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September 2006

Online at http://mpra.ub.uni-muenchen.de/252/ MPRA Paper No. 252, posted 9. October 2006

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Abstract

In this paper the long memory and non-linear properties of share prices in the UK's Stock Exchange and AIM are explored. The results suggest that the most commonly traded shares exhibit long memory thus raising interesting issues about the validity of normal assumptions of market efficiencies.

1 Introduction

Gil-Alana (2006) highlights the issue of whether the time series behaviour of stock market indexes can be described by a fractional representation. Such a possibility could be due to long memory or nonlinearities and has serious consequences for the efficient market hypothesis in stock markets. In this paper the issues are examined further using the time-series of individual share prices for one hundred companies from the London Main(FTSE) and Alternative Investment(AIM) markets. A simple fractional augmented Dickey-Fuller (FADF) test is used to test for long memory utilising a range of consistent estimates of the fractional differencing parameter d. The possibility of non-linearity is investigated by applying both random field based tests and traditional tests for structural breaks.

2 Background

Weak form market efficiency posits that the current share price fully reflects the information implicit in the share price history of a company. This implies that the share price time series, say $\{P_{it}\}_{t=0}^{T}$ for share *i*, does not exhibit serial correlation, or 'memory'. Recent literature puts forward the case for a generic random walk type model, albeit one with less stringent assumptions than those imposed in earlier literature on market efficiency. Traditionally the analysis has been whether $\{P_{it}\}_{t=0}^{T}$ is I(1) or I(0). In the I(0) case P_{it} is covariance stationary and displays the property of short-memory. In the I(1) case P_{it} is non-stationary and non-mean reverting. Fractional integration relaxes the assumption that the differencing parameter *d* is integer. A time-series process with 0 < d < 1 is defined as a long-memory process. If 0 < d < 0.5 the series P_{it} is covariance stationary, but the autocorrelation function is more persistent, with a longer decay period, than when d = 0. If $0.5 \le d < 1$ P_{it} is not covariance stationary, but is still mean reverting, with a persistent, slowly decaying autocorrelation function. Thus, if P_{it} is I(d) with 0 < d < 1 it exhibits memory and is thus not consistent with weak-form efficiency.

While studies focusing on the non-stationarity of time series have dominated the literature, increasingly attention is focusing on the issue of whether series that appear integrated of order > 0 are actually non-linear but stationary series. Diebold & Inoue (2001) showed that long memory and structural change are closely related in the I(d) case, but warned:

"... that the temptation to jump to the conclusion of "structural change producing spurious inferences of long memory" should be resisted."

(Diebold & Inoue 2001, p.157)

3 Methodology

To test whether share prices exhibit long memory and/or non-linearity, a variety of techniques are applied to the time-series of the top 25 and bottom 25 (by market capitalisation) companies in both the FTSE and the AIM from the 1^{st} January 1998 to 31^{st} December 2004, inclusive. The selection allows for the investigation of two hypotheses. The first is that there are differences in how information is processed between listed markets. The expectation being that AIM listed companies will be subject to less scrutiny by market participants than their FTSE counterparts and more prone to information inefficiencies. Therefore their share prices are more likely to exhibit long memory. The second hypothesis is the relative size hypothesis. Larger companies should be subject to greater scrutiny by market participants and therefore be less prone to information inefficiencies. The tendency towards long memory should thus be more prominent in smaller companies.

Firstly, long memory is investigated using Dolado, Gonzalo & Mayoral (2002) FADF test and consists of estimating the equation:

$$\Delta P_t = \theta \Delta^d P_{t-1} + \sum_{i=1}^p \varsigma_i \Delta P_{t-i} + \epsilon_t \tag{1}$$

The test of H_0 : $\theta = 0$ is a test of the hypothesis that d = 1. Using various consistent estimates of d (Shimotsu (2006) exact Whiittle estimators¹, semiparametric estimator of Geweke & Porter-Hudak (1983) and two parametric estimators, the exact maximum likelihood and nonlinear least squares, available in the ARFIMA package of Oxmetrics (Doornik & Ooms 1999) the test statistic is asymptotically normal under H_0 .

To investigate the possibility of non-linearity in the series both traditional parametric tests for structural breaks and random field based inference are used. In both cases the possibility of the share prices following a non-linear AR(1)process or some form of transition models was investigated. Using the suite of tests proposed by Bai & Perron (2003) the models

$$\begin{array}{l}
P_t = \alpha_0 + \sum_{i=1}^p \alpha_i \delta_{it} + \epsilon_t & and \\
P_t = \beta_0 P_{t-1} + \sum_{i=1}^p \beta_i P_{t-1} \delta_{it} + \epsilon_t & \end{array}
\right\} \quad \delta_{it} = \begin{cases}
0 & (t < t_i^*) \\
1 & (t \ge t_i^*)
\end{cases}$$

$$(2)$$

were investigated.² The random field tests for non-linearity are based on two special cases of the model introduced in Hamilton (2001).

$$P_{t} = \alpha + \beta P_{t-1} + \lambda m(g \odot \mathbf{P}) + \epsilon_{\mathbf{t}} \qquad and$$

$$P_{t} = \alpha + \beta P_{t-1} + \psi t + \lambda m(g \odot \mathbf{x}) + \epsilon_{\mathbf{t}} \qquad (3)$$

where $x_t = (P_t, t)$, and \odot denotes element by element multiplication. In (3) m(.) are random fields and g is a $(k \times 1)$ vector k = 1 or 2. The parameter g can be seen as a measure of the impact of either P or t on the non-linearity. The scalar parameter λ as a measure of the degree of non-linearity. A simple test for non-linearity is to test $H_0 : \lambda = 0$. Hamilton's approach was to assume a gaussian random field, meaning that the field is defined by its first two moments, and deriving a simple χ^2 test of the null hypothesis derived from a generalised linear regression model interpretation. In Dahl & Gonzalez-Rivera (2003) the basic test was extended by the introduction of three tests that took account, to various degrees, the problems of nuisance parameters. It was shown in Dahl & Gonzalez-Rivera (2003) that the four test were generally more powerful than other tests of non-linearity and fairly insensitive to model misspecification.³

4 Results

Tables 1 and 2 summarise the principal findings of the study. A series is classified as either fractional or I(1) when over 85% of the *FADF* tests indicate this otherwise the series has been classified as indeterminate. In all fractional cases d was non-stationary.

For the FTSE 64% of the largest companies exhibit strong signs of fractionality whilst none of the smallest companies do. Conversely, 32% of the largest companies and 92% of the smallest companies exhibit signs of weak form efficiency. The remaining 4% and 8% being indeterminate. These result seem contrary to the relative size hypothesis. The results of the non-linearity tests are generally inconclusive. The one exception is the single large company where the *FADF* was inconclusive - here there is a strong indication of some form of non-linearity. Similar, less strong, results follow for the 2 small companies where the *FADF* is inconclusive. Amongst the smaller companies there is more evidence that in some cases the non-stationarity could be due to non-linearity.

For the AIM the results are significantly different, 8% of the largest companies and 24% of the smallest companies exhibit signs of fractionality. The majority of large companies, 76%, are weak form efficient whilst only 44% of the smallest companies are. There are many more companies, 16% and 32% of the largest amd smallest respectively, being indeterminate. These result support the relative size hypothesis but when compared with the FTSE are contrary to the relative markets efficiency hypothesis. Again the results of the non-linearity tests are generally inconclusive. The increased probability of non-linearity amongst smaller companies together with the larger number of inconclusive fractionality results might be explained by thin trading.

5 Discussion and Conclusion

In this paper the work of Gil-Alana (2006) has been extended by investigating the time-series characteristics of individual share price series using a simple FADF test and a suite on non-linearity tests. Data from the top and bottom 25 firms in the FTSE 100 index and the AIM were investigated. The results of the study appear to cast doubt on whether the larger companies in the FTSE are weak form efficient. Thus someone using some form of trading rule with these companies, operating in 'real-time', may be capable of profiting from these inefficiencies. This is an interesting finding, as theory posits that stock price markets are information efficient. The concept of market efficiency is built upon the foundation that markets are at least weak form efficient. If this does not hold, then neither do the stricter definitions. If anything, the results provide evidence that information efficiency is more prevalent the smaller the company, and given the caveat re thin trading, the smaller the market.

The possibility of inefficiencies in the FTSE needs to be investigated further. For example, a possible explanation for the inefficiency in the FTSE may be the use from the late 1980's onwards of automated trading systems. Such systems may create self-perpetuating patterns, thereby creating long-memory, from which traders may profit.

Given the inconclusive nature of the Hamilton-Dahl and Bai-Perron tests the issue of whether the non-stationarity in share prices is in part due to nonlinearity needs to be investigated further. This could be by trying to model the possible non-linearity through random field or STR models. Or it could be by examining the behaviour of the consistent semi-parametric estimators of d in line with Diebold & Inoue (2001).

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Market	25 Largest Companies			25 Smallest Companies			
	Fractional	I(1)	Indeterminate	Fractional	I(1)	Indeterminate	
$\begin{array}{c} \text{FTSE} \\ Number \\ Percent \end{array}$	$rac{16}{64\%}$	$\frac{8}{32\%}$	$\frac{1}{4\%}$	0 0%	$23 \\ 92\%$	$^{2}_{8\%}$	
$\begin{array}{c} \text{AIM} \\ Number \\ Percent \end{array}$	$2 \\ 8\%$	$\underset{76\%}{19}$	$4_{16\%}$	${6\atop{24\%}}$	$\underset{44\%}{11}$	$\frac{8}{32\%}$	

Table 1: Stationarity Results

Table 2: Percentage of linearity tests significant at 5% levels

		Hamilton-Dahl		Bai-Perron	
		model		model	
		P_{t-1}	P_{t-1}, t	β	α
FTSE Largest		%	%	%	%
	$Fractional_{(n=16)}$	11	13	6	38
	I(1) = (n=8)	31	38	0	63
	$Indeterminate_{(n=1)}$	75	75	0	100
FTSE Smallest					
	$Fractional_{(n=0)}$	•	•	•	•
	I(1) (n=23)	8	33	30	30
	Indeterminate (n=2)	38	38	0	50
AIM Largest		%	%	%	%
-	$Fractional_{(n=2)}$	50	50	0	50
	I(1) = (n=19)	8	14	16	11
	$Indeterminate_{(n=4)}$	0	0	0	25
AIM Smallest					
	$Fractional_{(n=6)}$	16	46	0	50
	I(1)	7	25	27	45
	Indeterminate (n=11)	50	32	25	63

Note: Each variable has 4 tests applied for each method/model

Notes

 $^1 \rm using$ the MATLAB code supplied by the authors.

 $^2\mathrm{the}$ Gauss code provided by the authors was used.

 $^3{\rm The}$ Gauss code supplied by Hamilton (http://weber.ucsd.edu/~jhamilto) was used.