Abstract—This study documents validation of vertebral band-pair formation in spotted gully shark (Triakis megalopterus) with the use of fluorochrome injection and tagging of captive and wild sharks over a 21 -year period. Growth and mortality rates of T. megalopterus were also estimated and a demographic analysis of the species was conducted. Of the 23 OTC (oxytetracycline)-marked vertebrae examined ( 12 from captive and 11 from wild sharks), seven vertebrae (three from captive and four from wild sharks) exhibited chelation of the OTC and fluoresced under ultraviolet light. It was concluded that a single opaque and translucent band pair was deposited annually up to at least 25 years of age, the maximum age recorded. Reader precision was assessed by using an index of average percent error calculated at $5 \%$. No significant differences were found between male and female growth patterns ( $P>0.05$ ), and von Bertalanffy growth model parameters for combined sexes were estimated to be $L_{\infty}=1711.07 \mathrm{~mm} \mathrm{TL}, k=0.11 / \mathrm{yr}$ and $t_{0}=$ $-2.43 \mathrm{yr}(n=86)$. Natural mortality was estimated at $0.17 / \mathrm{yr}$. Age at maturity was estimated at 11 years for males and 15 years for females. Results of the demographic analysis showed that the population, in the absence of fishing mortality, was stable and not significantly different from zero and particularly sensitive to overfishing. At the current age at first capture and natural mortality rate, the fishing mortality rate required to result in negative population growth was low at $F>0.004 /$ yr. Elasticity analysis revealed that juvenile survival was the principal factor in explaining variability in population growth rate.

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# Age validation, growth, mortality, and demographic modeling of spotted gully shark (Triakis megalopterus) from the southeast coast of South Africa 

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Spotted gully shark (Triakis megalopterus, Smith, 1839), is one of five Triakis species and is endemic to southern Africa (Compagno, 1988). Its distribution range extends from northern Namibia (although anecdotal information indicates that it is caught as far north as Angola) southward around the coast to Coffee Bay in the Eastern Cape Province, South Africa (Compagno, 1988; Compagno et al., 2005). It is a shallow water ( $<50 \mathrm{~m}$ ) (Compagno et al., 1989; Smale and Goosen, 1999), demersal species that is recreationally important to shore and ski-boat anglers. With the exception of some information pertaining to its reproductive and feeding biology (Smale and Goosen, 1999) there is little information available to guide its management. Given its narrow distribution range and small population size, it could be vulnerable to overexploitation in a manner similar to its congener, T. semifasciata, that has declined in abundance and is now carefully managed (Smith and Abramson, 1990; Cailliet, 1992). An equivalent analysis is required for T. megalopterus to assess its vulnerability to fishing pressure.

Demographic modeling has been conducted on many elasmobranch populations when there are insufficient catch, effort, and abundance data available to conduct a full stock assessment (Simpfendorfer, 1998; Romine et al., 2009). Demographic modeling is a popular approach because it provides the best available description of the population being studied given several life history parameters. Demographic modeling therefore provides a compromise between simple life history tables and more detailed stock assessment models. Demographic models became popular in the 1990 s and are now the most widely used population models used to assess shark populations (Simpfendorfer, 2005). A fundamental requirement for the application of age-structured demographic models is that ages of sharks are available.

Correctly determining the age of fish, particularly elasmobranchs, is crucial if (unbiased) time-based life history rates such as growth, maturity, and mortality are to be estimated. Although numerous fish age and growth studies have been undertaken, remarkably few have included
validation (Campana, 2001; Cailliet and Goldman, 2004). Campana (2001) defines validation as either the validation of the periodicity of growth increments, or as the validation of the age estimate made by reader(s). In a recent review, Campana (2001) noted that of the 372 papers in which age validation was reported, only $15 \%$ of the papers actually incorporated validation of the absolute age of wild fish. Therefore, when conducting an age and growth study it is necessary to specify whether one or both validation goals are met and which validation method is used.

Three methods of validation are typically used. These are edge analysis (also known as marginal increment or zone analysis), carbon dating, and mark-recapture of tagged fish injected with a calcium-chelating fluorescing chemical such as the antibiotic, oxytetracycline hydrochloride (OTC). All of these methods have been used to validate elasmobranch ages (Campana et al., 2002; Smith et al., 2003; Cailliet and Goldman, 2004; Ardizzone et al., 2006; Chen et al., 2007). The most commonly used and most reliable validation method is tagging with OTC (Campana, 2001).

In studies of elasmobranchs by OTC-injection, sharks are measured, injected with OTC, tagged, and released. The date of release is then noted. Once the shark is recaptured, the number of vertebral band pairs distal to the fluorescent mark is compared with the time between release and recapture of the shark. This experimental approach can be problematic if recapture rates are low. Validating elasmobranch age estimates can be time consuming because recapture rates are generally low and there is a long period of time between collecting a sufficient number of samples and analyzing the vertebrae.

We developed a Leslie matrix-based demographic model to evaluate the ability of T. megalopterus to sustain increased levels and patterns of fishing pressure. We also validated the age of T. megalopterus using oxytetracycline to estimate growth and mortality rates, important demographic model parameters.

## Materials and methods

## General overview

In 1994, a tagging program incorporating researchers and trained volunteer fishermen was initiated at the Port Elizabeth Museum in order to obtain age validation, movement studies, and population dynamics of the spotted gully shark. A total of 402 wild sharks ( 113 male, 230 female, and 59 of undetermined sex) were tagged, injected with OTC at a dosage of $50 \mathrm{mg} / \mathrm{kg}$ (Tanaka, 1990), and released. A total of 53 sharks were recaptured once, and one shark twice. The date for the first recaptured fish was unfortunately unrecorded, but it was evident that it had been re-injected with OTC from the presence of an additional fluorescing mark in its vertebrae. An additional 12 display sharks were held in the Bayworld Aquarium in Port Elizabeth and injected with

OTC. All 12 display sharks and 11 tagged wild sharks were sacrificed for vertebral analysis.

In addition, a total of 129 spotted gully sharks were collected opportunistically from ski-boat fishermen, fishing competitions, and research cruises over a 21 -year period (1984-2009) between Cape St Francis and Coffee Bay, South Africa. Vertebrae were collected from a subsample of 96 sharks. Total length (TL) and sex were recorded for all these sharks.

## Age determination

Between five and eight vertebrae were removed from the trunk region in the vicinity of the first dorsal fin, soaked in $4.5 \%$ sodium hypochlorite for $15-45$ minutes to remove excess connective tissue, and were either stored in $70-80 \%$ ethyl alcohol or frozen (Yudin and Cailliet, 1990). Cleaned vertebrae were embedded in polyester casting resin and sectioned with a diamondbladed saw along the sagittal plane to a thickness of 0.6 mm (Natanson et al., 2006; Rizzo et al. ${ }^{1}$ ), and mounted on glass slides with DPX mountant (Lasec, South Africa).

Band pairs, defined as one optically opaque and one optically translucent band were counted by using a dissecting microscope with transmitted white light (460-490 nm). OTC-injected specimens were also viewed with an Olympus BX60 microscope (Olympus, Johannesburg, South Africa) under ultraviolet transmitted light ( $510-550 \mathrm{~nm}$ ). Each specimen was aged twice, three weeks apart by a single reader, without prior knowledge of the length or sex of the specimen. Counts were accepted only if both counts were in agreement. If the estimated number of bands differed by two or less, the specimen was recounted and the final count was accepted as the agreed upon number; if not, the specimen was discarded. If the third count did not concur with one of the previous two counts, the sample was rejected.

An age-bias plot was used to graphically assess the readings and their associated agreement (Campana, 2001; Natanson et al., 2006). A $t$-test was conducted on the slope of the age-bias plot (the linear regression of the second against the first age readings) to test the null hypothesis that the slope was equal to one. Comparisons of reader accuracy for each age were made by using a paired $t$-test, and a $\chi^{2}$-test of symmetry was used to test for systematic bias in the determination of age (Hoenig et al., 1995).

The variability of the within-reader age estimates was estimated with an index of average percent error

[^0](IAPE) (Beamish and Fournier, 1981) with the following equation:
$$
\operatorname{IAPE}(\%)=\frac{100}{N} \sum_{i=1}^{N}\left[\frac{1}{R} \sum_{j=1}^{R} \frac{\left|X_{i j}-X_{i}\right|}{X_{i}}\right]
$$
where $N=$ the number of fish aged;
$R=$ the number of readings;
$X_{i j}=j^{\text {th }}$ vertebral count of the $i^{\text {th }}$ fish; and
$X_{i}=$ the final agreed age of fish $i$.
As with Goosen and Smale (1997), an IAPE calculated to be less than $10 \%$ was considered acceptable.

Growth was modeled with the Schnute (1981) growth model. This four-parameter model is general and allows for the estimation of various nested models. The length of a shark at age $a, L_{a}$, is modeled as

$$
L_{a}=L_{1}^{\beta}+\left(L_{2}^{\beta}-L_{1}^{\beta}\right)\left[\frac{1-e^{-\alpha\left(\alpha-t_{1}\right)}}{1-e^{-\alpha\left(t_{2}-t_{1}\right)}}\right]^{1 / \beta},
$$

where $t_{1}=$ the youngest fish in the sample;
$t_{2}=$ the oldest fish in the sample;
$L_{1}=$ the estimated length of a fish at age $t_{1}$;
$L_{2}=$ the estimated length of fish at age $t_{2}$; and $\alpha$ and $\beta$ are the curvature parameters.

By setting the parameter $\beta$ to either 1 or -1 , the model reduces to either the von Bertalanffy or logistic growth model. Both of these nested models have three estimated parameters. The von Bertalanffy and logistic models are expressed as

$$
L_{a}=L_{\infty}\left(1-e^{-k\left(a-t_{0}\right)}\right) \text { and } L_{a}=\frac{L_{\infty}}{\left(1-e^{-k\left(a-t_{0}\right)}\right)}
$$

where $L_{\infty}=$ the theoretical maximum size;
$k=$ the growth coefficient; and $t_{0}=$ the theoretical age at zero length.

The von Bertalanffy ( $\beta=1$ ) and logistic ( $\beta=-1$ ) parameters were calculated with the following equations:

$$
\begin{gathered}
L_{\infty}=\left(\frac{e^{\alpha t_{2}} L_{2}^{\beta}-e^{\alpha t_{1}} L_{1}^{\beta}}{e^{\alpha t_{2}}-e^{\alpha t_{1}}}\right)^{1 / \beta}, \\
t_{0}=t_{1}+t_{2}-\frac{1}{\alpha} \ln \left(\frac{\beta\left(e^{\alpha t_{2}} L_{2}^{\beta}-e^{\alpha t_{1}} L_{1}^{\beta}\right)}{L_{2}^{\beta}-L_{1}^{\beta}}\right), \text { and } K=\alpha .
\end{gathered}
$$

The Schnute growth model was fitted to the com-bined-sex data and the parameters were estimated by nonlinear minimization of a negated normal log-likelihood function of the form

$$
-\ln L=\frac{n}{2} \ln \hat{\sigma}^{2},
$$

where $\hat{\sigma}^{2}=\frac{1}{n} \sum_{i=1}^{n}\left(L_{i a}-\hat{L}_{i a}\right)^{2}$ is the maximum likelihood esti$L_{i a}$ and $\hat{L}_{i a}$ are the observed and model predicted lengths of fish $i$ at age $a$; and $n$ is the number of observed data.

The most parsimonious model was selected with the lowest Akaike information criterion (Akaike, 1973) of the form

$$
A I C=2(-\ln L+p)
$$

where $p=$ the number of model parameters.
A likelihood ratio test was used to test the null hypotheses that there is no difference in growth parameters between males and females (Simpfendorfer et al., 2000; Neer et al., 2005; Natanson et al., 2006). Parameter variability was estimated by using parametric bootstrapping with 1000 iterations and the $95 \%$ confidence intervals were estimated from the bootstrap results by using the percentile method (Buckland, 1984).

## Age at maturity

Age at maturity, $t_{m}$, was estimated directly from the von Bertalanffy growth model as

$$
t_{m}=t_{0}-\frac{1}{k} \ln \left(1-\frac{l_{50}^{\psi}}{L_{\infty}}\right)
$$

where $l_{50}^{\psi}=$ the length at maturity obtained from Smale and Goosen (1999); and
$L_{\infty}, k$ and $t_{0}$ are the von Bertalanffy growth model parameters.

## Natural mortality

Natural mortality was estimated from the median of Pauly's (1980), Hoenig's (1983) and Jensen's (1996) empirical models of the form

$$
\begin{gathered}
M_{\text {Pauly }}= \\
\exp \left(-0.0152-0.279 \ln L_{\infty}+0.6543 \ln k+0.463 \ln T\right) \\
M_{\text {Hoenig }}=\exp \left(1.44-0.982 \ln t_{\max }\right) \\
M_{\text {Jensen }}=1.6 k
\end{gathered}
$$

where $L_{\infty}$ and $k$ are the Von Bertalanffy growth model parameters;
$T=$ the mean water temperature (estimated to be $16^{\circ} \mathrm{C}$ ); and
$t_{\max }=$ the age of the oldest fish sampled (Hoenig, 1983).

## Demographic modeling

The demographics of female T. megalopterus was modeled with an age-structured matrix model (Caswell, 2001) of the form $\mathbf{n}_{t+1}=\mathbf{A} \times \mathbf{n}_{t}$, where $\mathbf{n}_{\mathrm{t}}$ is a vector of numbers at age at time $t$ and $\mathbf{A}$ is the Leslie projection matrix. In matrix formulation the model is expressed as

$$
\left(\begin{array}{c}
N_{1, t+1} \\
N_{2, t+1} \\
N_{3, t+1} \\
\vdots \\
N_{t_{\max }, t+1}
\end{array}\right)=\left(\begin{array}{ccccc}
\phi_{1} & \phi_{2} & \phi_{3} & \cdots & \phi_{t_{\max }} \\
S_{1} & 0 & 0 & 0 & 0 \\
0 & S_{2} & 0 & 0 & 0 \\
0 & 0 & \ddots & 0 & 0 \\
0 & 0 & 0 & S_{t_{\max }-1} & 0
\end{array}\right) \times\left(\begin{array}{c}
N_{1, t} \\
N_{2, t} \\
N_{3, t} \\
\vdots \\
N_{t_{\max }, t}
\end{array}\right)
$$

where $S_{a}$ and $\phi_{a}=$ age-dependent survivals and fecundities, respectively; and
$t_{\text {max }}=$ the maximum age considered in the analysis.

The annual population growth rate ( $\lambda$ ), stable age distribution (w), and age-specific reproductive value (v) vectors were obtained by solving the equations $\mathbf{A w}=\lambda \mathbf{w}$ and $\mathbf{v}^{*} \mathbf{A}=\lambda \mathbf{w}$, where ${ }^{*}$ is the complex conjugate transpose function. In the solutions, $\lambda$ is the common dominant eigenvalue and $\mathbf{w}$ and $\mathbf{v}$ are the corresponding right and left eigenvectors. The reproductive value vector was normalized in relation to the age- 1 value.

The conditional intrinsic rate of increase (Gedamke et al., 2007) was calculated as $r=\ln \lambda$, the average age of mothers of newborn individuals in a population with a stable age distribution as $\bar{T}=\langle\mathbf{w}, \mathbf{v}\rangle=\mathbf{w}^{\prime} \mathbf{v}$, and the average number of female offspring per female during her lifespan as $R_{0}=\exp (r \bar{T})$.

The sensitivity of $\lambda$ to changes in the demographic parameters provides an indication of which parameter has the largest impact on the population growth rate. Sensitivity can be measured in either relative (as "sensitivity") or absolute terms (as "elasticity"). Both forms of sensitivity were calculated from the individual values of the Leslie matrix, $a_{i, j}$, the population growth rate, and the left-right eigenvectors as

$$
s_{i, j}=\frac{\partial \lambda}{\partial a_{i, j}}=\frac{v_{i} w_{j}}{\langle\mathbf{w}, \mathbf{v}\rangle}=\frac{\mathbf{v w}^{\prime}}{\langle\mathbf{w}, \mathbf{v}\rangle} .
$$

Elasticities, or

$$
\frac{\partial \ln \lambda}{\partial \ln a_{i, j}}
$$

were calculated as

$$
e_{i, j}=\frac{a_{i, j}}{\lambda} s_{i, j}
$$

such that

$$
\sum_{i} \sum_{j} e_{i, j}=1 .
$$

Elasticities were summarized by age

$$
E_{j}=\sum_{i} e_{i, j}
$$

fertilities

$$
E_{1}=\sum_{i=1}^{t_{\max }} e_{0, j}
$$

juvenile survival

$$
E_{2}=\sum_{i=2}^{t_{\text {max }}} \sum_{j=1}^{a_{50}} e_{i, j}
$$

and adult survival

$$
E_{2}=\sum_{i=2}^{t_{\text {max }}} \sum_{j=a_{50}}^{t_{\text {maxa }}} e_{i, j}
$$

(Mollet and Cailliet, 2002).

## Model implementation

Age-dependent survival was estimated as a function of both age-independent instantaneous natural mortality $(M)$, age-dependent selectivity $\left(\xi_{a}\right)$, and fully recruited fishing mortality $(F)$, such that $S_{a}=\exp \left(-M-\xi_{a} F\right)$. A maximum age of $t_{\max }=26$ was used in the analysis. All parameters used in the analysis are summarized in Table 1.

Age-dependent fertility was estimated as the number of embryos surviving to age-1 per female in a calendar year, $E_{a}$, and was weighted by the proportion of females that were mature at age $a, \psi_{a}$, such that $\phi_{a}=S_{0} \psi_{a} E_{a}$.

Because maturity, selectivity, and number of embryos are length- rather than age-dependent, all age-dependent values were calculated from the corresponding lengths predicted from the von Bertalanffy growth model for combined sexes.

Both maturity-at-length, $\psi_{l}$, and selection-at-length, $\xi_{l}$, were modeled as logistic ogives as

$$
\psi_{l}=\left(1+\exp \left(-\left(l-l_{50}^{\psi}\right) / \delta^{\psi}\right)\right)^{-1}
$$

and

$$
\xi_{l}=\left(1+\exp \left(-\left(l-l_{50}^{\xi}\right) / \delta^{\xi}\right)\right)^{-1}
$$

respectively, where $l_{50}^{\psi}$ and $l_{50}^{\xi}$ are the lengths at which shark were $50 \%$ mature or at which $50 \%$ of sharks were selected by the fishery. The inverse rates of maturation and selection are denoted as $\delta^{\psi}$ and $\delta^{\xi}$, respectively. Because these rates were not available from the literature, they were assumed to be $2 \%$ of their corresponding $50 \%$ values. These rates were considered reasonable given other maturity and selectivity studies (Booth, unpubl. data). Length at $50 \%$ selectivity was estimated as 1326 mm TL, which is the mean length of sharks ( $n=252$ ) measured from recreational anglers (Smale, unpubl. tag and release data). The number of female embryos per adult female at length $l$ in a calendar year, given a gestational period of 20 months and a sex ratio of $1: 1$, was calculated as

## Table 1

Fixed parameter values estimated during this study or obtained from Smale and Goosen's study (1999) that were used in the demographic analysis of Triakis megalopterus. TL=total length.

| Parameter | Description | Value | Source |
| :--- | :--- | :--- | :--- |
| $L_{\infty}$ | Theoretical maximum size | 1711.07 mm TL | This study |
| $k$ | Growth coefficient | $0.11 / \mathrm{yr}$ | This study |
| $t_{0}$ | Theoretical age at zero length | -2.43 yr | This study |
| $M$ | Natural mortality rate | $0.14 / \mathrm{yr}$ | This study |
| $F$ | Fishing mortality rate | $0 . \mathrm{yr}^{-1}$ | This study |
| $l_{50}^{\psi}$ | Length at $50 \%$ maturity | 1450 mm TL | Smale and Goosen (1999) |
| $l_{50}^{\xi}$ | Length at $50 \%$ selectivity | 1326 mm TL | This study |
| $\delta^{\psi}$ | Inverse rate of maturity | 29 mm TL | This study |
| $\delta^{\xi}$ | Inverse rate of selectivity | 26 mm TL | This study |
| $t_{\text {max }}$ | Maximum age | 26 years | This study |
| $E_{l}$ | Number of female embryos per adult female | $0.5 \times \frac{20}{12} \times(0.2 l-21.74)$ | Smale and Goosen (1999) |
|  | at length $l$ in a calendar year given a gestational |  |  |

$$
E_{l}=0.5 \times \frac{20}{12} \times(0.2 l-21.74)
$$

(Smale and Goosen, 1999).
Uncertainty in the model outputs was estimated with Monte-Carlo simulation (Cortés, 2002). For each simulation, $i$, random lengths were drawn around the predicted von Bertalanffy growth model with an estimated growth model standard deviation of 110 such that $L_{a}^{i}=L_{a}+\varepsilon_{p}$, and $\varepsilon_{p} \sim N\left(0,110^{2}\right)$. These length estimates were used to draw random variable for the length-atmaturity as $l_{50}^{i, \psi}=l_{50}^{\psi}+e_{\psi}$ and $\varepsilon_{\psi} \sim N\left(0,16^{2}\right)$. The standard deviation corresponded to that required to obtain the $1^{\text {st }}$ and $99^{\text {th }}$ percentiles at 1391 mm TL and 1500 mm TL, the lengths at first and $100 \%$ maturity reported by Smale and Goosen (1999). Natural mortality was assumed to be log-normally distributed with a coefficient of variation of $20 \%$, such that $M_{i}=M_{i} \exp \left(\varepsilon_{M}\right)$, and $\varepsilon_{M} \sim N\left(0,0.2^{2}\right)$. No adjustments were made for log-normal bias. A total of 1000 simulations were conducted, and standard error and $95 \%$ confidence intervals were calculated for $\lambda, \rho, \bar{T}, R_{0}, E_{a}, E_{1}, E_{2}$, and $E_{3}$ by using the percentile method (Buckland, 1984).

Three-dimensional isopleth plots were constructed to assess the response of the conditional intrinsic rate of population increase parameter, $r$, to different inputted combinations of fishing mortality and the age at which $50 \%$ of sharks were selected.

## Results

All aquarium specimens were either moribund or had died in the aquarium and had decreased in length by
the end of the study. These sharks were not included in the estimation of the growth parameters but were included in the validation aspect of the analysis. The vertebrae from 25 male ( $1123.2 \pm 404 \mathrm{~mm} \mathrm{TL}$ ) and 71 female ( $1258.0 \pm 397 \mathrm{~mm}$ TL) sharks were processed for age estimation. Vertebrae were interpreted without difficulty up to the margin of the corpus calcareum where magnification needed to be increased to accurately interpret the remaining band pairs (Fig. 1). No distinctive features were identified in reading the vertebrae of $T$. megalopterus.
Of the 96 vertebrae examined, 86 were considered suitable for aging. Age estimates ranged from 0+ to 25 years. Of the 15 OTC-injected specimens examined, only seven fluoresced (three captive and four wild specimens) under ultraviolet light and confirmed that one vertebral band pair was deposited annually (Table 2; Fig. 1). The maximum validated age was from a shark that was 25 years old. An age-length key is presented in Table 3.
There was no significant difference in the accuracy of age assessments between readings (paired $t$-test; $P>0.05$ ). There was a $45 \%$ agreement on all age assessments, $80 \%$ agreement between age readings within 1 year, and $94 \%$ agreement between readings within 2 years. There was a strong positive correlation between the first and second readings $\left(r^{2}=0.97\right)$ that was not statistically different from unity ( $t$-test on slopes; $P>0.05$ ) (Fig. 2). There was no evidence for systematic age bias between readings ( $\chi^{2}$-test; $P>0.05$ ). Band pair counts were considered to be reasonably precise with an estimated IAPE of $5.02 \%$.
Of the three models assessed, the von Bertalanffy was considered to be the most parsimonious (Schnute:


Figure 1
A sectioned vertebra of a 14 -year-old wild spotted gully shark (Triakis megalopterus) tagged and injected with oxytetracycline hydrochloride seven years before recapture.

Table 2
Sex, capture locations and dates, total lengths (TL), and estimated ages of specimens of Triakis megalopterus tagged and injected with oxytetracycline hydrochloride (OTC) that exhibited fluorescing zones on vertebrae. All specimens were sampled off the southeast coast of South Africa. "Zones distal to OTC"=zones that fluoresced and that were distal to the site of the OTC injection.

| Gender | Capture $\rightarrow$ recapture locations | Tagging date | Recapture date |  | Tagging length (mm TL) | $\begin{aligned} & \text { Recapture } \\ & \text { length } \\ & (\mathrm{mm} \mathrm{TL}) \end{aligned}$ | $\begin{gathered} \text { Zones } \\ \text { distal } \\ \text { to OTC } \end{gathered}$ | Total age |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Female | Bayworld aquarium | Feb 1999 | Dec 2000 | 1.8 | 1700 | 1580 | 1 | 20 |
| Female | Bayworld aquarium | Nov 2001 | Jan 2003 | 1.2 | 1660 | 1600 | 1 | 25 |
| Male | Bayworld aquarium | Apr 2002 | Dec 2002 | 0.7 | 1560 | 1428 | 1 | 11 |
| Female | De Hoop $\rightarrow$ De Hoop | Feb 1996 | Feb 1998 | 2.0 | 1111 | 1140 | 2 | 12 |
| Male | De Hoop $\rightarrow$ De Hoop | Sep 1996 | May 2003 | 6.7 | 1530 | 1550 | 7 | 14 |
| Female | Algoa Bay $\rightarrow$ Knysna | Mar 1998 | Mar 2005 | 7.0 | 950 | 1250 | 7 | 14 |
| Undetermined | De Hoop $\rightarrow$ De Hoop | Sep 2000 | May 2003 | 2.7 | 1020 | 1140 | 3 | 10 |

$p=4, \mathrm{AIC}=906.47$; von Bertalanffy: $p=3, \mathrm{AIC}=904.56$; logistic: $p=3$, AIC=912.89). The logistic model provided the worst fit of all three models considered. Although wild female sharks tended to grow larger and slower than males (Table 4, Fig. 3), there were no significant differences among any of the growth parameters (likelihood ratio test, $P>0.05$ ). Overall, the growth trajectories of the tagged and recaptured sharks were consistent with the predicted von Bertalanffy growth model for combined sexes (Fig. 4).

Age at maturity was calculated to be 10.9 years for males and 15.3 for females from the estimates of size
of maturity for males ( 1320 mm TL) and females (1450 mm TL) (Smale and Goosen, 1999). Natural mortality was estimated at $0.15 / \mathrm{yr}$, the median of $0.10 / \mathrm{yr}, 0.16 / \mathrm{yr}$ and $0.17 / \mathrm{yr}$, from the Pauly (1980), Hoenig (1983), and Jensen (1996) models, respectively.

Under the assumption of zero fishing mortality, the conditional intrinsic rate of increase parameter, $r$, was estimated at $0.00 \% / \mathrm{yr}$ and was not significantly different from zero based on the $95 \%$ confidence interval ( $P>0.05$ ) (Table 5). The average age of mothers of newborn individuals in the stable-age population was estimated at 18.76 years, and the average number of

## Table 3

Age-length key showing the number of fish aged within different length classes for Triakis megalopterus sampled on the southeast coast of South Africa. TL=total length. Sexes were combined.

|  | Age |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (mm TL) | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 | 21 | 22 | 23 | 24 | 25 |
| 200-399 | 6 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 400-599 | 5 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 600-799 |  |  | 3 | 2 |  |  | 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 800-999 |  |  |  | 1 | 3 | 1 | 2 | 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1000-1199 |  |  |  |  | 1 | 1 | 1 | 1 | 3 | 1 |  |  | 2 |  | 1 |  |  |  |  |  |  |  |  |  |  |  |
| 1200-1399 |  |  |  |  |  |  |  | 1 | 1 | 2 | 1 |  | 2 |  | 2 | 1 |  |  |  |  | 1 |  |  |  |  |  |
| 1400-1599 |  |  |  |  |  |  |  |  |  |  |  | 1 | 2 | 2 | 3 | 4 | 5 | 4 | 2 | 1 | 1 | 1 |  | 2 |  | 1 |
| 1600-1799 |  |  |  |  |  |  |  |  |  |  | 1 |  |  |  |  |  |  | 1 | 2 |  | 4 |  |  | 1 | 1 |  |

Table 4
Estimates for von Bertalanffy growth model parameters (and their associated variability and correlation matrices) for male, female, and combined-sexes of Triakis megalopterus sampled on the southeast coast of South Africa. $L_{\infty}=$ theoretical maximum size, $k=$ growth coefficient, $t_{0}=$ theoretical age at zero length, $\mathrm{SE}=$ standard error, $\mathrm{CV}=$ coefficient of variation, $95 \% \mathrm{CI}=95 \%$ confidence intervals, $\mathrm{TL}=$ total length.

| Parameter | Estimate | SE | CV | 95\% CI | $k$ | $t_{0}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Males ( $n=26$ ) |  |  |  |  |  |  |
| $L_{\infty}(\mathrm{mm}$ TL) | 1667.89 | 103.35 | 6.1\% | (1517.25 to 1909.17) | -0.95 | -0.63 |
| $k$ (/yr) | 0.12 | 0.02 | 17.4\% | (0.09 to 0.17) |  | 0.79 |
| $t_{0}$ (yrs) | -2.15 | 0.42 | 19.4\% | (-3.07 to -1.39) |  |  |
| Females ( $n=60$ ) |  |  |  |  |  |  |
| $L_{\infty}(\mathrm{mm} \mathrm{TL})$ | 1738.93 | 90.60 | 5.2\% | (1618.95 to 1969.19) | -0.95 | -0.70 |
| $k$ (/yr) | 0.10 | 0.02 | 16.2\% | (0.07 to 0.14 ) |  | 0.85 |
| $t_{0}$ (yrs) | -2.67 | 0.59 | 21.4\% | (-4.02 to -1.75) |  |  |
| Combined ( $n=86$ ) |  |  |  |  |  |  |
| $L_{\infty}(\mathrm{mm}$ TL) | 1711.07 | 56.42 | 3.3\% | (1615.79 to 1844.95) | $-0.93$ | -0.65 |
| $k$ (/yr) | 0.11 | 0.01 | 10.7\% | (0.09 to 0.13) |  | 0.82 |
| $t_{0}(\mathrm{yrs})$ | -2.43 | 0.35 | 14.2\% | (-3.18 to -1.80 ) |  |  |

female offspring per female during her lifespan was estimated at 1.00 (Table 5).

Over half (50.6\%) of the female stable-age distribution of this population was accounted for by sharks less than four years of age. Age- 1 sharks contributed $16 \%$ of the female stable-age distribution, and only $8 \%$ of the female stable-age distribution was mature ( $\geq 15$ years of age).

The elasticity of $\lambda$, the annual population growth rate, was relatively low for both fertility (5.5\%) and adult survival ( $14.4 \%$ ) and the highest for juvenile survival ( $80.2 \%$ ) (Table 5).

The conditional intrinsic rate of increase parameter, $r$, exhibited a nonlinear response when calculated for a variety of combinations of fishing mortality and age-at-capture scenarios (Fig. 5). In general, the conditional intrinsic rate of increase declined as fish-
ing mortality increased and the rate of decline was inversely proportional to age at first capture. When the conditional intrinsic rate of increase was modeled as a function of fishing mortality at the current natural mortality rate and age at $50 \%$ selectivity, a zero rate of increase was observed at $F=0.04 / \mathrm{yr}$.

## Discussion

In this study, T. megalopterus exhibited annular growth zone ftormation, depositing one band pair per year, as was also found in T. semifasciata (Smith, 1984; Kusher et al., 1992). The ratio of the number of sharks that exhibited fluorescence in their vertebrae ( $n=7$ ) to the number of sharks injected ( $n=23$ ) was unfortunately low and may (see Smith, 1984) or may

## Table 5

Summary of demographic parameters and elasticities estimated during this study for Triakis megalopterus and four other triakid shark species from Cortés (2002). Values in parenthesis denote lower and upper $95 \%$ confidence intervals calculated from 1000 Monte-Carlo simulations. $\lambda=$ annual population growth rate, $r=$ conditional intrinsic rate of increase (calculated as $r=\ln$ ), $\bar{T}=$ average age of mothers of newborn individuals in a population with a stable age distribution, $R_{0}=$ average number of female offspring per female during her lifespan (calculated as $R_{0}=\exp (r \bar{T})$ ).

| Species | $\lambda$ | $r$ | $\bar{T}$ | $R_{0}$ | Elasticity (\%) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Fecundity | Juvenile survival | Adult survival |
| Triakis megalopterus | $\begin{gathered} 1.000 \\ (0.976-1.041) \end{gathered}$ | $\begin{gathered} 0.000 \\ (-0.024-0.040) \end{gathered}$ | $\begin{gathered} 18.76 \\ (14.44-19.86) \end{gathered}$ | $\begin{gathered} 1.00 \\ (0.63-1.83) \end{gathered}$ | 5.4 | 78.3 | 16.4 |
| Mustelus californicus | 1.132 | 0.124 | 4.6 | 1.77 | 18.5 | 34.7 | 46.8 |
| Mustelus manazo | 1.096 | 0.092 | 6.6 | 1.83 | 13.3 | 52.4 | 34.3 |
| Mustelus antarcticus | 1.082 | 0.079 | 11.5 | 2.48 | 8.0 | 50.6 | 41.3 |
| Triakis semifasciata | 1.016 | 0.016 | 18.5 | 1.34 | 5.1 | 63.9 | 31.0 |



Age-bias plot of estimated ages from Triakis megalopterus vertebrae. Sharks were sampled on the southeast coast of South Africa. The age-bias plot regresses the age estimates from a second reading. IAPE=index of average percent error.
not (see McFarlane and King, 2009) be related to the slow growth and low levels of calcification of the vertebrae of these long lived sharks. Absolute age was validated up to 25 years. Oxytetracycline was shown to be an effective growth marker for T. megalopterus, which, like its congener T. semifasciata, exhibits slow growth and a maximum age of at least 25 years (Smith, 1984; Smith et al., 2003).

Female T. megalopterus were found to grow larger than males-a finding similar to that of Kusher et al. (1992) for T. semifasciata. In both studies, males grew to a similar size but female T. megalopterus were considerably larger. Kusher et al. (1992) found that


Figure 3
Observed total lengths of male and female Triakis megalopterus sampled from the southeast coast of South Africa. Length data were overlaid with the predicted von Bertalanffy growth model (solid line) and its 95\% confidence intervals (dotted lines) estimated by using parametric bootstrapping.
their growth coefficient estimates of $0.07 / \mathrm{yr}$ for females, $0.09 / \mathrm{yr}$ for males, and $0.08 / \mathrm{yr}$ for both sexes combined were lower than what would be predicted from Holden's (1974) method of estimation with the ratios of length at birth to maximum observed length (range $=0.1-0.2 / \mathrm{yr}$ ). Using a birth size of 435 mm and a female maximum observed length of 2075 mm TL together with a gestation period of 1.7 years (Smale and Goosen, 1999), Holden's (1974) estimated the growth coefficient at $0.12 / \mathrm{yr}$, which corresponds well with the growth curve estimate.

Our estimate of age at maturity for T. megalopterus exceeds that of T. semifasciata ( 7 years for males, 10


Figure 4
Observed length at age (dots) and von Bertalanffy predicted growth (line) of captive and wild Triakis megalopterus (sexes combined) sampled from the southeast coast of South Africa. The dotted lines between the dots represent the growth in length between successive ages after validation with oxytetracycline hydrochloride.
years for females) estimated by Kusher et al. (1992). In T. megalopterus, sexual maturation occurs at approximately $79 \%$ and $83 \%$ of asymptotic length for males and females, respectively. In terms of maximum observed length, sexual maturation occurs at $86 \%$ and $70 \%$ of asymptotic length. These ratios were higher than those of T. semifasciata, which are $63 \%$ and $72 \%$ for males and females, respectively (Kusher et al., 1992). Although our estimates are high, and could possibly be improved with larger sample sizes, they indicate that T. megalopterus is vulnerable to overexploitation and that management measures (particularly exemption of this species from commercial exploitation in South Africa) need to be rigorously enforced to protect this endemic species (Compagno et al., 1989).

Unfortunately all captive specimens examined in this study exhibited retarded growth. In some teleosts, this has been attributed to an effect of OTC (Monaghan, 1993). In both Japanese wobbegong (Orectolobus japonicas) and nurse sharks (Ginglymostoma cirratum), Tanaka (1990) and Gelsleichter et al. (1998) showed that OTC, at concentrations ranging from 20 to 80 $\mathrm{mg} / \mathrm{kg}$, had little adverse effects on growth or health. It is therefore unlikely that OTC, administered at 50 $\mathrm{mg} / \mathrm{kg}$, would have caused growth retardation, and there appears to be an alternative explanation for the negative growth observed -one that may be related to temperature stress.

Sustained water temperature, if beyond the optimal range of the species, can have an adverse affect on growth. Selong et al. (2001) showed that if an upper temperature threshold is exceeded, the result is inhibited feeding and later, death. Summer temperatures in the Bayworld Aquarium exceeded the upper threshold of


Figure 5
Isopleth of the conditional intrinsic rate of population increase ( $r$ ) for Triakis megalopterus at different combinations of age at $50 \%$ selectivity and fishing-induced mortality rate. The dashed line illustrates the current age at $50 \%$ selectivity of 11.13 years. The inset figure illustrates the conditional intrinsic rate of population increase as a function of fishing mortality at the current age at $50 \%$ selectivity.
the natural temperature range tolerated by spotted gully sharks in their natural habitat-occasionally spiking at $27^{\circ} \mathrm{C}$-whereas the normal temperature range off the Eastern Cape coast is approximately $12^{\circ} \mathrm{C}$ to $21^{\circ} \mathrm{C}$ in waters of $<20 \mathrm{~m}$. All the captive sharks were observed to reduce their food consumption and were noticeably thinner during summer than in winter, indicating that the summer elevated temperatures were exerting physiological stress on these display specimens. Whereas wild sharks are able to respond to high temperatures by swimming to deeper or cooler waters, this option is not available to captive individuals and they may not be able to avoid temperature stress. The aquarium animals used in this study were removed from display and were euthanized when they began to show signs of lethargy and exhibited marked weight loss and heat stress. It should be noted that some individuals of this species were able to cope with the high temperatures and survived multiple years, indicating that there is individual variation in susceptibility to heat stress.

The growth of the specimens over a seven-year period followed the growth curve and supports the robustness of the von Bertalanffy growth model for this species. Meaningful estimates of length between capture and recapture events in slow growing, long-lived species can only be attained if the time interval between the two events is large. If the time interval is too small, measurement error would be high in relation to the
gain in growth, particularly when measuring a robust, lively shark. The impact of this error would therefore be reduced when sampling intervals are years apart and more growth has occurred. Where growth is modest and measuring error occurs, the scope for inaccurate size estimates is greater. For this reason, size estimates of sharks from aquaria may be less reliable because of the relatively short time between measurements. However, this study would best be repeated in aquaria without unseasonably high summer temperatures.

The results from the demographic model indicate that T. megalopterus can sustain very limited fishing pressure, and therefore these results reinforce the biological interpretation of its life history parameters. It has been shown that species that are long lived, have a low rate of natural mortality, and produce few offspring per year cannot sustain high levels of fishing pressure (Holden, 1974; Cailliet, 1992; Simpfendorfer, 2005; Dulvey and Forrest, 2010). Even small increases in fishinginduced mortality, particularly at the current size of selection will negatively impact the population. At 11 years, sharks are harvested 4 years before the onset of sexual maturity. This age is of concern because the model predicts that an average age of a female shark is around 19 years and it will produce only a single female offspring over her lifetime. Any additional increase in fishing mortality would further decrease the number of adults and possibly contribute to recruitment overfishing. There are numerous examples in the literature of overexploitation and even extirpation of populations of chondrichthyans because their life history parameters are not understood and/or taken into consideration in management scenarios (Dulvey and Forrest, 2010).

The life history parameters of Triakis megalopterus were similar to those of its congener, T. semifasciata. Both sharks live to about 25 years of age, have similar natural mortality rates, and produce similar numbers of embryos per annum. As a result, it is not surprising that the demographic model applied to both species shows similar trends. Cailliet (1992) recommended that fishing mortality be reduced to 0.5 M , and that the size at capture be increased to prevent a decline in abundance. We determined in this study that, at current selection levels, the fishing mortality rate required for a stable population size would be $0.02 M$. Even if the age at $50 \%$ capture were increased significantly to 20 years, T. megalopterus would not be able to sustain even moderate levels of fishing mortality. Any possible increases in fishing mortality should, therefore, be closely monitored.

Triakis megalopterus is legislated as a noncommercial species with zero commercial harvest. This species is, however, targeted by recreational anglers. From personal observations, there has been a steady increase in recreational anglers targeting elasmobranchs because of a reduction in the availability of other favoured teleost species. Increased targeted fishing of this shark species could possibly result in higher levels of postrelease mortality from hooking and handling. Despite its noncommercial status, T. megalopterus is unfortunately
mistaken as Mustelus mustelus in a small, yet developing, inshore shark longline fishery. Even if bycatch rates were to remain constant, the increased catches in this developing fishery would result in obvious increases in mortality from commercial fishing.
Of the triakids that have been demographically modeled (Table 5), the Mustelus species appear to be the most resilient to harvesting pressure and have been shown to support sustainable fisheries (Walker, 1992; Chiaramonte, 1998; Francis and Shallard, 1999). The two reef-associated Triakis species have less habitat available and correspondingly smaller population sizes. Given their reduced habitat and life history characteristics, it is not surprising that their populations will decrease with commercial harvesting.

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