimated impact of DTR was assumed to be consistent overtime. However, this assumption might not be suitable in predicting the health impacts of 95 96 climate change because several factors, including intrinsic biological (e.g., disease/nutrition 97 status) and extrinsic factors (e.g., forecast and infrastructure improvements, local environment, or social system conditions), can modify the population's vulnerability to absolute temperature 98 99 and rapid temperature change within a day (Gasparrini et al. 2015a; Linares et al. 2014; Wu et al. 2014). Therefore, it is important to assess temporal change in the DTR-related mortality 100 relationship to examine whether people are adapted or mal-adapted to DTR. 101

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In this study, we assess the percent increases in risks and the attributable risk fraction of DTR
for 308 cities of 10 countries. We examine whether the excessive risks and attributable risk
fractions have changed during the study period. We used a Multi-Country Multi-City (MCC)
Collaborative Network to assess the impacts of weather on mortality using a multi-country data
set as referenced in previous papers (Gasparrini et al. 2015a; Gasparrini et al. 2016; Guo et al.
2014; Guo et al. 2016).

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110 2. Material and methods

111 **2.1. Data**

Time-series data covering mortality and weather variables were collected from 385 locations 112 in 10 countries: Canada (26 cities, 1986-2011), United States (USA) (135 cities, 1985-2006), 113 Brazil (18 cities, 1997-2011), Colombia (5 cities, 1998-2013), United Kingdom (UK) (10 114 regions, 1990-2012), Ireland (6 regions, 1984-2007), Spain (51 cities, 1990-2010), Japan (47 115 prefectures, 1972-2012), South Korea (7 cities, 1992-2010), and Australia (3 cities, 1988-2009). 116 117 For convenience of interpretation, the location is described as "city" in this study. The daily mortality count is the daily count of death for all causes. If a daily count of all causes of death 118 was not available for a city, then death for non-external causes (ICD-9: 0-799, ICD-10: A00-119 R99) was used instead. DTR was chosen as the exposure index, computed from monitoring 120 stations as the difference between the daily maximum and daily minimum temperatures. 121 Detailed information regarding data collection is provided in the Supplementary Material (Data 122 details). 123

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125 **2.2. First-stage time series model**

The first-stage time series model was divided into a two-step procedure. First, a time-series regression was applied, based on a generalized linear model using a quasi-Poisson distribution with parameters for: DTR, the day of week, seasonal long-term trend, inter-day temperature change (the change in mean temperature between two neighboring days), and absolute temperature. We modeled the DTR-response curve with a linear function and the lag-response curve with two internal knots placed at equally spaced values on a log scale using natural cubic 132 B-spline with 14 days of lag. Inter-day temperature change was adjusted in the same way as DTR. We also modeled the temperature-response relationship using a quadratic B-spline with 133 three internal knots (placed at the 10th, 75th, and 90th percentiles of location-specific 134 temperature distributions) and a lag-response (up to 21 days) curve with natural cubic B-spline 135 with three internal knots placed at equally spaced values on the log scale. This model approach 136 was used in a previous multi-country temperature-mortality study using a distributed lag non-137 linear model (DLNM) (Gasparrini et al. 2010; Gasparrini et al. 2015b). Seasonal trends were 138 139 adjusted using a natural cubic B-spline of time with 8 degrees of freedom (df) per year, and day of week was included as an indicator variable. Results of the first stage estimate the 140 association between DTR and mortality for each city. 141

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143 2.3. Time varying distributed lag non-liner model

The DLNMs, described in the first-stage analysis, assumed that the exposure-lag-response 144 associations between DTR and mortality in each location were constant across the whole study 145 146 period. We also applied a time-varying DLNM with a linear interaction (Gasparrini et al. 2015a; 147 Gasparrini et al. 2016) between DTR and year. Using the time-varying DLNM, we derived coefficients representing the exposure-lag-response association for the first and last year of the 148 149 study periods for each city. The set of four coefficients (the entire period, the first and the last year for each location) were reduced to one coefficient that modeled the overall cumulative 150 associations between DTR and mortality. The sets of four coefficients were used to determine 151 the lag-response relationships at the 99th percentile of DTR reference at $0 \square C$ DTR. 152

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154 2.4. Second stage meta-analysis

155 We pooled one parameter of the overall cumulative exposure-response relationship and the four parameters of the lag-response relation. Multivariate random-effect meta-regression was used 156 to pool the parameters by country. We used indicators of country as predictors in the meta-157 regression to country-pooled estimates and city-specific predicted parameters (Best Linear 158 Unbiased Prediction, BLUP). Overall pooled coefficient (only for calculating excessive 159 relative risk of overall countries) was estimated by meta-analysis without predictors. All 160 161 analyses were performed using R software (version 3.3.1) packages dlnm and mvmeta (Gasparrini 2011; Gasparrini et al. 2012; Gasparrini et al. 2010). 162

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164 **2.5. Attributable mortality risk faction**

Overall cumulative relative risk estimated from BLUP for each city was used to compute the 165 166 attributed number of deaths, and the fraction of deaths over the following 14 days at each location. The total number of deaths attributed to DTR was calculated as the sum of all days in 167 the series when DTR contributed to death and its ratio with the total number of deaths; this 168 169 provides the 'total attributable fraction' (Gasparrini and Leone 2014). We also computed the 170 time-varying attributable risk of DTR based on BLUP for each city from the time-varying DLNM. Although time-dependent distributions of DTR and death could be used to estimate 171 172 time-varying attributable risk, we used DTR and mortality distribution for the entire period because we did not find a clear difference between DTR distributions for the first and last three 173 years of the series for each city (Table 1). 174

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176 **2.6. Sensitivity analysis**

177 In order to test the sensitivity of our results to the modeling parameters and assumptions

described above, we changed lag days for DTR (21 days), inter-day temperature changes (10 and 21 days), and the degrees of freedom (df) of lag knots for DTR (df=5), and analyzed the first results. We also assessed sensitivity to controlling for humidity (only for 6 countries which include relative humidity data), air pollution (Korea, O3 and PM10), flexibility of long-term trend (df=7 and 9) and absolute temperature using various knot percentiles and changing lag days (14 and 28 days).

184

185 **3. Results**

186 Descriptive statistics of mortality, absolute temperature, and the distribution of DTR are in Table 1. Fig. 1 displays the geographical distributions of the 308 cities within the 10 countries 187 included in the analyses and the corresponding annual averaged DTR (\Box C). The data set 188 189 included 85,912,372 deaths. A variability in DTR was observed among countries over the entire study period, with mean values ranging from $6.7 \Box C$ (Ireland, 6 cities) to $10.9 \Box C$ (USA, 135) 190 cities). Table 1 also shows the DTRs and absolute temperature distributions during the first and 191 second halves of the time periods for each country. As expected, the mean temperature 192 increased slightly over time, although we did not detect a clear temporal pattern in the DTR 193 194 values. City-specific descriptive statistics are reported in Supplementary Table S1.

195

The percent increases in risks and attributable mortality risk fraction of DTR estimated from the model with no interaction (i.e. the average throughout the study period) are reported in Table 2. Percent increases in risks of DTR (per $10\Box C$) are highest in South Korea (6%, 95% CI: 3-9.1%), Spain (4.4%, 3–5.8%), and Brazil (4.2%, 1.7-6.7%). Colombia (-1.2%, -6.3–4.1%) and Ireland (0.3%, -3–3.8%) showed the lowest percent increases in risk of DTR, although

both were not significant. Table 2 also displays the total percentage of deaths attributable to
DTR (reference at minimum DTR of each city, 2.5% with 95% empirical confidence interval
(95% eCI): 2.3–2.7%). Similar with percent increases in risk, most of the country-specific
estimated attributable risks were statistically significant. The risk fraction was highest in Korea
(4.5%, 3-5.9%) and Spain (4.2%, 3.5-4.9%). The fractions were lowest in Colombia (-1.5%, 5.1–2.1%) and in Ireland (0.2%, -1.2–1.4%).

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Fig. 2 displays the country-pooled lag-response associations at the 99th percentile of DTR referenced at $0 \square C$. The coldest (Canada, Ireland, and UK) and warmest (Brazil, Colombia) countries showed the highest RR at lag 0 and lasted to a lag from 4–7 days. Other countries, which had moderate temperatures, have the highest RR in approximately 1–3 lag days and were limited to a lag of 7–14 days. The corresponding city-specific lag-response is displayed in Supplementary Fig. S1.

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215 Results from an analysis of the temporal variation in the percent increases in risk of DTR are illustrated in Fig. 3. Table 3 displays temporal variation of estimates per year and test results 216 for linear interaction (null hypothesis is the pooled interaction term is 0; the null hypothesis is 217 that no temporal change occurred). The percent increases in risk of DTR increased from 2.5% 218 219 (95% CI: 1.8–3.3%) to 3.8% (95% CI: 3.1–4.5%) between the first and last periods. Except for 220 Ireland and Japan, all countries showed patterns of increasing percent increases in DTR risk, with -2.9–5% in the first year and 1.5–13.8% in the last year of the series. The temporal increase 221 of percent increases in DTR risk were significant (P-value<0.05) in USA, UK, Spain, and South 222 223 Korea (Table 3). Country-pooled temporal changes in the lag-response relationship are displayed in Supplementary Fig. S2. A comparison between the curves suggests that a longer 224

225 lag-association and smaller harvesting effect were observed in most countries.

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Fig. 4 and Table 3 display the temporal variation in the attributable mortality risk fraction of 227 DTR. In overall, the attributable risk fraction of deaths increased from 2.4% (95% eCI: 2.1-228 2.7%) to 2.7% (2.4-2.9%) between the first and last periods. The increase in the attributable 229 risk fraction of death overtime was observed in all countries except Japan (0.07% decrease per 230 a year) and Ireland (0.23% decrease per a year). Korea (0.56% per year) and Colombia (0.31% 231 per year) showed the fastest increase of risk fraction, whereas Canada (0.03% per year), USA 232 (0.09% per year) and UK (0.11% per year) showed the slowest increasing patterns. 233 Corresponding city-specific estimates are reported in the Supplementary Material 234 (Supplementary Table S2). And main conclusions were robust to sensitivity analysis 235 (Supplementary Table S3). 236

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238 **4. Discussion**

Our findings show that DTR is responsible for a higher mortality risk increase (3.1%, 95% CI: 239 2.7–3.5%) and fraction of deaths (2.5%, 95% eCI: 2.3–2.7%) in all the countries studied. South 240 Korea and Spain showed the highest percent increase in risk (6% and 4.4%, respectively) and 241 attributable risk fractions (4.5% and 4.2%, respectively). This study also provides evidence of 242 the incremental health impact of DTR during the last few decades. With the exception of Japan 243 244 and Ireland, an increasing pattern of percentile increases in risks (3.8% in the last year of the study periods, compared with 2.5% in the first year) was observed, and the attributable risk 245 246 fraction showed the same temporal increasing pattern (2.7% in the last year of the study periods, compared with 2.4% in the first year). 247

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249 This study is comparable to a recent temperature variability-mortality association study in the MCC Collaborative Network (Guo et al. 2016). Both studies were based on a similar multi-250 251 country data set and addressed the significant association between temperature variability and mortality, even after controlling for the main effect of absolute temperature. Guo 252 et al. developed a new composite index of intra- and inter-day temperature variability using a 253 standard deviation of minimum and maximum temperatures during the exposure days, and 254 found the temperature variability-mortality relationship varied with exposure days (0-7 days), 255 countries (twelve countries/regions with 372 communities), and season (cold, hot, and 256 moderate). Meanwhile, our study only focused on the association between intra-day 257 258 temperature variability and mortality, using a classical meteorology index (DTR) and flexible statistical method, which considers a flexible lag-response structure of DTR. In addition, our 259 260 study included data from 308 cities in the 10 countries with more than 15 years of study data to estimate the temporal changes in the DTR-mortality association. It also described an overall 261 increase in the health burden of DTR on mortality during recent decades. 262

263

Interestingly, our finding suggested that the DTR effects on mortality were higher in warm 264 265 countries (Brazil, Australia, and Spain) compared to cold countries (Canada, Ireland, and UK), although Korea and Colombia were exceptions. These finding are consistent with previous 266 studies, such as multi-country studies that reported that the effect of temperature variability 267 with short exposure durations (0-1, and 0-2 days) on mortality is highest in hot area (>22.9°C) 268 than other areas (cold, moderate cold, and moderate hot areas) (Guo et al. 2016). U.S. studies 269 also showed that higher DTR effect in southern areas (percent change of non-accidental 270 mortality per one unit of DTR 0.24-0.31%) than other regions (0.22-0.27%) (Lim et al. 2014). 271

272 Studies in China also showed a similar trend of a relatively higher and more significant effect of DTR in warm cities (Guangzhou, and Shanghai) than cold cities (Anshan and Xi'an, 273 274 although Tangshan is an exception). To specifically assess the association between the DTR effects and annual mean of absolute temperature, we fitted a weighted regression model 275 (Supplementary Fig. S3); a city-specific BLUP of DTR coefficient (i.e. log of relative risk, 276 277 estimated from the second stage analysis) was used as a response variable, annual mean temperature was used as an explanatory variable, and inversed city-specific variances of the 278 279 DTR coefficient were used as weights. We observed a significantly positive linear association between the DTR effect and the annual mean temperature from the weighted regression model 280 (P-value=0.01). This result can be interpreted as evidence to support the hypothesis that there 281 may be an impact on mortality from the positive interactions between long-term temperature 282 and DTR. 283

284

The synergism effect of the DTR and long-term temperature on mortality may be due to a 285 number of factors. One mechanism may be aggravation. Hot temperature can disturb normal 286 physiological thermoregulation, including changes in blood viscosity, plasma cholesterol level, 287 and red blood cell count (Keatinge et al. 1986). Increasing DTR may also impact mortality 288 289 through lowering the thermoregulatory system and negatively affecting the heart rate, heart rate variability, blood platelets, red blood cells, and blood viscosity (Keatinge et al. 1984b; Lim et 290 al. 2012b). Because warm countries can be exposed to hot weather more often, the DTR effect 291 of warm areas can be amplified by the increase in biological burden. Another hypothesis is that 292 the effect of the DTR is higher in warm areas because people in warm and moderate areas are 293 more likely to keep their windows open and spend more time outdoors, which may increase 294 295 exposure to DTR, thus increasing the effect of DTR. However, our results only suggest the possible associations of the aggravation hypothesis; further research should be conducted tofind the causal relationship between DTR and long-term temperature on mortality.

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Additional questions should be raised as to why the effect of DTR on mortality changed across 299 time. We speculate several plausible explanations. The first hypothesis is deterioration by 300 climate change, suggested in the previous paragraph. We found the higher risk and sharper risk 301 302 increase in hot cities. This finding suggests that climate change may increase the risk of DTR. Secondly, an aging population may also be an important factor, as numerous studies have 303 304 revealed that elderly people are more susceptible to DTR (Kan et al. 2007a; Lim et al. 2012a; 305 Yang et al. 2013a), and the populations of developed and developing countries included in our research are aging (Börsch-Supan 2008; Faunce 2008). However, we could not identify the 306 exact reason of the temporal increase, hence additional research regarding this topic is needed 307 308 in further studies.

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310 Although not found in our study, prior studies have reported that climate change factors (greenhouse gases, urbanization, and aerosols) have led to global decline in the DTR during 311 twenty century, because the nocturnal minimum temperatures have increased faster than 312 maximum temperature (Braganza et al. 2004; Makowski et al. 2008). However, it is still unclear 313 314 how this decline in the DTR will affect human health. Also, since the increasing nocturnal 315 temperatures can affect mortality and distribution of DTR simultaneously, a confounding effect 316 of the nocturnal temperatures needs to be considered to estimate the effect of DTR on death. Even if the effect of increasing nocturnal minimum temperature is partly considered in our 317 318 study by controlling the daily averaged temperature, there is a limitation to control the effect of nocturnal temperature due to lack of data. Therefore, we expect more comprehensive studies 319

to be carried out under various climate conditions with longer study periods and more detailedweather data.

322

As described earlier, although DTR and absolute temperature may affect human health in 323 different ways, because both have a mechanism that negatively affects human health, the effects 324 of absolute temperature and temperature variability on mortality have been an interesting topic 325 326 of prior environmental research. In addition, comparing the health effects of two variables has important implications for understanding human health in a climate change context (which can 327 328 increase both the average values and the variability of temperature)(Guo et al. 2016; Stocker 329 2014; Vicedo-Cabrera et al. 2016). Recent studies asserted that DTR has a lesser effect than absolute temperature on mortality (Lee et al. 2017; Vicedo-Cabrera et al. 2016). Our results 330 also suggest a lesser effect of DTR on mortality when compared with the total attributable 331 mortality fraction of absolute temperature from a previous study (Gasparrini et al. 2015b). The 332 total fraction of DTR attributed to mortality (2.5%) was much smaller than the fraction for total 333 334 absolute temperature (7.71%)(Gasparrini et al. 2015b). However, our results may differ from the conclusions of previous studies (Chen et al. 2007; Kan et al. 2007a; Lim et al. 2015; Tam 335 et al. 2009; Yang et al. 2013a) that used modeling strategies that did not fully control the flexible 336 337 lag structures of absolute temperature. Because the effect of absolute temperature delayed up to several weeks of exposure, the estimates of DTR on mortality could be overestimated unless 338 the main effects of temperature are fully adjusted. Hence, we contend that our modelling 339 approach provides more appropriate results in estimating the health effects of DTR, in 340 comparison with prior studies. 341

342

343 In this study, more acute DTR–mortality relationships (highest RR at 0 lag days) were observed

in warm and cold countries (Brazil, Colombia, Canada, Ireland, and UK). In contrast, more delayed DTR-mortality relationships (highest RR at 2–4 lag days) were observed in other moderate temperature countries. Although this study does not explain the difference, we speculate that the exact factors are related to the physiological, technological, and behavioral adaptations to the climate.

349

350 A key strength of our study is the use of a large multi-country, multi-city data set with different demographic distributions, climate conditions, and socio-economic characteristics. To the best 351 352 of our knowledge, our study is the largest of its kind including 308 locations and more than 85 353 million deaths from 10 countries. Our study also is the first and the largest study of timevarying DTR-related mortality, and the use of a uniform statistical framework across all cities 354 makes the results directly comparable. In addition, unlike previous studies that have quantified 355 the association on terms of RR (Kan et al. 2007b; Vicedo-Cabrera et al. 2016; Yang et al. 2013b), 356 our study provides the attributable mortality burden of DTR. Because the attributable fraction 357 358 considers the distribution and risks of each variable, we contend that the attributable fraction is a suitable measure to estimate mortality burden of the exposure variable and to establish 359 corresponding public health strategies. 360

361

However, our study has some limitations that must be acknowledged. First, because regions of Africa, and large countries in Europe and Asia (such as France, Russia, and India) were not included in this study, our findings are not globally representative. Second, the data did not include age- or gender-specific mortality rates, which could be explored in future research. Third, we could only identify the suggestive association between DTR and all-causes mortality, and not the causal effect of DTR on mortality. Future studies should strive to overcome these 368 limitations by expanding the study populations and improving the study design.

369

5. Conclusions

371 In summary, this study finds that the significant DTR effect on mortality across all countries, and provides evidence that the effect of DTR was higher in warm regions. Although our 372 estimated attributable mortality fraction of DTR is smaller than fractions of absolute 373 temperature from a previous multi-country study (Gasparrini et al. 2015b), it is higher than 374 fractions of extreme heat and cold temperature. In addition, although the risks and contributions 375 376 of DTR on mortality varied for each country, it increased at the multi-country scale with significant increases estimated in USA, UK, Spain, and South Korea; and non-significant 377 increments in Canada, Brazil, Colombia, and Australia. The estimates of DTR-related mortality 378 379 increased throughout the study period in overall regions, which could be interpreted as maladaptation to DTR. Consequently, there is a possibility that the health burden of DTR will 380 increase in the near future. Hence, we suggest that public-health policies and climate change 381 research that have so far focused on the effects of extreme heat should be extended to account 382 for the health burden of DTR and its temporal variation. 383

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Table 1. Descriptive statistics by country. Including distribution of diurnal temperature range
in first 3 years (First) and last 3 years (Last) of country-specific study periods. USA: United
States of America, UK: United Kingdom.

| Country (# of city) | Time Period | Total Deaths | Study Period | Absolute temperature | Diurnal temperature range (°C) | | |) | | |
|------------------------|----------------|-----------------|-----------------|-------------------------|--------------------------------|-----|------------|------|------|------|
| | | | (year) | (\mathbf{C}) | | | | | | |
| | | | | Mean | Mean | 10% | 25% | 50% | 75% | 90% |
| Canada | Whole | 2,989,901 | 1986- | 6.8 | 10 | 4.4 | 6.6 | 9.7 | 13 | 15.9 |
| (26) | | | 2011 | _ | 10.1 | | < - | 0.0 | 10 | 1.6 |
| | First | | | 7 | 10.1 | 4.5 | 6.7 | 9.8 | 13 | 16 |
| | Last | | | 6.9 | 9.9 | 4.4 | 6.4 | 9.4 | 12.9 | 16.1 |
| USA | Whole | 22,896,409 | 1985- 2006 | 14.8 | 10.9 | 5.6 | 7.8 | 10.6 | 13.9 | 16.7 |
| (135) | First | | 2000 | 147 | 11 | 5.6 | 78 | 10.6 | 12.0 | 17.2 |
| | Let | | | 14.7 | 10.7 | 5.0 | 7.0 | 10.0 | 12.2 | 17.2 |
| | Last | | 1007 | 15.1 | 10.7 | 5.6 | /.8 | 10.6 | 13.3 | 10.1 |
| Brazil | Whole | 3,435,502 | 2011 | 24.2 | 9 | 5.4 | 6.8 | 8.6 | 10.7 | 13.2 |
| (18) | First | | | 24.1 | 8.8 | 5.1 | 6.6 | 8.4 | 10.6 | 13.1 |
| | Last | | | 24.3 | 9.1 | 5.5 | 7 | 8.6 | 10.7 | 13.4 |
| Colombia | Whole | 956,539 | 1998- | 23.4 | 9 | 5.8 | 7 | 8.8 | 10.8 | 12.4 |
| (5) | | | 2013 | | | | | | | |
| | First | | | 23.1 | 8.9 | 5.6 | 6.8 | 8.7 | 10.8 | 12.5 |
| | Last | | | 23.5 | 8.9 | 6.1 | 7.2 | 8.7 | 10.4 | 12.1 |
| UK | Whole | 1,2075,786 | 1990- | 10.3 | 7.3 | 3.8 | 5.2 | 6.9 | 9.1 | 11.3 |
| (10) | | | 2012 | | | • | | 6.0 | | |
| | First | | | 10.1 | 7.3 | 3.8 | 5.1 | 6.9 | 9.1 | 11.3 |
| | Last | | | 10.1 | 7.5 | 3.9 | 5.3 | 7 | 9.4 | 11.7 |
| Ireland | Whole | 1,058,215 | 1984- 2007 | 9.7 | 6.7 | 3.4 | 4.8 | 6.4 | 8.3 | 10.3 |
| (6) | First | | | 8.9 | 6.8 | 3.6 | 4.9 | 6.5 | 8.4 | 10.4 |
| | Last | | | 10.3 | 6.9 | 3.6 | 4.9 | 6.6 | 8.5 | 10.5 |
| Spain | Whole | 3,480,531 | 1990- | 15.5 | 10.6 | 4.9 | 7 | 10 | 13.8 | 17 |
| (51) | | | 2010 | | | | | | | |
| | First | | | 15.1 | 10.7 | 5 | 7.2 | 10.2 | 13.8 | 17.2 |
| | Last | | | 15.5 | 10.4 | 4.8 | 6.8 | 9.8 | 13.6 | 17 |
| Japan | Whole | 3,6113,897 | 1972- 2012 | 15.1 | 8.4 | 4.2 | 6 | 8.2 | 10.6 | 12.8 |

| (47) | First | | | 14.4 | 8.8 | 4.4 | 6.3 | 8.6 | 11 | 13.3 |
|------------------|-------|-----------|---------------|------|-----|-----|-----|-----|------|------|
| | Last | | | 15.5 | 8.2 | 4.1 | 5.9 | 8 | 10.2 | 12.4 |
| South Korea | Whole | 1,727,642 | 1992- 2010 | 13.7 | 8.2 | 4.1 | 5.9 | 8 | 10.2 | 12.7 |
| (7) | First | | | 13.5 | 8 | 3.8 | 5.6 | 7.7 | 10.1 | 12.5 |
| | Last | | | 13.8 | 8.3 | 4.3 | 5.9 | 8 | 10.3 | 12.7 |
| Australia (3) | Whole | 1,177,950 | 1988- 2009 | 18.1 | 8.2 | 4.4 | 5.9 | 7.8 | 10 | 12.4 |
| | First | | | 18.1 | 7.8 | 4.1 | 5.6 | 7.3 | 9.5 | 11.9 |
| | Last | | | 18.7 | 8.1 | 4.5 | 5.8 | 7.6 | 10 | 12.6 |

| 516 | Table 2. Percent increases in risk (per 10°C) and attributable risk fraction (%) of diurnal |
|-----|---------------------------------------------------------------------------------------------|
| 517 | temperature range on mortality by country. USA: United States of America, UK: United |
| 518 | Kingdom. |

| Country | Percent Increases in Risk (%, 95% CI) | Attributable Risk Fraction (%, 95% eCI) |
|-------------|------------------------------------------|--------------------------------------------|
| Canada | 2.6 % (0.9 , 4.2) | 2.7 % (1.8 , 3.5) |
| USA | 2.9 % (2.3 , 3.6) | 3.2 % (2.9 , 3.5) |
| Brazil | 4.2 % (1.7 , 6.7) | 3.7 % (2.6 , 4.9) |
| Colombia | -1.2 % (-6.3 , 4.1) | -1.5 % (-5.1 , 2.1) |
| UK | 2.9 % (1.5 , 4.4) | 2.1 % (1.6 , 2.7) |
| Ireland | 0.3 % (-3 , 3.8) | 0.2 % (-1.2 , 1.4) |
| Spain | 4.4 % (3,5.8) | 4.2 % (3.5 , 4.9) |
| Japan | 3.1 % (2.3 , 3.9) | 2.7 % (2.4 , 3) |
| South Korea | 6 % (3,9.1) | 4.5 % (3, 5.9) |
| Australia | 4.2 % (0.7 , 7.9) | 3.3 % (1.1 , 5.3) |
| Overall | 3.1 % (2.7 , 3.5) | 2.5 % (2.3 , 2.7) |

519 eCI: empirical confidence interval.

Table 3. Variation of excessive relative risk (per 10°C) and attributable fraction (%) of diurnal
temperature range on mortality per a year, and p-value of the test. USA: United States of
America, UK: United Kingdom.

| | Variation (per year) | | | | | | |
|-------------|-------------------------|---------------------------|----------------------------|----------|--|--|--|
| Country | Study Period (Years) | Percent Increases in Risk | Attributable Risk Fraction | p-value* | | | |
| Canada | 26 | 0.06 % | 0.03 % | 0.4384 | | | |
| USA | 22 | 0.09 % | 0.09 % | 0.0209 | | | |
| Brazil | 15 | 0.25 % | 0.23 % | 0.2243 | | | |
| Colombia | 16 | 0.29 % | 0.31 % | 0.6583 | | | |
| UK | 23 | 0.17 % | 0.11 % | 0.0344 | | | |
| Ireland | 24 | -0.35 % | -0.23 % | 0.0989 | | | |
| Spain | 21 | 0.28 % | 0.23 % | 0.0025 | | | |
| Japan | 41 | -0.09 % | -0.07 % | < 0.0001 | | | |
| South Korea | 19 | 0.78 % | 0.56 % | < 0.0001 | | | |
| Australia | 22 | 0.3 % | 0.26 % | 0.1543 | | | |

523 * Significant test on temporal change by Wald type test of the pooled reduced coefficient of the
524 year-interaction terms. The null hypothesis is that no change in year occurred.

525 Figure legends

526 **Fig. 1.** Geographical locations of study cities and their annual mean values of diurnal 527 temperature range (DTR, °C).

Fig. 2. Lag-response relationship between diurnal temperature range (DTR) and mortality
predicted for the overall study periods of 10 countries; RR: relative risk. USA: United States,
UK: United Kingdom.

- 531 Fig. 3: Temporal changes in percent increases in risk (%) between the first (First) and the last
- 532 (Last) year of country-specific study periods; USA: United States, UK: United Kingdom.
- **Fig. 4.** Temporal changes in attributable risk fraction (%) between the first (First) and the last
- 534 (Last) year of country-specific study periods; USA: United States, UK: United Kingdom.

- 535 Acknowledgements: This study was supported by the Global Research Lab (#K21004000001-
- 536 10A0500-00710) through the National Research Foundation of Korea.
- 537 **Competing financial interest:** The authors declare they have no actual or potential competing
- 538 financial interests.



Canada

USA

Brazil

Colombia









UK















South Korea









