

# The pseudo 2-D relaxation model for obtaining $T_1$ – $T_2$ relationships from 1-D $T_1$ and $T_2$ measurements of fluid in porous media

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## Abstract

NMR spin-lattice ( $T_1$ ) and spin-spin ( $T_2$ ) relaxation times and their inter-relation possess information on fluid behaviour in porous media. To elicit this information we utilise the pseudo 2-D relaxation model (P2DRM), which deduces the  $T_1$ – $T_2$  functional relationship from independent 1-D  $T_1$  and  $T_2$  measurements. Through model simulations we show empirically that the P2DRM accurately estimates  $T_1$ – $T_2$  relationships even when the marginal distributions of the true joint  $T_1$ – $T_2$  distribution are unknown or cannot be modeled. Estimates of the  $T_1$ : $T_2$  ratio for fluid interacting with pore surfaces remain robust when the P2DRM is applied to simulations of rapidly acquired data. Therefore, the P2DRM can be useful in situations where experimental time is limited.

## Keywords:

Relaxation correlation; Lognormal distribution; Inverse-gamma distribution; Magnetic resonance in porous media; Heterogeneity; Multidimensional distribution function

## 1. Introduction

Nuclear Magnetic Resonance (NMR) relaxation measurements provide a non-invasive means of studying fluid-saturated porous media. Heterogeneity of porous materials leads to distributions of spin-lattice ( $T_1$ ) and spin-spin ( $T_2$ ) relaxation times arising from the fluid within [1]. Since  $T_1$  and  $T_2$  are functions of the same material properties, e.g., surface-to-volume ratio [2], these quantities have a functional relationship [3]. The  $T_1$ – $T_2$  relationship provides information about surface interactions [4], which is unobtainable from a 1-D distribution alone. Venkataramanan et al. developed an efficient algorithm [5] for estimating a joint  $T_1$ – $T_2$  (probability) distribution from  $T_1$ – $T_2$  correlation experiment [6] data [3]. The  $T_1$ – $T_2$  distributions of porous media systems (e.g. fluid-saturated sandstones and carbonates [3]) confirm [7] that the observed relaxation rates often follow the Brownstein-Tarr [2] equations for the fast-diffusion limit, i.e., a sum of surface ( $\rho_1$  or  $\rho_2$ ) and bulk contributions, with the sum controlled by the surface-to-volume ratio:

$$\frac{1}{T_{1,2}} = \frac{1}{T_{1,2 \text{ bulk}}} + \rho_{1,2} \frac{S}{V}. \quad (1)$$

From this pair a  $T_1$ – $T_2$  relationship can be described by a single monotonic equation [8]

$$\frac{1}{T_1} = K + \frac{\varepsilon}{T_2} \quad (2)$$

where  $K = 1/T_{1 \text{ bulk}} - \varepsilon/T_{2 \text{ bulk}}$  and  $\varepsilon = \rho_1/\rho_2$ . If the  $T_1$ – $T_2$  relationship of a system follows Eq. (2) then it is obtainable from 1-D measurements. This was inferred as early as 1993 when Kleinberg et al. obtained single values of the  $T_1$ : $T_2$  ratio of rock cores by applying a cross-correlation function to the 1-D  $T_1$  and  $T_2$  distributions [9]. We developed the pseudo 2-D relaxation model (P2DRM): a method for obtaining  $T_1$ – $T_2$  distribution functions from 1-D  $T_1$  and  $T_2$  measurements [8]. The mathematical framework of relating distributions in this way was published by Röding et al. and is not specific to NMR [10]. The P2DRM is a 2-step parametric model fitting routine of independent 1-D  $T_1$  and  $T_2$  measurements. Parameter estimates from the  $T_2$  data fit in the first step are used to constrain the  $T_1$  data fit in the second step, making it possible to estimate the  $T_1$ – $T_2$  relationship without a 2-D data set. The utility of imposing constraints in multiple data fitting steps was also shown by Benjamini and Bassar [11], who found that constraining the distribution fit of 2-D relaxation and diffusion data by the estimated marginal distributions from fits to 1-D data led to a significant reduction in the amount of data required for a stable fit. In the P2DRM, utilizing Eq. (2) as prior knowledge allows for the parameters  $K$  and  $\varepsilon$  to be fit directly in the second step. A pseudo 2-D  $T_1$ – $T_2$  distribution results from mapping the independent 1-D probability distributions to 2-D space using the assumed  $T_1$ – $T_2$  relationship.

In our previous publication [8], when tested on simulated data for fluid in rock, the pseudo 2-D distributions estimated by

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the P2DRM were consistent with the known joint  $T_1$ - $T_2$  distribution. However, in that case, the parametric models which the P2DRM used were good choices in that they were capable of describing the known joint  $T_1$ - $T_2$  distribution. Though we gave physical justification for the parametric model sets used, in practice one does not have prior knowledge of the appropriate model. The first point of this publication is to test whether the P2DRM can accurately estimate the  $T_1$ - $T_2$  relationship when the parametric model sets are incapable of describing the marginal distributions of the true joint  $T_1$ - $T_2$  distribution. Utilizing a rapid  $T_1$  measurement along with the Carr-Purcell-Meiboom-Gill (CPMG)  $T_2$  measurement potentially offers a means of estimating the  $T_1$ - $T_2$  relationship in situations where experimental time is limited such as for investigation of time-sensitive processes. The second point of this publication is to test the capability of the P2DRM to utilize rapidly acquired data in estimating the  $T_1$ - $T_2$  relationship.

The theory section includes equations for using lognormal or inverse-gamma distributions and their associated  $T_1$  distributions as components in distribution models. Model sets are physically motivated and defined. The equations are reproduced from our previous publication [8] (and its corrigendum [12]). The methods section explains the data simulation and the fitting routine. Simulations give us access to the known joint  $T_1$ - $T_2$  distribution and allow us to test the limits of the P2DRM. The results and discussion section compares the pseudo 2-D distributions estimated by the P2DRM to the known joint  $T_1$ - $T_2$  distribution. We test the accuracy and precision by obtaining parameter estimates from 100 data simulations and fits.

## 2. Theory

The CPMG sequence [13, 14] measures the distribution of  $T_2$  relaxation times,  $f(T_2)$ , by acquiring the signal from the center of each echo in a train of  $180^\circ$  RF pulses. The signal,  $I(t_2)$ , as a function of the acquisition time,  $t_2$ , is related to  $f(T_2)$  by

$$I(t_2) = I_0 \int_0^\infty f(T_2) \exp(-t_2/T_2) dT_2, \quad (3)$$

where  $I_0$  is the signal intensity at  $t_2 = 0$ . A class of rapid  $T_1$  measurements built off the Look-Locker method [15] uses a series of short tip-angle RF pulses to linearly sample signal at tens to hundreds of points in the time domain with a single scan. In a double-shot implementation by Chandrasekera et al. [16], the signal from the FID following the  $n^{\text{th}}$  RF pulse of angle  $\theta$  is

$$I\{(n-1)\tau_1\} = M_0 \sin \theta (\cos \theta)^{n-1} \int_0^\infty f(T_1) (\exp(-(n-1)\tau_1/T_1)) dT_1 \quad (4)$$

where  $\tau_1$  is the time between RF pulses, the acquisition time is  $t_1 = (n-1)\tau_1$ , and  $M_0$  is the initial magnetization. The signal intensity at  $t_1 = 0$  is  $I_0 = M_0 \sin \theta$ .

A numerical inverse Laplace transform method is used to obtain  $f(T_2)$  or  $f(T_1)$ . The result is a non-unique estimate of the actual relaxation time distribution and is dependent on the choice of model [17, 18]. Parametric models are based on physically motivated information and use a pre-defined number of

components involving a commensurate number of pre-defined functions [18, 19, 20]. The P2DRM uses parametric models to allow for the utilization of Eq. (2) as prior knowledge. More specifically, for a given parametric distribution component of the  $f(T_2)$  model, Eq. (2) can be used in a change-of-variables to define the parametric form of the component as a function of  $T_1$ . First, for the lognormal distribution of  $T_2$ ,

$$P_{\log N}(T_2) = \frac{1}{\sigma T_2 \sqrt{2\pi}} \exp\left(-\frac{1}{2\sigma^2}(\ln T_2 - \mu)^2\right), \quad (5)$$

a change-of-variables using Eq. (2) results in

$$P_{\text{BT } \log N}(T_1) = \frac{1}{(1 - KT_1)(\sigma T_1 \sqrt{2\pi})} \times \exp\left(-\frac{1}{2\sigma^2}(\ln \varepsilon + \ln T_1 - \ln(1 - KT_1) - \mu)^2\right). \quad (6)$$

The subscript BT refers to Brownstein-Tarr. The parameters  $\mu$  and  $\sigma$  control the shape of Eq. (5) and  $K$  and  $\varepsilon$  control the transformation of Eq. (5) to Eq. (6).

Second, for the inverse-gamma distribution of  $T_2$ ,

$$P_{\Gamma^{-1}}(T_2) = \frac{\beta^\alpha}{\Gamma(\alpha)} T_2^{-\alpha-1} \exp\left(-\frac{\beta}{T_2}\right), \quad (7)$$

a change-of-variables using Eq. (2) results in

$$P_{\text{BT } \Gamma^{-1}}(T_1) = \frac{\varepsilon}{(1 - KT_1)^2} \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{\varepsilon T_1}{1 - KT_1}\right)^{-\alpha-1} \times \exp\left(-\frac{\beta(1 - KT_1)}{\varepsilon T_1}\right), \quad (8)$$

with parameters  $\alpha$  and  $\beta$  equal to their value in Eq. (7) and again the only free parameters are  $K$  and  $\varepsilon$ .

Model sets must include a  $T_2$  distribution model,  $f(T_2)$ , and an associated  $T_1$  distribution model,  $f(T_1)$ . Model sets can utilize any combination of component functions, including delta functions so long as the two associated  $P(T_2)$  and  $P(T_1)$  functions come as pairs in attempting to represent the same population of spins. We use two model sets, 'model set A' and 'model set B' which each include a distribution plus delta function. Model set A incorporates  $P_{\log N}(T_2)$  and  $P_{\text{BT } \log N}(T_1)$  as the distributed component. The distributed component in model set B is represented by  $P_{\Gamma^{-1}}(T_2)$  and  $P_{\text{BT } \Gamma^{-1}}(T_1)$ . The fact that  $T_1 \geq T_2$  is utilized to constrain the delta function. We have found model sets with a distributed component plus a delta function component to be more robust than model sets with a single distributed component and more stable than model sets with two distributed components. Inclusion of a delta function is physically motivated by the fact that as pore size increases, the Brownstein-Tarr fast diffusion limit will no longer apply and a significant portion of fluid in the interior of the pore can be described as having relaxation times equal to bulk values. Non-parametric, uniform-penalty inversions (UPEN) of relaxation experiments performed on porous media systems often show a sharp peak near the bulk relaxation time with a broad shoulder

extending to short relaxation times [21], perhaps indicating that such a physical motivation is well-founded.

The pseudo 2-D  $T_1$ - $T_2$  distribution exists along the  $T_1$ - $T_2$  relationship described by Eq. (2) and the estimated values of  $K$  and  $\varepsilon$ . Due to the probability distribution being infinitely thin in all directions other than along the  $T_1$ - $T_2$  relationship, the function for the pseudo 2-D  $T_1$ - $T_2$  distribution is related to the 1-D marginal  $T_1$  or  $T_2$  distribution by a line integral. When parameterized by  $T_2$ , the resulting function is

$$P_c(T_2) = \frac{P(T_2)}{\sqrt{1 + \frac{\varepsilon^2}{(\varepsilon + KT_2)^4}}} \quad (9)$$

where  $P(T_2)$  is the 1-D marginal  $T_2$  distribution, either  $P_{\log N}(T_2)$  or  $P_{\Gamma^{-1}}(T_2)$ . The subscript c refers to  $P_c(T_2)$  being a distribution along a curve in  $T_1$ - $T_2$  space.

### 3. Methods

Paramagnetic species along pore walls determine surface relaxivities [1, 7]. Foley et al. measured  $T_1$  and  $T_2$  of water in packed calcium silicate powders synthesized with known concentrations of iron paramagnetic ions [22]. The iron concentration had a stronger effect on  $\rho_2$  than  $\rho_1$ ,

$$\rho_1 = 4.05 \mu\text{m/s} + (0.000819 \mu\text{m/s/ppm})[\text{Fe}]_{\text{ppm}}, \quad (10a)$$

$$\rho_2 = 3.96 \mu\text{m/s} + (0.00227 \mu\text{m/s/ppm})[\text{Fe}]_{\text{ppm}}, \quad (10b)$$

and therefore a distribution of iron concentrations can lead to a distribution of  $T_1$ - $T_2$  relationships. We simulated data sets by randomly sampling a large number of discrete radii and iron concentration values from a prescribed pore size and iron distribution, assigning  $T_1$  and  $T_2$  values to each radii and iron concentration pair using Eqs. (10) and Eqs. (1) and modeling signal relaxation as a sum of contributions from all radii and iron concentration pairs. The known joint  $T_1$ - $T_2$  distribution is a 2-D histogram from the discrete  $T_1$  and  $T_2$  values. The prescribed pore radii distribution was a sum of two lognormal distributions (the first with  $w = 0.8$ , mean = 27  $\mu\text{m}$ , and standard deviation (std) = 36  $\mu\text{m}$ , the second with mean = 380  $\mu\text{m}$  and std = 200  $\mu\text{m}$ ), which creates a bimodal  $T_1$ - $T_2$  distribution. The prescribed paramagnetic iron distribution was lognormal with mean =  $3.6 \times 10^4$  ppm and std =  $4.8 \times 10^4$  ppm, similar to concentrations found in many sandstones and carbonates [22]. Through Eqs. (10) the iron distribution corresponded to a  $\rho_2/\rho_1$  distribution with mean = 2.37 and std = 0.28. Bulk relaxation times were  $T_{1 \text{ bulk}} = T_{2 \text{ bulk}} = 2$  s. Discrete relaxation time values found using Eqs. (1) were used in Eqs. (3) and (4) to simulate CPMG data (with 4000 echoes and a 400  $\mu\text{s}$  echo time) and rapid  $T_1$  data (with 100 RF pulses of  $\theta = 5^\circ$  and  $\tau_1 = 30$  ms). The signal to noise ratio ( $\text{SNR} = I_0/\sigma_{\text{noise}}$ ) was set to 340 for the  $T_2$  data and 34 for the  $T_1$  data by adding zero mean Gaussian noise.

The parametric distribution model sets A and B were employed in Eqs. (3) and (4) for a 2-step fitting of the 1-D data sets. The estimated values of  $K$  and  $\varepsilon$  complete the  $T_1$ - $T_2$  relationship described by Eq. (1). The pseudo 2-D  $T_1$ - $T_2$  distribution

was determined from the results and Eq. (9). A least-squares fitting routine was implemented in MATLAB R2015a (Mathworks, Natick, USA) and was made available as supplementary material to ref. [8].

### 4. Results and discussion

The fits of distribution model sets A and B to simulated  $T_2$  and  $T_1$  relaxation data and the residuals of the fits are shown in Fig. 1. The factor  $(\cos \theta)^{n-1}$  seen in Eq. (4) was divided from the  $T_1$  signal intensity.

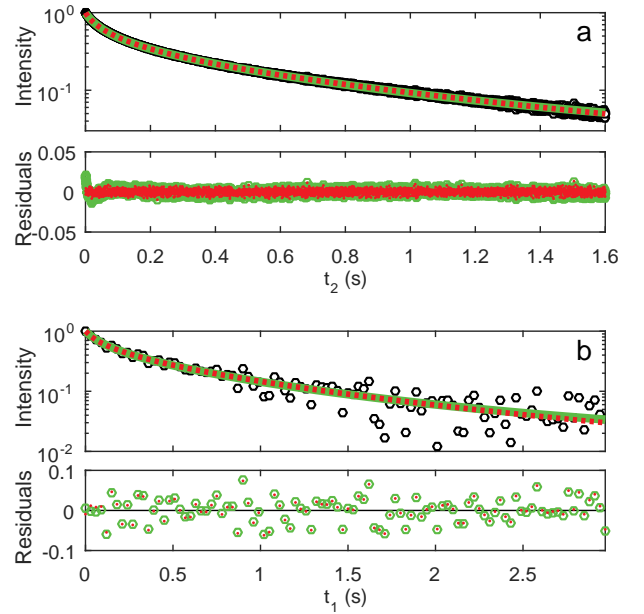


Figure 1: Results from using the P2DRM model sets A (red dotted line) and B (green solid line) to fit the simulated 1-D CPMG  $T_2$  data (a) and rapid  $T_1$  data (b), showing signal intensity (black circles), fits, and residuals.

The pseudo 2-D distributions estimated by the P2DRM are compared to the known joint 2-D distribution in Fig. 2. Both model sets estimated essentially the same  $T_1$ - $T_2$  relationship and both trace the known  $T_1$ - $T_2$  relationship. The delta functions are near the bulk relaxation limit, though inconsistent with one another and at  $T_1$  values below the known  $T_1$ - $T_2$  relationship. The inability of the model sets to describe the marginal distributions of the known joint  $T_1$ - $T_2$  distribution results in the distributed components being shifted to shorter relaxation times. Even so, the models still show good agreement with the known  $T_1$ - $T_2$  relationship. With respect to fitting 1-D data, we have found parametric distribution models which under-fit the data, e.g., using a single distributed component to model the decay from a multimodal distribution, are still capable of estimating the means,  $\langle T_2 \rangle$  or  $\langle T_1 \rangle$ . We can extend this to the P2DRM because obtaining the  $T_1$ - $T_2$  relationship in essence involves estimating the conditional  $\langle T_1 \rangle$  given a value of  $T_2$  (this is what the P2DRM  $T_1$ - $T_2$  relationship curve describes). For these reasons, P2DRM estimations of  $T_1$ - $T_2$  relationships are robust with respect to model assumptions.

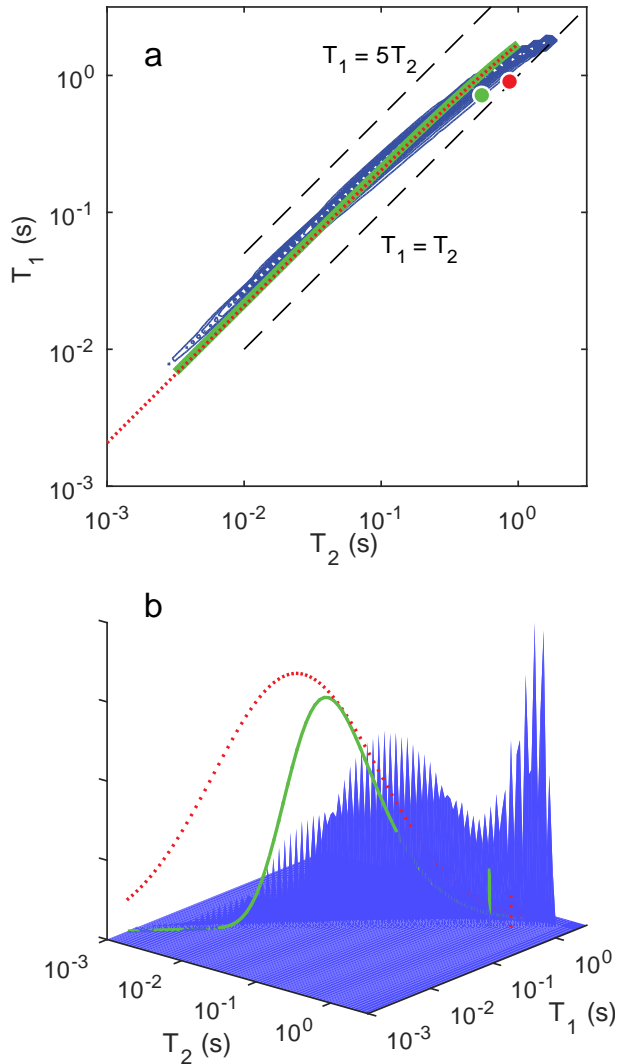


Figure 2: Results of the pseudo-2-D  $T_1$ - $T_2$  distributions from using P2DRM model sets A, based on the lognormal, (red dotted line) and B, based on the inverse-gamma, (green solid line) to fit the 1-D data sets compared to the known joint  $T_1$ - $T_2$  distribution (blue intensity map) highlighting (a) the  $T_1$ - $T_2$  relationships and (b) intensity of the distributions. The maximum intensities of the P2DRM distribution components are scaled by their weights.

Information about surface interactions can be gleaned from the ratio  $T_1:T_2$  for fluid interacting with surfaces and  $\varepsilon^{-1}$  is therefore the key parameter for the P2DRM to obtain. From fitting 100 data sets simulated with unique instances of random noise, estimates of  $\varepsilon^{-1}$  for model set A (mean=2.54, std=0.43) and model set B (mean=2.41, std=0.43) were not significantly different from the known mean of the distribution of  $\rho_2/\rho_1 = 2.37$ . Data with this SNR could be acquired in a timescale of minutes, versus a timescale of hours for the  $T_1$ - $T_2$  correlation experiment. Therefore, the P2DRM should indeed be useful in situations where experimental time is limited.

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